# Verifying Snapping Mechanism - Ideal Version

In order to verify the differential privacy proeprty of the snapping mechanism[3], we follow the logic rules designed from [1].

Some new rules are added into this logic in Figure 1 following with correctness proof. Then we formalized the snapping mechanism and verified its differential privacy property under these logic rules.

# 1 Extended Programming Logic[1]

## **Definition 1 (Laplce mechanism [2])**

Let  $\epsilon > 0$ . The Laplace mechanism  $\mathcal{L}_{\epsilon} : \mathbb{R} \to \mathsf{Distr}(\mathbb{R})$  is defined by  $\mathcal{L}(t) = t + \nu$ , where  $\nu \in \mathbb{R}$  is drawn from the Laplace distribution  $\mathsf{laplce}(\frac{1}{\epsilon})$ .

### **Definition 2**

Let  $\epsilon \leq 0$ . The  $\epsilon$ -DP divergence  $\Delta_{\epsilon}(\mu_1, \mu_2)$  between two sub-distributions  $\mu_1 \in \mathsf{Distr}(U)$ ,  $\mu_2 \in \mathsf{Distr}(U)$  is defined as:

$$\sup_{E \in U} \left( \Pr_{x \leftarrow \mu_1} [x \in E] - \exp(\epsilon) \Pr_{x \leftarrow \mu_2} [x \in E] \right]$$

### **Definition 3** $((\epsilon, \delta)$ - **lifting** [1])

Two sub-distributions  $\mu_1 \in \mathsf{Distr}(U_1)$ ,  $\mu_2 \in \mathsf{Distr}(U_2)$  are related by the  $(\varepsilon, \delta)$  - dilation lifting of  $\Psi \subseteq U_1 \times U_2$ , written  $\mu_1 \Psi^{\#(\varepsilon, \delta)} \mu_2$ , if there exist two witness sub-distributions  $\mu_L \in \mathsf{Distr}(U_1 \times U_2)$  and  $\mu_R \in \mathsf{Distr}(U_1, U_2)$  s.t.:

- 1.  $\pi_1(\mu_L) = \mu_1$  and  $\pi_2(\mu_R) = \mu_2$ ;
- 2.  $supp(\mu_L) \subseteq \Psi$  and  $supp(\mu_R) \subseteq \Psi$ ; and
- 3.  $\Delta_{\epsilon}(\mu_L, \mu_R) \leq \delta$ .

The logic rules we are using in our work is presented in Figure 1. The correctness of rules is proved in Theorem 1 and Theorem 2

### Theorem 1

Let  $\mu_1 \in \mathsf{Distr}(\mathbb{R})$ ,  $\mu_2 \in \mathsf{Distr}(\mathbb{R})$  are defined:

$$\mu_1(x) = \operatorname{unif}(x)$$

$$\mu_2(y) = \text{unif}(y)$$

where unif is uniform distribution over [0, 1) whoes pdf. is defined as:

$$\mathsf{pdf}_{\mathsf{unif}}(x) = \begin{cases} 1 & x \in [0,1) \\ 0 & o.w. \end{cases}.$$

$$\frac{-1}{ \vdash u_1 \stackrel{\$}{\leftarrow} \mu \sim_{\epsilon,0} u_2 \stackrel{\$}{\leftarrow} \mu : \top \Rightarrow e^{-\epsilon} u_2 \leq u_1 \leq e^{\epsilon} u_2} \text{ AXUNIF}$$

$$\frac{-1}{ \vdash u_1 \stackrel{\$}{\leftarrow} \mathcal{L}_{\epsilon}(e_1) \sim_{k' \cdot \epsilon,0} y_2 \stackrel{\$}{\leftarrow} \mathcal{L}_{\epsilon}(e_2) : |k + e_1 - e_2| \leq k' \Rightarrow y_1 + k = y_2} \text{ LAPGEN}$$

$$\frac{-1}{ \vdash u_1 \stackrel{\$}{\leftarrow} \mathcal{L}_{\epsilon}(e_1) \sim_{0,0} y_2 \stackrel{\$}{\leftarrow} \mathcal{L}_{\epsilon}(e_2) : \top \Rightarrow y_1 - y_2 = e_1 - e_2} \text{ LAPNULL}$$

$$\frac{-1}{ \vdash u_1 \stackrel{\$}{\leftarrow} \mu \sim_{0,0} y_2 \stackrel{\$}{\leftarrow} \mu : \top \Rightarrow y_1 = y_2} \text{ AXNULL}$$

$$\frac{-1}{ \vdash u_1 \stackrel{\$}{\leftarrow} \mu \sim_{0,0} y_2 \stackrel{\$}{\leftarrow} \mu : \top \Rightarrow y_1 = y_2} \text{ AXNULL}$$

$$\frac{-1}{ \vdash u_1 \stackrel{\$}{\leftarrow} \mu \sim_{0,0} \mu_2 : \Phi_1 \Rightarrow \Phi_1' \qquad c_1 \sim_{k',0} c_2 : \Phi_1' \Rightarrow \Phi_2} \text{ COMP}$$

$$\frac{-1}{ \vdash u_1 \stackrel{\$}{\leftarrow} \mu \sim_{0,0} \mu_2 : \Phi_1 \Rightarrow \Phi_1' \qquad c_1 \sim_{k',0} c_2 : \Phi_1' \Rightarrow \Phi_2} \text{ COMP}$$

$$\frac{-1}{ \vdash u_1 \stackrel{\$}{\leftarrow} \mu \sim_{0,0} \mu_2 : \Phi_1 \Rightarrow \Phi_1' \qquad c_1 \sim_{k',0} c_2 : \Phi_1' \Rightarrow \Phi_2} \text{ COMP}$$

$$\frac{-1}{ \vdash u_1 \stackrel{\$}{\leftarrow} \mu \sim_{0,0} \mu_2 : \Phi_1 \Rightarrow \Phi_1' \qquad c_1 \sim_{k',0} c_2 : \Phi_1' \Rightarrow \Phi_2} \text{ COMP}$$

$$\frac{-1}{ \vdash u_1 \stackrel{\$}{\leftarrow} \mu \sim_{0,0} \mu_2 : \Phi_1 \Rightarrow \Phi_1' \qquad c_1 \sim_{k',0} c_2 : \Phi_1' \Rightarrow \Phi_2} \text{ COMP}$$

$$\frac{-1}{ \vdash u_1 \stackrel{\$}{\leftarrow} \mu \sim_{0,0} \mu_2 : \Phi_1 \Rightarrow \Phi_1' \qquad c_1 \sim_{k',0} c_2 : \Phi_1' \Rightarrow \Phi_2} \text{ COMP}$$

Figure 1: Logic Rules Extended from [1]

Then,  $\mu_1 \Psi^{\#(\epsilon,0)} \mu_2$ , where

$$\Psi = \{(x, y) \in \mathbb{R} \times \mathbb{R} | x \cdot e^{-\epsilon} \le y \le x \cdot e^{\epsilon} \}$$

### Theorem 2

For any distributions  $\mu_1 \in \mathsf{Distr}(\mathbb{R}), \ \mu_2 \in \mathsf{Distr}(\mathbb{R}), \ \mu_1 \Psi^{\#(0,0)} \mu_2$ , where

$$\Psi = \{(x, y) \in \mathbb{R} \times \mathbb{R} | x = y\}$$

*Proof of Theorem 1.* Existing  $\mu_L, \mu_R \in \mathsf{Distr}(\mathbb{R} \times \mathbb{R})$ :

$$\mu_L(x,y) = \begin{cases} \mathsf{unif}(x) & x \cdot e^{-\epsilon} = y \land x \in [0,1) \\ 0 & o.w. \end{cases} \\ \mu_R(x,y) = \begin{cases} \mathsf{unif}(y) & x \cdot e^{-\epsilon} = y \land y \in [0,1) \\ 0 & o.w. \end{cases}.$$

Their pdf. are defined:

$$\mathsf{pdf}_{\mu_L}(x,y) = \begin{cases} \mathsf{pdf}_{\mathsf{unif}}(x) & x \cdot e^{-\epsilon} = y \land x \in [0,1) \\ 0 & o.w. \end{cases}$$

$$\mathsf{pdf}_{\mu_R}(x,y) = \begin{cases} \mathsf{pdf}_{\mathsf{unif}}(y) & x \cdot e^{-\epsilon} = y \land y \in [0,1) \\ 0 & o.w. \end{cases}.$$

- $supp(\mu_L) \in \Psi \land supp(\mu_R) \in \Psi$ 
  - supp $(\mu_L) \subseteq \Psi$ By definition of the pdf of  $\mu_L$ , we have:  $\Pr_{(x,y) \stackrel{\$}{\leftarrow} \mu_L} [(x,y) \notin \Psi] = 0.$

Then we can derive  $supp(\mu_L) \in \Psi$ 

- supp $(\mu_R) \subseteq \Psi$ By definition of the pdf of  $\mu_R$ , we have:  $\Pr_{(x,y) \overset{\$}{\leftarrow} \mu_R} [(x,y) \notin \Psi] = 0$ .

Then we can derive  $supp(\mu_L) \in \Psi$ 

• 
$$\pi_1(\mu_L) = \mu_1 \wedge \pi_2(\mu_R) = \mu_2$$

$$- \pi_1(\mu_L) = \mu_1$$

By definition of the  $\pi_1$  and pdf of  $\mu_L$ , we have  $\forall x \in \mathbb{R}$ :

$$\mathsf{pdf}_{\pi_1(\mu_L)}(x) = \begin{cases} \int_{\mathcal{Y}} \mathsf{pdf}_{\mathsf{unif}}(x) & (x,y) \in \Psi \land x \in [0,1) \\ 0 & o.w. \end{cases} = \begin{cases} \mathsf{pdf}_{\mathsf{unif}}(x) & x \in [0,1) \\ 0 & o.w. \end{cases} = \mathsf{pdf}_{\mu_1}(x)$$

$$- \pi_1(\mu_R) = \mu_2$$

Equivalent to showpdf<sub> $\pi_2(\mu_R)$ </sub> = pdf<sub> $\mu_2$ </sub>.

By definition of the  $\pi_2$  and pdf of  $\mu_R$ , we have  $\forall y \in \mathbb{R}$ :

$$\mathsf{pdf}_{\pi_{2}(\mu_{R})}(y) = \begin{cases} \int_{x} \mathsf{pdf}_{\mathsf{unif}}(y) & (x,y) \in \Psi \land y \in [0,1) \\ 0 & o.w. \end{cases} = \begin{cases} \mathsf{pdf}_{\mathsf{unif}}(y) & y \in [0,1) \\ 0 & o.w. \end{cases} = \mathsf{pdf}_{\mu_{2}}(y)$$

•  $\Delta_{\epsilon}(\mu_L, \mu_R) \leq 0$ 

By definition of  $\epsilon$ -DP divergence, we have:

$$\Delta_{\epsilon}(\mu_{L}, \mu_{R}) = \sup_{S} \left( \Pr_{(x,y) \overset{\$}{\leftarrow} \mu_{L}} [(x,y) \in S] - e^{\epsilon} \Pr_{(x,y) \overset{\$}{\leftarrow} \mu_{R}} [(x,y) \in S] \right)$$
$$= \sup_{S} \left( \int_{(x,y) \in S} \mathsf{pdf}_{\mu_{L}}(x,y) - e^{\epsilon} \int_{(x,y) \in S} \mathsf{pdf}_{\mu_{R}}(x,y) \right)$$

**case**  $S \subseteq \{(x, y) | x \in [0, 1) \land x \cdot e^{-\epsilon} = y\}$ :

$$\begin{array}{ll} \Delta_{\epsilon}(\mu_L,\mu_R) &= \int_{(x,y)\in S} \mathsf{pdf}_{\mathsf{unif}}(x) - e^{\epsilon} \int_{(x,y)\in S} \mathsf{pdf}_{\mathsf{unif}}(y) \\ &= \int_{(x,y)\in S} \mathsf{pdf}_{\mathsf{unif}}(x) - e^{\epsilon} \int_{(x,y)\in S} \mathsf{pdf}_{\mathsf{unif}}(x * e^{-\epsilon}) \\ &= \int_{(x,y)\in S} \mathsf{pdf}_{\mathsf{unif}}(x) - e^{\epsilon} * e^{-\epsilon} \int_{(x,y)\in S} \mathsf{pdf}_{\mathsf{unif}}(x) \\ &= 0 \end{array}$$

**case**  $S \subseteq \{(x, y) | x \in [1, e^{\epsilon}) \land x \cdot e^{-\epsilon} = y\}$ :

$$\Delta_{\epsilon}(\mu_L, \mu_R) = 0 - e^{\epsilon} \int_{(x,y) \in S} \mathsf{pdf}_{\mathsf{unif}}(y)$$
  
< 0

case o.w.

$$\Delta_{\epsilon}(\mu_L, \mu_R) = 0 - 0 = 0$$

# 2 Formalization of Snap Mechanism in Probabilistic Logic

**Definition 4** (Snap(a):  $A \rightarrow Distr(B)$ )

The ideal Snapping mechanism Snap(a) is defined as:

$$u \stackrel{\$}{\leftarrow} \mu; y = \frac{\ln(u)}{\epsilon}; s \stackrel{\$}{\leftarrow} \{-1, 1\}; z = s * y; x = f(a); w = x + z; w' = \lfloor w \rceil_{\Lambda}; r = \mathsf{clamp}_{B}(w')$$

where f is the query function over input  $a \in A$ ,  $\epsilon$  is the privacy budget, B is the clampping bound and  $\Lambda$  is the rounding argument satisfying  $\lambda = 2^k$  where  $2^k$  is the smallest power of 2 greater or equal to the  $\frac{1}{\epsilon}$ .

3

$$\overline{u_1 \overset{\$}{\leftarrow} \mu \sim_{\epsilon,0} u_2 \overset{\$}{\leftarrow} \mu : \top \Rightarrow e^{-\epsilon} u_2 \leq u_1 \leq e^{\epsilon} u_2} \xrightarrow{\text{AXNULL}} \xrightarrow{\text{AXNULL}}$$

$$\overline{y_1 = \frac{\ln(u_1)}{\epsilon} \sim_{0,0} y_2 = \frac{\ln(u_2)}{\epsilon} : e^{-\epsilon} u_2 \leq u_1 \leq e^{\epsilon} u_2 \Rightarrow y_2 - 1 \leq y_1 \leq 1 + y_2} \xrightarrow{\text{AXNULL}} \xrightarrow{\text{AXNULL}}$$

$$\overline{s_1 \overset{\$}{\leftarrow} \{-1, 1\} \sim_{0,0} s_2 \overset{\$}{\leftarrow} \{-1, 1\} : \top \Rightarrow s_1 = s_2} \xrightarrow{\text{AXNULL}}$$

$$\overline{z_1 = s_1 * y_1 \sim_{0,0} z_2 = s_2 * y_2 : s_1 = s_2 \wedge y_2 - 1 \leq y_1 \leq 1 + y_2 \Rightarrow |z_1 - z_2| \leq 1} \xrightarrow{\text{AXNULL}}$$

$$\overline{x_1 = f(a_1) \sim_{0,0} x_2 = f(a_2) : a_1 = a_2 + 1 \wedge f(a_1) = f(a_2) + 1 \Rightarrow x_1 = x_2 + 1} \xrightarrow{\text{AXNULL}}$$

$$\overline{w_1 = x_1 + z_1 \sim_{0,0} w_2 = x_2 + z_2 : x_1 = x_2 + 1 \wedge |z_1 - z_2| \leq 1 \wedge -2 \leq k \leq 0 \Rightarrow w_1 + k = w_2} \xrightarrow{\text{AXNULL}}$$

$$\overline{w_1' = \lfloor w_1 \rfloor_{\Lambda} \sim_{0,0} w_2' = \lfloor w_2 \rfloor_{\Lambda} : w_1 + k = w_2 \wedge -2 \leq k \leq 0 \Rightarrow w_1' + k = w_2'} \xrightarrow{\text{AXNULL}}$$

$$\overline{r_1 = \text{clamp}_B(w_1') \sim_{0,0} r_2 = \text{clamp}_B(w_2') : w_1' + k = w_2' \wedge -2 \leq k \leq 0 \Rightarrow r_1 + k = r_2} \xrightarrow{\text{AXNULL}}$$

$$\cdots$$

$$r_1 = \text{Snap}(a_1) \sim_{\epsilon,0} r_2 = \text{Snap}(a_1) : a_1 = a_2 + 1 \wedge f(a_1) = f(a_2) + 1 \wedge |k + f(a_1) - f(a_2)| \leq 1 \Rightarrow r_1 + k = r_2} \xrightarrow{\text{Comp}}$$

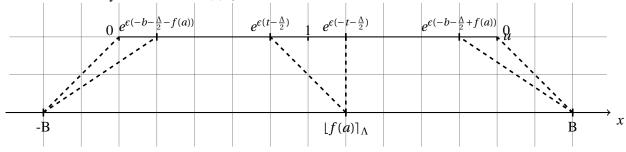
Figure 2: Coupling Derivation of two Snap mechanisms:  $Snap(a_1)$ ,  $Snap(a_2)$ 

### Theorem 3 (The Snap mechanism is $\epsilon$ -differentially praivate)

*Proof.* The proof follows the derivation in Figure 2.

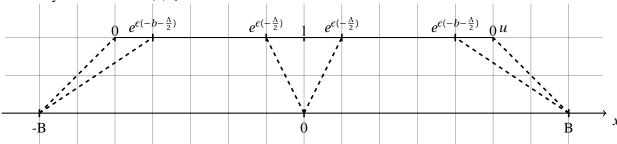
# 3 Proof of Differential Privacy for Snap Mechanism

Assume x be the output of Snap mechanism, we have following maps from the output of Snap mechanism to uniformly distributed  $u \in (0,1]$ .



where *b* is the greatest rounding of  $\Lambda$  that is smaller than *B* and  $t = \lfloor f(a) \rceil_{\Lambda} - f(a)$ .

Given the  $f(a) = \lfloor f(a) \rceil_{\Lambda} = 0$ , we have following maps from the output of Snap mechanism to uniformly distributed  $u \in (0,1]$ .



Assuming that  $f(a) \in [-B, B]$ , otherwise we can always redefine the f(a) restricting its output in this range. The probability of obtaining output x from Snap mechanism can be calculated by cases of x:

case x = -B

In this case, we know s = 1.

We have:  $\lfloor f(a) + \frac{1}{\epsilon} \ln(u) \rceil_{\Lambda} \le -B$ .

Since b is the greatest rounding of  $\Lambda$  that is smaller than B, then -b is the smallest rounding of  $\Lambda$  that is greater than -B, we have:  $f(a) + \frac{1}{\epsilon} \ln(u) < -b - \frac{\Lambda}{2}$ .

Then we get:  $u \in (0, e^{\epsilon(-b-\frac{\Lambda}{2}-f(a))})$ 

**case**  $x \in (-B, \lfloor f(a) \rceil_{\Lambda})$ 

In this case, we know s = 1 and  $x \in [-b, \lfloor f(a) \rceil_{\Lambda} - \Lambda]$ .

We have:  $\lfloor f(a) + \frac{1}{6} \ln(u) \rceil_{\Lambda} = x$ .

By the rule of rounding, we get:  $u \in \left[e^{\epsilon(x-\frac{\Lambda}{2}-f(a))}, e^{\epsilon(x+\frac{\Lambda}{2}-f(a))}\right]$ .

By the range of x, we get:  $u \in \left[e^{\epsilon(-b-\frac{\Lambda}{2}-f(a))}, e^{\epsilon(t-\frac{\Lambda}{2})}\right]$ .

case  $x = \lfloor f(a) \rceil_{\Lambda}$ 

**subcase** s = 1

In this case, we have:  $\lfloor f(a) + \frac{1}{\epsilon} \ln(u) \rceil_{\Lambda} = \lfloor f(a) \rceil_{\Lambda}$ .

Then we can get:  $u \in \left[e^{\epsilon(t-\frac{\Lambda}{2})}, e^{\epsilon(t+\frac{\Lambda}{2})}\right]$ .

Since  $t = \lfloor f(a) \rceil_{\Lambda} - f(a)$ , we know:  $-\frac{\Lambda}{2} \le t \le \frac{\Lambda}{2}$ . So we can get:  $e^{\epsilon(t + \frac{\Lambda}{2})} > 1$ .

Since  $u \in (0,1]$ , we have:  $u \in \left[e^{\epsilon(t-\frac{\Lambda}{2})}, 1\right]$ .

**subcase** s = -1

By the symmetric of the range, we can get:  $u \in [e^{\epsilon(-t-\frac{\Lambda}{2})}, 1]$ .

case  $x \in (\lfloor f(a) \rfloor_{\Lambda}, B)$ 

In this case, we know s = -1 and  $x \in [\lfloor f(a) \rfloor_{\Lambda} + \Lambda, b]$ .

We have:  $\lfloor f(a) - \frac{1}{\epsilon} \ln(u) \rceil_{\Lambda} = x$ .

By the rule of rounding, we get:  $u \in \left[e^{\epsilon(f(a) - \frac{\Lambda}{2} - x)}, e^{\epsilon(f(a) + \frac{\Lambda}{2} - x)}\right]$ .

By the range of x, we get:  $u \in \left(e^{\epsilon(-b-\frac{\Lambda}{2}+f(a))}, e^{\epsilon(-t-\frac{\Lambda}{2})}\right]$ .

case x = B

In this case, we know s = -1.

We have:  $\lfloor f(a) - \frac{1}{\epsilon} \ln(u) \rceil_{\Lambda} \ge B$ .

Since b is the greatest rounding of  $\Lambda$  that is smaller than B, we have:  $f(a) - \frac{1}{\epsilon} \ln(u) \ge b + \frac{\Lambda}{2}$ .

Then we get:  $u \in (0, e^{\epsilon(-b-\frac{\Lambda}{2}+f(a))})$ 

## Theorem 4 (Snap mechanism is $\epsilon$ -differentially private.)

*Proof.* Consider two arbitrary adjacent database a and a', we have  $|f(a) - f(a')| \le 1$ . Without loss of generalization, we assume f(a) + 1 = f(a') ( $\diamond$ ). The proof is developed by cases of the output space E of Snap mechanism, where  $x = \operatorname{Snap}(a)$ ,  $y = \operatorname{Snap}(a')$ .

case E = -B

$$\frac{Pr[x \in E]}{Pr[y \in E]} = \frac{\frac{1}{2}(e^{(-b-\frac{\Lambda}{2}-f(a))\epsilon})}{\frac{1}{2}(e^{(-b-\frac{\Lambda}{2}-f(a'))\epsilon})} = \frac{\frac{1}{2}(e^{(-b-\frac{\Lambda}{2}-f(a))\epsilon})}{\frac{1}{2}(e^{(-b-\frac{\Lambda}{2}-f(a)-1)\epsilon})} = e^{\epsilon}$$

case  $E = (-B, \lfloor f(a) \rceil_{\Lambda})$ 

From  $(\diamond)$ , we have  $0 \le \lfloor f(a') \rceil_{\Lambda} - \lfloor f(a) \rceil_{\Lambda} \le 1 + \Lambda$ . So we have  $(-B, \lfloor f(a) \rceil_{\Lambda}) \subset (-B, \lfloor f(a') \rceil_{\Lambda})$ .

$$\frac{Pr[x \in E]}{Pr[y \in E]} = \frac{\frac{1}{2}(e^{\lfloor f(a) \rceil_{\Lambda} - f(a) - \frac{\Lambda}{2}\varepsilon} - e^{(-b - \frac{\Lambda}{2} - f(a))\varepsilon})}{\frac{1}{2}(e^{(\lfloor f(a) \rceil_{\Lambda} - \frac{\Lambda}{2} - f(a'))\varepsilon} - e^{(-b - \frac{\Lambda}{2} - f(a'))\varepsilon})} = \frac{\frac{1}{2}(e^{\lfloor f(a) \rceil_{\Lambda} - f(a) - \frac{\Lambda}{2}\varepsilon} - e^{(-b - \frac{\Lambda}{2} - f(a))\varepsilon})}{\frac{1}{2}(e^{(\lfloor f(a) \rceil_{\Lambda} - \frac{\Lambda}{2} - f(a) - 1)\varepsilon} - e^{(-b - \frac{\Lambda}{2} - f(a) - 1)\varepsilon})} = e^{\varepsilon}$$

case  $E = \lfloor f(a) \rfloor_{\Lambda}$ 

$$\frac{Pr[x \in E]}{Pr[y \in E]} = \frac{(1 - \frac{1}{2}(e^{(\lfloor f(a) \rceil_{\Lambda} - f(a) - \frac{\Lambda}{2})\varepsilon} + e^{(f(a) - \lfloor f(a) \rceil_{\Lambda} - \frac{\Lambda}{2})\varepsilon})}{\frac{1}{2}(e^{(\lfloor f(a) \rceil_{\Lambda} - f(a') + \frac{\Lambda}{2})\varepsilon} - e^{(\lfloor f(a) \rceil_{\Lambda} - f(a') - \frac{\Lambda}{2})\varepsilon})} \ (\star)$$

Let  $t = \lfloor f(a) \rceil_{\Lambda} - f(a)$ , we have  $-\frac{\Lambda}{2} \le t \le \frac{\Lambda}{2}$  and:

$$(\star) = \frac{1 - \frac{1}{2}(e^{(t - \frac{\Lambda}{2})\epsilon} + e^{-t - \frac{\Lambda}{2})\epsilon})}{\frac{1}{2}(e^{(t - 1 + \frac{\Lambda}{2})\epsilon} - e^{(t - 1 - \frac{\Lambda}{2})\epsilon})} = \frac{2 - (e^{(t - \frac{\Lambda}{2})\epsilon} + e^{(-t - \frac{\Lambda}{2})\epsilon})}{e^{(t - 1 + \frac{\Lambda}{2})\epsilon} - e^{(t - 1 - \frac{\Lambda}{2})\epsilon}} < \frac{e^{(t + \frac{\Lambda}{2})\epsilon} - e^{(t - \frac{\Lambda}{2})\epsilon}}{e^{(t - 1 + \frac{\Lambda}{2})\epsilon} - e^{(t - 1 - \frac{\Lambda}{2})\epsilon}} = e^{\epsilon}$$

The inequality holds because given  $1 \le \Lambda \epsilon < 2$  and  $-\frac{\Lambda}{2} \le t \le \frac{\Lambda}{2}$ , we have:

$$2 - e^{(-t - \frac{\Lambda}{2})\epsilon} < e^{(t + \frac{\Lambda}{2})\epsilon}$$

case  $E = (\lfloor f(a) \rceil_{\Lambda}, \lfloor f(a') \rceil_{\Lambda})$ 

$$\frac{Pr[x \in E]}{Pr[y \in E]} = \frac{\frac{1}{2} (e^{(f(a) - \frac{\Lambda}{2} - \lfloor f(a) \rceil_{\Lambda})\epsilon} - e^{(f(a) + \frac{\Lambda}{2} - \lfloor f(a') \rceil_{\Lambda})\epsilon})}{\frac{1}{2} (e^{(\lfloor f(a') \rceil_{\Lambda} - \frac{\Lambda}{2} - f(a'))\epsilon} - e^{(\lfloor f(a) \rceil_{\Lambda} + \frac{\Lambda}{2} - f(a'))\epsilon})}$$
 (\*\*)

subcase  $\Lambda \ge 1$ 

Because f(a) + 1 = f(a'), we have  $E = \emptyset$ , i.e.:

$$Pr[x \in E] = Pr[y \in E] = 0$$

### subcase $\Lambda < 1$

Because f(a) + 1 = f(a'), we have:

$$\lfloor f(a) + 1 \rceil_{\Lambda} = \lfloor f(a') \rceil_{\Lambda}$$

Since  $\Lambda$  is power of 2 and  $\Lambda < 1$ , so we have:  $\lfloor 1 \rfloor_{\Lambda} = 1$ . Then we have:

$$\lfloor f(a) + 1 \rceil_{\Lambda} = \lfloor f(a) \rceil_{\Lambda} + 1 = \lfloor f(a') \rceil_{\Lambda}$$

. Substitute  $[f(a')]_{\Lambda}$  and f(a') with  $[f(a)]_{\Lambda}$  and f(a) in  $(\star)$ , we can get:

$$\frac{Pr[x \in E]}{Pr[y \in E]} = \frac{\frac{1}{2} (e^{(f(a) - \frac{\Lambda}{2} - \lfloor f(a) \rceil_{\Lambda})\epsilon} - e^{(f(a) + \frac{\Lambda}{2} - \lfloor f(a) \rceil_{\Lambda} - 1)\epsilon})}{\frac{1}{2} (e^{(\lfloor f(a) \rceil_{\Lambda} + 1 - \frac{\Lambda}{2} - f(a) - 1)\epsilon} - e^{(\lfloor f(a) \rceil_{\Lambda} + \frac{\Lambda}{2} - f(a) - 1)\epsilon})}$$
  $(\diamond)$ 

Let  $t = f(a) - \lfloor f(a) \rceil_{\Lambda}$ , we have  $-\frac{\Lambda}{2} < t < \frac{\Lambda}{2}$ . We also have  $\lambda \le \frac{1}{\epsilon}$ , so we can get:

$$(\diamond) = \frac{\frac{1}{2}(e^{(t-\lfloor f(a) \rceil_{\Lambda})\epsilon} - e^{(t+\frac{\Lambda}{2}-1)\epsilon})}{\frac{1}{2}(e^{(-t-\frac{\Lambda}{2})\epsilon} - e^{(-t+\frac{\Lambda}{2}-1)\epsilon})} = e^{2t\epsilon} \in (e^{-1}, e^{1})$$

Because  $\Lambda$  is the smallest power of 2 where  $\Lambda \ge \frac{1}{\epsilon}$  and  $\Lambda < 1$ , we have  $\epsilon > 1$ . Then we can conclude that  $(e^{-1}, e^1) \subset (e^{-\epsilon}, e^{\epsilon})$ , i.e.:

$$\frac{Pr[x \in E]}{Pr[y \in E]} \in (e^{-\epsilon}, e^{\epsilon})$$

case  $E = \lfloor f(a') \rfloor_{\Lambda}$ 

$$\frac{Pr[x \in E]}{Pr[y \in E]} = \frac{\frac{1}{2} (e^{(f(a) - \lfloor f(a') \rceil_{\Lambda} + \frac{\Lambda}{2})\epsilon} - e^{(f(a) - \lfloor f(a') \rceil_{\Lambda} - \frac{\Lambda}{2})\epsilon})}{1 - \frac{1}{2} (e^{(\lfloor f(a) \rceil_{\Lambda} - f(a') - \frac{\Lambda}{2})\epsilon} + e^{(f(a') - \lfloor f(a') \rceil_{\Lambda} - \frac{\Lambda}{2})\epsilon})} \ (\star)$$

Let  $t = f(a') - \lfloor f(a') \rceil_{\Lambda}$ , we have  $-\frac{\Lambda}{2} \le t \le \frac{\Lambda}{2}$  and:

$$(\star) = \frac{\frac{1}{2}(e^{(t+\frac{\Lambda}{2}-1)\epsilon} + e^{t-\frac{\Lambda}{2}-1)\epsilon})}{1 - \frac{1}{2}(e^{(-t-\frac{\Lambda}{2})\epsilon} + e^{(t-\frac{\Lambda}{2})\epsilon})} = \frac{(e^{(t+\frac{\Lambda}{2}-1)\epsilon} + e^{(t-\frac{\Lambda}{2}-1)\epsilon})}{2 - (e^{(-t-\frac{\Lambda}{2})\epsilon} + e^{(t-\frac{\Lambda}{2})\epsilon})} > \frac{e^{(t+\frac{\Lambda}{2}-1)\epsilon} - e^{(t-\frac{\Lambda}{2}-1)\epsilon}}{e^{(t+\frac{\Lambda}{2})\epsilon} - e^{(t-\frac{\Lambda}{2})\epsilon}} = e^{-\epsilon}$$

The inequality holds because given  $1 \le \Lambda \epsilon < 2$  and  $-\frac{\Lambda}{2} \le t \le \frac{\Lambda}{2}$ , we have:

$$2 - e^{(-t - \frac{\Lambda}{2})\epsilon} < e^{(t + \frac{\Lambda}{2})\epsilon}$$

case  $E = (\lfloor f(a') \rceil_{\Lambda}, B)$ 

From  $(\star)$ , we have  $0 \le \lfloor f(a') \rceil_{\Lambda} - \lfloor f(a) \rceil_{\Lambda} \le 1 + \Lambda$ . So we have  $(\lfloor f(a') \rceil_{\Lambda}, B) \subset (\lfloor f(a) \rceil_{\Lambda}, B)$ .

$$\frac{Pr[x \in E]}{Pr[y \in E]} = \frac{\frac{1}{2}(e^{(f(a) - \frac{\Lambda}{2} - \lfloor f(a') \rceil_{\Lambda})\epsilon} - e^{(f(a) - b - \frac{\Lambda}{2})\epsilon})}{\frac{1}{2}(e^{f(a') - \lfloor f(a') \rceil_{\Lambda} - \frac{\Lambda}{2}\epsilon} - e^{(f(a') - b - \frac{\Lambda}{2})\epsilon})} = \frac{\frac{1}{2}(e^{(f(a) - \frac{\Lambda}{2} - \lfloor f(a') \rceil_{\Lambda})\epsilon} - e^{(f(a) - b - \frac{\Lambda}{2})\epsilon})}{\frac{1}{2}(e^{f(a) + 1 - \lfloor f(a') \rceil_{\Lambda} - \frac{\Lambda}{2}\epsilon} - e^{(f(a) + 1 - b - \frac{\Lambda}{2})\epsilon})} = e^{-\epsilon}$$

case E = B

$$\frac{Pr[x \in E]}{Pr[y \in E]} = \frac{\frac{1}{2}(e^{(-b - \frac{\Lambda}{2} + f(a))\epsilon})}{\frac{1}{2}(e^{(-b - \frac{\Lambda}{2} + f(a'))\epsilon})} = \frac{\frac{1}{2}(e^{(-b - \frac{\Lambda}{2} + f(a))\epsilon})}{\frac{1}{2}(e^{(-b - \frac{\Lambda}{2} + f(a) + 1)\epsilon})} = e^{-\epsilon}$$

# References

- [1] Gilles Barthe, Marco Gaboardi, Benjamin Grégoire, Justin Hsu, and Pierre-Yves Strub. Proving differential privacy via probabilistic couplings. In *LICS 2016*.
- [2] Cynthia Dwork, Frank McSherry, Kobbi Nissim, and Adam Smith. Calibrating Noise to Sensitivity in Private Data Analysis. In *TCC*, 2016.
- [3] Ilya Mironov. On significance of the least significant bits for differential privacy. In *CCS 2012*, 2012.