

# Kalman Filtering and Smoothing Calculations

Using the preferred timing for the state space model,

$$\begin{aligned} y_t &= Z_t \alpha_t + d_t + \epsilon_t & \epsilon_t &\sim N(0, H_t) \\ \alpha_t &= T_t \alpha_{t-1} + c_t + R_t \eta_t & \eta_t &\sim N(0, Q_t) \end{aligned}$$

the filtered estimates are given by

$$\begin{aligned} a_{t+1} &= T_{t+1} a_t + c_{t+1} + K_t v_t & P_{t+1} &= T_{t+1} P_t L'_t + R_{t+1} Q_{t+1} R'_{t+1} \\ v_t &= y_t - Z_t a_t - d_t & K_t &= T_{t+1} P_t Z'_t F_t^{-1} \\ F_t &= Z_t P_t Z'_t + H_t & L_t &= T_{t+1} - K_t Z_t \end{aligned}$$

and the smoothed estimates are given by

$$\begin{aligned} \hat{\alpha}_t &= a_t + P_t r_t & V_t &= P_t - P_t N_t P_t \\ r_t &= Z'_t F_t^{-1} v_t + L'_t r_{t+1} & N_t &= Z'_t F_t^{-1} Z_t + L'_t N_{t+1} L_t \end{aligned}$$

where  $r_{n+1} = 0$  and  $N_{n+1} = 0$ .

## Univariate Filter

When any  $H_t$  is non-diagonal, the observation equation is transformed by taking the LDL factorization of the  $H_t$  matrices:

$$\begin{aligned} y_t^* &= Z_t^* \alpha_t + d_t^* + \epsilon_t^* & \epsilon_t &\sim N(0, H_t^*) \\ y_t^* &= C_t^{-1} y_t & Z_t^* &= C_t^{-1} Z_t & d_t^* &= C_t^{-1} d_t & \epsilon_t^* &= C_t^{-1} \epsilon_t & H_t &= C_t H_t^* C'_t \end{aligned}$$

To avoid taking the inverse of  $F_t$  in the above recursion, several quantities are computed using the univariate filter that can be used on this transformed system:

$$\begin{aligned} a_{t,i+1} &= a_{t,i} + K_{t,i} v_{t,i}^* & P_{t,i+1} &= P_{t,i} - K_{t,i} F_{t,i} K'_{t,i} \\ v_{t,i}^* &= y_{t,i}^* - Z_{t,i}^* a_{t,i} - d_{t,i}^* & F_{t,i} &= Z_{t,i}^* P_{t,i} Z_{t,i}^{*'} + H_{t,i}^* & K_{t,i} &= P_{t,i} Z_{t,i}^{*'} F_{t,i}^{-1} \\ a_{t+1,1} &= T_{t+1} a_{t,p+1} + c_{t+1} & P_{t+1,1} &= T_{t+1} P_{t,p+1} T'_{t+1} + R_{t+1} Q_{t+1} R'_{t+1} \end{aligned}$$

where  $Z_{t,i}^*$  is the  $i$ th row of  $Z_t^*$ ,  $d_{t,i}^*$  is the  $i$ th element of  $d_t^*$ , and  $H_{t,i}^*$  is the  $i$ th diagonal element of  $H_t^*$ . The filtered estimates of the state  $a_t = a_{t,1}$  and  $P_t = P_{t,1}$  are equivalent to those computed above. Similarly, the univariate smoother provides  $r_t = r_{t,0}$  and  $N_t = N_{t,0}$ :

$$\begin{aligned} r_{t,i-1} &= Z_{t,i}^{*'} F_{t,i}^{-1} v_{t,i}^* + L'_{t,i} r_{t,i} & N_{t,i-1} &= Z_{t,i}^{*'} F_{t,i}^{-1} Z_{t,i}^* + L'_{t,i} N_{t,i} L_{t,i} & L_{t,i} &= I_m - K_{t,i} Z_{t,i}^{*'} F_{t,i}^{-1} \\ r_{t-1,p} &= T'_{t-1} r_{t,0} & N_{t-1,p} &= T'_{t-1} N_{t,0} T_{t-1} \end{aligned}$$

where  $r_{t+1,p} = 0$  and  $N_{t+1,p} = 0$ .

## Exact Initial Kalman Filter

When the state  $\alpha_t$  is stationary, the initial values  $a_0$  and  $P_0$  can be computed as the unconditional mean and variance of the state given the system parameters:

$$\begin{aligned} a_0 &= (I_m - T_0)^{-1} c_0 \\ \text{vec}(P_0) &= (I_{m^2} - T_0 \otimes T_0)^{-1} \text{vec}(R_0 Q_0 R_0') \end{aligned}$$

In cases where the state  $\alpha_t$  is nonstationary, the state is separated into stationary components and nonstationary components via the eigenvalues of the  $T_0$  matrix – states associated with an eigenvalue less than one in absolute value are considered stationary while those with an eigenvalue greater than one are diffuse. The initial values are given by

$$\begin{aligned} a_0 &= a + R_0 \eta_0 + A \delta & \eta_0 &\sim N(0, Q_0) & \delta &\sim N(0, \kappa I) \\ P_0 &= P_{*,0} + \kappa P_{\infty,0} & P_{*,0} &= R_0 Q_0 R_0' & P_{\infty,0} &= A A' \end{aligned}$$

The unconditional mean of the stationary states is computed as  $a_0$  was above and placed in  $a$  with any elements of  $a$  associated with diffuse states set to 0. The unconditional variance of the stationary states,  $Q_0$ , is computed as  $P_0$  was in the stationary case. The selection matrix  $R_0$  is composed of columns of the identity matrix such that the initial shock  $\eta_0$  is applied to the stationary states. The selection matrix  $A$  is composed of the columns of the identity matrix associated with the diffuse states such that taking the limit as  $\kappa \rightarrow \infty$  allows for states with infinite initial variance.

The univariate filter recursions must be altered to separate the states with finite v. infinite variances:

$$\begin{aligned} F_{*,t,i} &= Z_{t,i}^* P_{*,t,i} Z_{t,i}^{*'} + H_{t,i}^* & F_{\infty,t,i} &= Z_{t,i}^* P_{\infty,t,i} Z_{t,i}^{*'} \\ K_{*,t,i} &= P_{*,t,i} Z_{t,i}^{*'} & K_{\infty,t,i} &= P_{\infty,t,i} Z_{t,i}^{*'} \\ a_{t,i+1} &= \begin{cases} a_{t,i} + K_{*,t,i} F_{*,t,i}^{-1} v_{t,i}^* & F_{\infty,t,i} = 0 \\ a_{t,i} + K_{\infty,t,i} F_{\infty,t,i}^{-1} v_{t,i}^* & F_{\infty,t,i} \neq 0 \end{cases} \\ P_{*,t,i+1} &= \begin{cases} P_{*,t,i} - K_{*,t,i} K_{*,t,i}' F_{*,t,i}^{-1} & F_{\infty,t,i} = 0 \\ P_{*,t,i} + K_{\infty,t,i} K_{\infty,t,i}' F_{\infty,t,i}^{-2} - (K_{*,t,i} K_{\infty,t,i}' + K_{\infty,t,i} K_{*,t,i}') F_{\infty,t,i}^{-1} & F_{\infty,t,i} \neq 0 \end{cases} \\ P_{\infty,t,i+1} &= \begin{cases} P_{\infty,t,i} & F_{\infty,t,i} = 0 \\ P_{\infty,t,i} - K_{\infty,t,i} K_{\infty,t,i}' F_{\infty,t,i}^{-1} & F_{\infty,t,i} \neq 0 \end{cases} \\ a_{t+1,1} &= T_{t+1} a_{t,p+1} + c_{t+1} \\ P_{\infty,t+1,1} &= T_{t+1} P_{\infty,t,p+1} T_{t+1}' & P_{*,t+1,1} &= T_{t+1} P_{*,t,p+1} T_{t+1}' + R_{t+1} Q_{t+1} R_{t+1}' \end{aligned}$$

For any set of system parameters where the state can be identified, there exists some time  $d$  such that  $F_{\infty,d,i} = 0$  for all  $i$ . For time  $t > d$ , the simpler Kalman filter recursion above can be employed.

The smoother must be similarly altered so that beginning at  $t = d$ , the computation of  $r_{t,i}$  is expanded to account for the initialization:

$$\begin{aligned}
L_{\infty,t,i} &= I_m - K_{\infty,t,i} Z_{t,i}^* F_{\infty,t,i}^{-1} & L_{*,t,i} &= I_m - K_{*,t,i} Z_{t,i}^* F_{*,t,i}^{-1} \\
L_{t,i}^{(0)} &= (K_{\infty,t,i} F_{*,t,i} F_{\infty,t,i}^{-1} - K_{*,t,i}) Z_{t,i}^* F_{\infty,t,i}^{-1} \\
r_{t,i-1}^{(0)} &= \begin{cases} Z_{t,i}' F_{*,t,i}^{-1} v_{t,i}^* + L_{*,t,i}' r_{t,i}^{(0)} & F_{\infty,t,i} = 0 \\ L_{\infty,t,i}' r_{t,i}^{(0)} & F_{\infty,t,i} \neq 0 \end{cases} \\
r_{t,i-1}^{(1)} &= \begin{cases} r_{t,i}^{(1)} & F_{\infty,t,i} = 0 \\ Z_{t,i}' F_{\infty,t,i}^{-1} v_{t,i}^* + L_{t,i}^{(0)'} r_{t,i}^{(0)} + L_{\infty,t,i}' r_{t,i}^{(1)} & F_{\infty,t,i} \neq 0 \end{cases} \\
r_{t-1,p}^{(0)} &= T_t' r_{t,0}^{(0)} & r_{t-1,p}^{(1)} &= T_t' r_{t,0}^{(1)} \\
\hat{a}_t &= a_t + P_{*,t,1} r_{t,0}^{(0)} + P_{\infty,t,1} r_{t,0}^{(1)}
\end{aligned}$$

where  $r_{d,p}^{(0)} = r_{d,p}$  and  $r_{d,p}^{(1)} = 0$ . Note that  $L_{\infty,t,i}$  and  $L_{t,i}^{(0)}$  only need to be computed when  $F_{\infty,t,i} \neq 0$  and  $L_{*,t,i}$  only needs to be computed when  $F_{\infty,t,i} = 0$ .

For the smoothed variance of the state,

$$\begin{aligned}
V_t &= P_{*,t,1} - P_{*,t,1} N_{t,0}^{(0)} P_{*,t,1} - \left( P_{\infty,t,1} N_{t,0}^{(1)} P_{*,t,1} \right)' - P_{\infty,t,1} N_{t,0}^{(1)} P_{*,t,1} - P_{\infty,t,1} N_{t,0}^{(2)} P_{\infty,t,1} \\
N_{t-1,p}^{(0)} &= T_t' N_{t,0}^{(0)} T_t & N_{t-1,p}^{(1)} &= T_t' N_{t,0}^{(1)} T_t & N_{t-1,p}^{(2)} &= T_t' N_{t,0}^{(2)} T_t
\end{aligned}$$

where when  $F_{\infty,t,i} = 0$ ,

$$N_{t,i-1}^{(0)} = Z_{t,i}' F_{*,t,i}^{-1} Z_{t,i}^* + L_{*,t,i}' N_{t,i}^{(0)} L_{*,t,i} \quad N_{t,i-1}^{(1)} = N_{t,i}^{(1)} \quad N_{t,i-1}^{(2)} = N_{t,i}^{(2)}$$

and when  $F_{\infty,t,i} \neq 0$ ,

$$\begin{aligned}
N_{t,i-1}^{(0)} &= L_{\infty,t,i}' N_{t,i}^{(0)} L_{\infty,t,i} \\
N_{t,i-1}^{(1)} &= Z_{t,i}' F_{\infty,t,i}^{-1} Z_{t,i}^* + L_{\infty,t,i}' N_{t,i}^{(0)} L_{t,i}^{(0)} + L_{\infty,t,i}' N_{t,i}^{(1)} L_{\infty,t,i} \\
N_{t,i-1}^{(2)} &= Z_{t,i}' F_{\infty,t,i}^{-2} Z_{t,i}^* F_{*,t,i} + L_{t,i}^{(0)'} N_{t,i}^{(1)} L_{t,i}^{(0)} + L_{\infty,t,i}' N_{t,i}^{(1)} L_{t,i}^{(0)} + L_{t,i}^{(0)'} N_{t,i}^{(1)} L_{\infty,t,i} + L_{\infty,t,i}' N_{t,i}^{(2)} L_{\infty,t,i}
\end{aligned}$$

where  $N_{d,p}^{(0)} = N_{d,p}$  and  $N_{d,p}^{(1)} = N_{d,p}^{(2)} = 0$ .

## Likelihood Calculation

The likelihood of of data  $y_1, \dots, y_n$  is given by

$$\log L(Y_n) \equiv -\frac{np}{2} \log(2\pi) - \frac{1}{2} \sum_{t=1}^n \log |F_t| + v_t' F_t^{-1} v_t$$

## Univariate Gradient

From Jungbacker, Koopman & van der Wel (2011), the gradient of the likelihood can be computed without taking the inverse of  $F_t$  provided the smoothed estimates are computed. The gradient of the likelihood  $L_n$  with respect to each parameter matrix is (note, wrong timing, corrected for naming differences):

$$\begin{aligned} \frac{\partial L_n}{\partial Z_t} &= H_t^{-1} [(y_t - d_t) \hat{\alpha}_t' - Z_t M_{Z_t}] \\ \frac{\partial L_n}{\partial d_t} &= H_t^{-1} (y_t - Z_t \hat{\alpha}_t - d_t) \\ \frac{\partial L_n}{\partial H_t} &= H_t^{-1} M_{H_t} H_t^{-1} - \frac{1}{2} \text{diag} \{ H_t^{-1} M_{H_t} H_t^{-1} \} \\ \frac{\partial L_n}{\partial T_t} &= \bar{R}_t' Q_t^{-1} \bar{R}_t (M_{T_t} - T_t M_{Z_t}) \\ \frac{\partial L_n}{\partial c_t} &= \bar{R}_t' Q_t^{-1} \bar{R}_t (\hat{\alpha}_{t+1} - T_t \hat{\alpha}_t - c_t) \\ \frac{\partial L_n}{\partial Q_t} &= Q_t^{-1} M_{Q_t} Q_t^{-1} - \frac{1}{2} \text{diag} \{ Q_t^{-1} M_{Q_t} Q_t^{-1} \} \end{aligned}$$

where

$$\begin{aligned} M_{Z_t} &= \hat{\alpha}_t \hat{\alpha}_t' + V_t \\ M_{H_t} &= (y_t - Z_t \hat{\alpha}_t)(y_t - Z_t \hat{\alpha}_t)' + Z_t V_t Z_t' - H_t \\ M_{T_t} &= \hat{\alpha}_t \hat{\alpha}_t' + J_t \\ M_{Q_t} &= \mathbb{E}(\eta_t \eta_t' | y_1, \dots, y_n) - Q_t \end{aligned}$$

Jungbacker & Koopman (2015), also have a formulation of the gradient but do so with time-invariant matrices without mean adjustments. Starting from the likelihood function

$$L_n = c - \frac{n}{2} \log |H| - \frac{1}{2} \text{tr} M_H - \frac{n-1}{2} \log |Q| - \frac{1}{2} \text{tr} M_Q \\ - \frac{1}{2} \log |P| - \frac{1}{2} \text{tr} (P^{-1}((\hat{\alpha}_1 - a)(\hat{\alpha}_1 - a)' + V_1))$$

they produce parameter gradients (with a fair amount of guessing on my part for the timing)

$$\frac{\partial L_n}{\partial Z_t} = H_t^{-1} [y_t \hat{\alpha}'_t - Z_t M_{Z_t}] \\ \frac{\partial L_n}{\partial H_t} = H_t^{-1} M_{H_t} H_t^{-1} - \frac{1}{2} \text{diag} \{ H_t^{-1} M_{H_t} H_t^{-1} \} \\ \frac{\partial L_n}{\partial T_t} = Q_t^{-1} (M_{T_t} - T_t M_{Z_t}) \quad (\text{or is this } M_{Z_{t-1}}?) \\ \frac{\partial L_n}{\partial Q_t} = Q_t^{-1} M_{Q_t} Q_t^{-1} - \frac{1}{2} \text{diag} \{ Q_t^{-1} M_{Q_t} Q_t^{-1} \}$$

where

$$M_{Z_t} = \hat{\alpha}_t \hat{\alpha}'_t + V_t \\ M_{H_t} = \hat{\epsilon}_t \hat{\epsilon}'_t + \text{Var}(\epsilon_t | Y_n, \theta) \\ M_{T_t} = \hat{\alpha}_t \hat{\alpha}'_t + J_t \\ M_{Q_t} = \hat{\eta}_t \hat{\eta}'_t + \text{Var}(\eta_t | Y_n, \theta)$$

## Multivariate Gradient

From Nagakura (working paper), we can get the gradient:

$$\begin{aligned}
G_\theta(\ell_t) &= G_\theta(a_t)Z_t'w_t \\
&+ \frac{1}{2}G_\theta(P_t)\text{vec}(Z_t'w_tw_t'Z_t - Z_t'F_t^{-1}Z_t) \\
&+ G_\theta(d_t)w_t \\
&+ G_\theta(Z_t)\text{vec}(w_ta_t' + w_tv_t'M_t' - M_t') \\
&+ \frac{1}{2}G_\theta(H_t)\text{vec}(w_tw_t' - F_t^{-1})
\end{aligned}$$

where

$$\begin{aligned}
G_\theta(a_{t+1}) &= G_\theta(a_t)L_t' \\
&+ G_\theta(P_t)(Z_t'w_t \otimes L_t') \\
&+ G_\theta(c_{t+1}) \\
&- G_\theta(d_t)K_t' \\
&+ G_\theta(Z_t)[P_tL_t' \otimes w_t - (a_t + M_tv_t) \otimes K_t'] \\
&- G_\theta(H_t)(w_t \otimes K_t') \\
&+ G_\theta(T_{t+1})[(a_t + M_tv_t) \otimes I_m]
\end{aligned}$$

$$\begin{aligned}
G_\theta(P_{t+1}) &= G_\theta(P_t)(L_t' \otimes L_t') \\
&+ G_\theta(H_t)(K_t' \otimes K_t') \\
&+ G_\theta(Q_{t+1})(R_{t+1}' \otimes R_{t+1}') \\
&+ [G_\theta(T_{t+1})(P_tL_t' \otimes I_m) \\
&\quad - G_\theta(Z_t)(P_tL_t' \otimes K_t') \\
&\quad + G_\theta(R_{t+1})(Q_{t+1}R_{t+1}' \otimes I_m)]N_m
\end{aligned}$$

The initial conditions for the recursion are given by a simplification of the expressions above (which can also be easily derived from the explicit expressions for  $a_1$  and  $P_1$ ):

$$\begin{aligned}
G_\theta(a_1) &= G_\theta(a_0)T_1' \\
&+ G_\theta(c_1) \\
&+ G_\theta(T_1)[a_0 \otimes I_m]
\end{aligned}$$

$$\begin{aligned}
G_\theta(P_1) &= G_\theta(P_0)(T_1' \otimes T_1') \\
&+ G_\theta(Q_1)(R_1' \otimes R_1') \\
&+ [G_\theta(T_1)(P_0T_1' \otimes I_m) \\
&\quad + G_\theta(R_1)(Q_1R_1' \otimes I_m)]N_m
\end{aligned}$$

To determine  $G_\theta(a_0)$  and  $G_\theta(P_0)$  when  $a_0$  and  $P_0$  are set as the unconditional mean and variance of the state (i.e., when they are not explicitly provided and the system is stationary) use the definitions of the unconditional state:

$$\begin{aligned}
a_0 &= (I_m - T)^{-1}c \\
G_\theta(a_0) &= G_\theta([I_m - T]^{-1}c) \\
&= G_\theta([I_m - T]^{-1})(c \otimes I_m) + G_\theta(c)(I_1 \otimes [(I_m - T)^{-1}]') \\
&= -G_\theta(I_m - T)[(I_m - T)^{-1} \otimes (I_m - T)^{-1}](c \otimes I_m) + G_\theta(c)(I_m - T)'^{-1} \\
&= G_\theta(T)[(I_m - T)^{-1} \otimes (I_m - T)^{-1}](c \otimes I_m) + G_\theta(c)(I_m - T)'^{-1}
\end{aligned}$$

$$\begin{aligned}
\text{vec}(P_0) &= (I_{m^2} - T \otimes T)^{-1}\text{vec}(RQR') \\
G_\theta(P_0) &= G_\theta(\text{vec}(P_0)) \\
&= G_\theta(S \text{vec}(RQR')) \\
&= G_\theta(S)[\text{vec}(RQR') \otimes I_{m^2}] + G_\theta(\text{vec}(RQR'))(I_1 \otimes S') \\
&= -G_\theta(I_{m^2} - T \otimes T)(S \otimes S')[\text{vec}(RQR') \otimes I_{m^2}] \\
&\quad + [G_\theta(R)(QR' \otimes I_m)N_m + G_\theta(Q)(R' \otimes R')]S' \\
&= G_\theta(T \otimes T)(S \otimes S')[\text{vec}(RQR') \otimes I_{m^2}] \\
&\quad + [G_\theta(R)(QR' \otimes I_m)N_m + G_\theta(Q)(R' \otimes R')]S'
\end{aligned}$$

where  $S = (I_{m^2} - T \otimes T)^{-1}$ . Note that  $G_\theta(T \otimes T)$  must be computed separately.