

## Tutorial Session

# **Solution and Estimation Methods for Nonlinear DSGE Models**

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# Introduction

- Dynamic stochastic general equilibrium (DSGE) models have been one of the primary tools in macroeconomic analysis.
  - Microfoundation
  - Immune to the Lucas Critique
  - Suitable for welfare analysis
  - Rational expectations
- Following the development of Bayesian estimation techniques, many economists have estimated DSGE models.
  - Cross-equation restrictions
  - Avoid weak-instrument problems that would arise in GMM estimation.
  - Priors for structural parameters
    - Deal with identification issues.
    - Estimate parameters with bounded domains.
    - Make the likelihood function well-shaped.

# Introduction

- While DSGE models are inherently nonlinear, the nonlinearities are often small.
- Equilibrium conditions are linearly approximated.
  - Rational expectations solution (policy function) can be obtained using a standard solution method.
    - Blanchard and Kahn (1980); Sims (2002)
  - Likelihood can be evaluated using the Kalman filter.
    - Additional assumption: Normality of shocks

# Steps for Estimating Linearized DSGE Models

- DSGE model:

$$\Gamma_0(\theta)s_t = \Gamma_1(\theta)s_{t-1} + \Psi_0(\theta)\varepsilon_t + \Pi_0(\theta)\eta_t$$

- Rational expectations solution (state transition equations):

$$s_t = \Phi_1(\theta)s_{t-1} + \Phi_\varepsilon(\theta)\varepsilon_t \quad \varepsilon_t \sim N(0, \Sigma_\varepsilon)$$

- Relation between model variables and data (observation equations):

$$y_t = \Psi_0(\theta) + \Psi_1(\theta)s_t + u_t \quad u_t \sim N(0, \Sigma_u)$$

- State transition equations & observation equations  $\Rightarrow$  Kalman filter  
 $\Rightarrow$  Likelihood function:  $L(\theta|Y)$
- Prior distribution:  $p(\theta)$
- Bayes' Theorem  $\Rightarrow$  Posterior distribution

$$p(\theta|Y) \propto L(\theta|Y)p(\theta)$$

# Needs for Nonlinear DSGE models

- However, linearized DSGE models cannot deal with features that generate pronounced nonlinearities:
  - Occasionally binding constraints
  - Stochastic volatilities
  - Markov switching coefficients
  - Asymmetric adjustment costs
- Nonlinear solution and estimation methods are required.

# Session Plan

- ① Sunakawa: How to solve nonlinear DSGE models
  - ② Hirose: How to estimate nonlinear DSGE models
- A survey paper will be published in *Japanese Economic Review*.

# 1. How to solve nonlinear DSGE models

# Plan for the first part

- What is the time iteration method?
  - Time iteration is a method to solve the Collman (1990) operator.
- Applications to New Keynesian models
  - Need to deal with *the curse of dimensionality*.
  - Smolyak's method with sparse grid points (Fernández-Villaverde et al., 2015)
  - Simulation-based method with EDS grid (Maliar and Maliar, 2015)
- MATLAB codes available at <https://github.com/tkksnk/NKZLB>



# What is the time iteration method?

- Neoclassical growth model
  - Univariate Interpolation for  $k$
  - Nonlinear optimization (or how to avoid it)
- Stochastic neoclassical growth model
  - Multivariate interpolation for  $(k, z)$
  - Nonlinear optimization (or how to avoid it)
  - Numerical integration (or how to avoid it)

# Neoclassical growth model

- We consider an example of neoclassical growth model. An individual maximizes life-time utility

$$\sum_{t=0}^{\infty} \beta^t u(c_t)$$

subject to

$$c_t + k_{t+1} \leq f(k_t).$$

where  $u(\cdot)$  and  $f(\cdot)$  satisfy standard conditions.

- The first-order necessary condition is given by

$$u_c(c_t) = \beta u_c(c_{t+1}) f_k(k_{t+1})$$

where  $u_c(\cdot)$  denotes the derivative of  $u$  wrt  $c$  and  $f_k(\cdot)$  denotes the derivative of  $f$  wrt  $k$ .

# Collman operator

- There is a mapping  $\sigma = K\sigma$  that solves

$$u_c(c) = \beta u_c(\sigma(f(k) - c)) f_k(f(k) - c)$$

for  $c = \sigma(k)$ .  $\sigma$  is called policy function.

- Note that  $k' = f(k) - c$  and  $c' = \sigma(k') = \sigma(f(k) - c)$ .
- Bellman operator

$$V(k) = \max_{c \in (0, f(k)]} \{u(c) + \beta V(f(k) - c)\}.$$

is to solve the Bellman equation, whereas the Collman operator is helpful to solve the Euler equation.

## Collman operator, cont'd

- Collman (1990) proves the existence of the fixed point of  $K$  in a stochastic neoclassical growth model with distortionary tax.
  - Greenwood and Huffman (1995) extend it to several cases. Also see Richter, Throckmorton and Walker (2014) and Sargent and Stachurski (2018).

# Time iteration: Algorithm

- Time iteration is a method to solve the Collman operator.
- The time iteration method takes the following steps:
  - 1 Make an initial guess for the policy function  $\sigma^{(0)}$ .
  - 2 Given the policy function previously obtained  $\sigma^{(i-1)}$  ( $i$  is an index for the number of iteration), solve

$$u_c(c) = \beta u_c \left( \sigma^{(i-1)}(f(k) - c) \right) f_k(f(k) - c)$$

for  $c$ .

- 3 Update the policy function by setting  $c = \sigma^{(i)}(k)$ .
- 4 Repeat 2-3 until  $\|\sigma^{(i)} - \sigma^{(i-1)}\|$  is small enough.

# Optimization and interpolation

- We discretize the state space of  $k$  by grid points:

$$k_j \in \{k_1, k_2, \dots, k_N\},$$

where  $j$  is an index for grid points.

- Then, given  $\sigma^{(i-1)}$ , we solve

$$\tilde{R}(c; k_j, \sigma^{(i-1)}) \equiv -u_c(c) + \beta u_c \left( \sigma^{(i-1)}(f(k_j) - c) \right) f_k(f(k_j) - c) = 0$$

for  $c$  at each grid point  $k_j$ . That is, the residual function is equal to zero at each grid point. This is called **collocation**.

- We need to know the value of  $\sigma^{(i-1)}(f(k_j) - c)$ , which may be off the grid points.

# Optimization and interpolation, cont'd

- More generally, we want to:
- solve  $f(x) = 0$  for  $x$  (optimization),
- when we know only the values of  $f(x_j)$  at  $x_j \in \{x_1, \dots, x_N\}$  (interpolation).

# Polynomial function

- The policy functions are approximated by a higher-order polynomial function

$$\hat{\sigma}(x; \boldsymbol{\theta}) = \theta_0 + \theta_1 x + \theta_2 x^2 + \cdots + \theta_{N-1} x^{N-1}.$$

- We need to know the values of  $\sigma(x)$  (at least) at  $N$  grid points to fit the polynomial.



# Chebyshev polynomial

- We define the basis functions  $T(x) : [-1, 1] \rightarrow [-1, 1]$ , and have an univariate polynomial

$$\hat{\sigma}(x; \boldsymbol{\theta}) = \theta_0 + \theta_1 T_1(x) + \theta_2 T_2(x) + \cdots + \theta_{N-1} T_{N-1}(x).$$

- An example of  $T(x)$ :

$$T_0(x) = 1,$$

$$T_1(x) = x,$$

$$T_2(x) = 2x^2 - 1,$$

$$\vdots$$

$$T_n(x) = 2xT_{n-1}(x) - T_{n-2}(x).$$

where  $x \in [-1, 1]$ . These are called Chebyshev polynomials, or Chebyshev basis functions.

# Chebyshev collocation points

- The polynomial is evaluated at the collocation points
- Chebyshev zeros:

$$x_j = \cos \left( \frac{(2j+1)\pi}{2(N-1)} \right) \text{ for } j = 0, 1, \dots, N-1.$$

- Chebyshev extrema:

$$x_j = \cos \left( \frac{j\pi}{N-1} \right) \text{ for } j = 0, 1, \dots, N-1.$$

# Transforming grid points

- For a general function which takes a value  $k_j \in [k_1, k_N]$  at each grid point, we have to transform  $k_j$  to  $x_j \in [-1, 1]$  to (or  $x_j$  to  $k_j$ ) by applying

$$x_j = \varphi(k_j) = \frac{2(k_j - k_1)}{k_N - k_1} - 1,$$

or

$$k_j = \varphi^{-1}(x_j) = k_1 + 0.5(1 + x_j)(k_N - k_1).$$

## Fitting polynomial

- Once we have the collocation points  $\{x_j\}$  and the function values  $\{\sigma(x_j)\}$  evaluated at  $x_j$  for  $j = 1, 2, \dots, N$ , we can fit  $\hat{\sigma}(x; \theta)$  to the data to obtain  $\theta$ .

$$\begin{bmatrix} \sigma(x_1) \\ \sigma(x_2) \\ \vdots \\ \sigma(x_N) \end{bmatrix} = \begin{bmatrix} 1 & T_1(x_1) & T_2(x_1) & \dots & T_{N-1}(x_1) \\ 1 & T_1(x_2) & T_2(x_2) & \dots & T_{N-1}(x_2) \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 1 & T_1(x_N) & T_2(x_N) & \dots & T_{N-1}(x_N) \end{bmatrix} \begin{bmatrix} \theta_0 \\ \theta_1 \\ \vdots \\ \theta_{N-1} \end{bmatrix},$$

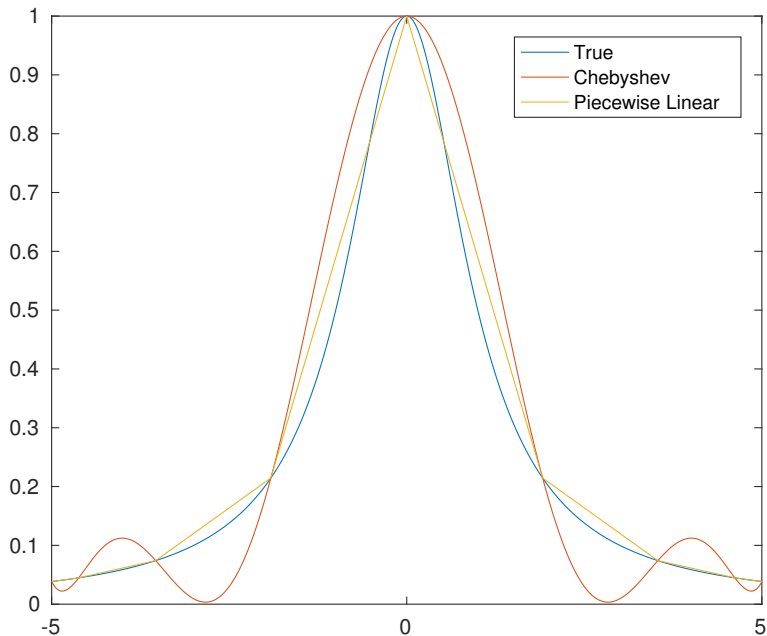
or

$$\sigma(\mathbf{x}) = T(\mathbf{x})\boldsymbol{\theta}.$$

Then we have  $\boldsymbol{\theta} = T(\mathbf{x})^{-1}\sigma(\mathbf{x})$ .  $T(\mathbf{x})$  needs to be nonsingular.

- The basis functions with Chebyshev zeros/extrema satisfy orthogonality property. That is, each column of  $T(\mathbf{x})$  is uncorrelated to each other.

Example:  $1/(1+x^2)$  with  $N = 9$



# Nonlinear optimization

- Given  $\hat{\sigma}^{(i-1)}(k; \theta) = T(\varphi(k))\theta$ , we solve a nonlinear equation

$$\tilde{R}(c; k_j, \hat{\sigma}^{(i-1)}) \approx -u_c(c) + \beta u_c \left( \hat{\sigma}^{(i-1)}(f(k_j) - c; \theta) \right) f_k(f(k_j) - c) = 0$$

for  $c$  at each grid point  $k_j$ .

- We use Newton's method to solve the nonlinear equation. For example, Matlab's command `fsolve` or Chris Sims' `csolve` does such a job.
- But, such a nonlinear optimization can be costly when the number of grid points and/or the number of nonlinear equations is large.

# Parameterized expectation

- By applying a version of the parameterized expectation algorithm (PEA), **we can avoid solving nonlinear equations** (Maliar and Maliar, 2015; Gust et al., 2017, Hirose and Sunakawa, 2015; 2017).
  - Marcet (1988) uses a stochastic approach based on Monte Carlo simulations.
  - Christiano and Fisher (2000) propose a non-stochastic approach called **PEA collocation**.
  - Endogenous grid-point method (EGM) is another popular method to avoid nonlinear optimization.

# PEA collocation

- There are two ways for applying PEA collocation (Christiano and Fisher, 2000):
  - One is to fit polynomials to future variables.
  - The other is to fit polynomials to current variables.



# Fitting future variables

- We define

$$e(k) \equiv \beta u_c(c') f_k(k').$$

- Then, given the values of  $e^{(i-1)}(k_j)$  at each grid point, we have

$$\begin{aligned} c &= u_c^{-1}(e^{(i-1)}(k_j)), \\ k' &= f(k_j) - c, \end{aligned}$$

and an intermediate policy function  $c = \sigma^{(i)}(k_j)$ . Note that we don't have to solve the nonlinear equation here.

## Fitting future variables, cont'd

- We also update

$$e^{(i)}(k_j) = \beta u_c \left( \sigma^{(i)}(k') \right) f_k(k')$$

where  $k'$  is obtained in the previous step.

- Note that  $c' = \sigma^{(i)}(k')$ , or equivalently  $e^{(i-1)}(k')$ , needs to be interpolated here. That is,

$$\begin{aligned} c' &= \sigma^{(i)}(k') \\ &\approx u_c^{-1}(\hat{e}^{(i-1)}(k'; \boldsymbol{\theta})) \end{aligned}$$

where  $\hat{e}^{(i-1)}(k; \boldsymbol{\theta}) = \theta_0 + \theta_1 T_1(k) + \theta_2 T_2(k) + \cdots + \theta_{N-1} T_{N-1}(k)$ .  
 $\boldsymbol{\theta}$  is obtained by fitting the polynomial to the data of future variables  $e^{(i-1)}(k_j)$  at each grid point  $k_j$ .

## Fitting current variables

- Or, we define

$$v(k) \equiv \beta u_c(c) f_k(k),$$

- Then, given the function  $v^{(i-1)}(k)$ , we have

$$\begin{aligned} c &= u_c^{-1}(v^{(i-1)}(k')), \\ &\approx u_c^{-1}(\hat{v}^{(i-1)}(f(k_j) - c; \theta)) \end{aligned}$$

- Note that  $v^{(i-1)}(k)$  needs to be interpolated by using its approximation  $\hat{v}^{(i-1)}(k; \theta)$ .  $\theta$  is obtained by fitting the polynomial to the data of current variables  $v^{(i-1)}(k_j)$  at each grid point  $k_j$ .
- We can use a successive approximation  $c = \sigma^{(i-1)}(k_j)$  to avoid nonlinear optimization.
- We also update

$$v^{(i)}(k_j) = \beta u_c(c) f_k(k_j),$$

where  $c$  is obtained in the previous step.

# What is the time iteration method?

- Neoclassical growth model
  - Univariate Interpolation  $k$
  - Nonlinear optimization (or how to avoid it)
- Stochastic neoclassical growth model
  - Multivariate interpolation  $(k, z)$
  - Nonlinear optimization (or how to avoid it)
  - Numerical integration (or how to avoid it)

# Stochastic neoclassical growth model

- Now we extend the earlier example with stochastic technology. An individual maximizes expected life-time utility

$$\mathbb{E}_0 \sum_{t=0}^{\infty} \beta^t u(c_t)$$

subject to

$$c_t + k_{t+1} \leq f(k_t, z_t).$$

$\mathbb{E}_0$  is expectation operator at time 0.

- $z_t$  follows an AR(1) process

$$z_{t+1} = \rho z_t + \epsilon_{t+1}, \quad \epsilon_{t+1} \sim N(0, \sigma_{\epsilon}^2)$$

# Collman operator

- The first-order necessary condition is given by

$$u_c(c_t) = \beta \mathbb{E}_t \{ u_c(c_{t+1}) f_k(k_{t+1}, z_{t+1}) \} .$$

- There is a mapping  $\sigma = K\sigma$  that solves

$$u_c(c) = \beta \int u_c(\sigma(f(k, z) - c, z')) f_k(f(k, z) - c, z') p(z'|z) dz'$$

for  $c = \sigma(k, z)$ .

# Time iteration

- The time iteration method takes the following steps:

- 1 Make an initial guess for the policy function  $\sigma^{(0)}$ .
- 2 Given the policy function previously obtained  $\sigma^{(i-1)}$ , solve

$$u_c(c) = \beta \int u_c \left( \sigma^{(i-1)}(f(k, z) - c, z') \right) f_k(f(k, z) - c, z') p(z'|z) dz'$$

for  $c$ .

- 3 Update the policy function by setting  $c = \sigma^{(i)}(k, z)$ .
- 4 Repeat 2-3 until  $\|\sigma^{(i)} - \sigma^{(i-1)}\|$  is small enough.

# Optimization, interpolation and integration

- We discretize the state space of  $(k, z)$  by grid points:

$$k_j \in \{k_1, k_2, \dots, k_N\}, \quad z_m \in \{z_1, z_2, \dots, z_N\},$$

where  $(j, m)$  is an index for the set of grid points.

- Then, given  $\sigma^{(i-1)}$ , we solve

$$\begin{aligned} & \tilde{R}(c; k_j, z_m, \sigma^{(i-1)}) \\ & \approx -u_c(c) \\ & + \beta \int \left[ u_c \left( \sigma^{(i-1)}(f(k_j, z_m) - c, z') \right) f_k(f(k_j, z_m) - c, z') p(z'|z_m) \right] dz' \\ & = 0 \end{aligned}$$

for  $c$  at each grid point  $(k_j, z_m)$  (optimization).



# Optimization, interpolation and integration, cont'd

- We need to know the value of  $\sigma^{(i-1)}(f(k_j, z_m) - c, z')$ , which may be off the grid points ([interpolation](#)).
- $\sigma(k, z)$  is a two-dimensional object, which can be approximated by two-dimensional Chebyshev polynomial.
- Also, we compute an integral with regard to  $z'$  for the next period's expectation ([integration](#)).

## 2-D Chebyshev polynomial

- The policy functions are approximated by basis functions. For example, if  $N_x = N_y = 3$ ,

$$\begin{aligned}\hat{\sigma}(x, y; \boldsymbol{\theta}) = & \theta_{0,0} + \theta_{1,0}T_1(x) + \theta_{2,0}T_2(x) + \theta_{0,1}T_1(y) + \theta_{0,2}T_2(y) \\ & + \theta_{1,1}T_1(x)T_1(y) + \theta_{1,2}T_1(x)T_2(y) \\ & + \theta_{2,1}T_2(x)T_1(y) + \theta_{2,2}T_2(x)T_2(y)\end{aligned}$$

There are 9 coefficients, so we need at least  $N = N_x N_y = 9$  collocation points.

- Chebyshev extrema is used as collocation points

$$(x, y) \in \{(0, 0), (-1, 0), (1, 0), (0, -1), (0, 1), (-1, -1), (1, -1), (-1, 1), (1, 1)\}.$$

## Fitting 2-D Chebyshev polynomial

- Once we have the collocation points  $\{x_i, y_j\}$  and the function values  $\{\sigma(x_i, y_j)\}$  evaluated at  $(x_i, y_j)$  for  $i = 1, 2, \dots, N_x$  and  $j = 1, 2, \dots, N_y$ , we can fit  $\hat{\sigma}(x, y; \boldsymbol{\theta})$  to the data to obtain  $\boldsymbol{\theta}$ .
- For example, if  $N_x = N_y = 2$ ,

$$\begin{bmatrix} \sigma(x_1, y_1) \\ \sigma(x_2, y_1) \\ \sigma(x_1, y_2) \\ \sigma(x_2, y_2) \end{bmatrix} = \begin{bmatrix} 1 & T_1(x_1) & T_1(y_1) & T_1(x_1)T_1(y_1) \\ 1 & T_1(x_2) & T_1(y_1) & T_1(x_2)T_1(y_1) \\ 1 & T_1(x_1) & T_1(y_2) & T_1(x_1)T_1(y_2) \\ 1 & T_1(x_2) & T_1(y_2) & T_1(x_2)T_1(y_2) \end{bmatrix} \begin{bmatrix} \theta_{0,0} \\ \theta_{1,0} \\ \theta_{1,0} \\ \theta_{1,1} \end{bmatrix}$$

or  $\sigma(\mathbf{x}, \mathbf{y}) = T(\mathbf{x}, \mathbf{y})\boldsymbol{\theta}$ . Then we have  $\boldsymbol{\theta} = T(\mathbf{x}, \mathbf{y})^{-1}\sigma(\mathbf{x}, \mathbf{y})$ .

- More generally, we have a tensor product for  $T(\mathbf{x}, \mathbf{y}) = T(\mathbf{x}) \otimes T(\mathbf{y})$ . The total number of grid points is exponentially increasing in the number of variables.

# Approximating stochastic process

- How to compute an integral with regard to  $z'$ ?
  - Tauchen's method (Rouwenhorst's method): AR(1) process is approximated by a Markov chain.
  - Gaussian Quadrature: The quadrature nodes  $\{x_i\}_{i=1}^M$  and quadrature weights  $\{w_i\}_{i=1}^M$  approximate

$$\int f(x)w(x)dx \approx \sum w_i f(x_i).$$

- The total number of quadrature points is exponentially increasing in the number of exogeneous variables.

# PEA collocation

- There are two ways for applying PEA collocation (Christiano and Fisher, 2000):
  - One is to fit polynomials to future variables.
  - The other is to fit polynomials to current variables.
  - In the latter approach, we can also avoid computing numerical integration in the expectation terms by precomputation technique of Judd, Maliar, Maliar and Tsener (2017).

# Numerical examples I

- Solve the stochastic neoclassical model by the time iteration method
  - Interpolation: Chebyshev polynomial (either  $N_d = 3$  or 5 for each  $d \in \{k, z\}$ ).
  - Optimization:
    - Chris Sims' csolve is used in Time iteration with Newton's method (TI).
    - No optimization is required in PEA collocation.
  - Integration:
    - Gaussian-Hermite quadrature is used in TI and PEA collocation with fitting future variables (future PEA).
    - Precomputation technique in Judd et al. (2017) is used in PEA collocation with fitting current variables (current PEA).

# Euler equation errors

- Accuracy and computation speed are compared. Look at the Euler equation errors:

$$\mathcal{E}(k, z) \equiv 1 - \beta \int \left\{ \left( \frac{\sigma_c(k', z')}{\sigma_c(k, z)} \right)^{-\tau} (1 - \delta + \alpha z' k'^{\alpha-1}) \right\} p(z'|z) dz'$$

where  $k' = f(k, z) - \sigma_c(k, z)$ .

# Summary

$(N_d, \tau)$	TI			future PEA			current PEA		
	$L_1$	$L_\infty$	CPU	$L_1$	$L_\infty$	CPU	$L_1$	$L_\infty$	CPU
(3, 1.0)	-5.12	-4.60	4.03	-4.23	-3.69	0.04	-3.13	-2.44	0.02
(5, 1.0)	-7.08	-6.72	9.76	-5.92	-5.59	0.09	-3.13	-2.44	0.04
(3, 2.0)	-4.82	-4.35	0.86	-3.99	-3.53	0.03	-2.95	-2.26	0.01
(5, 2.0)	-6.76	-6.45	2.88	-5.63	-5.36	0.11	-2.96	-2.27	0.04
(3, 5.0)	-4.48	-3.87	0.62	-3.57	-2.88	0.05	-2.67	-1.99	0.02
(5, 5.0)	-6.43	-5.38	1.91	-5.10	-3.90	0.15	-2.69	-2.00	0.05

Notes:  $L_1$  and  $L_\infty$  are, respectively, the average and maximum of absolute Euler errors (in log 10 units) on a 10,000 period stochastic simulation. CPU is the elapsed time for computing equilibrium (in seconds).



# Applications to New Keynesian models

- We solve small-scale nonlinear New Keynesian DSGE model with
  - Smolyak's method with sparse grid points
  - Simulation-based method with EDS grid
  - The techniques we mentioned earlier (PEA collocation and precomputing integrals) are also applied.

# Time iteration in the New Keynesian literature

- Time iteration is a popular method to solve nonlinear New Keynesian models.
  - We have to look at the decentralized economy, as the second welfare theorem fails to hold.
  - An incomplete list includes: Fernández-Villaverde, Gordon, Guerrón-Quintana, and Rubio-Ramírez, 2015; Maliar and Maliar, 2015; Gavin, Keen, Richter, and Throckmorton, 2015; Gust, Herbst, López-Salido, and Smith, 2017; Iliboshi, Ueda, and Shintani, 2018; Nakata, 2016a, 2016b; Hills, Nakata, and Schmitt, 2016; Dennis, 2016; Ngo, 2014; Hirose and Sunakawa, 2015, 2017; Hills, Nakata, and Sunakawa, 2018.

# A small scale New Keynesian DSGE model

- The example here is taken from An and Schorfheide (2007) and Herbst and Schorfheide (2015).
- The model economy consists of
  - Final-good and intermediate-good producing firms
  - Households
  - Monetary and fiscal authorities
- Prices are sticky due to Rotemberg-type (1982) adjustment cost.

# Model overview

- Equilibrium conditions (after detrending):

$$\begin{aligned}1 &= \beta \mathbb{E}_t \left[ \left( \frac{c_{t+1}}{c_t} \right)^{-\tau} \frac{R_t}{\gamma_{t+1} \pi_{t+1}} \right] \\0 &= (1 - \nu^{-1}) + \nu^{-1} \chi_H c_t^\tau - \phi (\pi_t - \bar{\pi}) \left[ \pi_t - \frac{1}{2\nu} (\pi_t - \bar{\pi}) \right] \\&+ \beta \phi \mathbb{E}_t \left[ \left( \frac{c_{t+1}}{c_t} \right)^{-\tau} \frac{y_{t+1}}{y_t} (\pi_{t+1} - \bar{\pi}) \pi_{t+1} \right], \\R_t^* &= \left( r \bar{\pi} \left( \frac{\pi}{\bar{\pi}} \right)^{\psi_1} \left( \frac{y_t}{y_t^*} \right)^{\psi_2} \right)^{1-\rho_R} R_{t-1}^{*\rho_R} e^{\epsilon_{R,t}}, \\c_t + \frac{\phi}{2} (\pi_t - \pi)^2 y_t &= g_t^{-1} y_t, \\R_t &= \max \{ R_t^*, 1 \}.\end{aligned}$$

There are 5 equations and 5 endogenous variables  $\{c_t, \pi_t, R_t^*, R_t, y_t\}$  and 3 exogenous variables  $\{\gamma_t, g_t, \epsilon_{R,t}\}$ . The natural level of output is given by  $y_t^* = (1 - \nu)^{1/\tau} g_t$ .

## Model overview, cont'd

- The solution has a form of

$$\begin{aligned}c &= \sigma_c(R_{-1}^*, s), & \pi &= \sigma_\pi(R_{-1}^*, s), \\ R^* &= \sigma_{R^*}(R_{-1}^*, s), & y &= \sigma_y(R_{-1}^*, s),\end{aligned}$$

where  $s = (\gamma, g, \epsilon_R)$ . Note that  $R = \max\{\sigma_{R^*}(R_{-1}^*, s), 1\}$ .

# Collman operator

- The mapping  $\sigma = K\sigma$  solves

$$0 = -c^{-\tau} + \beta R \int \left[ \frac{\sigma_c(R, s')^{-\tau}}{\gamma' \sigma_\pi(R, s')} \right] p(s'|s) ds',$$

$$0 = \left( (1 - \nu^{-1}) + \nu^{-1} c^\tau - \phi(\pi - \bar{\pi}) \left[ \pi - \frac{1}{2\nu} (\pi - \bar{\pi}) \right] \right) c^{-\tau} y \\ + \beta \phi \int [\sigma_c(R, s')^{-\tau} \sigma_y(R, s') (\sigma_\pi(R, s') - \bar{\pi}) \sigma_\pi(R, s')] p(s'|s) ds',$$

$$R = \left( r \bar{\pi} \left( \frac{\pi}{\bar{\pi}} \right)^{\psi_1} \left( \frac{y}{y^*} \right)^{\psi_2} \right)^{1-\rho_R} R_{-1}^{\rho_R} e^{\epsilon_R},$$

$$c + \frac{\phi}{2} (\pi - \bar{\pi})^2 y = g^{-1} y.$$

$$R = \max \{R^*, 1\},$$

for  $c, \pi, R^*, R, y$ .

# Time iteration

- The time iteration method takes the following steps:
  - 1 Make an initial guess for the policy function  $\sigma^{(0)}$ .
  - 2 Given the policy function previously obtained  $\sigma^{(i-1)}$ , solve the relevant equations for  $(c, \pi, R^*, y)$ .
  - 3 Update the policy function by setting  $c = \sigma_c^{(i)}(R_{-1}^*, s)$ ,  
 $\pi = \sigma_\pi^{(i)}(R_{-1}^*, s)$ ,  $R^* = \sigma_{R^*}^{(i)}(R_{-1}^*, s)$ , and  $y = \sigma_y^{(i)}(R_{-1}^*, s)$ .
  - 4 Repeat 2-3 until  $\|\sigma^{(i)} - \sigma^{(i-1)}\|$  is small enough.

# Optimization, interpolation, and integration

- Here, we need to solve the system of nonlinear equations (optimization).
  - PEA collocation can also be used here to avoid costly nonlinear optimization.
- Also, we need to evaluate the function  $\sigma(R^*, s')$  off the grid points (interpolation).
  - $\sigma(R_{-1}^*, s)$  is a high dimensional object (our model have 4 state variables,  $(R_{-1}^*, s) = (R_{-1}^*, \gamma, g, \epsilon_R)$ ), which can be dealt with Smolyak's method with sparse grid points or simulation based method with EDS grid.
- We compute an integral wrt  $s'$  for the next period's expectation (integration).
  - Precomputing integrals can also be applied here to avoid numerical integration.



# Smolyak's method

- We introduce Smolyak's (1963) method with sparse grid points to handle such a high-dimensional object.
  - Krueger and Kubler (2004) first introduced Smolyak sparse grid points to solve heterogeneous OLG models with aggregate uncertainty.
  - Applications of Smolyak's method to New Keynesian models are found in Fernández-Villaverde et al. (2015), Gust et al. (2017), Hirose and Sunakawa (2015; 2017).
- We look at simple cases with second-order polynomials with  $N_d = 3$  for each dimension.
  - Cases with higher-order polynomials (for example  $N_d = 5$  or 9) are a bit more complicated but manageable. See Judd, Maliar, Maliar and Valero (2014).

## Smolyak's method: Simple cases

- The policy functions are approximated by basis functions. For example, if we use second-order polynomials and  $N_x = N_y = 3$ ,

$$\hat{\sigma}(x, y; \boldsymbol{\theta}) = \theta_{0,0} + \theta_{1,0}T_1(x) + \theta_{2,0}T_2(x) + \theta_{0,1}T_1(y) + \theta_{0,2}T_2(y).$$

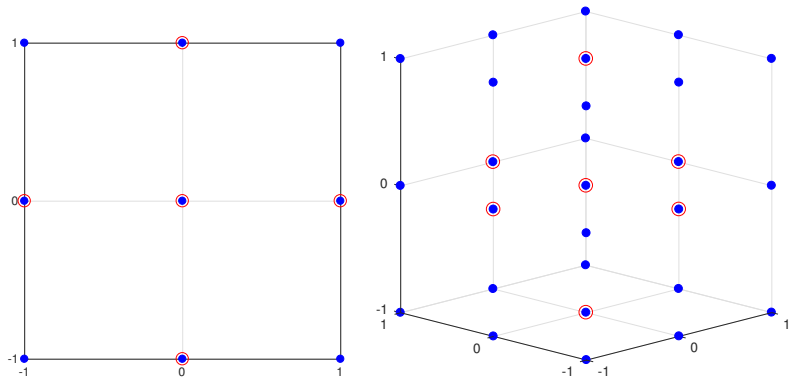
There are 5 coefficients, so we need 5 collocation points. We have just eliminated all the cross terms!

# Simple cases with second-order polynomials

- Chebyshev extrema is used as collocation points:

$$N_d = 2 : (x, y) \in \{(0, 0), (1, 0), (-1, 0), (0, 1), (0, -1)\}$$

$$N_d = 3 : (x, y, z) \in \{(0, 0, 0), (1, 0, 0), (-1, 0, 0), (0, 1, 0), (0, -1, 0), (0, 0, 1), (0, 0, -1)\}$$



## Smolyak's method: Simple cases, cont'd

- The total number of the grid points is  $1 + 2N_d$ , whereas it is  $3^{N_d}$  with the standard Chebyshev polynomials.

$N_d$	$1 + 2N_d$	$3^{N_d}$
2	5	9
3	7	27
4	9	81
5	11	243
$\vdots$	$\vdots$	$\vdots$
10	21	59,049
20	41	3,486,784,401

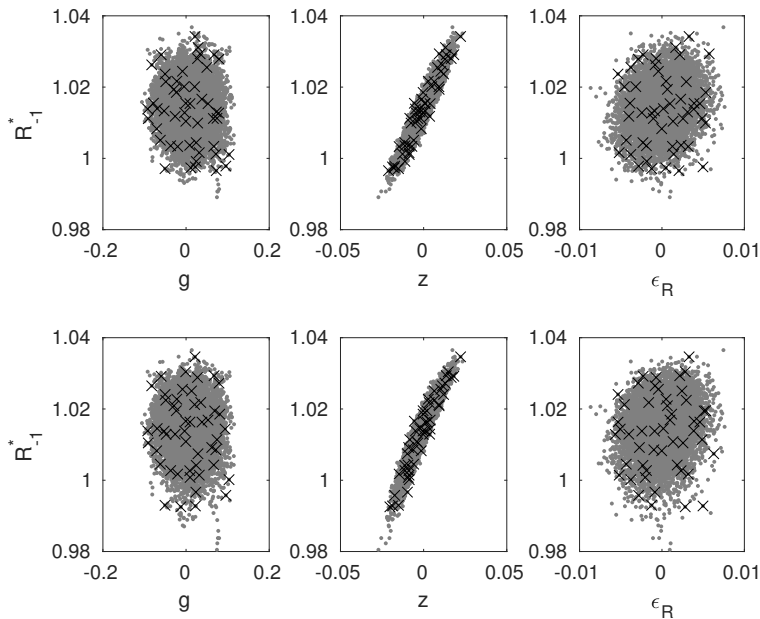
# Simulation-based method

- We solve for the policy functions on simulated grid points based on ergodic distribution of the state variables.
  - Judd, Maliar and Maliar (2011) and Maliar and Maliar (2015, MM hereafter) developed simulation-based method, based on the original work of Marcet's (1988) parameterized expectation algorithm.
  - Applications to New Keynesian models are found in Maliar and Maliar (2015), Lepetuyk, Maliar, and Maliar (2017), Aruoba, Cuba-borda, and Schorfheide (2018), and Hills, Nakata, and Sunakawa (2018).

# Constructing EDS grid

- In constructing an EDS grid from the ergodic set, we do the following two step procedure (See MM for more details):
  - 1 Selecting points within an essentially ergodic set (called Algorithm  $\mathcal{A}^\eta$  in MM)
  - 2 Constructing a uniformly spaced set of points that covers the essentially ergodic set (called Algorithm  $P^\epsilon$  in MM)

# Constructing EDS grid



# Numerical examples II

- Solve the nonlinear NK model by the time iteration method
  - Interpolation:
    - Non-stochastic method: Chebyshev polynomial with Smolyak sparse grid points (in the latter two).  $(N_d, N) = (3, 81), (3, 9), (5, 41)$ .
    - Simulation-based method: Second-order polynomials with cross terms. The number of grid points  $N = 25, 50$  or  $100$  EDS grid points. As we have 4 state variables, the number of coefficients is 15.
  - Optimization:
    - Chris Sims' csolve is used in Time iteration with Newton's method (TI).
    - No optimization is required in PEA collocation.
  - Integration:
    - Gaussian-Hermite quadrature is used with  $M = 3^3 = 27$  in TI and PEA collocation with fitting future variables (future PEA).
    - Precomputation technique in Judd et al. (2017) is used in PEA collocation with fitting current variables (current PEA, only with non-stochastic method).



# Non-stochastic method with ZLB

$(N_d, N)$	TI								
	$L_{1,c}$	$L_{1,\pi}$	$L_{\infty,c}$	$L_{\infty,\pi}$	$\sigma_{\Delta y}$	$\sigma_\pi$	$\sigma_R$	$\text{Pr}_{\text{ZLB}}$	CPU
(3, 81)	-3.73	-2.62	-2.07	-1.42	0.76	2.05	2.53	1.53	1127.3
(3, 9)	-3.40	-2.38	-2.06	-1.09	0.76	2.02	2.50	1.79	12.98
(5, 41)	-3.97	-3.14	-2.07	-1.73	0.76	2.04	2.50	1.40	270.65

$(N_d, N)$	future PEA								
	$L_{1,c}$	$L_{1,\pi}$	$L_{\infty,c}$	$L_{\infty,\pi}$	$\sigma_{\Delta y}$	$\sigma_\pi$	$\sigma_R$	$\text{Pr}_{\text{ZLB}}$	CPU
(3, 81)	-3.73	-2.67	-2.08	-1.44	0.76	2.11	2.59	1.71	82.28
(3, 9)	-3.26	-2.66	-1.92	-1.48	0.76	2.19	2.68	3.42	0.96
(5, 41)	-4.04	-3.54	-2.13	-1.49	0.76	2.03	2.48	1.18	14.66

$(N_d, N)$	current PEA								
	$L_{1,c}$	$L_{1,\pi}$	$L_{\infty,c}$	$L_{\infty,\pi}$	$\sigma_{\Delta y}$	$\sigma_\pi$	$\sigma_R$	$\text{Pr}_{\text{ZLB}}$	CPU
(3, 81)	-4.05	-2.71	-2.12	-1.53	0.76	2.01	2.45	1.03	4.05
(3, 9)	-3.35	-2.42	-1.97	-1.25	0.76	2.01	2.47	1.92	0.19
(5, 41)	-4.17	-2.95	-2.12	-1.44	0.76	2.03	2.47	1.07	0.99

Notes:  $L_{1,c}$ ,  $L_{1,\pi}$ ,  $L_{\infty,c}$ , and  $L_{\infty,\pi}$  are, respectively, the average and maximum of absolute Euler errors (in log 10 units) on a 10,000 period stochastic simulation. CPU is the elapsed time for computing equilibrium (in seconds).  $\sigma_{\Delta y}$ ,  $\sigma_\pi$ , and  $\sigma_R$  are the standard deviation of output growth, inflation, and the policy rate.  $\text{Pr}_{\text{ZLB}}$  is the probability of binding the ZLB (in percent).

# Simulation-based method with ZLB

$N$					TI				
	$L_{1,c}$	$L_{1,\pi}$	$L_{\infty,c}$	$L_{\infty,\pi}$	$\sigma_{\Delta y}$	$\sigma_{\pi}$	$\sigma_R$	$\text{Pr}_{\text{ZLB}}$	CPU
25	-3.39	-2.93	-2.21	-1.66	0.77	2.14	2.63	1.21	44.41
50	-3.00	-2.56	-1.83	-1.48	0.77	2.21	2.71	1.92	65.37
100	-3.28	-2.80	-2.08	-1.69	0.76	2.10	2.57	1.25	106.17

$N$	future PEA								
	$L_{1,c}$	$L_{1,\pi}$	$L_{\infty,c}$	$L_{\infty,\pi}$	$\sigma_{\Delta y}$	$\sigma_{\pi}$	$\sigma_R$	$\text{Pr}_{\text{ZLB}}$	CPU
25	-2.44	-1.24	-1.58	-0.61	0.77	2.18	2.68	1.31	3.37
50	-2.44	-1.24	-1.69	-0.63	0.77	2.15	2.63	1.46	5.08
100	-2.46	-1.26	-1.71	-0.67	0.76	2.11	2.58	1.31	7.93

Notes:  $L_{1,c}$ ,  $L_{1,\pi}$ ,  $L_{\infty,c}$ , and  $L_{\infty,\pi}$  are, respectively, the average and maximum of absolute Euler errors (in log 10 units) on a 10,000 period stochastic simulation. CPU is the elapsed time for computing equilibrium (in seconds).  $\sigma_{\Delta y}$ ,  $\sigma_{\pi}$ , and  $\sigma_R$  are the standard deviation of output growth, inflation, and the policy rate.  $\text{Pr}_{\text{ZLB}}$  is the probability of binding the ZLB (in percent).

## Parameter values

Parameter	Value
$\nu$ Inverse of demand elasticity	1/6
$\bar{g}$ Steady state government expenditure	1.25
$\gamma$ Steady state technology growth	1.0052
$\beta$ Discount factor	0.9990
$\bar{\pi}$ Steady state inflation	1.0083
$\tau$ CRRA parameter	2.83
$\phi$ Price adjustment cost	17.85
$\psi_1$ Interest rate elasticity to inflation	1.80
$\psi_2$ Interest rate elasticity to output gap	0.63
$\rho_r$ Interest rate smoothing	0.77
$\rho_g$ Persistence of government shock	0.98
$\rho_z$ Persistence of technology growth shock	0.88
$\sigma_r$ Std. dev. of monetary policy shock	0.0022
$\sigma_g$ Std. dev. of government shock	0.0071
$\sigma_z$ Std. dev. of technology growth shock	0.0031

Notes: Taken from Schorfheide and Herbst (2015). The observations used in the estimation range from 1983:I to 2002:IV, giving us a total of  $T = 80$  observations.

# index-function approach for the ZLB

- To deal with the ZLB, we adapt an index-function approach as in Aruoba et al. (2017), Gust et al. (2017), Nakata (2017), Hirose and Sunakawa (2017).
- Given  $\sigma_{\text{NZLB}}$  and  $\sigma_{\text{ZLB}}$ , use an index function to have

$$\begin{aligned}\sigma_c(R_{-1}^*, s_m) &= I_{(R^* > 1)} \sigma_{c, \text{NZLB}}(R_{-1}^*, s_m) + (1 - I_{(R^* > 1)}) \sigma_{c, \text{ZLB}}(R_{-1}^*, s_m), \\ \sigma_\pi(R_{-1}^*, s_m) &= I_{(R^* > 1)} \sigma_{\pi, \text{NZLB}}(R_{-1}^*, s_m) + (1 - I_{(R^* > 1)}) \sigma_{\pi, \text{ZLB}}(R_{-1}^*, s_m),\end{aligned}$$

where

$$I_{(R^* > 0)} = \begin{cases} 1 & \text{when } R^* = \sigma_{R^*, \text{NZLB}}(R_{-1}^*, s_m) > 1, \\ 0 & \text{otherwise.} \end{cases}$$

$\sigma_{\text{NZLB}}(R_{-1}^*, s_m)$  is the policy function assuming that ZLB *always* does not bind and  $\sigma_{\text{ZLB}}(R_{-1}^*, s_m)$  is the policy function assuming that ZLB *always* binds.

## 2. How to estimate nonlinear DSGE models

# State-Space Representation

- Linear case

- ① State transition equations:

$$s_t = \Phi_1(\theta)s_{t-1} + \Phi_\varepsilon(\theta)\varepsilon_t, \quad \varepsilon_t \sim N(0, \Sigma_\varepsilon)$$

- ② Observation equations:

$$y_t = \Psi_0(\theta) + \Psi_1(\theta)s_t + u_t, \quad u_t \sim N(0, \Sigma_u)$$

- Nonlinear case

- ① State transition equations:

$$s_t = \Phi(s_{t-1}, \varepsilon_t; \theta), \quad \varepsilon_t \sim F_\varepsilon(\cdot; \theta)$$

- ② Observation equations:

$$y_t = \Psi(s_t; \theta) + u_t, \quad u_t \sim F_u(\cdot; \theta)$$

# Kalman Filter

- In a linear case, the *Kalman filter* is available to evaluate likelihood.
- Distributional assumption about the initial state  $s_0$ :

$$s_0 \sim N(\bar{s}_{0|0}, P_{0|0})$$

- It is common to set  $\bar{s}_{0|0}$  and  $P_{0|0}$  equal to the unconditional first and second moments of implied by the law of motion of  $s_t$ .
- Given  $\bar{s}_{0|0}$  and  $P_{0|0}$ , the one-period-ahead forecasts of these moments are

$$\begin{aligned}\bar{s}_{1|0} &= \Phi_1 \bar{s}_{0|0}, \\ P_{1|0} &= \Phi_1 P_{0|0} \Phi_1' + \Phi_\varepsilon \Sigma_\varepsilon \Phi_\varepsilon' .\end{aligned}$$

- Given  $\bar{s}_{1|0}$  and  $P_{1|0}$ , the conditional mean and variances of the observables  $y_1$  are

$$\begin{aligned}\bar{y}_{1|0} &= A + B \bar{s}_{1|0}, \\ F_{1|0} &= B P_{1|0} B' .\end{aligned}$$

- Then, the forecast error of the observables  $y_1$  is

$$\nu_{1|0} = y_1 - \bar{y}_{1|0}.$$

- It has been known that the optimal updating leads to

$$\bar{s}_{1|1} = \bar{s}_{1|0} + P_{1|0}B'F_{1|0}^{-1}\nu_{1|0},$$

$$P_{1|1} = P_{1|0} - P_{1|0}B'F_{1|0}^{-1}BP_{1|0},$$

where  $\bar{s}_{1|1}$  and  $P_{1|1}$  are the mean and variances of  $s_1$  conditional on  $y_1$ .

- In the same way as above, given  $\bar{s}_{1|1}$  and  $P_{1|1}$ , we can obtain  $\bar{s}_{2|1}$ ,  $P_{2|1}$ ,  $\bar{y}_{2|1}$ ,  $F_{2|1}$ , and  $\nu_{2|1}$ .

- Also, updating gives  $\bar{s}_{2|2}$  and  $P_{2|2}$ , conditional on  $\{y_1, y_2\}$ .

- Repeating these steps yields the sequences of  $\bar{s}_{t|t-1}$ ,  $P_{t|t-1}$ ,  $\bar{y}_{t|t-1}$ ,  $F_{t|t-1}$ , and  $\nu_{t|t-1}$  conditional on  $\{y_1, y_2, \dots, y_{t-1}\}$  for  $t = 1, 2, \dots, T$ .



- Since  $\varepsilon_t \sim N(0, \Sigma_\varepsilon)$ ,

$$y_t|Y^{t-1} \sim N(A + B\hat{s}_{t|t-1}, F_{t|t-1}),$$

where  $Y^{t-1} = \{y_1, y_2, \dots, y_{t-1}\}$ .

- Its probability density function is

$$p(y_t|Y^{t-1}) = (2\pi)^{-\frac{n}{2}} |F_{t|t-1}|^{-\frac{1}{2}} \exp\left(-\frac{1}{2}\nu'_{t|t-1}F_{t|t-1}^{-1}\nu_{t|t-1}\right).$$

- Therefore, the log-likelihood is given by

$$\begin{aligned}\ln L(\theta|Y) &= \sum_{t=1}^T \ln p(y_t|Y^{t-1}) \\ &= -\frac{nT}{2} \ln 2\pi - \frac{1}{2} \sum_{t=1}^T \ln |F_{t|t-1}| - \frac{1}{2} \sum_{t=1}^T \nu'_{t|t-1} F_{t|t-1}^{-1} \nu_{t|t-1}.\end{aligned}$$

# Particle Filter

- In a nonlinear case, the *Kalman filter* is NOT available because the distribution of  $y_t|Y^{t-1}$  is non-normal.
- A particle filter can approximate the likelihood function.
- While there are many particle filters, we will focus on the *bootstrap particle filter*.
  - Gordon, Salmond, and Smith (1993)

# Bootstrap Particle Filter

- Idea: Particles representing  $s_t$  are propagated according to state space representation, so that the distribution of  $y_t|Y^{t-1}$  can be approximated.
- Draw the initial particles  $\{s_0^j\}_{j=1}^M$  from the distribution:

$$s_0^j \sim N(\bar{s}_0, P_0).$$

- $M$  denotes the number of particles.
- Set  $\bar{s}$  and  $P$  equal to the unconditional first and second moments of implied by the law of motion of  $s_t$ .
- Set particle weights  $W_0^j = 1$ , for  $j = 1, \dots, M$ .

For  $t = 1, \dots, T$ :

- 1 Draw shocks in period  $t$ :

$$\varepsilon_t^j \sim F_\varepsilon(\cdot; \theta),$$

and propagate particles  $\{s_{t-1}^j\}$  using the state-transition equation,

$$\tilde{s}_t^j = \Phi(s_{t-1}^j, \varepsilon_t; \theta).$$

- 2 Define the incremental weights:

$$\tilde{w}_t^j = p(y_t | \tilde{s}_t^j, \theta).$$

- The predictive density  $p(y_t | Y_{1:t-1}, \theta)$  can be approximated by

$$\hat{p}(y_t | Y_{1:t-1}, \theta) = \frac{1}{M} \sum_{j=1}^M \tilde{w}_t^j W_{t-1}^j.$$

- If the measurement errors  $u_t \sim N(0, \Sigma_u)$ , then the incremental weights are evaluated as

$$\tilde{w}_t^j = (2\pi)^{-\frac{n}{2}} |\Sigma_u|^{-\frac{1}{2}} \exp \left\{ -\frac{1}{2} (y_t - \Psi(\tilde{s}_t^j; \theta))' \Sigma_u^{-1} (y_t - \Psi(\tilde{s}_t^j; \theta)) \right\},$$

where  $n$  denotes the number of observables.

- 3 Define the normalized weights:

$$\tilde{W}_t^j = \frac{\tilde{w}_t^j W_{t-1}^j}{\frac{1}{M} \sum_{j=1}^M \tilde{w}_t^j W_{t-1}^j}$$

- An approximation of  $E[s_t | Y_{1:t}, \theta]$  is given by

$$\hat{E}[s_t | Y_{1:t}, \theta] = \frac{1}{M} \sum_{j=1}^M \tilde{s}_t^j \tilde{W}_t^j.$$

- 4 If resampling is needed,  $M$  iid draws  $\{s_t^j\}_{j=1}^M$  from a multinomial distribution characterized by support points and weights  $\{\tilde{s}_t^j, \tilde{W}_t^j\}$  and set  $W_t^j = 1$ , for  $j = 1, \dots, M$ .

- If resampling is not needed, let  $s_t^j = \tilde{s}_t^j$  and  $W_t^j = \tilde{W}_t^j$  for  $j = 1, \dots, M$ .

- 5 Repeat steps 1–4 for next  $t$ .

# Resampling

- A resampling step is necessary to avoid the degeneracy of the distribution of particle weights.
  - A situation in which all but a few of the weights are near zero
- Since resampling is done with replacement, a particle with a large weight is likely to be drawn many times and particles with small weights are not likely to be drawn at all.
  - Resampling effectively deals with the degeneracy problem by eliminating the particles with very small weights.
- Resampling is done whenever the effective sample size

$$\widehat{ESS}_t = \frac{M}{\frac{1}{M} \sum_{j=1}^M (\tilde{W}_t^j)^2}$$

falls below a threshold, e.g.,  $M/2$ .

# Likelihood Approximation

- Repeating steps 1–4 for  $t = 1, \dots, T$  gives particle weights  $\{\tilde{w}_t^j, W_{t-1}^j\}_{j=1}^M$  for each  $t$ .
- The approximation of the log-likelihood function is given by

$$\ln \hat{p}(Y_{1:T}|\theta) = \sum_{t=1}^T \ln \left( \frac{1}{M} \sum_{j=1}^M \tilde{w}_t^j W_{t-1}^j \right).$$

# Central Difference Kalman Filter

- Likelihood approximation using a particle filter causes a huge computational cost.
- If a DSGE model is approximated by a 2nd- or 3rd-order perturbation method, the Central Difference Kalman Filter (CDFK) can approximate the likelihood more efficiently.
  - Andreasen (2013)
- Idea: Approximate the filtering equations that compute and update first and second moments of state variables by 2nd-order multivariate Stirling interpolations.
- Andreasen (2013) show that a quasi maximum likelihood estimator based on the CDFK can be consistent and asymptotically normal for DSGE models solved up to third order.



# Bayesian Estimation

- Once we approximate and evaluate the likelihood function  $\hat{p}(Y|\theta)$ , the Bayesian likelihood approach is applicable.
  - Prior distribution:  $p(\theta)$
  - Bayes' Theorem  $\Rightarrow$  Posterior distribution

$$\hat{p}(\theta|Y) \propto \hat{p}(Y|\theta)p(\theta)$$

- Generate draws from the posterior distribution using Markov Chain Monte Carlo (MCMC) algorithm.
  - Random-Walk Metropolis Hasting (RWMH) algorithm is widely used.

# Sequential Monte Carlo Algorithm

- Issues in the RWMH algorithm:
  - The posterior distribution is possibly multimodal.
  - The RWMH algorithm can get stuck near a local mode and fail to find the entire posterior distribution.
  - It is often very difficult to find a model.
- Herbst and Schorfheide (2014, 2015) propose the Sequential Monte Carlo (SMC) algorithm.
  - Particles representing  $\theta$  are propagated, similar to a particle filter.
  - Overcome the issues by building a particle approximation to the posterior gradually through tempering the likelihood function.
  - Sequence of tempered posteriors:

$$\pi_n(\theta) = \frac{[p(Y|\theta)]^{\phi_n} p(\theta)}{\int [p(Y|\theta)]^{\phi_n} p(\theta) d\theta}, \quad n = 0, \dots, N_\phi.$$

- Tempering schedule:  $\phi_n = (n/N_\phi)^\chi$

# Parallelization

- Both the particle filter and the SMC algorithm can be parallelized.
  - Nonlinear solution methods can be also parallelized.
  - Massive speed gains in estimation
- Matlab parallel toolbox is very easy to use, but not so fast.
- Explicit parallelization:
  - OpenMP
  - MPI
  - GPU programming: CUDA and OpenCL
- Hardware:
  - Workstation with multi-core and multi-processors
  - Computer cluster
  - Cloud computing

# Faster Programming Languages

- Aruoba and Fernández-Villaverde (2015): “A Comparison of Programming Languages in Macroeconomics,” *Journal of Economic Dynamics and Control*, 58, 265–273.
  - Solve the stochastic neoclassical growth model using C++, Fortran, Java, Julia, Python, Matlab, Mathematica, and R.
  - Report the execution time of the codes.
- Aruoba and Fernández-Villaverde (2018) update results for new versions of each language.

## Aruoba and Fernández-Villaverde (2018)

Language	Compiler	Time	Rel. Time
C++	GCC	1.60	1.00
	Intel C++	1.67	1.04
	Clang	1.64	1.03
Fortran	GCC	1.61	1.01
	Intel Fortran	1.74	1.09
Java		3.20	2.00
Julia		2.35	1.47
	fast	2.14	1.34
Matlab		4.80	3.00
Python	CPython	145.27	90.79
	CPython	166.75	104.22
R		57.06	35.66
Mathematica	base	1634.94	1021.84

## Aruoba and Fernández-Villaverde (2018)

Language	Compiler	Time	Rel. Time
Matlab, Mex		2.01	1.26
Rcpp		6.60	4.13
Python	Numba	2.31	1.44
Cython	Cython	2.13	1.33
Mathematica	idiomatic	4.42	2.76

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