

Workbook for Introduction to TTE modeling

Integrating the hazard function in Stan

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Preliminaries for R examples

```
library(tidyverse)
library(stringr)
library(survival)
library(survminer)
library(texreg)
library(mgcv)
library(muhaz)
library(rstan)
library(bayesplot)
library(tidybayes)

theme_set(theme_bw())
bayesplot::color_scheme_set("viridis")

# OS data
d <- read_csv('../data/source/DDmodel0218_Simulated_OS.csv', na = c('.', '-99'))

# Add week 12 (Day 84) predicted tumor size
d84 <- d %>%
  filter(TIME <= 84) %>%
  group_by(ID) %>%
  mutate(rate = KG/1000 - KD0/1000*AUC0 - KD1/100*AUC1,
         prevTIME = lag(TIME, default = 0),
         change = exp(rate * (TIME-prevTIME)),
         ipred = IBASE * 1000 * cumprod(change)
        ) %>%
  arrange(ID, TIME) %>%
  slice(n()) %>%
  mutate(ipred84 = ipred * exp(rate * (84-TIME)),
         rts84 = ipred84 / ( IBASE * 1000 ) )

dos <- d %>%
  filter(TIME>0) %>%
```

```

group_by(ID) %>%
mutate(meanGem = mean(AUC1),
       Group = if_else(meanGem > 0, "Cb+G", "Cb")) %>%
ungroup() %>%
filter(CMT==2, EVID==0) %>%
left_join(d84 %>% select(ID, ipred84, rts84)) %>%
mutate(rts84_f = paste0("Q", ntile(rts84, n = 4)))

dos84 <- dos %>%
  filter(TIME>84) %>%
  mutate(TIME = TIME-84)

```

Working with traditional pharmacometric parameterization of Weibull model

$$h_i(t) = h_0(t) \times \exp(\beta \times \text{RTS}_i)$$

$$h_0(t) = \lambda_0 \times \alpha \times t^{\alpha-1}$$

$$H_i(t) = \lambda_i t^\alpha$$

$$S_i(t) = \exp(-\lambda_i t^\alpha)$$

Under this parameterization,

- median OS = $\left(\frac{\log 2}{\lambda_0 \exp(\beta \times \text{RTS})} \right)^{1/\alpha}$
- mean OS = $\{ \lambda_0 \exp(\beta \times \text{RTS}) \}^{-1/\alpha} \Gamma(1 + 1/\alpha)$

Set-up the data

```

tte_emax_data <- list(
  N = nrow(dos84),
  Y = dos84$TIME/365, # Use years instead of days
  cens = 1-dos84$DV,
  K = 1,
  X = model.matrix(~-1 + ECOG, data=dos84),
  RTS = dos84$rts84,
  prior_only = 0
)

fit_weibull_int <- stan(file = '../model/stan/weibull_integrated_hazard.stan',
                      data = tte_emax_data,
                      chains = 4,
                      iter=1000,
                      cores = 4,
                      seed = 76341
                      )

print(fit_weibull_int, pars=c('slope','lambda0','b','shape'))

```

```

. Inference for Stan model: weibull_integrated_hazard.
. 4 chains, each with iter=1000; warmup=500; thin=1;
. post-warmup draws per chain=500, total post-warmup draws=2000.
.

```

```

.           mean se_mean   sd 2.5% 25% 50% 75% 97.5% n_eff Rhat
. slope    0.86      0 0.16 0.53 0.75 0.86 0.96 1.15 1091 1
. lambda0  0.13      0 0.02 0.09 0.12 0.13 0.15 0.19 975 1
. b[1]     0.68      0 0.16 0.38 0.58 0.68 0.79 0.98 1340 1
. shape    2.15      0 0.13 1.91 2.07 2.15 2.24 2.40 1091 1
.
. Samples were drawn using NUTS(diag_e) at Wed Aug 31 22:19:28 2022.
. For each parameter, n_eff is a crude measure of effective sample size,
. and Rhat is the potential scale reduction factor on split chains (at
. convergence, Rhat=1).

```

```

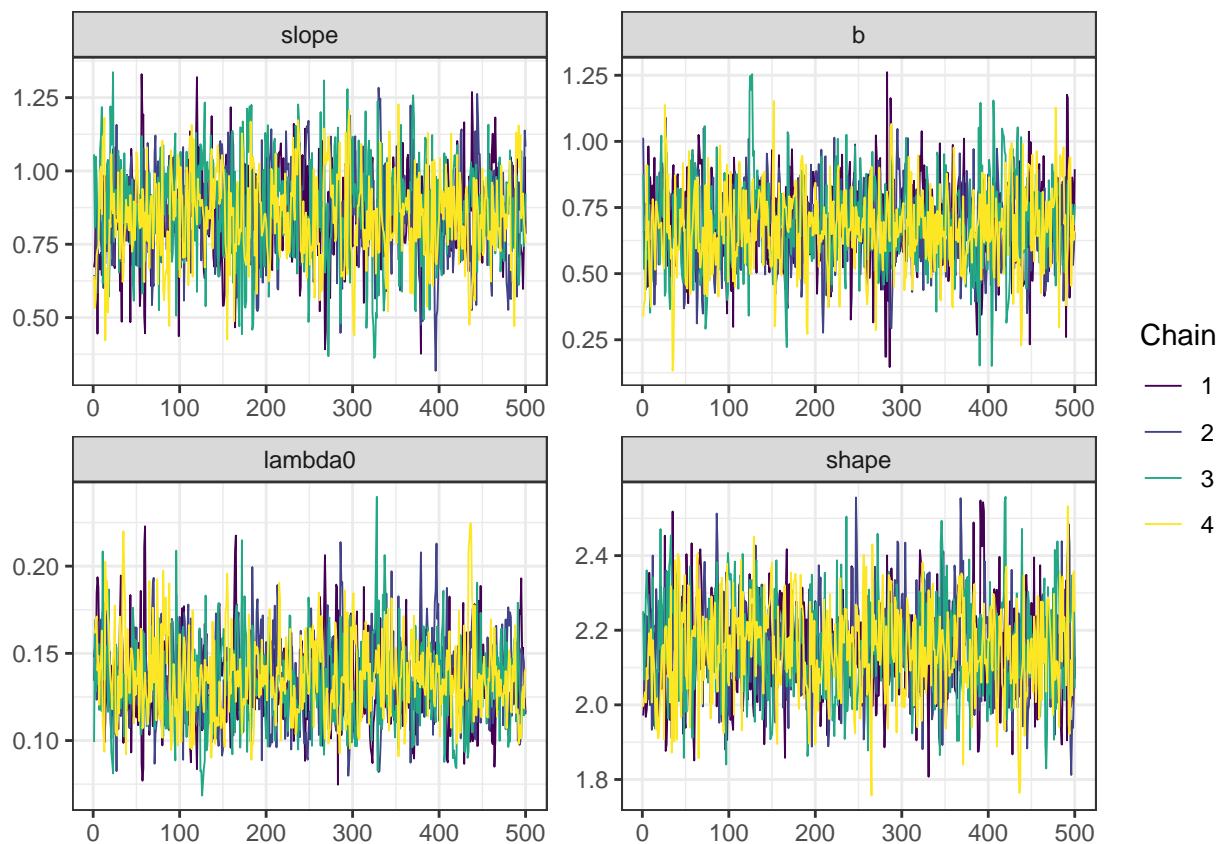
samples_int <- spread_draws(fit_weibull_int, slope,b[i],lambda0,shape) %>%
  rename(Chain=.chain)

```

```

mcmc_trace(samples_int, pars=c('slope','b','lambda0','shape'))

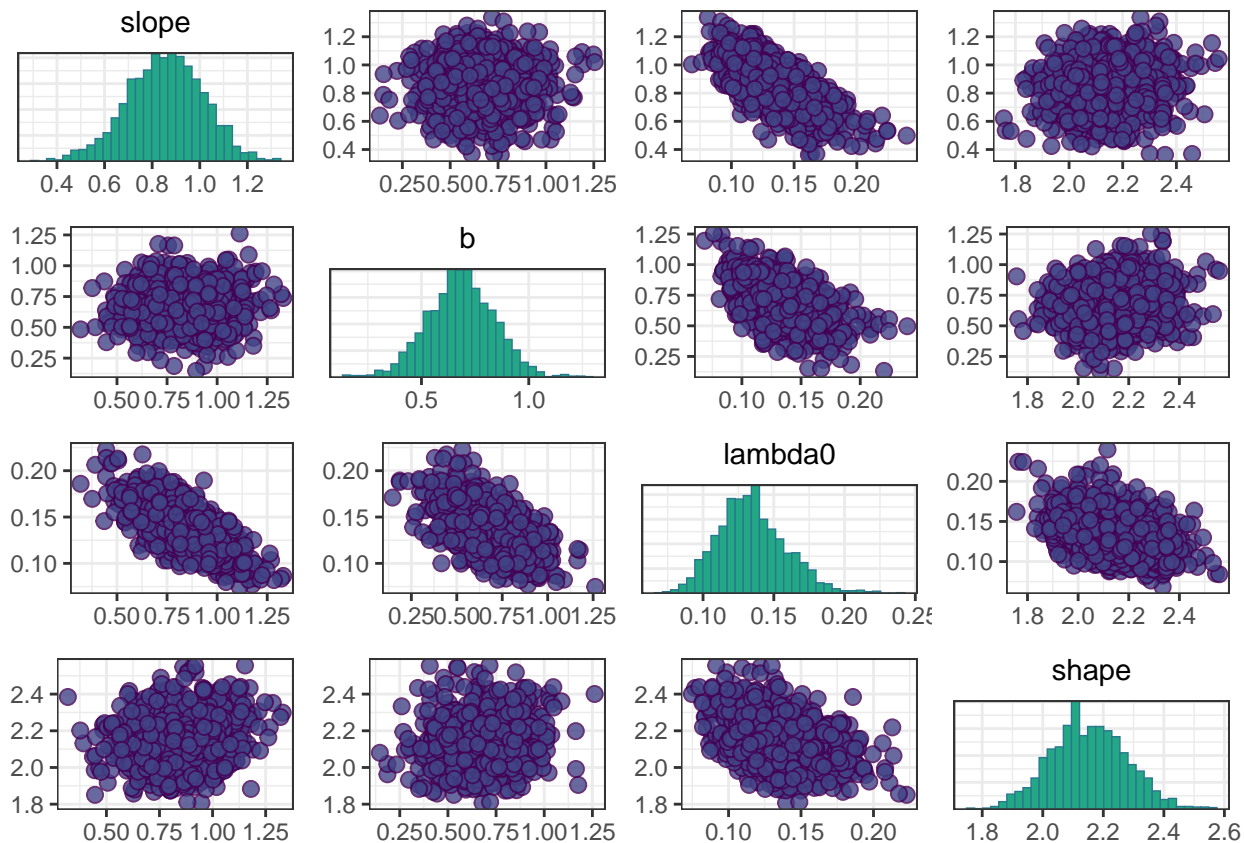
```



```

mcmc_pairs(samples_int, pars=c('slope','b','lambda0','shape'))

```



**** Exercise: ****

How does this compare to the closed-form model? Re-fit the closed-form model and compare parameter estimates.

Try coding a model with a 'hockey stick' hazard:

$$h(t) = \begin{cases} \alpha_1 + \beta_1 t & 0 \leq t < \tau_1 \\ \alpha_1 + \beta_1 \tau_1 + \beta_2 (t - \tau_1) & t \geq \tau_1 \end{cases}$$

where you can either pick a fixed value of τ_1 or estimate it.

How does this compare to the Weibull model fit?