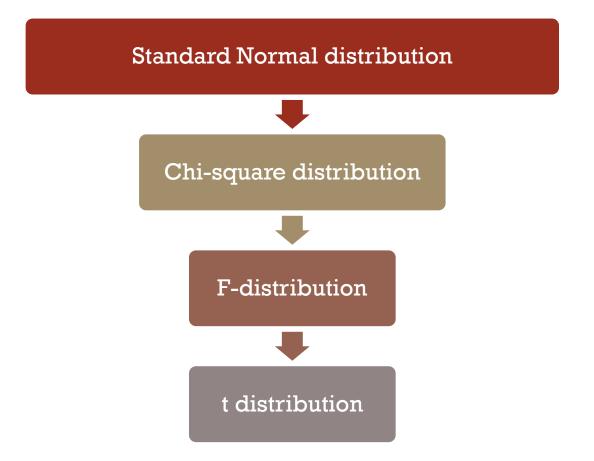
LAB: RELATED DISTRIBUTIONS

Lab 5.1



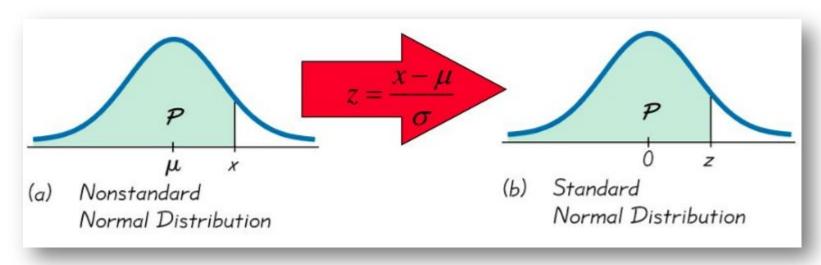
CONTENTS







Standard Normal Distribution



PDF:

$$f(z) = \frac{1}{\sigma \sqrt{2\pi}} e^{-z^2/2}, -\infty < z < +\infty$$

CDF:

$$f(z) = \frac{1}{\sigma\sqrt{2\pi}}e^{-z^2/2}, -\infty < z < +\infty \qquad \Phi(z) = \frac{1}{\sigma\sqrt{2\pi}}\int_{-\infty}^{z} e^{-z^2/2}dz, -\infty < z < +\infty$$





Chi-square Distribution

Chi-square goodness of fit test

Chi-square test of independence

Imagine taking a random sample of a standard normal distribution (Z). If you squared all the values in the sample, you would have the chi-square distribution with k = 1.

$$x_1^2 = (z)^2$$

Now imagine taking samples from two standard normal distributions (Z1 and Z2). If each time you sampled a pair of values, you squared them and added them together, you would have the chi-square distribution with k = 2.

$$x_2^2 = (z_1)^2 + (z_2)^2$$



More generally, if you sample from k independent standard normal distributions and then square and sum the values, you'll produce a chi-square distribution with k degrees of freedom.

$$x_k^2 = (z_1)^2 + (z_2)^2 + \dots + (z_k)^2$$

$$Y \square x_n^2$$



PDF of chi-square distribution

$$f(x) = \begin{cases} \frac{1}{2^{k/2} \Gamma(k/2)} x^{k/2-1} e^{-x/2} & \text{for } x > 0, \\ 0 & \text{for } x \le 0 \end{cases}$$



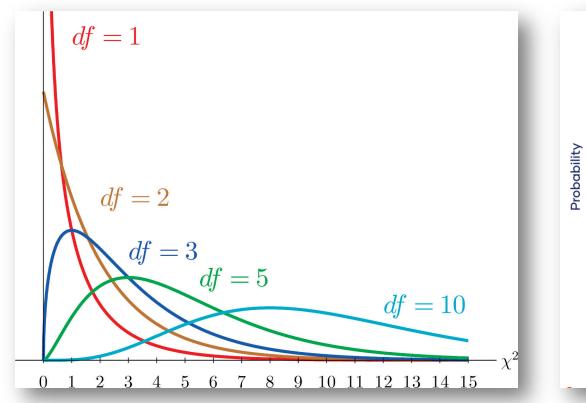
where:

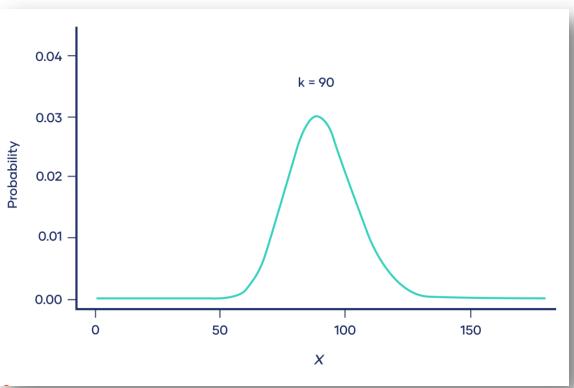
k = Positive Integer Denoting Degrees of Freedom, Karl Pearson (1857 ~ 1936)

 $\Gamma = Gamma Function$









exponential distribution is a special chi-square distribution





(1) If
$$Y \square x_n^2$$
, then $E(Y)=n$, and $D(Y)=2n$

Prove

$$Y = x_1^2 + x_2^2 + x_3^2 + \dots + x_n^2$$

$$Ex_i^2 = Dx_i + (Ex_i)^2 = 1$$

$$EY = E(\sum_{i=1}^{n} x_{i}^{2}) = \sum_{i=1}^{n} Ex_{i}^{2} = n$$

$$D(X) = E \{ [X - E(X)]^{2} \}$$

$$= E \{ X^{2} + E^{2}(X) - 2X E(X) \}$$

$$= E(X^{2}) + E^{2}(X) - 2E(X) \cdot E(X)$$

$$= E(X^{2}) - E^{2}(X)$$

$$\therefore E(X^{2}) = D(X) + E^{2}(X)$$



$$egin{aligned} Var(x) &= E((x-E(x))^2) \ &= E(x^2-2xE(x)+E(x)^2) \ &= E(x^2)-E(2xE(x))+E(E(x)^2) \ &= E(x^2)-\sum 2E(x)xf(x)+\sum E(x)^2f(x) \ &= E(x^2)-2E(x)^2+E(x)^2 \sum f(x) \ &= E(x^2)-2E(x)^2+E(x)^2 \ &= E(x^2)-E(x)^2 \end{aligned}$$

because:
$$Dx_i^2 = Ex_i^4 - (Ex_i^2)^2 = 3 - 1 = 2$$

thus:
$$DY = D(\sum_{i=1}^{n} x_i^2) = \sum_{i=1}^{n} Dx_i^2 = 2n$$

$$EX_i^4 = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{+\infty} x^4 e^{-\frac{x^2}{2}} dx = 3$$





(2) If $Y_1 \square x_m^2, Y_2 \square x_n^2$, and Y1 and Y2 are independent

Then
$$Y_1 + Y_2 \square x_{m+n}^2$$





(3) If $X \square N(\mu, \sigma^2), (X_1, ..., X_n)$ are randomly selected from population X

The distribution of
$$Y \square \frac{1}{\sigma^2} \sum_{i=1}^n (X_i - \mu)^2$$

because
$$X \square N(\mu, \sigma^2)$$
, so $\frac{x_i - \mu}{\sigma} \sim N(0, 1)$

According to the definition of chi-square distribution

$$Y = \sum_{i=1}^{n} \left(\frac{x_i - \mu}{\sigma}\right)^2 \sim x_n^2$$





(4)

$$x_n^2 = \sum_{i=1}^n (\frac{x_i - \mu}{\sigma})^2 \approx \sum_{i=1}^n \frac{(x_i - \overline{X})^2}{\sigma^2}$$

$$S^{2} = \sum_{i=1}^{n} \frac{(x_{i} - \overline{X})^{2}}{n-1}$$

$$x^{2} \approx \sum_{i=1}^{n} \frac{(n-1) * S^{2}}{\sigma^{2}}$$

$$\sim x_{n-1}^{2}$$

$$\sum_{i=1}^{n} (x_i - \overline{X})^2 = (n-1) * S^2$$





ANOVA



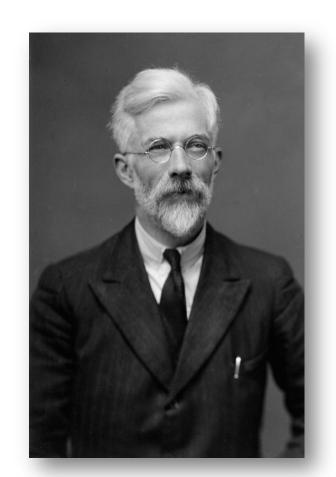
if $X \square x_n^2, Y \sim x_m^2$. X and Y are independent.

$$F = \frac{X/n}{Y/m}, then$$

$$F \sim F_{n,m}, or F \sim F(n,m)$$

The PDF of an *F* random variable with r1 numerator degrees of freedom and **r2** denominator degrees of freedom is:

$$f(w) = rac{(r_1/r_2)^{r_1/2}\Gamma[(r_1+r_2)/2]w^{(r_1/2)-1}}{\Gamma[r_1/2]\Gamma[r_2/2][1+(r_1w/r_2)]^{(r_1+r_2)/2}}$$



Ronald A. Fisher





Properties of an F-distribution

(1) If $F \sim F(n, m)$, then $1/F \sim F(m, n)$

Prove

$$F_{n,m} = \frac{X}{m}, then$$

$$1/F = \frac{Y}{M} \sim F_{m,n}$$

$$X \square x_n^2, Y \sim x_m^2$$





Properties of an F-distribution

(2) If $T \sim t_n$, then $T^2 \sim F(1, n)$

Prove

According to the definition of t distribution

$$T = \frac{x}{\sqrt{\frac{Y}{n}}}$$

 $T = \frac{\Lambda}{\sqrt{Y/n}}$ Of which, $X \sim N(0,1), Y \sim X_n^2$, and X and Y are independent

$$T^2 = \frac{x^2}{\frac{Y}{n}} = \frac{\frac{x^2}{1}}{\frac{Y}{n}}$$
 Of which, $X^2 \sim X^2_1$

t Distribution

t-test

If $X\sim N(0,1), Y\sim X_n^2$, and X and Y are independent

$$T = \frac{x}{\sqrt{Y/n}}$$

Then T follows a t distribution with df=n, we write

$$T \sim t_n$$



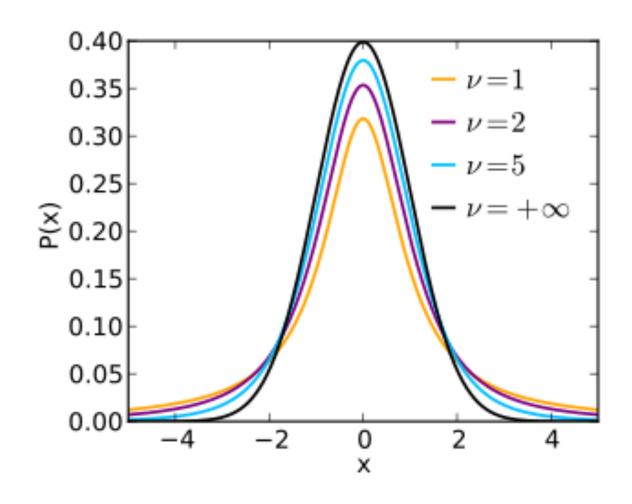
$$f(t) = \frac{\Gamma[(\nu+1)/2]}{\sqrt{\pi\nu}\Gamma(\nu/2)} (1 + \frac{t^2}{\nu})^{-\frac{(\nu+1)}{2}}$$

$$-\infty < \nu < +\infty$$

William Sealy Gosset



Student in 1908







Properties of t-distribution

(1) If n=1, then it follows a Cauchy distribution

(2) If
$$n > 2$$
, then ET=0, DT= $\frac{n}{n-2}$

(3) If n becomes large, the t distribution converges to an N(0, 1) distribution.

(4) t distributions have heavier tails than N(0,1).





Exercise, Using R

- 1. plot standard normal curve, and return the cumulative probability of the standard normal distribution at 1.96;
- •2. construct a χ 2 distribution (df=1), make a histogram of x2;
- 3. construct a t-distribution (with df=5) from the normal and χ^2 distributions, generate a histogram for the data;
- 4. randomly select 6 values from N(0,1) then replete 10000 times. Plot a histogram of its sample mean.

