Homework 1 Statistical Methods in Applied Computer Science DD2447

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Exercise 2.16 Mean, mode, variance for the beta distribution Suppose $\theta \sim Beta(a, b)$. Derive the mean, mode and variance.

Solution. Denote $Beta(a,b) = Beta_{a,b}$ to simplify the expressions later on. The mean of a random variable sampled from Beta(a,b) is defined as:

$$E\theta = \int_{\Omega} \theta' Beta_{a,b}(\theta') d\theta' \tag{1}$$

where

$$Beta_{a,b}(\theta) = \frac{\theta^{a-1}(1-\theta)^{b-1}}{B(a,b)}, B(a,b) = \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)}$$
(2)

and Ω is the support [0,1]

Lemma 0.1.

$$B(a+1,b) = B(b,a) \bigwedge B(a+1,b) = \frac{a}{a+b} B(a,b)$$
 (3)

Proof. The first relation is given by the symmetry in the definition of B together with a commutative multiplication. For the second relation we first need to show the well known identity

$$\Gamma(z+1) = z\Gamma(z) \tag{4}$$

By the definition of $\Gamma(z) \stackrel{\triangle}{=} \int\limits_0^\infty u^{z-1} e^{-u} du$ and by using partial integration we get

$$\Gamma(z+1) = \int u^z e^{-u} du = \left[-u^z e^{-u} \right]_{u=0}^{\infty} + \int_0^{\infty} z u^{z-1} e^{-u} du = z \int_0^{\infty} u^{z-1} e^{-u} du = z \Gamma(z)$$
 (5)

Having this relation we can show that

$$B(a+1,b) = \frac{\Gamma(a+1)\Gamma(b)}{\Gamma(a+1+b)} = \frac{a\Gamma(a)\Gamma(b)}{(a+b)\Gamma(a+b)} = \frac{a}{a+b}B(a,b)$$
 (6)

Starting out by deriving the mean:

$$E \theta = \int \theta' Beta_{a,b}(\theta') d\theta' = \int \frac{B(a+1,b)Beta_{a+1,b}(\theta')}{B(a,b)} d\theta' =$$

$$= \frac{B(a+1,b)}{B(a,b)} \int Beta_{a+1,b}(\theta') d\theta' = \frac{B(a+1,b)}{B(a,b)} = \{\text{Using Lemma 0.1}\} =$$

$$= \frac{aB(a,b)}{(a+b)B(a,b)} = \frac{a}{a+b}$$

$$(7)$$

Next we derive the mode which is the most occurring θ

$$Mode(\theta) = \underset{\theta'}{argmax}(Beta_{a,b}(\theta'))$$
 (8)

To find it we use

$$\frac{\partial Beta_{a,b}(\theta)}{\partial \theta} = 0 \tag{9}$$

$$0 = B(a, b) \frac{\partial Beta_{a,b}(\theta)}{\partial \theta} = \{ \text{prod. rule} \} =$$

$$= (a - 1)\theta^{a-2} (1 - \theta)^{b-1} - (b - 1)\theta^{a-1} (1 - \theta)^{b-2}$$
(10)

which gives us the equation

$$(a-1)\theta^{a-2}(1-\theta)^{b-1} = (b-1)\theta^{a-1}(1-\theta)^{b-2}$$
(11)

$$(a-1)(1-\theta) = (b-1)\theta \tag{12}$$

$$\theta = \frac{1}{1 + \frac{b-1}{a-1}} = \frac{a-1}{a-1+b-1} = \frac{a-1}{a+b-2} \tag{13}$$

Finally we deal with the variance:

$$\operatorname{Var}(\theta) = \operatorname{E}\theta^{2} - (\operatorname{E}\theta)^{2} = \int \theta'^{2} Beta_{a,b}(\theta') d\theta' - (\operatorname{E}\theta)^{2} =$$

$$= \frac{B(a+2,b)}{B(a,b)} - (\operatorname{E}\theta)^{2} = \frac{(a+1)}{(a+b+1)} \frac{a}{(a+b)} - \left(\frac{a}{a+b}\right)^{2} =$$

$$= \frac{a}{a+b} \left(\frac{(a+1)(a+b) - a(a+b+1)}{(a+b)(a+b+1)}\right) = \frac{ab}{(a+b)^{2}(a+b+1)}$$
(14)

Exercise 3.6 MLE for the Poisson distribution

The Poisson pmf is defined as $Poi(x|\lambda) = e^{-\lambda} \frac{\lambda^x}{x!}$ for $x \in \{0, 1, 2, ...\}$ where $\lambda > 0$ is the rate parameter. Derive the MLE.

Solution. The data $D = \{x_i\}$ is $i.i.d \land x_i \sim Poi(\lambda)$

$$p(x_i|\lambda) = e^{-\lambda} \frac{\lambda^{x_i}}{x_i!} \tag{15}$$

The probability for D:

$$L(\lambda) = p(\lbrace x_i \rbrace | \lambda) = \prod_i p(x_i | \lambda) = \prod_i e^{-\lambda} \frac{\lambda^{x_i}}{x_i!}$$
 (16)

The log-likelihood for the parameter λ

$$\ell(\lambda) = \log(\prod_{i} e^{-\lambda} \frac{\lambda^{x_i}}{x_i!}) = \sum_{i} (-\lambda + x_i \log \lambda - \log x_i)$$
 (17)

Find the maximum likely λ by finding the maximum log-likely λ which will coincide since log is a strictly increasing function.

$$\frac{\partial \ell(\lambda)}{\partial \lambda} = 0 \tag{18}$$

$$\frac{\partial \ell(\lambda)}{\partial \lambda} = \sum_{i} \left(-1 + \frac{x_i}{\lambda} \right) = -|\{x_i\}| + \frac{\sum x_i}{\lambda} = 0 \tag{19}$$

Which gives MLE for λ :

$$\lambda = \frac{\sum x_i}{|\{x_i\}|} = \mathbf{E} x \tag{20}$$

Exercise 3.7 Bayesian analysis of the Poisson distribution

In exercise 3.6, we defined the Poisson distribution with rate λ and derived its MLE. Here we perform a conjugate Bayesian analysis.

- **a.** Derive the posterior $p(\lambda|D)$ assuming a conjugate prior $p(\lambda) = Ga(\lambda|a,b) \propto \lambda^{a-1}e^{-\lambda b}$. Hint: the posterior is also a Gamma distribution.
- **b.** What does the posterior mean tend to as $a \to 0$ and $b \to 0$? (Recall that the mean of a Ga(a, b) distribution is a/b.)

Solution.