Advanced Data Analysis -HW2

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1

```
library(openxlsx)
data.1 <- read.xlsx("~/Desktop/2019 Spring/Advanced Data Analysis/hw/HW2/HW2(Diets).xlsx")
a)
model.1.1 <- lm(Response~as.factor(Diet),data=data.1)</pre>
anova(model.1.1)
## Analysis of Variance Table
##
## Response: Response
                    Df Sum Sq Mean Sq F value Pr(>F)
## as.factor(Diet) 3 0.52652 0.175506 4.4891 0.01185 *
                   25 0.97740 0.039096
## Residuals
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
H_0: \mu_1 = \mu_2 = \mu_3 = \mu_4 \text{ vs } H_1: \text{Not } H_0
From the anova table above, we have p-value = 0.012 < 0.05, so we reject H_0 under significance level
\gamma = 0.05
In conclusion, there are at least two means that are not equal.
b)
t.test(Response~Diet, data=data.1[1:15,],conf.level=0.95)
##
##
   Welch Two Sample t-test
##
## data: Response by Diet
## t = 1.1625, df = 9.1239, p-value = 0.2745
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
## -0.1325412 0.4139698
## sample estimates:
## mean in group 1 mean in group 2
          3.745714
                            3.605000
##
Our test : H_0: \mu_1 - \mu_2 = 0
According to the test, we have 95% confidence interval for \mu 1 - \mu 2 is (-0.13,0.41)
```

```
c) & d)
```

```
mu_1 <- mean(data.1$Response[data.1$Diet==1])</pre>
mu_2 <- mean(data.1$Response[data.1$Diet==2])</pre>
mu_3 <- mean(data.1$Response[data.1$Diet==3])</pre>
mu_4 <- mean(data.1$Response[data.1$Diet==4])</pre>
n1 <- sum(data.1$Diet==1)</pre>
n2 <- sum(data.1$Diet==2)
n3 <- sum(data.1$Diet==3)
n4 <- sum(data.1$Diet==4)
L \leftarrow (mu_1+mu_2-mu_3-mu_4)/2
## [1] -0.08505952
# From the anova table in part a), we know that MSE is 0.0391
SE_L \leftarrow sqrt(0.0391*sum(0.25/n1+0.25/n2+0.25/n3+0.25/n4))
t \leftarrow qt(1-0.05/2,df=29-4)
# hyphothesis test
abs(L) > t*SE_L
## [1] FALSE
# confidence interval
L-SE_L*t
## [1] -0.2373727
L+SE L*t
## [1] 0.06725369
part c):
The confidence interval for L is: (-0.237, 0.067)
part d):
Our test: H_0: L = \frac{\mu_1 + \mu_2 - \mu_3 - \mu_4}{2} = 0 vs. H_1: L \neq 0
\hat{L} = \frac{\hat{\mu_1} + \hat{\mu_2} - \hat{\mu_3} - \hat{\mu_4}}{2} = -0.08
Since |\hat{L}| < t_{\gamma/2} (df = n - k = 29 - 4), we fail to reject H_0
2
a)
# Define dummy variables
D <- function(x,i) {</pre>
  temp <- rep(NA,length(x))
  for(j in 1:length(x)) {
    if (x[j] == i) \{temp[j] <- 1\}
    else if(x[j] == 4) {temp[j] <- -1}
```

```
else \{temp[j] \leftarrow 0\}
  }
  return(temp)
D1 <- D(data.1$Diet,1)
D2 <- D(data.1$Diet,2)
D3 <- D(data.1$Diet,3)
model.2.1 <- lm(data.1$Response~D1+D2+D3)</pre>
summary(model.2.1)
##
## Call:
## lm(formula = data.1$Response ~ D1 + D2 + D3)
##
## Residuals:
##
                  1Q Median
       Min
                                    3Q
                                           Max
## -0.3857 -0.0950 -0.0525 0.1250 0.4443
##
## Coefficients:
##
                Estimate Std. Error t value Pr(>|t|)
## (Intercept) 3.71789
                              0.03698 100.549
                                                  <2e-16 ***
## D1
                 0.02783
                              0.06450
                                         0.431
                                                  0.6698
## D2
                -0.11289
                              0.06173 -1.829
                                                  0.0794 .
## D3
                -0.11955
                              0.06801 - 1.758
                                                  0.0910 .
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.1977 on 25 degrees of freedom
## Multiple R-squared: 0.3501, Adjusted R-squared: 0.2721
## F-statistic: 4.489 on 3 and 25 DF, p-value: 0.01185
From the defination and the table above, we have:
                                              \hat{\mu} = 3.72
                                             \hat{\alpha}_1 = 0.028
                                            \hat{\alpha_2} = -0.113
                                            \hat{\alpha}_3 = -0.120
                      \hat{\alpha}_4 = -(\hat{\alpha}_1 + \hat{\alpha}_2 + \hat{\alpha}_3) = -(0.028 - 0.113 - 0.120) = 0.205
b)
model.2.2 <- lm(data.1$Response~1)
anova(model.2.2, model.2.1)
## Analysis of Variance Table
##
## Model 1: data.1$Response ~ 1
## Model 2: data.1$Response ~ D1 + D2 + D3
     Res.Df
                RSS Df Sum of Sq
                                     F Pr(>F)
```

```
## 1 28 1.5039  
## 2 25 0.9774 3 0.52652 4.4891 0.01185 *  
## ---  
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1  
Our test: H_0: \alpha_1 = \alpha_2 = \alpha_3 = 0 vs. H_\alpha: Not H_0
```

From the anova table, we have p-value=0.012 < 0.05, so reject H_0

The p-value is exactly the same as the value in Problem 1 and we reached the same conclusion.

3

a)

Source	df	SS	MS	F
feed	2	23.43	11.715	17.564
error	7	4.67	0.667	
total	9	28.10		

b)

```
p.value <- 1-pf(17.69,df1=2,df2=7)
p.value</pre>
```

[1] 0.001831386

Our test: $H_0: \mu_1 = \mu_2 = \mu_3$ vs. H_α : Not H_0

From the calculation above, we know that p-value = 0.0018 < 0.05, so reject H_0

Therefore, we know that different chichen feeds have different effect.