

OLS Properties

Name: Jikhan Jeong

February 5, 2020

1. Positive Definite, Trace, Idempotent, Inverse, Matrix Derivative

1.1 Positive Definite

- $X'X$ = Square Matrix = Positive Definite = Convex = Can be minimized
- In OLS assumptions, we require $(X'X)^{-1} \rightarrow 0$ to derive consistency of $\text{Cov}(\beta)$

1.2 Trace

- $\text{Trace}(\text{scalar} \times X) = \text{scalar} \times \text{Trace}(X)$
- $\text{Trace}(X+Z) = \text{Trace}(X) + \text{Trace}(Z)$
- $\text{Trace}(X') = \text{Trace}(X)$
- $\text{Trace}(AB) = \text{Trace}(BA)$
- We need it when $\text{Cov}(Y) = \sigma^2 I = \text{Cov}(\varepsilon) = E(\varepsilon\varepsilon')$, where $E[\text{Trace}(e'e)] = \text{Trace}[E(e'e)] = (n-k)E(e'e)$, where n = sample number, k = variable numbers = column $\dim(X)$

1.3 Idempotent Matrix

- We need it when $\text{Cov}(Y) = \sigma^2 I = \text{Cov}(\varepsilon) = E(\varepsilon\varepsilon')$, where $E[\text{Trace}(e'e)] = E[\text{Trace}(e'e)]$
- P Idempotent Matrix, then $P'P = P$
- P Idempotent Matrix, then $I - P = M$, M Idempotent Matrix

1.4 Inverse Matrix

- $A = (A^{-1})^{-1}$
- $(A')^{-1} = (A^{-1})'$
- $(AB)^{-1} = B^{-1}A^{-1}$
- $|A^{-1}| = \frac{1}{|A|}$

1.5 Matrix Derivative

- From Prof. Ron's Book p.440, Useful for OLS Loss minimization
- \mathbf{z} : (k x 1) vector, \mathbf{A} : (k x j) matrix, \mathbf{w} : (j x 1) vector
- $\frac{\partial \mathbf{z}' \mathbf{A} \mathbf{w}}{\partial \mathbf{z}} = \mathbf{A} \mathbf{w}$
- $\frac{\partial \mathbf{z}' \mathbf{A} \mathbf{w}}{\partial \mathbf{w}} = \mathbf{A}' \mathbf{z}$
- $\frac{\partial \mathbf{z}' \mathbf{A} \mathbf{z}}{\partial \mathbf{z}} = 2 \mathbf{A} \mathbf{z}$, where k=j and A is a symmetric ($\mathbf{A}' = \mathbf{A}$)
- $\frac{\partial \mathbf{z}' \mathbf{w}}{\partial \mathbf{z}} = \mathbf{w}$ (k=j)
- $\frac{\partial \mathbf{z}' \mathbf{w}}{\partial \mathbf{w}} = \mathbf{z}$, (k=j)

2. Summarize the classical assumptions of GLM

- $E(\mathbf{Y}) = \mathbf{X}\beta$ and $E(\varepsilon) = 0$
- $\text{Cov}(\mathbf{Y}) = \sigma^2 \mathbf{I} = \text{Cov}(\varepsilon) = E(\varepsilon \varepsilon')$
- \mathbf{X} is a fixed matrix of values with $\text{rank}(\mathbf{X}) = k$, where k = unknown parameters

3. OLS objective function

- $S = \mathbf{e}'\mathbf{e}$, OLS minimize S
- $S = \mathbf{e}'\mathbf{e} = (\mathbf{y} - \mathbf{X}\beta)'(\mathbf{y} - \mathbf{X}\beta)$
- FOC β : $-2\mathbf{X}'\mathbf{y} + 2\mathbf{X}'\mathbf{X}\hat{\beta} = 0 \mapsto \mathbf{X}'\mathbf{y} = \mathbf{X}'\mathbf{X}\hat{\beta} \mapsto (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y} = \hat{\beta}$
- SOC β : $2\mathbf{X}'\mathbf{X}$ is positive definite

4. Gauss-Markov

- $\hat{\beta}$ in OLS is unbiased and efficient within the class of linear, unbiased estimator
- $\hat{\beta} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'(\mathbf{X}\beta + \mathbf{e}) = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{X}\beta + (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{e} = \beta + (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{e}$
- $E(\hat{\beta}) = E(\beta + (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{e}) = \beta + 0 = \beta$, unbiased
- $\text{Cov}(\hat{\beta}) = E[(\hat{\beta} - E(\hat{\beta}))(\hat{\beta} - E(\hat{\beta}))'] = E[(\hat{\beta} - \beta)(\hat{\beta} - \beta)'] = E[(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{e}\mathbf{e}'\mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}] = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'E(\mathbf{e}\mathbf{e}')\mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}$, where $E(\mathbf{e}\mathbf{e}') = \sigma^2 \mathbf{I}$
 $= \sigma^2 (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{X}(\mathbf{X}'\mathbf{X})^{-1} = \sigma^2 (\mathbf{X}'\mathbf{X})^{-1}$
- $\text{Cov}(\mathbf{Y}) = E(\mathbf{e}'\mathbf{e}) = E(\mathbf{e}'\mathbf{M}'\mathbf{M}\mathbf{e}) = E(\mathbf{e}'\mathbf{M}\mathbf{e}) = E(\text{tr}(\mathbf{e}'\mathbf{M}\mathbf{e})) = E(\text{tr}(\mathbf{M})\mathbf{e}'\mathbf{e}) = \text{tr}(\mathbf{M})E(\mathbf{e}'\mathbf{e}) = (n-k)\sigma^2$
- $\frac{\hat{\mathbf{e}}'\hat{\mathbf{e}}}{n-k} = s^2$, sample variance to make it unbiased estimator of $\sigma^2 \mapsto E(\frac{\hat{\mathbf{e}}'\hat{\mathbf{e}}}{n-k}) = E(s^2) = \sigma^2$

5. Consistent estimator of θ

- If $\hat{\theta}$ converges in probability to θ , then we say it is a consistent estimator
- $\text{plim } \hat{\beta} \rightarrow \beta$, consistent estimator as follows:

$$\text{plim } \hat{\beta} = \beta + \text{plim}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{e} = \beta + \text{plim}(\mathbf{X}'\mathbf{X}_n^{-1})^{-1}\mathbf{X}'\mathbf{e} = \beta + \text{plim}(\mathbf{X}'\mathbf{X}_n^{-1})^{-1} \frac{\mathbf{X}'\mathbf{e}}{n}$$

$$= \beta + \text{plim}(\mathbf{X}'\mathbf{X}_n^{-1}) \cdot 0 = \beta$$
- $\text{plim } s^2 = \frac{\hat{e}'\hat{e}}{n-k} \rightarrow \sigma^2$, consistent estimator as follows:

$$s^2 = \frac{\hat{e}'\hat{e}}{n-k} = \frac{\mathbf{e}'\mathbf{M}\mathbf{e}}{n-k} = \frac{\mathbf{e}'(\mathbf{I}-\mathbf{P})\mathbf{e}}{n-k} = \frac{\mathbf{e}'\mathbf{e} - \mathbf{e}'\mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{e}}{n-k} = \frac{\mathbf{e}'\mathbf{e}}{n-k} - \frac{\mathbf{e}'\mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{e}}{n-k}$$

$$\text{plim } s^2 \rightarrow \text{plim}\left(\frac{\mathbf{e}'\mathbf{e}}{n-k}\right) - \text{plim}\left(\frac{\mathbf{e}'\mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{e}}{n-k}\right) \rightarrow \sigma^2 - \text{plim}\left(\frac{\mathbf{e}'\mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{e}}{n-k}\right)$$

$$\rightarrow \sigma^2, \text{ where } \text{plim}\left(\frac{\mathbf{e}'\mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{e}}{n-k}\right) = 0 \text{ as } \frac{(\mathbf{X}'\mathbf{X})^{-1}}{n} \rightarrow 0$$

6. R^2

- Simple regression case, 1 exogenous variable“
- R^2 measures the percentage of variations in \hat{Y} can be explained by the regression model
- SST = total sum of squares = $(\mathbf{Y}-\bar{Y})(\mathbf{Y}-\bar{Y})'$
- SSE = error sum of squares = $(\mathbf{Y}-\hat{\mathbf{Y}})(\mathbf{Y}-\hat{\mathbf{Y}})'$
- SSR = sum of square residuals = $\text{SST} - \text{SSE} = (\hat{\mathbf{Y}}-\bar{Y})(\hat{\mathbf{Y}}-\bar{Y})'$
- $R^2 = \frac{\text{SSR}}{\text{SST}} = \frac{\text{SST}-\text{SSE}}{\text{SST}} = 1 - \frac{\text{SSE}}{\text{SST}}$

Reference

- Mittelhammer, R.C. and Mittelhammer, R.C., 1996. Mathematical statistics for economics and business (Vol. 78). New York: Springer.
- Fomby, T.B., Hill, R.C. and Johnson, S.R., 2012. Advanced econometric methods. Springer Science & Business Media.