

Efficient Variational Inference for Sparse Deep Learning with Theoretical Guarantee

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Introduction

Motivation:

- **Compress** deep learning model for hardware limited devices.
- **Recover** potential sparsity structure of the target function.
- Variable **selection**.

Existing approaches:

Pruning methods (no theoretical guarantee on choosing the threshold):

- Frequentist: Zhu and Gupta 2018; Frankle and Carbin 2018, etc.
- Bayesian: Molchanov et al. 2017; Ghosh et al. 2018, etc.

Theoretical work (no efficient implementation):

- Polson and Rockova 2018; Cherief-Abdellatif 2020, etc.

Our contribution:

A complete package of both **theory** and **computation** for sparse DNN from a **Bayesian** perspective.

Model setup

Nonparametric regression.

Consider a nonparametric regression model with random covariates:

$$Y_i = f_0(X_i) + \epsilon_i, \quad i = 1, \dots, n,$$

where $X_i \sim \mathcal{U}([-1, 1]^p)$ and $\epsilon_i \sim \mathcal{N}(0, \sigma^2)$.

Deep Neural Network

The DNN $f_\theta(X)$ is used to approximate $f_0(X)$, where θ denotes all the coefficients in the neural network. The network configuration includes

- L: number of hidden layers
- N: width of the network (assume equal width)
- s: sparsity level ($\|\theta\|_0 \leq s$)

Variational Inference

Sparse Bayesian deep learning with spike-and-slab prior

- Prior distribution $\pi(\theta)$:

$$\theta_i | \gamma_i \sim \gamma_i \mathcal{N}(0, \sigma_0^2) + (1 - \gamma_i) \delta_0, \quad \gamma_i \sim \text{Bern}(\lambda), \quad i = 1, \dots, T,$$

Note: the theoretical guarantees will be established under proper deterministic choices of hyperparameters λ and σ_0 .

- Variational distribution $q(\theta)$:

$$\theta_i | \gamma_i \sim \gamma_i \mathcal{N}(\mu_i, \sigma_i^2) + (1 - \gamma_i) \delta_0, \quad \gamma_i \sim \text{Bern}(\phi_i), \quad i = 1, \dots, T.$$

The variational parameters are $(\mu_i, \sigma_i, \phi_i)$, the transformed variational parameters are denoted as ω .

Stochastic estimator of negative ELBO and its gradient

$$\begin{aligned} \tilde{\Omega}^m(\omega) &= -\frac{n}{m} \frac{1}{K} \sum_{i=1}^m \sum_{k=1}^K \log p_{g(\omega, \nu_k)}(D_i) + \text{KL}(q_\omega(\theta) || \pi(\theta)), \\ \nabla_\omega \tilde{\Omega}^m(\omega) &= -\frac{n}{m} \frac{1}{K} \sum_{i=1}^m \sum_{k=1}^K \nabla_\omega \log p_{g(\omega, \nu_k)}(D_i) + \nabla_\omega \text{KL}(q_\omega(\theta) || \pi(\theta)), \end{aligned} \quad (1)$$

where D_i 's are randomly sampled data points and $g(\omega, \nu_k)$ are the reparameterized version of θ with some auxiliary noise ν_k , m and K are minibatch size and Monte Carlo sample size.

Theoretical Results

Optimal sparsity level:

Define s^* as

$$s^* = \text{argmin}_s \{r_n(L, N, s) + \xi_n(L, N, s)\},$$

such that s^* strikes the **balance (trade-off)** between the variational error r_n and the approximation error ξ_n . Note that s^* is generally unknown, but our modeling is capable of automatically attaining the same rate of convergence as if s^* is known.

Conditions:

The main theorems are under the following conditions.

- The activation function is 1-Lipschitz continuous.
- The hyperparameter σ_0^2 is set to be some constant, and λ satisfies $\log(1/\lambda) = O\{(L+1) \log N + \log(p\sqrt{n/s^*})\}$ and $\log(1/(1-\lambda)) = O((s^*/T)\{(L+1) \log N + \log(p\sqrt{n/s^*})\})$.
- $\max\{s^* \log(p\sqrt{n/s^*}), (L+1)s^* \log N\} = o(n)$ and $r_n(L, N, s^*) \asymp \xi_n(L, N, s^*)$.

Main Theorems:

Denote the log-likelihood ratio between p_0 and p_θ as $l_n(P_0, P_\theta) = \log(p_0(D)/p_\theta(D)) = \sum_{i=1}^n \log(p_0(D_i)/p_\theta(D_i))$. Given some constant $B > 0$, we define

$$\begin{aligned} r_n^* &:= r_n(L, N, s^*) = ((L+1)s^*/n) \log N + (s^*/n) \log(p\sqrt{n/s^*}), \\ \xi_n^* &:= \xi_n(L, N, s^*) = \inf_{\theta \in \Theta(L, \mathbf{p}, s^*), \|\theta\|_\infty \leq B} \|f_\theta - f_0\|_\infty^2. \end{aligned}$$

The Hellinger distance is defined as

$$d^2(P_\theta, P_0) = \mathbb{E}_X \left(1 - \exp\{-[f_\theta(X) - f_0(X)]^2 / (8\sigma_\epsilon^2)\} \right).$$

In addition, let $s_n = s^* \log^{2\delta-1}(n)$ for any $\delta > 1$.

Lemma

With dominating probability,

$$\inf_{q(\theta) \in \mathcal{Q}} \left\{ \text{KL}(q(\theta) || \pi(\theta | \lambda)) + \int_{\Theta} l_n(P_0, P_\theta) q(\theta) d\theta \right\} \leq Cn(r_n^* + \xi_n^*)$$

where C is either some positive constant if $\lim n(r_n^* + \xi_n^*) = \infty$, or any diverging sequence if $\limsup n(r_n^* + \xi_n^*) \neq \infty$.

Lemma

If σ_0^2 is set to be constant and $\lambda \leq T^{-1} \exp\{-Mnr_n^*/s_n\}$ for any positive diverging sequence $M \rightarrow \infty$, then with dominating probability,

$$\int_{\Theta} d^2(P_\theta, P_0) \hat{q}(\theta) d\theta \leq C\epsilon_n^{*2} + \frac{3}{n} \inf_{q(\theta) \in \mathcal{Q}} \left\{ \text{KL}(q(\theta) || \pi(\theta | \lambda)) + \int_{\Theta} l_n(P_0, P_\theta) q(\theta) d\theta \right\},$$

where C is some constant, and

$$\epsilon_n^* := \epsilon_n(L, N, \mathbf{s}^*) = \sqrt{r_n(L, N, \mathbf{s}^*)} \log^\delta(n), \quad \text{for any } \delta > 1,$$

which is the estimation error from the statistical estimator for P_0 .

The above two lemmas together imply the following guarantee for VB posterior:

Theorem

Let σ_0^2 be a constant and $-\log \lambda = \log(T) + \delta[(L+1) \log N + \log \sqrt{np}]$ for any constant $\delta > 0$. We have with high probability

$$\int_{\Theta} d^2(P_\theta, P_0) \hat{q}(\theta) d\theta \leq C\epsilon_n^{*2} + C'(r_n^* + \xi_n^*),$$

where C is some positive constant and C' is any diverging sequence.

Implementation

Since it is impossible to reparameterize the discrete variable γ by a continuous system, we apply the Gumbel-softmax approximation (Maddison et al. 2017), and $\gamma_i \sim \text{Bern}(\phi_i)$ is approximated by $\tilde{\gamma}_i \sim \text{Gumbel-softmax}(\phi_i, \tau)$, where

$$\begin{aligned} \tilde{\gamma}_i &= (1 + \exp(-\eta_i/\tau))^{-1}, \quad \eta_i = \log \frac{\phi_i}{1 - \phi_i} + \log \frac{u_i}{1 - u_i}, \\ u_i &\sim \mathcal{U}(0, 1), \quad \tau > 0 \text{ is the temperature.} \end{aligned}$$

Algorithm 1 Variational inference for sparse BNN with normal slab distribution.

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1: parameters:  $\omega = (\mu, \sigma', \phi')$ ,
2: where  $\sigma_i = \log(1 + \exp(\sigma'_i))$ ,  $\phi_i = (1 + \exp(\phi'_i))^{-1}$ , for  $i = 1, \dots, T$ 
3: repeat
4:    $D^m \leftarrow$  Randomly draw a minibatch of size  $m$  from  $D$ 
5:    $\epsilon_i, u_i \leftarrow$  Randomly draw  $K$  samples from  $\mathcal{N}(0, 1)$  and  $\mathcal{U}(0, 1)$ 
6:    $\tilde{\Omega}^m(\omega) \leftarrow$  Use (1) with  $(D^m, \omega, \epsilon, u)$ ; Use  $\gamma$  in the forward pass
7:    $\nabla_\omega \tilde{\Omega}^m(\omega) \leftarrow$  Use (1) with  $(D^m, \omega, \epsilon, u)$ ; Use  $\tilde{\gamma}$  in the backward pass
8:    $\omega \leftarrow$  Update with  $\nabla_\omega \tilde{\Omega}^m(\omega)$  using gradient descent algorithms (e.g. SGD,
9:     RMSprop or Adam)
10: until convergence of  $\tilde{\Omega}^m(\omega)$ 
11: return  $\omega$ 

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Experiment

Consider the following sparse target function f_0 :

$$f_0(x_1, \dots, x_{200}) = \frac{7x_2}{1+x_2^2} + 5 \sin(x_3 x_4) + 2x_5, \quad \epsilon \sim \mathcal{N}(0, 1)$$

The λ_{opt} is chosen according to the main theorem.

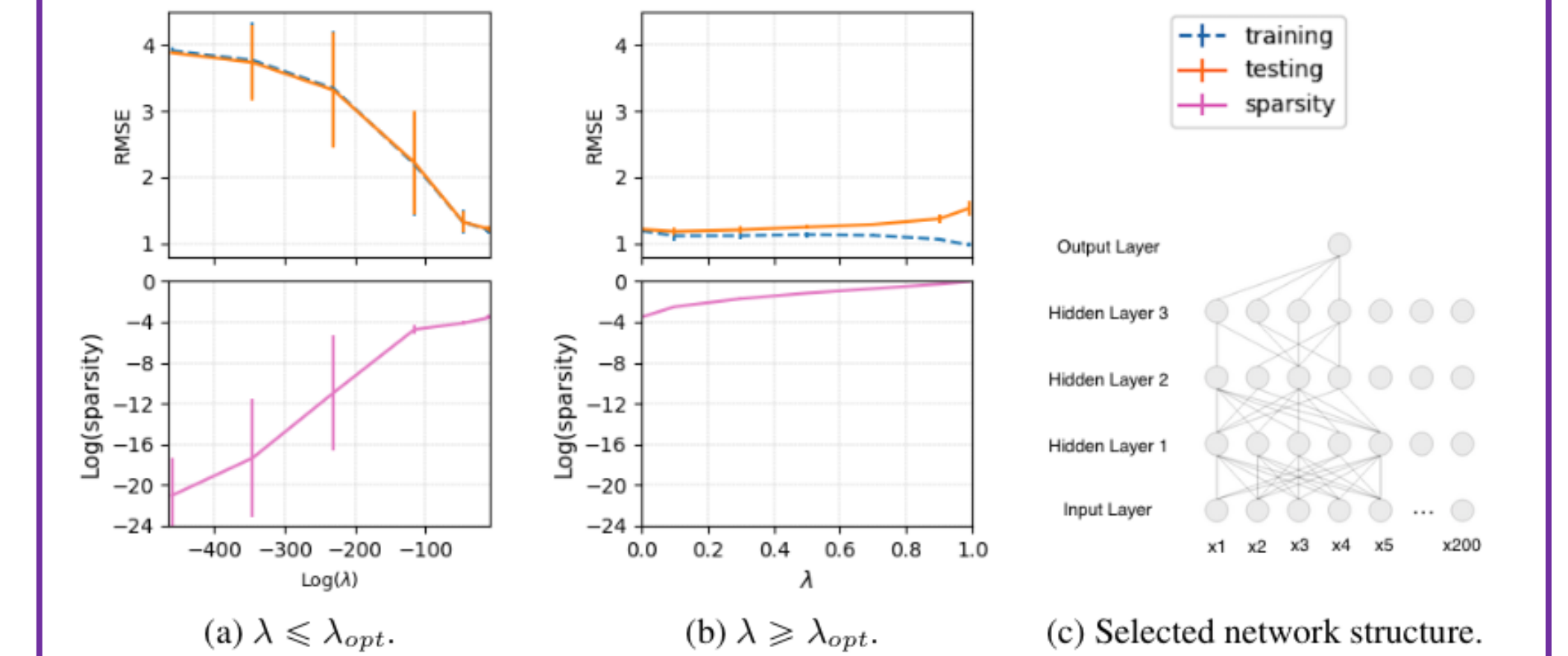


Figure 1. Nonlinear regression example.

Figure 1(a) and 1(b) show λ_{opt} is a reasonable choice and a possible network structure is provided in 1(c).

Selected references

- Ghosh, S., Yao, J., and Doshi-Velez, F. (2018). Structured variational learning of Bayesian neural networks with horseshoe priors. In ICML 2018.
- Maddison, C., Mnih, A., and Teh, Y. W. (2017). The concrete distribution: A continuous relaxation of discrete random variables. In ICLR 2017.
- Polson, N. and Rockova, V. (2018). Posterior concentration for sparse deep learning. In NeurIPS 2018.
- Zhu, M. and Gupta, S. (2018). To prune, or not to prune: Exploring the efficacy of pruning for model compression. In ICLR 2018.