Verifying Snapping Mechanism

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In order to verify the differential privacy proeprty of the snapping mechanism[3], we follow the logic rules designed from [1].

Some new rules are added into this logic in Figure 1 following with correctness proof. Then we formalized the snapping mechanism and verified its differential privacy property under these logic rules.

1 Extended Programming Logic[1]

Definition 1 (Laplce mechanism [2])

Let $\epsilon > 0$. The Laplace mechanism $\mathcal{L}_{\epsilon} : \mathbb{R} \to \mathsf{Distr}(\mathbb{R})$ is defined by $\mathcal{L}(t) = t + v$, where $v \in \mathbb{R}$ is drawn from the Laplace distribution $\mathsf{Laplce}(\frac{1}{\epsilon})$.

Definition 2

Let $\epsilon \leq 0$. The ϵ -DP divergence $\Delta_{\epsilon}(\mu_1, \mu_2)$ between two sub-distributions $\mu_1 \in \mathsf{Distr}(U)$, $\mu_2 \in \mathsf{Distr}(U)$ is defined as:

$$\sup_{E \in U} \left(\Pr_{x \leftarrow \mu_1} [x \in E] - \exp(\epsilon) \Pr_{x \leftarrow \mu_2} [x \in E] \right)$$

Definition 3 $((\epsilon, \delta)$ - **lifting [1]**)

Two sub-distributions $\mu_1 \in \mathsf{Distr}(U_1)$, $\mu_2 \in \mathsf{Distr}(U_2)$ are related by the (ε, δ) - dilation lifting of $\Psi \subseteq U_1 \times U_2$, written $\mu_1 \Psi^{\#(\varepsilon, \delta)} \mu_2$, if there exist two witness sub-distributions $\mu_L \in \mathsf{Distr}(U_1 \times U_2)$ and $\mu_R \in \mathsf{Distr}(U_1, U_2)$ s.t.:

- 1. $\pi_1(\mu_L) = \mu_1$ and $\pi_2(\mu_R) = \mu_2$;
- 2. $supp(\mu_L) \subseteq \Psi$ and $supp(\mu_R) \subseteq \Psi$; and
- 3. $\Delta_{\epsilon}(\mu_L, \mu_R) \leq \delta$.

The logic rules we are using in our work is presented in Figure 1. The correctness of rules is proved in Theorem 1 and Theorem 2

Theorem 1

Let $\mu_1 \in \mathsf{Distr}(\mathbb{R})$, $\mu_2 \in \mathsf{Distr}(\mathbb{R})$ are defined:

$$\mu_1(x) = \text{unif}(x)$$

$$\mu_2(y) = \mathsf{unif}(y)$$

Figure 1: Logic Rules Extended from [1]

where unif is uniform distribution over [0, 1) whoes pdf. is defined as:

$$pdf_{unif}(x) = \begin{cases} 1 & x \in [0,1) \\ 0 & o.w. \end{cases}.$$

Then, $\mu_1 \Psi^{\#(\epsilon,0)} \mu_2$, where

$$\Psi = \{(x, y) \in \mathbb{R} \times \mathbb{R} | x \cdot e^{-\epsilon} \le y \le x \cdot e^{\epsilon} \}$$

Theorem 2

For any distributions $\mu_1 \in \mathsf{Distr}(\mathbb{R}), \ \mu_2 \in \mathsf{Distr}(\mathbb{R}), \ \mu_1 \Psi^{\#(0,0)} \mu_2$, where

$$\Psi = \{(x, y) \in \mathbb{R} \times \mathbb{R} | x = y\}$$

Proof of Theorem 1. Existing $\mu_L, \mu_R \in \mathsf{Distr}(\mathbb{R} \times \mathbb{R})$:

$$\mu_L(x,y) = \begin{cases} \operatorname{unif}(x) & x \cdot e^{-\epsilon} = y \wedge x \in [0,1) \\ 0 & o.w. \end{cases} \\ \mu_R(x,y) = \begin{cases} \operatorname{unif}(y) & x \cdot e^{-\epsilon} = y \wedge y \in [0,1) \\ 0 & o.w. \end{cases}.$$

Their pdf. are defined:

$$\mathsf{pdf}_{\mu_L}(x,y) = \begin{cases} \mathsf{pdf}_{\mathsf{unif}}(x) & x \cdot e^{-\epsilon} = y \land x \in [0,1) \\ 0 & o.w. \end{cases}$$

$$\mathsf{pdf}_{\mu_R}(x,y) = \begin{cases} \mathsf{pdf}_{\mathsf{unif}}(y) & x \cdot e^{-\epsilon} = y \land y \in [0,1) \\ 0 & o.w. \end{cases}.$$

- $supp(\mu_L) \in \Psi \land supp(\mu_R) \in \Psi$
 - $\sup (\mu_L) \subseteq \Psi$ By definition of the pdf of μ_L , we have: $\Pr_{(x,y) \stackrel{\$}{\leftarrow} \mu_L} [(x,y) \notin \Psi] = 0$. Then we can derive $\sup (\mu_L) \in \Psi$

- supp(μ_R) ⊆ Ψ

By definition of the pdf of μ_R , we have: $\Pr_{(x,y) \stackrel{\$}{\leftarrow} \mu_R} [(x,y) \notin \Psi] = 0.$

Then we can derive $supp(\mu_L) \in \Psi$

•
$$\pi_1(\mu_L) = \mu_1 \wedge \pi_2(\mu_R) = \mu_2$$

$$- \pi_1(\mu_L) = \mu_1$$

By definition of the π_1 and pdf of μ_L , we have $\forall x \in \mathbb{R}$:

$$\mathsf{pdf}_{\pi_1(\mu_L)}(x) = \begin{cases} \int_{\mathcal{Y}} \mathsf{pdf}_{\mathsf{unif}}(x) & (x,y) \in \Psi \land x \in [0,1) \\ 0 & o.w. \end{cases} = \begin{cases} \mathsf{pdf}_{\mathsf{unif}}(x) & x \in [0,1) \\ 0 & o.w. \end{cases} = \mathsf{pdf}_{\mu_1}(x)$$

 $- \pi_1(\mu_R) = \mu_2$

Equivalent to showpdf_{$\pi_2(\mu_R)$} = pdf_{μ_2}.

By definition of the π_2 and pdf of μ_R , we have $\forall y \in \mathbb{R}$:

$$\mathsf{pdf}_{\pi_2(\mu_R)}(y) = \begin{cases} \int_{\mathcal{X}} \mathsf{pdf}_{\mathsf{unif}}(y) & (x,y) \in \Psi \land y \in [0,1) \\ 0 & o.w. \end{cases} = \begin{cases} \mathsf{pdf}_{\mathsf{unif}}(y) & y \in [0,1) \\ 0 & o.w. \end{cases} = \mathsf{pdf}_{\mu_2}(y)$$

• $\Delta_{\epsilon}(\mu_L, \mu_R) \leq 0$

By definition of ϵ -DP divergence, we have:

$$\Delta_{\epsilon}(\mu_{L}, \mu_{R}) = \sup_{S} \left(\Pr_{(x,y) \stackrel{\$}{\leftarrow} \mu_{L}} [(x,y) \in S] - e^{\epsilon} \Pr_{(x,y) \stackrel{\$}{\leftarrow} \mu_{R}} [(x,y) \in S] \right)$$
$$= \sup_{S} \left(\int_{(x,y) \in S} \mathsf{pdf}_{\mu_{L}}(x,y) - e^{\epsilon} \int_{(x,y) \in S} \mathsf{pdf}_{\mu_{R}}(x,y) \right)$$

case $S \subseteq \{(x, y) | x \in [0, 1) \land x \cdot e^{-\epsilon} = y\}$:

$$\begin{array}{ll} \Delta_{\epsilon}(\mu_L,\mu_R) &= \int_{(x,y)\in S} \mathsf{pdf}_{\mathsf{unif}}(x) - e^{\epsilon} \int_{(x,y)\in S} \mathsf{pdf}_{\mathsf{unif}}(y) \\ &= \int_{(x,y)\in S} \mathsf{pdf}_{\mathsf{unif}}(x) - e^{\epsilon} \int_{(x,y)\in S} \mathsf{pdf}_{\mathsf{unif}}(x*e^{-\epsilon}) \\ &= \int_{(x,y)\in S} \mathsf{pdf}_{\mathsf{unif}}(x) - e^{\epsilon} * e^{-\epsilon} \int_{(x,y)\in S} \mathsf{pdf}_{\mathsf{unif}}(x) \\ &= 0 \end{array}$$

case $S \subseteq \{(x, y) | x \in [1, e^{\epsilon}) \land x \cdot e^{-\epsilon} = y\}$:

$$\begin{array}{ll} \Delta_{\epsilon}(\mu_L, \mu_R) &= 0 - e^{\epsilon} \int_{(x,y) \in S} \mathsf{pdf}_{\mathsf{unif}}(y) \\ &< 0 \end{array}$$

case o.w.

$$\Delta_{\epsilon}(\mu_L, \mu_R) = 0 - 0 = 0$$

$$\overline{u_1 \overset{\$}{\leftarrow} \mu \sim_{\epsilon,0} u_2 \overset{\$}{\leftarrow} \mu : \top \Rightarrow e^{-\epsilon} u_2 \leq u_1 \leq e^{\epsilon} u_2} \xrightarrow{\text{AXNULL}} \xrightarrow{\text{AXNULL}}$$

$$\overline{y_1 = \frac{\ln(u_1)}{\epsilon} \sim_{0,0} y_2 = \frac{\ln(u_2)}{\epsilon} : e^{-\epsilon} u_2 \leq u_1 \leq e^{\epsilon} u_2 \Rightarrow y_2 - 1 \leq y_1 \leq 1 + y_2} \xrightarrow{\text{AXNULL}}$$

$$\overline{z_1 = s_1 * y_1 \sim_{0,0} z_2 = s_2 * y_2 : s_1 = s_2 \wedge y_2 - 1 \leq y_1 \leq 1 + y_2 \Rightarrow |z_1 - z_2| \leq 1} \xrightarrow{\text{AXNULL}}$$

$$\overline{x_1 = f(a_1) \sim_{0,0} x_2 = f(a_2) : a_1 = a_2 + 1 \wedge f(a_1) = f(a_2) + 1 \Rightarrow x_1 = x_2 + 1} \xrightarrow{\text{AXNULL}}$$

$$\overline{w_1 = x_1 + z_1 \sim_{0,0} w_2 = x_2 + z_2 : x_1 = x_2 + 1 \wedge |z_1 - z_2| \leq 1 \wedge -2 \leq k \leq 0 \Rightarrow w_1 + k = w_2} \xrightarrow{\text{AXNULL}}$$

$$\overline{w'_1 = \lfloor w_1 \rceil_{\Lambda} \sim_{0,0} w'_2 = \lfloor w_2 \rceil_{\Lambda} : w_1 + k = w_2 \wedge -2 \leq k \leq 0 \Rightarrow w'_1 + k = w'_2} \xrightarrow{\text{AXNULL}}$$

$$\overline{r_1 = \text{clamp}_B(w'_1) \sim_{0,0} r_2 = \text{clamp}_B(w'_2) : w'_1 + k = w'_2 \wedge -2 \leq k \leq 0 \Rightarrow r_1 + k = r_2} \xrightarrow{\text{AXNULL}}$$

$$\overline{r_1 = \text{Snap}(a_1) \sim_{\epsilon,0} r_2 = \text{Snap}(a_1) : a_1 = a_2 + 1 \wedge f(a_1) = f(a_2) + 1 \wedge |k + f(a_1) - f(a_2)| \leq 1 \Rightarrow r_1 + k = r_2} \xrightarrow{\text{COMP}}$$

Figure 2: Coupling Derivation of two Snap mechanisms: $Snap(a_1)$, $Snap(a_2)$

2 Formalization of Snap Mechanism in Probabilistic Logic

Definition 4 (Snap(a): $A \rightarrow Distr(B)$)

The ideal Snapping mechanism Snap(a) is defined as:

$$u \stackrel{\$}{\leftarrow} \mu; y = \frac{\ln(u)}{\epsilon}; s \stackrel{\$}{\leftarrow} \{-1, 1\}; z = s * y; x = f(a); w = x + z; w' = \lfloor w \rceil_{\Lambda}; r = \mathsf{clamp}_{B}(w')$$

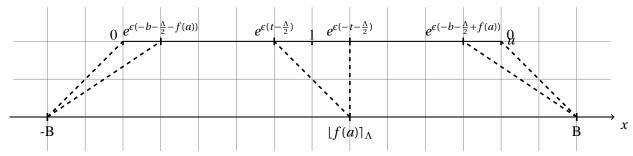
where f is the query function over input $a \in A$, ϵ is the privacy budget, B is the clampping bound and Λ is the rounding argument satisfying $\lambda = 2^k$ where 2^k is the smallest power of 2 greater or equal to the $\frac{1}{\epsilon}$.

Theorem 3 (The Snap mechanism is ϵ -differentially praivate)

Proof. The proof follows the derivation in Figure 2.

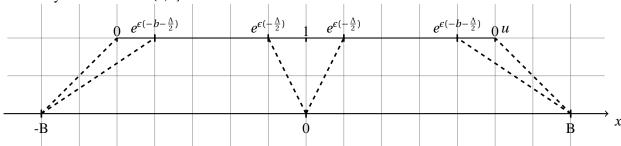
3 Proof of Differential Privacy for Snap Mechanism

Assume x be the output of Snap mechanism, we have following maps from the output of Snap mechanism to uniformly distributed $u \in (0,1]$.



where *b* is the greatest rounding of Λ that is smaller than *B* and $t = \lfloor f(a) \rceil_{\Lambda} - f(a)$.

Given the $f(a) = \lfloor f(a) \rceil_{\Lambda} = 0$, we have following maps from the output of Snap mechanism to uniformly distributed $u \in (0,1]$.



Assuming that $f(a) \in [-B, B]$, otherwise we can always redefine the f(a) restricting its output in this range. The probability of obtaining output x from Snap mechanism can be calculated by cases of x:

case x = -B

In this case, we know s = 1.

We have: $\lfloor f(a) + \frac{1}{\epsilon} \ln(u) \rceil_{\Lambda} \le -B$.

Since b is the greatest rounding of Λ that is smaller than B, then -b is the smallest rounding of Λ that is greater than -B, we have: $f(a) + \frac{1}{\epsilon} \ln(u) < -b - \frac{\Lambda}{2}$.

Then we get: $u \in (0, e^{\epsilon(-b - \frac{\Lambda}{2} - f(a))})$

case $x \in (-B, \lfloor f(a) \rceil_{\Lambda})$

In this case, we know s=1 and $x \in [-b, \lfloor f(a) \rceil_{\Lambda} - \Lambda]$. We have: $\lfloor f(a) + \frac{1}{\epsilon} \ln(u) \rceil_{\Lambda} = x$.

By the rule of rounding, we get: $u \in \left[e^{\epsilon(x-\frac{\Lambda}{2}-f(a))}, e^{\epsilon(x+\frac{\Lambda}{2}-f(a))}\right]$.

By the range of x, we get: $u \in \left[e^{\epsilon(-b-\frac{\Lambda}{2}-f(a))}, e^{\epsilon(t-\frac{\Lambda}{2})}\right]$.

case $x = \lfloor f(a) \rceil_{\Lambda}$

subcase s = 1

In this case, we have: $\lfloor f(a) + \frac{1}{\epsilon} \ln(u) \rceil_{\Lambda} = \lfloor f(a) \rceil_{\Lambda}$.

Then we can get: $u \in [e^{\epsilon(t-\frac{\Lambda}{2})}, e^{\epsilon(t+\frac{\Lambda}{2})}].$

Since $t = \lfloor f(a) \rceil_{\Lambda} - f(a)$, we know: $-\frac{\Lambda}{2} \le t \le \frac{\Lambda}{2}$. So we can get: $e^{\epsilon(t + \frac{\Lambda}{2})} > 1$. Since $u \in (0,1]$, we have: $u \in \left[e^{\epsilon(t - \frac{\Lambda}{2})}, 1\right]$.

subcase s = -1

By the symmetric of the range, we can get: $u \in [e^{\epsilon(-t-\frac{\Lambda}{2})}, 1]$.

case $x \in (\lfloor f(a) \rceil_{\Lambda}, B)$

In this case, we know s = -1 and $x \in [\lfloor f(a) \rceil_{\Lambda} + \Lambda, b]$.

We have: $\lfloor f(a) - \frac{1}{\epsilon} \ln(u) \rceil_{\Lambda} = x$.

By the rule of rounding, we get: $u \in \left(e^{\epsilon(f(a) - \frac{\Lambda}{2} - x)}, e^{\epsilon(f(a) + \frac{\Lambda}{2} - x)}\right]$.

By the range of x, we get: $u \in \left(e^{\varepsilon(-b-\frac{\lambda}{2}+f(a))}, e^{\varepsilon(-t-\frac{\lambda}{2})}\right]$.

case x = B

In this case, we know s = -1.

We have: $\lfloor f(a) - \frac{1}{\epsilon} \ln(u) \rceil_{\Lambda} \ge B$.

Since b is the greatest rounding of Λ that is smaller than B, we have: $f(a) - \frac{1}{6} \ln(u) \ge b + \frac{\Lambda}{2}$.

Then we get: $u \in (0, e^{\epsilon(-b-\frac{\Lambda}{2}+f(a))})$

Theorem 4 (Snap mechanism is ϵ -differentially private.)

Proof. Consider two arbitrary adjacent database a and a', we have $|f(a) - f(a')| \le 1$. Without loss of generalization, we assume f(a) + 1 = f(a') (\diamond). The proof is developed by cases of the output space E of Snap mechainism, where $x = \operatorname{Snap}(a)$, $y = \operatorname{Snap}(a')$.

case E = -B

$$\frac{Pr[x \in E]}{Pr[y \in E]} = \frac{\frac{1}{2}(e^{(-b - \frac{\Lambda}{2} - f(a))\epsilon})}{\frac{1}{2}(e^{(-b - \frac{\Lambda}{2} - f(a'))\epsilon})} = \frac{\frac{1}{2}(e^{(-b - \frac{\Lambda}{2} - f(a))\epsilon})}{\frac{1}{2}(e^{(-b - \frac{\Lambda}{2} - f(a) - 1)\epsilon})} = e^{\epsilon}$$

case $E = (-B, \lfloor f(a) \rfloor_{\Lambda})$

From (\diamond) , we have $0 \le \lfloor f(a') \rceil_{\Lambda} - \lfloor f(a) \rceil_{\Lambda} \le 1 + \Lambda$. So we have $(-B, \lfloor f(a) \rceil_{\Lambda}) \subset (-B, \lfloor f(a') \rceil_{\Lambda})$.

$$\frac{Pr[x \in E]}{Pr[y \in E]} = \frac{\frac{1}{2}(e^{\lfloor f(a) \rceil_{\Lambda} - f(a) - \frac{\Lambda}{2}\epsilon} - e^{(-b - \frac{\Lambda}{2} - f(a))\epsilon})}{\frac{1}{2}(e^{(\lfloor f(a) \rceil_{\Lambda} - \frac{\Lambda}{2} - f(a'))\epsilon} - e^{(-b - \frac{\Lambda}{2} - f(a'))\epsilon})} = \frac{\frac{1}{2}(e^{\lfloor f(a) \rceil_{\Lambda} - f(a) - \frac{\Lambda}{2}\epsilon} - e^{(-b - \frac{\Lambda}{2} - f(a))\epsilon})}{\frac{1}{2}(e^{(\lfloor f(a) \rceil_{\Lambda} - \frac{\Lambda}{2} - f(a) - 1)\epsilon} - e^{(-b - \frac{\Lambda}{2} - f(a) - 1)\epsilon})} = e^{\epsilon}$$

case $E = \lfloor f(a) \rceil_{\Lambda}$

$$\frac{Pr[x \in E]}{Pr[y \in E]} = \frac{(1 - \frac{1}{2}(e^{(\lfloor f(a) \rceil_{\Lambda} - f(a) - \frac{\Lambda}{2})\varepsilon} + e^{(f(a) - \lfloor f(a) \rceil_{\Lambda} - \frac{\Lambda}{2})\varepsilon})}{\frac{1}{2}(e^{(\lfloor f(a) \rceil_{\Lambda} - f(a') + \frac{\Lambda}{2})\varepsilon} - e^{(\lfloor f(a) \rceil_{\Lambda} - f(a') - \frac{\Lambda}{2})\varepsilon})} \ (\star)$$

Let $t = \lfloor f(a) \rceil_{\Lambda} - f(a)$, we have $-\frac{\Lambda}{2} \le t \le \frac{\Lambda}{2}$ and:

$$(\star) = \frac{1 - \frac{1}{2}(e^{(t - \frac{\Lambda}{2})\varepsilon} + e^{-t - \frac{\Lambda}{2})\varepsilon})}{\frac{1}{2}(e^{(t - 1 + \frac{\Lambda}{2})\varepsilon} - e^{(t - 1 - \frac{\Lambda}{2})\varepsilon})} \leq \frac{\frac{1}{2}(e^{(t + \frac{\Lambda}{2})\varepsilon} + e^{(-t + \frac{\Lambda}{2})\varepsilon}) - \frac{1}{2}(e^{(t - \frac{\Lambda}{2})\varepsilon} + e^{(-t - \frac{\Lambda}{2})\varepsilon})}{\frac{1}{2}(e^{(t - 1 + \frac{\Lambda}{2})\varepsilon} - e^{(t - 1 - \frac{\Lambda}{2})\varepsilon})} \leq \frac{e^{(t + \frac{\Lambda}{2})\varepsilon} - e^{(t - \frac{\Lambda}{2})\varepsilon}}{e^{(t - 1 + \frac{\Lambda}{2})\varepsilon} - e^{(t - 1 - \frac{\Lambda}{2})\varepsilon}} \leq e^{\varepsilon}$$

case $E = (\lfloor f(a) \rceil_{\Lambda}, \lfloor f(a') \rceil_{\Lambda})$

$$\frac{Pr[x \in E]}{Pr[y \in E]} = \frac{\frac{1}{2} (e^{(f(a) - \frac{\Lambda}{2} - \lfloor f(a) \rceil_{\Lambda})\epsilon} - e^{(f(a) + \frac{\Lambda}{2} - \lfloor f(a) \rceil_{\Lambda})\epsilon})}{\frac{1}{2} (e^{(\lfloor f(a') \rceil_{\Lambda} - \frac{\Lambda}{2} - f(a'))\epsilon} - e^{(\lfloor f(a) \rceil_{\Lambda} + \frac{\Lambda}{2} - f(a'))\epsilon})}$$
 (*)

Let $t = \lfloor f(a') \rceil_{\Lambda} - f(a')$, we have $-\frac{\Lambda}{2} \le t \le \frac{\Lambda}{2}$. We also have $\lfloor f(a') \rceil_{\Lambda} - 1 - \Lambda \le \lfloor f(a) \rceil_{\Lambda} \le \lfloor f(a') \rceil_{\Lambda}$ by adjacency of a and a'. So we can get:

$$(\star) \geq \frac{\frac{1}{2}(e^{(-t-\frac{\Lambda}{2}-1)\epsilon} - e^{(-t+\frac{\Lambda}{2}-1)\epsilon})}{\frac{1}{2}(e^{(t-\frac{\Lambda}{2})\epsilon} - e^{(t+\frac{\Lambda}{2})\epsilon})} \geq e^{-\epsilon} \quad and \quad (\star) \leq \frac{\frac{1}{2}(e^{(-t+\frac{\Lambda}{2})\epsilon} - e^{(-t+\frac{\Lambda}{2}-1)\epsilon})}{\frac{1}{2}(e^{(t-\frac{\Lambda}{2})\epsilon} - e^{(t-\frac{\Lambda}{2}-1)\epsilon})} \leq e^{\epsilon}$$

case $E = \lfloor f(a') \rceil_{\Lambda}$

$$\frac{Pr[x \in E]}{Pr[y \in E]} = \frac{\frac{1}{2} (e^{(f(a) - \lfloor f(a') \rceil_{\Lambda} + \frac{\Lambda}{2})\epsilon} - e^{(f(a) - \lfloor f(a') \rceil_{\Lambda} - \frac{\Lambda}{2})\epsilon})}{1 - \frac{1}{2} (e^{(\lfloor f(a) \rceil_{\Lambda} - f(a') - \frac{\Lambda}{2})\epsilon} + e^{(f(a) - \lfloor f(a') \rceil_{\Lambda} - \frac{\Lambda}{2})\epsilon})} (\star)$$

Let $t = f(a') - \lfloor f(a') \rceil_{\Lambda}$, we have $-\frac{\Lambda}{2} \le t \le \frac{\Lambda}{2}$ and:

$$(\star) = \frac{\frac{1}{2}(e^{(t+\frac{\Lambda}{2}-1)\epsilon} + e^{t-\frac{\Lambda}{2}-1)\epsilon})}{1-\frac{1}{2}(e^{(-t+\frac{\Lambda}{2})\epsilon} - e^{(t-\frac{\Lambda}{2})\epsilon})} \geq \frac{\frac{1}{2}(e^{(t-\frac{\Lambda}{2}-1)\epsilon} + e^{(t-\frac{\Lambda}{2}-1)\epsilon})}{\frac{1}{2}(e^{(-t+\frac{\Lambda}{2})\epsilon} - e^{(t-\frac{\Lambda}{2})\epsilon}) - \frac{1}{2}(e^{(-t+\frac{\Lambda}{2})\epsilon} - e^{(t-\frac{\Lambda}{2})\epsilon})} \geq \frac{e^{(t+\frac{\Lambda}{2}-1)\epsilon} - e^{(t-\frac{\Lambda}{2}-1)\epsilon}}{e^{(t+\frac{\Lambda}{2})\epsilon} - e^{(t-\frac{\Lambda}{2}-1)\epsilon}} \geq e^{-\epsilon}$$

case $E = (\lfloor f(a') \rceil_{\Lambda}, B)$

From (\star) , we have $0 \le \lfloor f(a') \rceil_{\Lambda} - \lfloor f(a) \rceil_{\Lambda} \le 1 + \Lambda$. So we have $(\lfloor f(a') \rceil_{\Lambda}, B) \subset (\lfloor f(a) \rceil_{\Lambda}, B)$.

$$\frac{Pr[x \in E]}{Pr[y \in E]} = \frac{\frac{1}{2}(e^{(f(a) - \frac{\Lambda}{2} - \lfloor f(a') \rceil_{\Lambda})\epsilon} - e^{(f(a) - b - \frac{\Lambda}{2})\epsilon})}{\frac{1}{2}(e^{f(a') - \lfloor f(a') \rceil_{\Lambda} - \frac{\Lambda}{2}\epsilon} - e^{(f(a') - b - \frac{\Lambda}{2})\epsilon})} = \frac{\frac{1}{2}(e^{(f(a) - \frac{\Lambda}{2} - \lfloor f(a') \rceil_{\Lambda})\epsilon} - e^{(f(a) - b - \frac{\Lambda}{2})\epsilon})}{\frac{1}{2}(e^{f(a) + 1 - \lfloor f(a') \rceil_{\Lambda} - \frac{\Lambda}{2}\epsilon} - e^{(f(a) + 1 - b - \frac{\Lambda}{2})\epsilon})} = e^{-\epsilon}$$

case E = B

$$\frac{Pr[x \in E]}{Pr[y \in E]} = \frac{\frac{1}{2}(e^{(-b - \frac{\Lambda}{2} + f(a))\epsilon})}{\frac{1}{2}(e^{(-b - \frac{\Lambda}{2} + f(a'))\epsilon})} = \frac{\frac{1}{2}(e^{(-b - \frac{\Lambda}{2} + f(a))\epsilon})}{\frac{1}{2}(e^{(-b - \frac{\Lambda}{2} + f(a) + 1)\epsilon})} = e^{-\epsilon}$$

References

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