

Chapter 5: Matrix approach to simple linear regression analysis

Need to understand matrix algebra for multiple regression.

5.1 Matrices

What is a matrix?

A matrix is a rectangular array of elements arranged in rows and columns.

Example:

$$\begin{bmatrix} 1 & 2 & 3 \\ 4 & 5 & 6 \end{bmatrix}$$

Dimension – Size of matrix: # rows \times # columns = $r \times c$

Example: 2×3

Symbolic representation of a matrix:

Example:

$$\mathbf{A} = \begin{bmatrix} a_{11} & a_{12} & a_{13} \\ a_{21} & a_{22} & a_{23} \end{bmatrix}$$

where a_{ij} is the row i and column j element of \mathbf{A}

$a_{11}=1$ from the above example

Notice that the matrix \mathbf{A} is in bold.

a_{11} is often called the “(1,1) element” of \mathbf{A} , a_{12} is called the “(1,2) element” of \mathbf{A} ,...

Example: $r \times c$ matrix

$$\mathbf{A} = \begin{bmatrix} a_{11} & a_{12} & \cdots & a_{1j} & \cdots & a_{1c} \\ a_{21} & a_{22} & \cdots & a_{2j} & \cdots & a_{2c} \\ \vdots & \vdots & & \vdots & & \vdots \\ a_{i1} & a_{i2} & \cdots & a_{ij} & \cdots & a_{ic} \\ \vdots & \vdots & & \vdots & & \vdots \\ a_{r1} & a_{r2} & \cdots & a_{rj} & \cdots & a_{rc} \end{bmatrix}$$

Example: Square matrix is $r \times c$ where $r=c$

Vector – a $r \times 1$ (column vector) or $1 \times c$ (row vector) matrix – special case of a matrix

Example: Symbolic representation of a 3×1 column vector

$$\mathbf{A} = \begin{bmatrix} a_1 \\ a_2 \\ a_3 \end{bmatrix}$$

Example:

$$\mathbf{Y} = \begin{bmatrix} 3.10 \\ 2.30 \\ 3.00 \\ \vdots \\ 2.20 \\ 1.60 \end{bmatrix}$$

Transpose: Interchange the rows and columns of a matrix or vector

Example:

$$\mathbf{A} = \begin{bmatrix} a_{11} & a_{12} & a_{13} \\ a_{21} & a_{22} & a_{23} \end{bmatrix} \text{ and } \mathbf{A}' = \begin{bmatrix} a_{11} & a_{21} \\ a_{12} & a_{22} \\ a_{13} & a_{23} \end{bmatrix}$$

\mathbf{A} is 2×3 and \mathbf{A}' is 3×2

Example:

$$\mathbf{Y}' = [3.10 \quad 2.30 \quad 3.00 \quad \dots \quad 2.20 \quad 1.60]$$

Equality of matrices – Two matrices are equal if all of their elements are equal.

5.2 Matrix addition and subtraction

Add or subtract the corresponding elements of matrices with the same dimension.

Example:

$$\begin{aligned} \text{Suppose } \mathbf{A} &= \begin{bmatrix} 1 & 2 & 3 \\ 4 & 5 & 6 \end{bmatrix} \text{ and } \mathbf{B} = \begin{bmatrix} -1 & 10 & -1 \\ 5 & 5 & 8 \end{bmatrix}. \text{ Then} \\ \mathbf{A} + \mathbf{B} &= \begin{bmatrix} 0 & 12 & 2 \\ 9 & 10 & 14 \end{bmatrix} \text{ and } \mathbf{A} - \mathbf{B} = \begin{bmatrix} 2 & -8 & 4 \\ -1 & 0 & -2 \end{bmatrix}. \end{aligned}$$

Example: Simple linear regression model

$Y_i = E(Y_i) + \varepsilon_i$ for $i=1, \dots, n$ can be represented as

$\mathbf{Y} = \mathbf{E}(\mathbf{Y}) + \boldsymbol{\varepsilon}$ where

$$\mathbf{Y} = \begin{bmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_n \end{bmatrix}, \mathbf{E}(\mathbf{Y}) = \begin{bmatrix} E(Y_1) \\ E(Y_2) \\ \vdots \\ E(Y_n) \end{bmatrix}, \text{ and } \boldsymbol{\varepsilon} = \begin{bmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \vdots \\ \varepsilon_n \end{bmatrix}$$

5.3 Matrix multiplication

Scalar - 1×1 matrix

Example: Matrix multiplied by a scalar

$$c\mathbf{A} = \begin{bmatrix} ca_{11} & ca_{12} & ca_{13} \\ ca_{21} & ca_{22} & ca_{23} \end{bmatrix} \text{ where } c \text{ is a scalar}$$

$$\text{Let } \mathbf{A} = \begin{bmatrix} 1 & 2 & 3 \\ 4 & 5 & 6 \end{bmatrix} \text{ and } c=2. \text{ Then } 2\mathbf{A} = \begin{bmatrix} 2 & 4 & 6 \\ 8 & 10 & 12 \end{bmatrix}.$$

Multiplying two matrices

Suppose you want to multiply the matrices \mathbf{A} and \mathbf{B} ; i.e., $\mathbf{A} * \mathbf{B}$ or \mathbf{AB} . In order to do this, you need the number of columns of \mathbf{A} to be the same as the number of rows as \mathbf{B} . For example, suppose \mathbf{A} is 2×3 and \mathbf{B} is 3×10 . You can multiply these matrices. However if \mathbf{B} is 4×10 instead, these matrices could NOT be multiplied.

The resulting dimension of $\mathbf{C} = \mathbf{AB}$

1. The number of rows of \mathbf{A} is the number of rows of \mathbf{C} .
2. The number of columns of \mathbf{B} is the number of columns of \mathbf{C} .
3. In other words, $\mathbf{C} = \mathbf{A} \mathbf{B}$ where the dimension of the

matrices are shown below them.

How to multiply two matrices – an example

$$\text{Suppose } \mathbf{A} = \begin{bmatrix} 1 & 2 & 3 \\ 4 & 5 & 6 \end{bmatrix} \text{ and } \mathbf{B} = \begin{bmatrix} 3 & 0 \\ 1 & 2 \\ 0 & 1 \end{bmatrix}. \text{ Notice that } \mathbf{A} \text{ is } 2 \times 3 \text{ and } \mathbf{B} \text{ is}$$

3×2 so $\mathbf{C} = \mathbf{AB}$ can be done.

$$\begin{aligned}
\mathbf{C} = \mathbf{AB} &= \begin{bmatrix} 1 & 2 & 3 \\ 4 & 5 & 6 \end{bmatrix} \begin{bmatrix} 3 & 0 \\ 1 & 2 \\ 0 & 1 \end{bmatrix} \\
&= \begin{bmatrix} 1*3+2*1+3*0 & 1*0+2*2+3*1 \\ 4*3+5*1+6*0 & 4*0+5*2+6*1 \end{bmatrix} \\
&= \begin{bmatrix} 5 & 7 \\ 17 & 16 \end{bmatrix}
\end{aligned}$$

The “cross product” of the rows of **A** and the columns of **B** are taken to form **C**

In the above example, **D=BA**≠**AB** where **BA** is:

$$\begin{aligned}
\mathbf{BA} &= \begin{bmatrix} 3 & 0 \\ 1 & 2 \\ 0 & 1 \end{bmatrix} \begin{bmatrix} 1 & 2 & 3 \\ 4 & 5 & 6 \end{bmatrix} \\
&= \begin{bmatrix} 3*1+0*4 & 3*2+0*5 & 3*3+0*6 \\ 1*1+2*4 & 1*2+2*5 & 1*3+2*6 \\ 0*1+1*4 & 0*2+1*5 & 0*3+1*6 \end{bmatrix} \\
&= \begin{bmatrix} 3 & 6 & 9 \\ 9 & 12 & 15 \\ 4 & 5 & 6 \end{bmatrix}
\end{aligned}$$

In general for a 2×3 matrix times a 3×2 matrix:

$$\begin{aligned}
\mathbf{C} = \mathbf{AB} &= \begin{bmatrix} a_{11} & a_{12} & a_{13} \\ a_{21} & a_{22} & a_{23} \end{bmatrix} \begin{bmatrix} b_{11} & b_{12} \\ b_{21} & b_{22} \\ b_{31} & b_{32} \end{bmatrix} \\
&= \begin{bmatrix} a_{11}b_{11} + a_{12}b_{21} + a_{13}b_{31} & a_{11}b_{12} + a_{12}b_{22} + a_{13}b_{32} \\ a_{21}b_{11} + a_{22}b_{21} + a_{23}b_{31} & a_{21}b_{12} + a_{22}b_{22} + a_{23}b_{32} \end{bmatrix}
\end{aligned}$$

Example:

$$\mathbf{X} = \begin{bmatrix} 1 & 3.04 \\ 1 & 2.35 \\ 1 & 2.70 \\ \vdots & \vdots \\ 1 & 2.28 \\ 1 & 1.88 \end{bmatrix} \quad \mathbf{Y} = \begin{bmatrix} 3.10 \\ 2.30 \\ 3.00 \\ \vdots \\ 2.20 \\ 1.60 \end{bmatrix}$$

Find $\mathbf{X}'\mathbf{X}$, $\mathbf{X}'\mathbf{Y}$, and $\mathbf{Y}'\mathbf{Y}$

Notes:

$$1. \quad \mathbf{Y}'\mathbf{Y} = [y_1, y_2, \dots, y_n] \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{bmatrix} = \sum_{i=1}^n y_i^2$$

$$2. \quad \mathbf{X}'\mathbf{Y} = \begin{bmatrix} x_{11} & x_{21} & \cdots & x_{n1} \\ x_{12} & x_{22} & \cdots & x_{n2} \end{bmatrix} \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{bmatrix} = \begin{bmatrix} \sum_{i=1}^n x_{i1} y_i \\ \sum_{i=1}^n x_{i2} y_i \end{bmatrix} = \begin{bmatrix} \sum_{i=1}^n y_i \\ \sum_{i=1}^n x_{i2} y_i \end{bmatrix} \text{ since } x_{11} = \dots = x_{n1} = 1$$

3. $\mathbf{X}'\mathbf{X}$

$$= \begin{bmatrix} x_{11} & x_{21} & \cdots & x_{n1} \\ x_{12} & x_{22} & \cdots & x_{n2} \end{bmatrix} \begin{bmatrix} x_{11} & x_{12} \\ x_{21} & x_{22} \\ \vdots & \vdots \\ x_{n1} & x_{n2} \end{bmatrix}$$

$$= \begin{bmatrix} 1 & 1 & \cdots & 1 \\ x_{12} & x_{22} & \cdots & x_{n2} \end{bmatrix} \begin{bmatrix} 1 & x_{12} \\ 1 & x_{22} \\ \vdots & \vdots \\ 1 & x_{n2} \end{bmatrix} = \begin{bmatrix} n & \sum_{i=1}^n x_{i2} \\ \sum_{i=1}^n x_{i2} & \sum_{i=1}^n x_{i2}^2 \end{bmatrix}$$

5.4 Special types of matrices

Symmetric matrix: If $\mathbf{A} = \mathbf{A}'$, then \mathbf{A} is symmetric.

$$\text{Example: } \mathbf{A} = \begin{bmatrix} 1 & 2 \\ 2 & 3 \end{bmatrix}, \mathbf{A}' = \begin{bmatrix} 1 & 2 \\ 2 & 3 \end{bmatrix}$$

Diagonal matrix: A square matrix whose “off-diagonal” elements are 0.

$$\text{Example: } \mathbf{A} = \begin{bmatrix} a_{11} & 0 & 0 \\ 0 & a_{22} & 0 \\ 0 & 0 & a_{33} \end{bmatrix}$$

Identity matrix: A diagonal matrix with 1's on the diagonal.

Example:

$$\mathbf{I} = \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{bmatrix}$$

Note that “**I**” (the letter **I**, not the number one) usually denotes the identity matrix.

Vector and matrix of 1's

A column vector of 1's:

$$\mathbf{1}_{r \times 1} = \mathbf{j}_{r \times 1} = \begin{bmatrix} 1 \\ 1 \\ \vdots \\ 1 \end{bmatrix}$$

A matrix of 1's:

$$\mathbf{J}_{r \times r} = \begin{bmatrix} 1 & 1 & \cdots & 1 \\ 1 & 1 & \cdots & 1 \\ \vdots & \vdots & \ddots & \vdots \\ 1 & 1 & \cdots & 1 \end{bmatrix}$$

Notes:

1. $\mathbf{j}'_{r \times 1} \mathbf{j}_{r \times 1} = r$
2. $\mathbf{j}_{r \times 1} \mathbf{j}'_{r \times 1} = \mathbf{J}_{r \times r}$
3. $\mathbf{J}'_{r \times r} \mathbf{J}_{r \times r} = r \mathbf{J}_{r \times r}$

Vector of 0's:

$$\mathbf{0}_{r \times 1} = \begin{bmatrix} 0 \\ 0 \\ \vdots \\ 0 \end{bmatrix}$$

5.5 Linear dependence and rank of matrix

Let $\mathbf{A} = \begin{bmatrix} 1 & 2 & 6 \\ 3 & 4 & 12 \\ 5 & 6 & 18 \end{bmatrix}$. Think of each column of **A** as a vector; i.e., $\mathbf{A} = [\mathbf{A}_1, \mathbf{A}_2, \mathbf{A}_3]$.

Note that $3\mathbf{A}_2 = \mathbf{A}_3$. This means the columns of **A** are “linearly dependent.”

Formally, a set of column vectors are **linearly dependent** if there exists constants $\lambda_1, \lambda_2, \dots, \lambda_n$ (not all zero) such that $\lambda_1\mathbf{A}_1 + \lambda_2\mathbf{A}_2 + \dots + \lambda_n\mathbf{A}_n = \mathbf{0}$. A set of column vectors are **linearly independent** if $\lambda_1\mathbf{A}_1 + \lambda_2\mathbf{A}_2 + \dots + \lambda_n\mathbf{A}_n = \mathbf{0}$ only for $\lambda_1 = \lambda_2 = \dots = \lambda_n = 0$.

The rank of a matrix is the maximum number of linearly independent columns in the matrix.

$$\text{rank}(\mathbf{A}) = 2$$

5.6 Inverse of a matrix

Note that the inverse of a scalar, say b , is b^{-1} . For example, the inverse of $b=3$ is $3^{-1}=1/3$. Also, $b*b^{-1}=1$. In matrix algebra, the inverse of a matrix is another matrix. For example, the inverse of \mathbf{A} is \mathbf{A}^{-1} , and $\mathbf{A}*\mathbf{A}^{-1}=\mathbf{A}^{-1}*\mathbf{A}=\mathbf{I}$. Note that \mathbf{A} must be a square matrix.

Example: $\mathbf{A} = \begin{bmatrix} 1 & 2 \\ 3 & 4 \end{bmatrix}$, $\mathbf{A}^{-1} = \begin{bmatrix} -2 & 1 \\ 1.5 & -0.5 \end{bmatrix}$

Check: $\mathbf{A}\mathbf{A}^{-1} = \begin{bmatrix} 1*(-2) + 2*1.5 & 1*1 + 2*(-0.5) \\ 3*(-2) + 4*1.5 & 3*1 + 4*(-0.5) \end{bmatrix} = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$

Finding the inverse

For a general way, see a matrix algebra book (such as: Kolman, 1988). For a 2×2 matrix, there is a simple formula. Let $\mathbf{A} = \begin{bmatrix} a_{11} & a_{12} \\ a_{21} & a_{22} \end{bmatrix}$.

Then $\mathbf{A}^{-1} = \frac{1}{a_{11}a_{22} - a_{12}a_{21}} \begin{bmatrix} a_{22} & -a_{12} \\ -a_{21} & a_{11} \end{bmatrix}$.

Verify $\mathbf{A}\mathbf{A}^{-1}=\mathbf{I}$ on your own.

Example: College and HS GPA (GPA_example5.sas)

Remember that $\mathbf{X} = \begin{bmatrix} 1 & 3.04 \\ 1 & 2.35 \\ 1 & 2.70 \\ \vdots & \vdots \\ 1 & 2.28 \\ 1 & 1.88 \end{bmatrix}$ and $\mathbf{Y} = \begin{bmatrix} 3.10 \\ 2.30 \\ 3.00 \\ \vdots \\ 2.20 \\ 1.60 \end{bmatrix}$

Find $(\mathbf{X}'\mathbf{X})^{-1}$ and $(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y}$

Note that $(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y} = \mathbf{b} = \begin{bmatrix} b_0 \\ b_1 \end{bmatrix}$!!!

Read “Uses of Inverse Matrix” on p. 192-193 of KNN.

Look at some R codes: Chapter5_Ch6_PR05.R

5.7 Some basic theorems of matrices (Page 193)

1. $A + B = B + A$
2. $(A + B) + C = A + (B + C)$
3. $(AB)C = A(BC)$
4. $C(A + B) = CA + CB$
5. $\lambda(A + B) = \lambda A + \lambda B$
6. $(A')' = A$
7. $(A + B)' = A' + B'$
8. $(AB)' = B'A'$
9. $(ABC)' = C'B'A'$
10. $(AB)^{-1} = B^{-1}A^{-1}$
11. $(ABC)^{-1} = C^{-1}B^{-1}A^{-1}$
12. $(A^{-1})^{-1} = A$
13. $(A')^{-1} = (A^{-1})'$

5.8 Random vectors and matrices

A random vector or random matrix contains elements that are random variables.

Example: Simple linear regression model

$Y_i = E(Y_i) + \varepsilon_i$ for $i=1, \dots, n$ can be represented as

$Y = E(Y) + \varepsilon$ where

$$Y = \begin{bmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_n \end{bmatrix} \text{ and } \varepsilon = \begin{bmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \vdots \\ \varepsilon_n \end{bmatrix} \text{ are random vectors.}$$

Expectation of random vector or matrix: Find the expected value of the individual elements

Example:

$$E(Y) = E\left(\begin{bmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_n \end{bmatrix}\right) = \begin{bmatrix} E(Y_1) \\ E(Y_2) \\ \vdots \\ E(Y_n) \end{bmatrix} = \begin{bmatrix} \beta_0 + \beta_1 X_1 \\ \beta_0 + \beta_1 X_2 \\ \vdots \\ \beta_0 + \beta_1 X_n \end{bmatrix}$$

$$\mathbf{E}(\boldsymbol{\varepsilon}) = \mathbf{E} \begin{pmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \vdots \\ \varepsilon_n \end{pmatrix} = \begin{bmatrix} \mathbf{E}(\varepsilon_1) \\ \mathbf{E}(\varepsilon_2) \\ \vdots \\ \mathbf{E}(\varepsilon_n) \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \\ \vdots \\ 0 \end{bmatrix} \text{ since we assume } \varepsilon_i \sim N(0, \sigma^2)$$

Variance-covariance matrix of a random vector

Let Z_1 and Z_2 be random variables. Remember that $\sigma^2\{Z_1\} = \mathbf{E}(Z_1 - \mu_1)^2$ where $\mathbf{E}(Z_1) = \mu_1$.

The covariance of Z_1 and Z_2 is defined as $s\{Z_1, Z_2\} = \mathbf{E}[(Z_1 - \mu_1)(Z_2 - \mu_2)]$.
The covariance measures the relationship between Z_1 and Z_2 .

Notes:

- $\sigma\{Z_1, Z_1\} = \mathbf{E}[(Z_1 - \mu_1)(Z_1 - \mu_1)] = \mathbf{E}(Z_1 - \mu_1)^2 = \sigma^2\{Z_1\}$
- If Z_1 and Z_2 are independent, then $\sigma\{Z_1, Z_2\} = 0$.

A variance-covariance matrix (most often just called the covariance matrix) is a matrix whose elements are the variances and covariances of random variables.

Example: Simple linear regression model

\mathbf{Y} is a $n \times 1$ random vector

$$\sigma^2\{\mathbf{Y}\} = \begin{bmatrix} \sigma^2\{Y_1\} & \sigma\{Y_1, Y_2\} & \cdots & \sigma\{Y_1, Y_n\} \\ \sigma\{Y_2, Y_1\} & \sigma^2\{Y_2\} & \cdots & \sigma\{Y_2, Y_n\} \\ \vdots & \vdots & \ddots & \vdots \\ \sigma\{Y_n, Y_1\} & \sigma\{Y_n, Y_2\} & \cdots & \sigma^2\{Y_n\} \end{bmatrix}$$

$$\sigma^2\{\boldsymbol{\varepsilon}\} = \begin{bmatrix} \sigma^2 & 0 & \cdots & 0 \\ 0 & \sigma^2 & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & \sigma^2 \end{bmatrix} = \sigma^2 \mathbf{I}$$

Note that all covariance matrices are symmetric!

Some basic theorems

Let $\mathbf{W} = \mathbf{A}\mathbf{Y}$ where \mathbf{Y} is a random vector and \mathbf{A} is a matrix of constants (no random variables in it).

The following results follow:

1. $E(\mathbf{A}) = \mathbf{A}$

Remember this is like saying $E(3) = 3$

2. $E(\mathbf{W}) = E(\mathbf{A}\mathbf{Y}) = \mathbf{A}E(\mathbf{Y})$

Again, this is like saying $E(3Y_1) = 3E(Y_1)$

$$\sigma^2\{\mathbf{W}\} = \sigma^2(\mathbf{A}\mathbf{Y}) = \mathbf{A} \sigma^2\{\mathbf{Y}\} \mathbf{A}'$$

You may have seen before that $\sigma^2(aY_1) = a^2 \sigma^2\{Y_1\}$ where a is a constant.

Example:

Let $\mathbf{A} = \begin{bmatrix} 1 & -1 \\ 1 & 0 \end{bmatrix}$ and $\mathbf{Y} = \begin{bmatrix} Y_1 \\ Y_2 \end{bmatrix}$. Let $\mathbf{W} = \mathbf{A}\mathbf{Y}$, then $\mathbf{W} = \begin{bmatrix} Y_1 - Y_2 \\ Y_1 \end{bmatrix}$.

What is $E(\mathbf{W})$ and $\sigma^2\{\mathbf{W}\}$?

5.9 Simple linear regression model in matrix terms

$$Y_i = \beta_0 + \beta_1 X_i + \varepsilon_i$$

where $\varepsilon_i \sim \text{independent } N(0, \sigma^2)$ and $i = 1, \dots, n$

The model can be rewritten as

$$Y_1 = \beta_0 + \beta_1 X_1 + \varepsilon_1$$

$$Y_2 = \beta_0 + \beta_1 X_2 + \varepsilon_2$$

$$Y_3 = \beta_0 + \beta_1 X_3 + \varepsilon_3$$

$$\vdots$$

$$Y_n = \beta_0 + \beta_1 X_n + \varepsilon_n$$

In matrix terms, let

$$\mathbf{Y} = \begin{bmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_n \end{bmatrix}, \mathbf{X} = \begin{bmatrix} 1 & X_1 \\ 1 & X_2 \\ \vdots & \vdots \\ 1 & X_n \end{bmatrix}, \boldsymbol{\beta} = \begin{bmatrix} \beta_0 \\ \beta_1 \end{bmatrix} \text{ and } \boldsymbol{\varepsilon} = \begin{bmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \vdots \\ \varepsilon_n \end{bmatrix}$$

Then $\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon}$, which is

$$\begin{bmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_n \end{bmatrix} = \begin{bmatrix} 1 & X_1 \\ 1 & X_2 \\ \vdots & \vdots \\ 1 & X_n \end{bmatrix} \begin{bmatrix} \beta_0 \\ \beta_1 \end{bmatrix} + \begin{bmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \vdots \\ \varepsilon_n \end{bmatrix} = \begin{bmatrix} \beta_0 + X_1\beta_1 + \varepsilon_1 \\ \beta_0 + X_2\beta_1 + \varepsilon_2 \\ \vdots \\ \beta_0 + X_n\beta_1 + \varepsilon_n \end{bmatrix}$$

Note that $E(\mathbf{Y}) = E(\mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon}) = E(\mathbf{X}\boldsymbol{\beta}) = \mathbf{X}\boldsymbol{\beta}$ since $E(\boldsymbol{\varepsilon}) = \mathbf{0}$

5.10 Least squares estimation of regression parameters

From Chapter 1: The least squares method tries to find the b_0 and b_1 such that $SSE = \sum (Y_i - \hat{Y}_i)^2 = \sum (\text{residual})^2$ is minimized. Formulas for b_0 and b_1 are derived using calculus.

It can be shown that b_0 and b_1 can be found from solving the “normal equations”:

$$-2 \sum_{i=1}^n (Y_i - b_0 - b_1 X_i) = 0 \Rightarrow \sum_{i=1}^n Y_i = nb_0 + b_1 \sum_{i=1}^n X_i \quad (1a)$$

$$-2 \sum_{i=1}^n (Y_i - b_0 - b_1 X_i) X_i = 0 \Rightarrow \sum_{i=1}^n X_i Y_i = b_0 \sum_{i=1}^n X_i + b_1 \sum_{i=1}^n X_i^2 \quad (2a)$$

The normal equations can be rewritten as $\mathbf{X}'\mathbf{X}\mathbf{b} = \mathbf{X}'\mathbf{Y}$, where $\mathbf{b}_{2 \times 1} = \begin{bmatrix} b_0 \\ b_1 \end{bmatrix}$. Recall in Section 5.3, it was shown that

$$\mathbf{X}'\mathbf{Y} = \begin{bmatrix} 1 & 1 & \dots & 1 \\ X_1 & X_2 & \dots & X_n \end{bmatrix} \begin{bmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_n \end{bmatrix} = \begin{bmatrix} \sum_{i=1}^n Y_i \\ \sum_{i=1}^n X_i Y_i \end{bmatrix} \text{ and}$$

$$\mathbf{X}'\mathbf{X} = \begin{bmatrix} 1 & 1 & \dots & 1 \\ X_1 & X_2 & \dots & X_n \end{bmatrix} \begin{bmatrix} 1 & X_1 \\ 1 & X_2 \\ \vdots & \vdots \\ 1 & X_n \end{bmatrix} = \begin{bmatrix} n & \sum_{i=1}^n X_i \\ \sum_{i=1}^n X_i & \sum_{i=1}^n X_i^2 \end{bmatrix}.$$

Therefore, we have a way to find \mathbf{b} using matrix algebra:

$$\mathbf{b}_{s \times 1} = (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\mathbf{Y}$$

Example: Toluca example (Page 200).

5.11 Fitted values and residuals

Let $\hat{\mathbf{Y}} = \begin{bmatrix} \hat{Y}_1 \\ \vdots \\ \hat{Y}_n \end{bmatrix}$, Then $\hat{\mathbf{Y}} = \mathbf{X}\mathbf{b}$.

Hat matrix: $\hat{\mathbf{Y}} = \mathbf{X}\mathbf{b} = \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y} = \mathbf{H}\mathbf{Y}$, where $\mathbf{H} = \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'$ is the hat matrix.

The matrix is symmetric and has the special property (called idempotency): $\mathbf{H}\mathbf{H} = \mathbf{H}$.

We will use this in Chapter 9 to measure the influence of observations on the estimated regression line.

Residuals:

Let $\mathbf{e} = \begin{bmatrix} e_1 \\ \vdots \\ e_n \end{bmatrix} = \begin{bmatrix} Y_1 - \hat{Y}_1 \\ \vdots \\ Y_n - \hat{Y}_n \end{bmatrix}$, then

$$\mathbf{e} = \mathbf{Y} - \hat{\mathbf{Y}} = \mathbf{Y} - \mathbf{X}\mathbf{b} = \mathbf{Y} - \mathbf{H}\mathbf{Y} = (\mathbf{I} - \mathbf{H})\mathbf{Y}.$$

Covariance matrix of the residuals:

$\sigma^2\{\mathbf{e}\} = \sigma^2(\mathbf{I} - \mathbf{H})$ (see p. 204 for derivation) and is estimated by

$$\mathbf{s}^2\{\mathbf{e}\} = MSE(\mathbf{I} - \mathbf{H}).$$

5.12 Analysis of variance results

Sums of Squares

From Chapter 2: $SSTO = \sum_{i=1}^n (Y_i - \bar{Y})^2$, $SSE = \sum_{i=1}^n (Y_i - \hat{Y}_i)^2$, and

$$SSR = SSTO - SSE = \sum_{i=1}^n (\hat{Y}_i - \bar{Y})^2$$

These can be rewritten using matrices:

$$\begin{aligned}
SSTO &= \mathbf{Y}'\mathbf{Y} - \frac{1}{n}\mathbf{Y}'\mathbf{J}\mathbf{Y} \\
&= \begin{bmatrix} Y_1 & \dots & Y_n \end{bmatrix} \begin{bmatrix} Y_1 \\ \vdots \\ Y_n \end{bmatrix} - \frac{1}{n} \begin{bmatrix} Y_1 & \dots & Y_n \end{bmatrix} \begin{bmatrix} 1 & \dots & 1 \\ \vdots & \ddots & \vdots \\ 1 & \dots & 1 \end{bmatrix} \begin{bmatrix} Y_1 \\ \vdots \\ Y_n \end{bmatrix} \\
&= \sum_{i=1}^n Y_i^2 - \frac{1}{n} \left[\sum_{i=1}^n Y_i \quad \dots \quad \sum_{i=1}^n Y_i \right] \begin{bmatrix} Y_1 \\ \vdots \\ Y_n \end{bmatrix} \\
&= \sum_{i=1}^n Y_i^2 - \frac{1}{n} \left(\sum_{i=1}^n Y_i \right)^2 \\
&= \sum_{i=1}^n (Y_i - \bar{Y})^2
\end{aligned}$$

$$SSE = \mathbf{e}'\mathbf{e} = (\mathbf{Y} - \mathbf{X}\mathbf{b})'(\mathbf{Y} - \mathbf{X}\mathbf{b}) = \mathbf{Y}'\mathbf{Y} - \mathbf{b}'\mathbf{X}'\mathbf{Y}$$

$$SSR = \mathbf{b}'\mathbf{X}'\mathbf{Y} - \left(\frac{1}{n}\right)\mathbf{Y}'\mathbf{J}\mathbf{Y}$$

5.13 Inferences in Regression Analysis

Covariance matrix of \mathbf{b} :

$$\sigma^2\{\mathbf{b}\} = \sigma^2\{(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y}\} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\sigma^2\{\mathbf{Y}\}[(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}']' = \sigma^2(\mathbf{X}'\mathbf{X})^{-1}$$

Estimated covariance matrix of \mathbf{b} :

$$\sigma^2\{\mathbf{b}\} = MSE(\mathbf{X}'\mathbf{X})^{-1}$$

Estimated variance used in the C.I. for $E(Y_h)$ (mean response):

Define the vector

$$\mathbf{X}_h = \begin{bmatrix} 1 \\ X_h \end{bmatrix},$$

The fitted value in matrix notation is $\hat{Y}_h = \mathbf{X}_h'\mathbf{b}$, then

$$\sigma^2\{\hat{Y}_h\} = \sigma^2(\mathbf{X}_h'(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}_h), \quad s^2\{\hat{Y}_h\} = MSE(\mathbf{X}_h'(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}_h)$$

Estimated variance used in the P.I. for $Y_{h(\text{new})}$ (predcition of new observation):

$$s^2\{\text{pred}\} = MSE(1 + \mathbf{X}_h'(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}_h)$$

R tips: Remember that the C.I. and P.I. can also be found using the built in functions in R. Also, the estimated covariance matrix of **b** can be found using the VCOV function.

```
#####  
## You can use some built-in R functions  
Data <-  
read.table(file="C:/Hongmei/Teaching/stat350/DataSet/CH06PRO  
5.txt", header=FALSE)  
  
# Give names to the columns of data set  
colnames(Data) <- c("Y", "X1", "X2")  
fit <- lm(Y ~ X1, data=Data)  
# or use fit <- lm(Data$Y ~ Data$X1)  
fit  
anova(fit)  
  
# The estimated variance-covariance matrix for b  
cov.b <- vcov(fit)  
cov.b
```