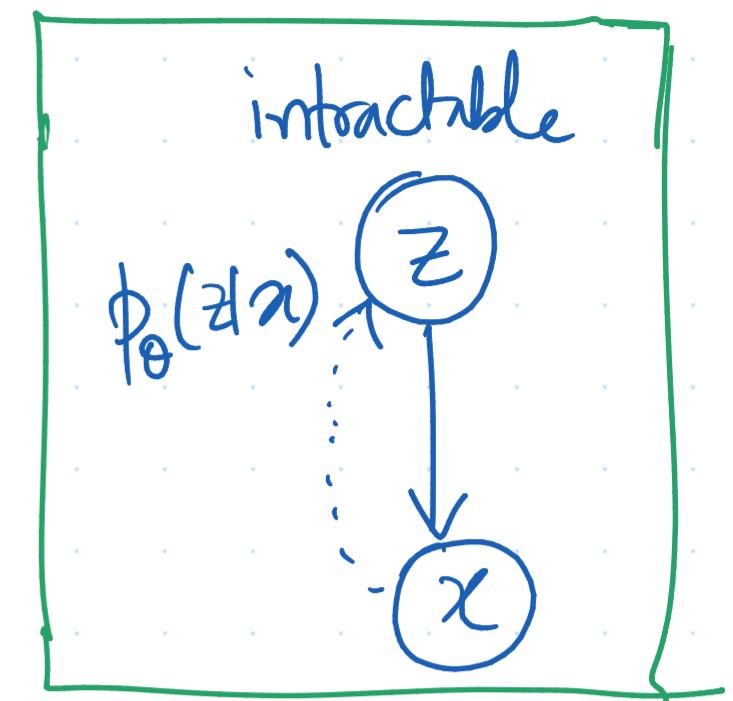


VAE using Expectation Maximization strategy. 16/3/25

M.C.M.C. Expectation Maximization → MCMC Sampling.
Assumptions :-

We can calculate $p(z|x^{(i)})$ → this may not hold in practice.



E-step : For all i

$$\text{set } Q_i^{(t)}(z) = p(z|x^{(i)}; \theta^{(t)})$$

M-step

$$\theta^{(t+1)} = \arg \max \sum_{i=1}^n \sum_z Q_i^{(t)}(z^{(i)}) \log \frac{p(x^{(i)}, z^{(i)}; \theta)}{Q_i^{(t)}(z)}$$

← doesn't depend on θ .

$$= \arg \max_{\theta} \sum_{i=1}^n \sum_z Q_i^{(t)}(z^{(i)}) \log p(x^{(i)}, z^{(i)}; \theta)$$

(same as maximizing
this quantity).

$$z \sim N(0, I_{k \times k})$$

$$x|z \sim \text{NeuralNet}_\theta(z^{(i)}) \sim \text{Complex}(z^{(i)}, \theta)$$

$p(z|x) \rightarrow$ impossible, cannot be calculated, very complex

intractable

$$\log\left(\frac{A}{B}\right) = \log A - \cancel{\log B} \xrightarrow{\text{constant}}$$

w.r.t. θ

hence, we get rid of this form.

$$\Rightarrow \arg \max_{\theta} \sum_{i=1}^n \sum_z Q_i^{(t)}(z^{(i)}) \log p(x^{(i)}, z^{(i)}; \theta).$$

$$\Rightarrow \arg \max_{\theta} \sum_{i=1}^n \underbrace{E_{z^{(i)} \sim Q_i^{(t)}} [\log p(x^{(i)}, z^{(i)}; \theta)]}$$

$Q_i^{(t)}(z)$ in E-step is only used to calculate this expectation.

We are not interested in the density of Q . We are only interested in the density value of Q to perform this expectation.

$$\approx \arg \max_{\theta} \sum_{i=1}^n \frac{1}{M} \sum_{m=1}^M \log p(x^{(i)}, z^{(m)}; \theta)$$

$$z_i^{(m)} \sim Q_i^{(t)}$$

(replace with Monte-Carlo estimation of
the estimation)

$z_i^{(m)}$ = sampled from $Q_i^{(t)}$.

Q -posterior

Sample from posterior of a complex probability distribution even though we don't know how to evaluate it.

Law of large numbers tells us that
 as the M goes to ∞ the Monte-Carlo
 estimate will tend to converge
 towards the true expectation.

} Gibbs Sampling
 } Metropolis Hastings.

We don't know the posterior, but we can approximate
 the posterior using the Monte-Carlo Technique / Sampling
 Technique.

Here the guarantee of the likelihood increase after
 every step doesn't hold anymore, because this is an
 approximation of the lower bound and not the exact
 lower bound.

So, there are three ways to do this :-

- * Exact posterior \rightarrow maths / calculation etc. \rightarrow (diff)
- * Gibbs / Sampling methods \rightarrow ✓
- * Variational Inference \rightarrow optimization \rightarrow (previous class).

Jensen's Inequality

$$\log p(x) \geq \text{ELBO}(\alpha; \theta)$$

\hookrightarrow lower bound of the evidence.

$$\log p(x) = \text{ELBO}(x; \theta) + \underbrace{\quad}_{?}$$

$$D_{KL}(\theta \| p(z|x)) = \log p(x) - \text{ELBO}(x; \theta)$$

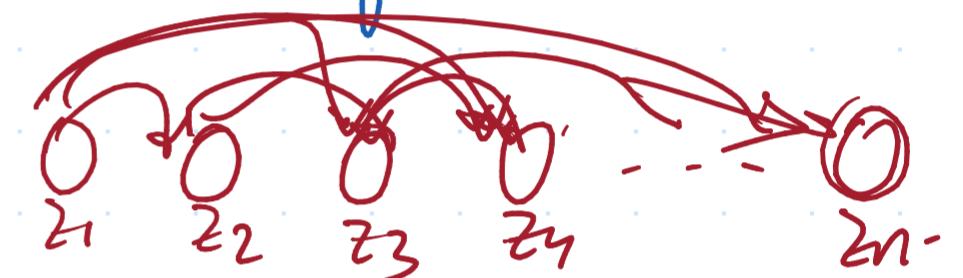
$$\log p(x) = \underbrace{\text{ELBO}(x; \theta)}_{\substack{\text{const. w.r.t.} \\ \theta}} + D_{KL}(\theta \| p_{z|x})$$

maximize
w.r.t. θ

≥ 0

$$P_{z|x} = \underset{q \in Q}{\operatorname{argmax}} \text{ELBO}(x; q) \leftarrow \begin{array}{l} \text{Variational} \\ \text{Inference} \end{array}$$

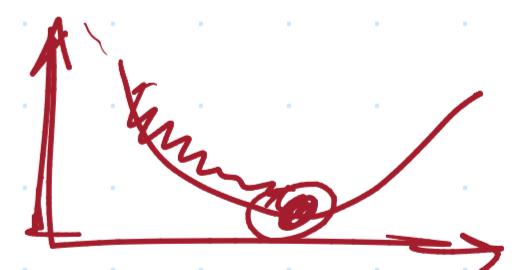
Mean field Assumption:-



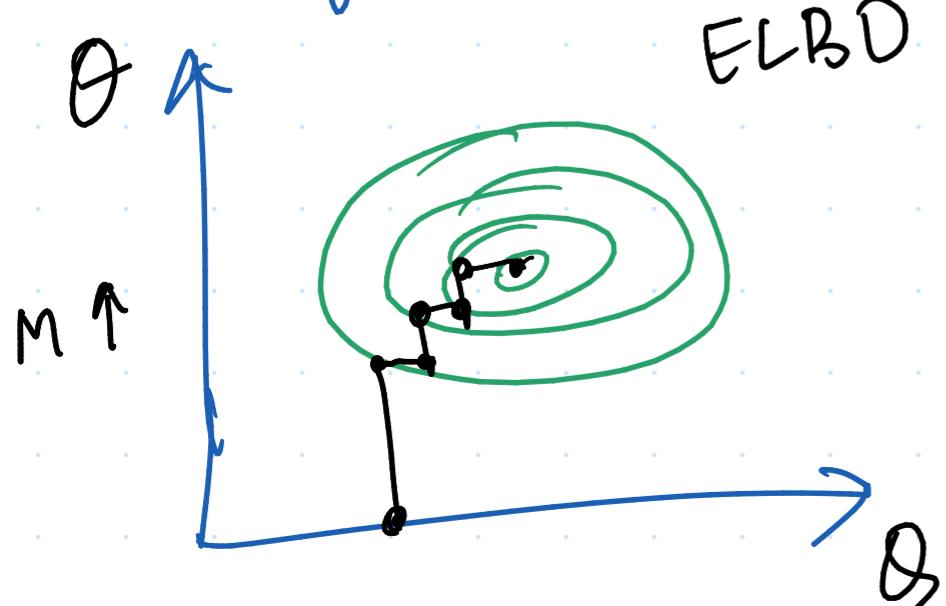
$$Q(z) = Q_1(z_1) Q_2(z_2) Q_3(z_3) \dots Q_K(z_K) \quad \begin{matrix} 0 & 0 & 0 \end{matrix}$$

$z \in \mathbb{R}^K$

If we assume that the component of z vector can be factored into K -independent scalar probability distribution, then, this is called mean field assumption.



Mean field inference makes computation easier -



co-ordinate ascent - multiple
variables • Start with random initialization, all except some fixed, optimize the ones that are not held fixed.

Worked well with classic EM.

Once updated estimates are got, hold them fixed and update the other ones & so on.

Can we do gradient ascent? → Variational A.E..

$z \sim N(0, I_{k \times k}) \rightarrow$ Latent variable / prior → sampled from some normal distribution.

Important:

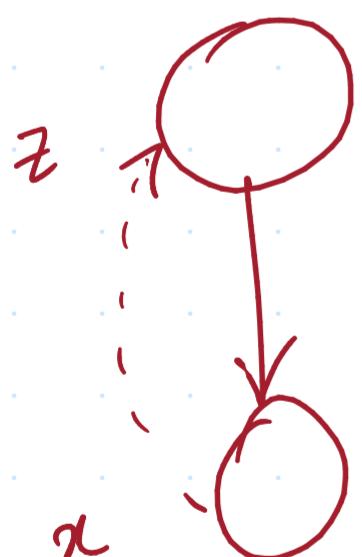
$$p(z|x) = \frac{p(z|x) p(x)}{p(z)}$$

posterior evidence

likelihood prior

$x|z \sim N(g(z;\theta), \sigma^2 I)$

Likelihood -



$p(z|x)$; $g \rightarrow$ Neural Network with param, θ , we are never going to obtain.

Approximate using variational inference $p(z|x)$

$\theta \rightarrow$ family for Variational Inference (V.I.).

$$\tilde{q}_i(z) = N\left(q(z^{(i)}; \phi), \text{diag}(V(x^{(i)}; \psi))^2\right)$$

$\in \mathbb{R}^k$ $\in \mathbb{R}^{k \times k}$

μ

Σ

We will get different $Q \rightarrow \text{dist}$ per example. In E.M.
 E-step is performed separately per example.

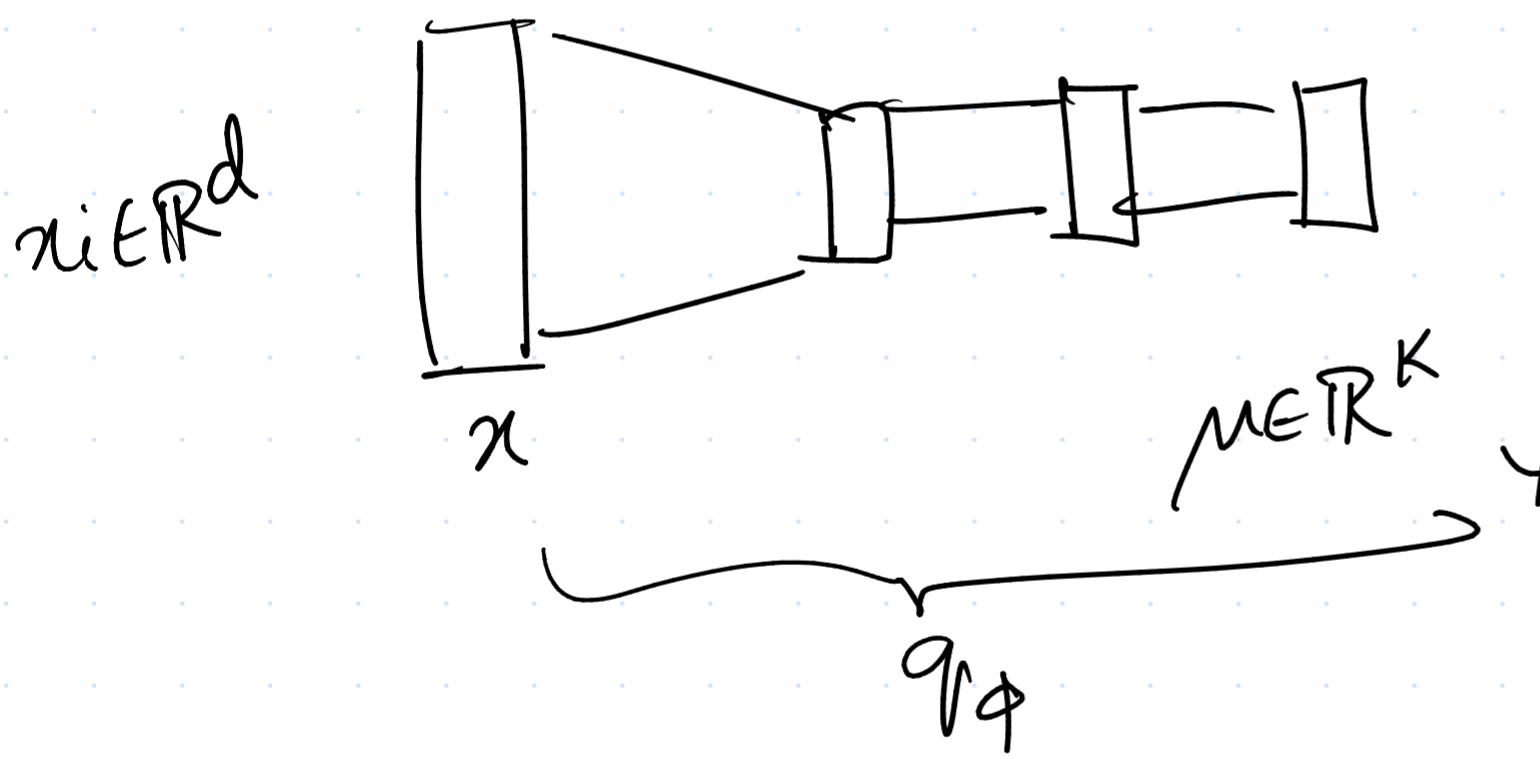
$Q_i \rightarrow \text{depends on } x \quad Q_i = p(z|x)$ N.N. \rightarrow takes x as input
 ("Amortized Influence":

function of x , and those function is the N.N., so we don't have to optimize for each example independently

as the E-M Step.

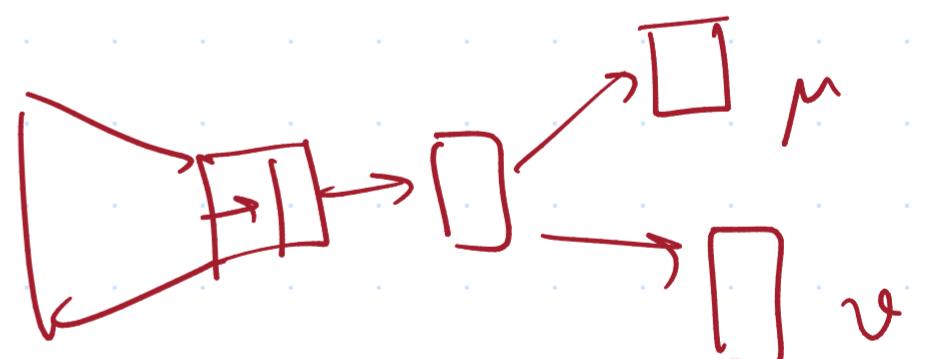
$$q_{\phi}(x^{(i)}) = \mu^{(i)}$$

why is $k < d$, $\mu \rightarrow$ smaller dim.

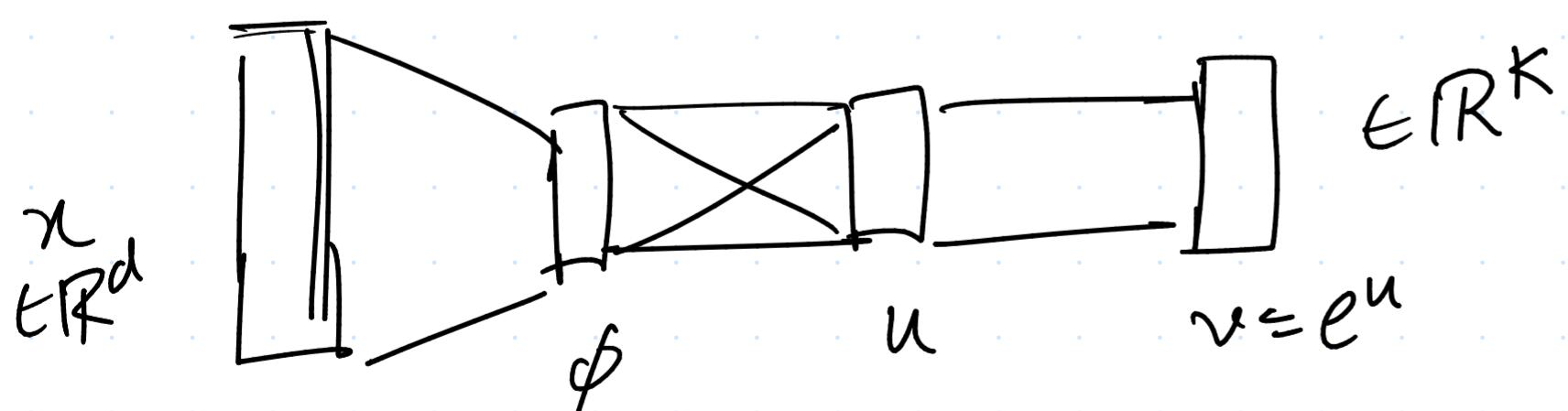


$Q \rightarrow \text{dist over } z$.

and $z = k$ dim, $\mu = k$ -dim.



Covariance:



to get +ve s.d.

positive = Square them, exponentiate each element.

$Q_i \rightarrow$ mean & covariance.

Normal Dist \rightarrow vector & a matrix
in high dimensional space.

$$k \begin{bmatrix} v_1 \\ v_2 \\ \vdots \\ v_k \end{bmatrix}$$

✓

Gaussian dist :- & diagonal covariance matrix. If there is no correlation b/w two components of a joint Gaussian, then, they necessarily must be independent.

Diagonal covariance matrix \rightarrow making the mean field assumption. Each of z_i 's \rightarrow independent.

ϕ & ψ \rightarrow shared across all the examples.

Decoder
 $p(x|z) \cdot p(z)$

$$\text{ELBO } (\phi, \psi, \theta) = \sum_{i=1}^n E_{z^{(i)} \sim Q_i} \left[\log \frac{p(x^{(i)}, z^{(i)}; \theta)}{Q_i(z^{(i)})} \right]$$

where, $Q_i = N(\underline{q}(x^{(i)}; \phi), \underline{\Sigma} \underline{\Sigma})$, $\text{diag}(\underline{v}(x^{(i)}, \psi))$.

$$\text{ELBO } (\underline{Q}, \theta) = \sum_{i=1}^n E_{z^{(i)} \sim Q^{(i)}} \left[\log \frac{p(x, z; \theta)}{Q_i(z^{(i)})} \right]$$

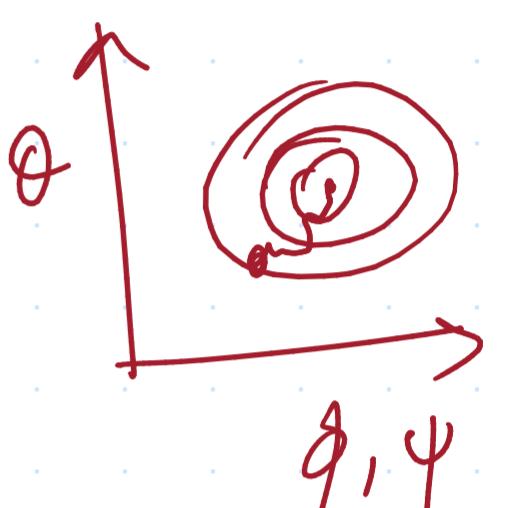
ELBO in EM, where each $\theta \rightarrow$ separately calculated for each example in parallel.

Optimize ϕ, ψ, θ with gradient descent:-

$$\theta := \theta + \eta \nabla_{\theta} \text{ELBO}(\phi, \psi, \theta)$$

$$\phi := \phi + \eta \nabla_{\phi} \text{ELBO}(\phi, \psi, \theta)$$

$$\psi := \psi + \eta \nabla_{\psi} \text{ELBO}(\phi, \psi, \theta)$$



$$\nabla_{\theta} \text{ELBO}(\phi, \psi, \theta)$$

$$= \nabla_{\theta} \sum_{i=1}^N \mathbb{E}_{z^{(i)} \sim q_i} \left[\log \frac{p(x^{(i)}, z^{(i)}; \theta)}{q_i(z^{(i)})} \right]$$

$$= \sum_{i=1}^N \mathbb{E}_{z^{(i)} \sim q_i} \left[\nabla_{\theta} \log p(x^{(i)}, z^{(i)}; \theta) \right]$$

$$= \sum_{i=1}^N \mathbb{E}_{z^{(i)} \sim q_i} \left[\underbrace{\nabla_{\theta} \log p(x^{(i)} | z^{(i)}; \theta)}_{\text{Generator Decoder}} + \nabla_{\theta} \log p(z^{(i)}) \right]$$

Monte Carlo (M.C.)

Generator Decoder

$$\nabla_{\phi} \text{ELBO}(\phi, \psi, \theta) = \nabla_{\phi} \sum_{i=1}^N \mathbb{E}_{z^{(i)} \sim Q_i} \left[\log \frac{p(x^{(i)}, z^{(i)}; \theta)}{Q(z^{(i)})} \right]$$

The distribution $z^{(i)} \sim Q_i$ depends on ϕ , hence we cannot swap the expectations.

$$z \sim N(\mu, \sigma)$$

$$z = \varepsilon \cdot \sigma + \mu$$

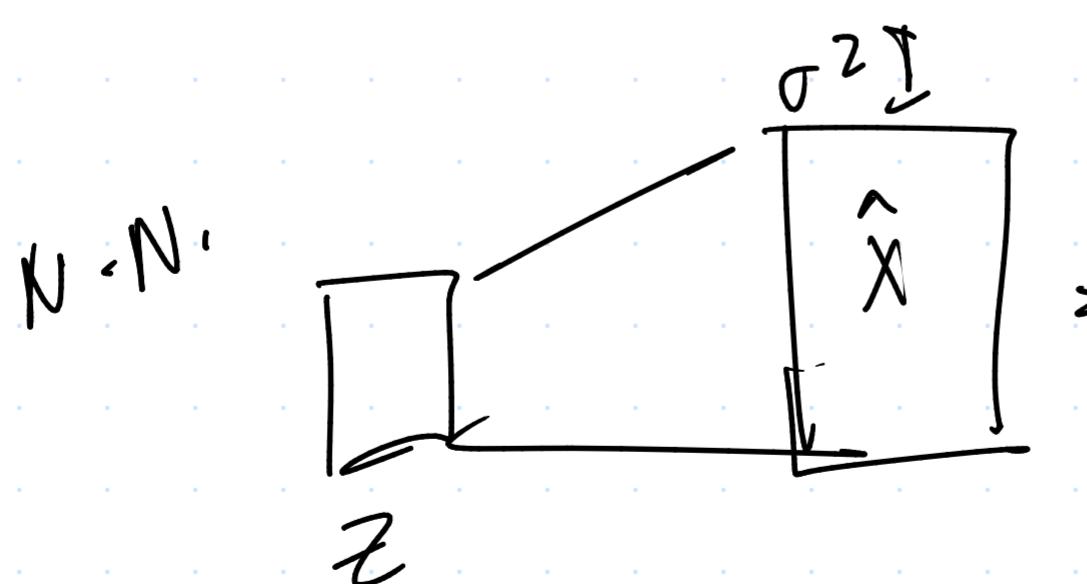
\rightarrow decoupled into this.

$$\varepsilon \sim N(0, 1)$$

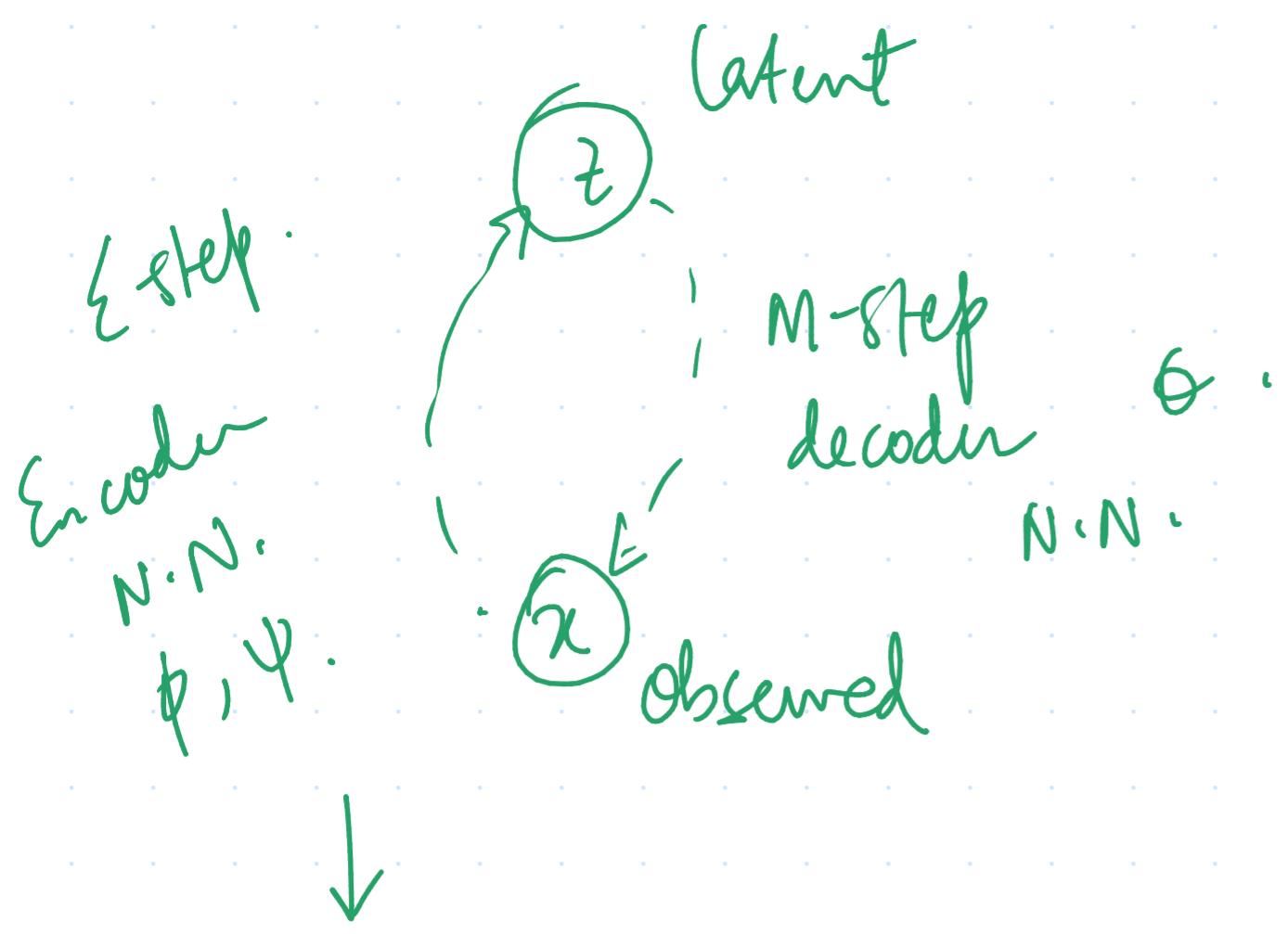
$$\nabla_{\phi} \sum_{i=1}^N \mathbb{E}_{\varepsilon^{(i)} \sim N(0, 1)} \left[\log \frac{p(x^{(i)}, \underbrace{\varepsilon^{(i)} \odot \Sigma^{(i)} + \mu^{(i)}}_{z^{(i)}}; \theta)}{Q(\underbrace{\varepsilon^{(i)} \odot \Sigma^{(i)} + \mu^{(i)}}_{z^{(i)}})} \right]$$

$$\text{where, } \mu^{(i)} = q_r(x^{(i)}; \phi)$$

$$\Sigma^{(i)} = \text{diag}(V(x^{(i)}); \psi)$$



$$\log p(x^{(i)}; g(z^{(i)}), \sigma^2 I)$$



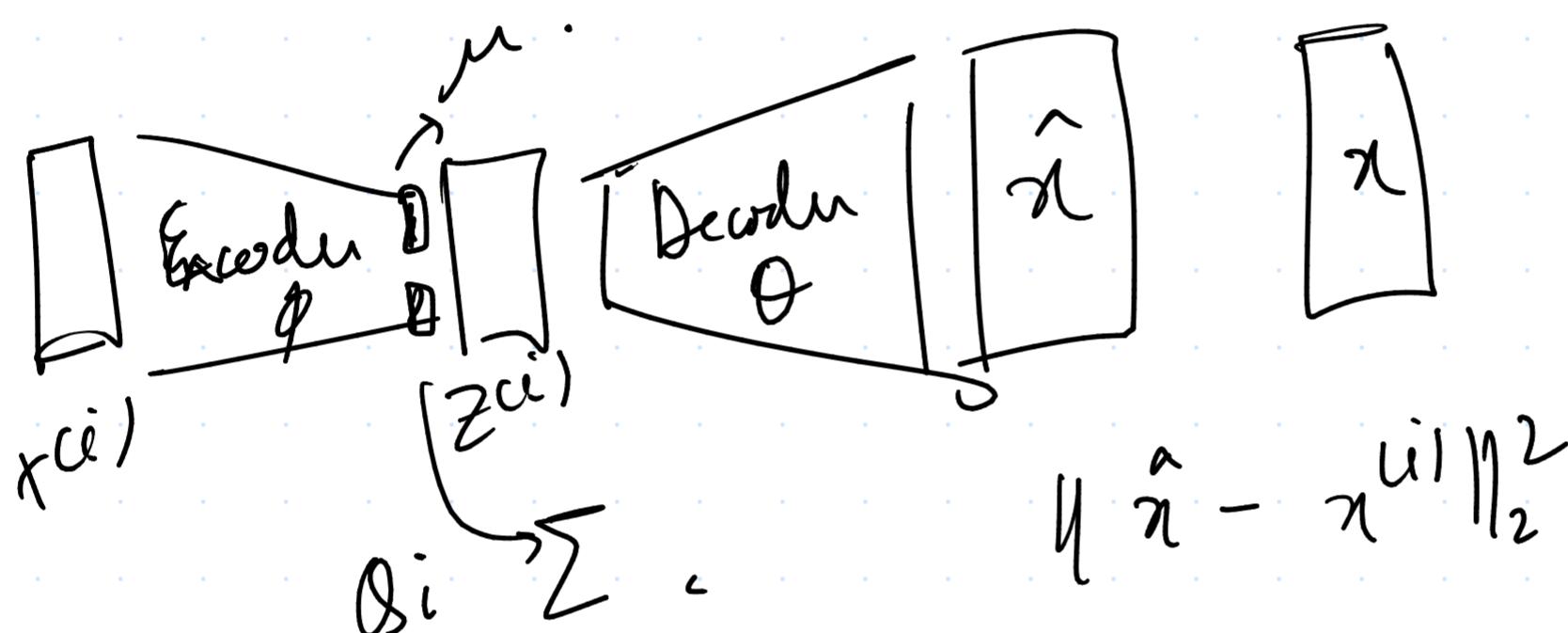
Reparameterization trick

Maximizing the ELBO,

$$\text{ELBO}(\phi, \psi, \theta)$$

Easy gradient:

Expectation of the gradients are approximated using Monte-Carlo estimate.



Maximizing the ELBO - calculating the gradients and then gradient ascent steps. Alternate to EM → fit a latent variable model.