Recall Monte-Carlo:

$$\frac{\prod_{x \in P(x)} f(x) = \int f(x) p(x) dx}{x e^{-\frac{1}{n}} \int_{i=1}^{\infty} f(x_i)} \int_{i=1}^{\infty} \frac{1}{n} \int_{i=1}^{\infty} f(x_i) \int_{i=1}^{\infty} \frac{1}{n} \int_{i=1}^{\infty} f(x_i) \int_{i=1}^{\infty} \frac{1}{n} \int_{i=1}^{\infty} \frac{1}{n} \int_{i=1}^{\infty} f(x_i) \int_{i=1}^{\infty} \frac{1}{n} \int_{i=1}^{\infty} \frac{$$

. f GN = n : expectation of 2

Setup: Assume p(x) is a density, i.e. [p(x) dx=1. Then:

$$\mathbb{E}\left[f(x)\right] = \int f(x) p(x) dx = \int \left[f(x) \frac{p(x)}{q(x)}\right] q(x) dx$$

$$= \underbrace{\mathbb{E}}_{x \sim q \cdot \Theta x} \left[f \cdot \Theta x \frac{p \cdot (x_1)}{q \cdot (x_1)} \right] \approx \frac{1}{n} \sum_{i=1}^{n} f \cdot (x_i) \frac{p \cdot (x_i)}{q \cdot (x_i)}, x_i \overset{\text{i.id}}{\sim}$$

$$\simeq \frac{1}{m} \sum_{i=1}^{n} f(x_i) W(x_i)$$
, $W(x_i) := \frac{P(x_i)}{q_i(x_i)}$ Theorem ce weights.

 $\frac{9(91)}{5}: \text{ proposal distribution}$ $\frac{1}{5}: \text{ Easy to sample from}$ $\frac{2}{5}: \text{ Easy to evaluate for a given x.}$

POXT: \$1. May not be easy to somple 2. Should be easy to evaluate

@ Y donsity g(x) s.t. g(x)=0=7p(x)=0

Romarks:

1.)
$$\mathbb{E}_{x \sim q(x_1)} \left[\prod_{n=1}^{n} \frac{1}{n} \sum_{i=1}^{n} \mathbb{E}_{x \sim q(x_1)} \left[f(x_i) \frac{p(x_i)}{q(x_i)} \right] \right]$$

$$= \frac{1}{n} \sum_{i=1}^{n} \int f(x_i) \frac{p(x_i)}{q(x_i)} g(x_i) dx =$$

$$= \frac{1}{n} \sum_{i=1}^{n} \mathbb{E}_{x \sim p(x_i)} \left[f(x_i) \right] \xrightarrow{y \to \infty} \mathbb{E}_{x \sim p(x_i)} \left[f(x_i) \right]$$

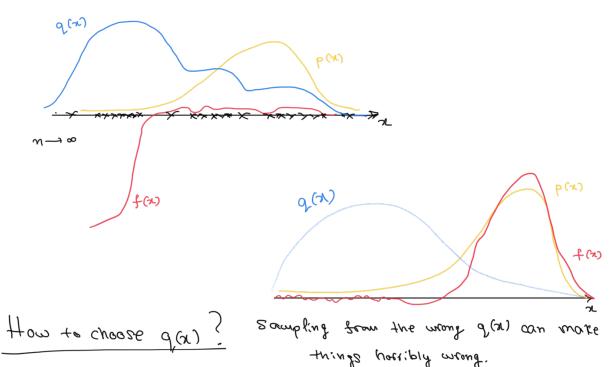
=> (Honce, IS estimator is unbiased, consistent.

In fact, it is straightforward to solve for optimal $q^*(x)$, but in practice it may be difficult to sample from $q^*(x)$.

Sconarios:

- I.) Can't sample from p(x), use Is.
- 2.) Whether or not we sample from p(x), use Is to improve upon the convergence of the vanilla MC estimator.

Example: Suppose f(x) is the return of some investment x, we want to approximate the expected returns: [[f(x)] x~p(x)



Choose d'es to pe pordé

where |f(x)|p(x) is large $\frac{1}{2} = \sum_{i=1}^{\infty} \frac{p(x_i)}{p(x_i)} + \sum_{i=1}^{\infty} \frac{1}{2} \frac{$

n = 1 q(ni)

Limitations: I. Our intuition on how to choose a good que breaks down in high dimensions.

2. It is difficult to assess the acanacy of our estimator.

Importance sampling for un-normalized distribution:

Often we may know pox or gow only up to a normalizing cons

$$\frac{e.g.}{p(\mathfrak{d})}$$
; Posterior inferece: $p(g|\mathcal{D}) = \frac{p(\mathfrak{d}|g)p(g)}{p(\mathfrak{d})p(g)dg} \propto p(\mathfrak{D}|g)p(g)$

Setup: Assume:
$$p(x) = \frac{p(x)}{z_p}$$
, $q(x) = \frac{2q}{q(x)}$, $\int_{0}^{\infty} (x) dx = z_p$

$$\left[\left[\int_{\mathbb{R}^{2}} \left[\int_{\mathbb{R$$

$$= \int f(x) \frac{\tilde{p}(x)}{\tilde{q}(x)} \frac{z_{q}}{z_{p}} q(x) dx$$

Assume that we can efficiently sample from q(x), and we come evaluate $\tilde{p}(x)$, $\tilde{q}(x)$.

But we still don't know the normalizing constants ...

Then let's try to apportunate them via Monte-Carlo:

$$\begin{cases}
\frac{\overline{Z_p}}{\overline{Z_q}} - \frac{1}{\overline{Z_q}} \int \widetilde{p}(x) dx = \int \frac{\widetilde{p}(x)}{\widetilde{q}(x)} q(x) dx = \underbrace{\mathbb{E}}_{x \sim q(x)} \left(\frac{\widetilde{p}(x)}{\widetilde{q}(x)} \right) \\
\approx \frac{1}{\pi} \sum_{i=1}^{\infty} \widetilde{u}(x_i) \qquad (2)$$

$$\frac{1}{n} \sum_{i=1}^{m} f(x_i) \widetilde{W}(x_i) = \frac{1}{n} \sum_{i=1}^{n} f(x_i) \widetilde{W}(x_i)$$

$$\frac{1}{n} \sum_{i=1}^{m} \widetilde{W}(x_i)$$

$$= \frac{1}{n} \sum_{i=1}^{n} f(x_i) \widetilde{W}(x_i)$$

$$\frac{1}{n} \sum_{i=1}^{m} \widetilde{W}(x_i)$$

(2) Keep in mind that we made two MC approximations.

... honce, this is expected to perform worse than regular I.S.