

report

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Introduction

Our client, Reshma Satish, a PhD student at Boston University's Department of Forensic Anthropology, administered a project involves fracturing 135 long deer bones using a standardized fracturing apparatus over a period of 540 days. The bones are broken into 9 groups with 15 bones in each group and each group is fractured at different times (0 days, 30 days, 60 days, 90 days, 180 days, 270 days, etc.) Her goal was to investigate how time of exposure of long bones affects variables such as the type of fracture angle, the type of fracture surface morphology, etc. which are created on the bones. Currently, our client only has data until day 360 since this experiment still keeps going.

This report will start with a description of the EDA, as well as Data cleaning/Data analysis/Data structure. Then we will describe the methods we will use on modeling part, followed by our analysis.

Data Cleaning and EDA

Data structure

We are provided the data in an excel document with 7 groups of drying days and a total of 105 bones. There are 10 variables in the dataset, but we will only focus on 8 of them, with Bisect Fracture and Fracture-like as binary responses, Number of fragments as a numerical response, and Fracture Angle, Fracture Type, Fracture Surface, and Fracture Outline as unordered categorical responses. We will create dummy variables for Fractographic Features, considering there can be multiple features at the same time.

Data Cleaning

At a first glance of the raw dataset, there are many NA inputs. After discussed with our client, we decide to change NA in Fracture Angle, Fracture Type, Fracture Surface, and Fracture Outline to non applicable and change NA in Fractographic Features to No Feature. Moreover, we create numerical variable drying days from variable Bone Specimen for future analysis.

Create Dummy Variable

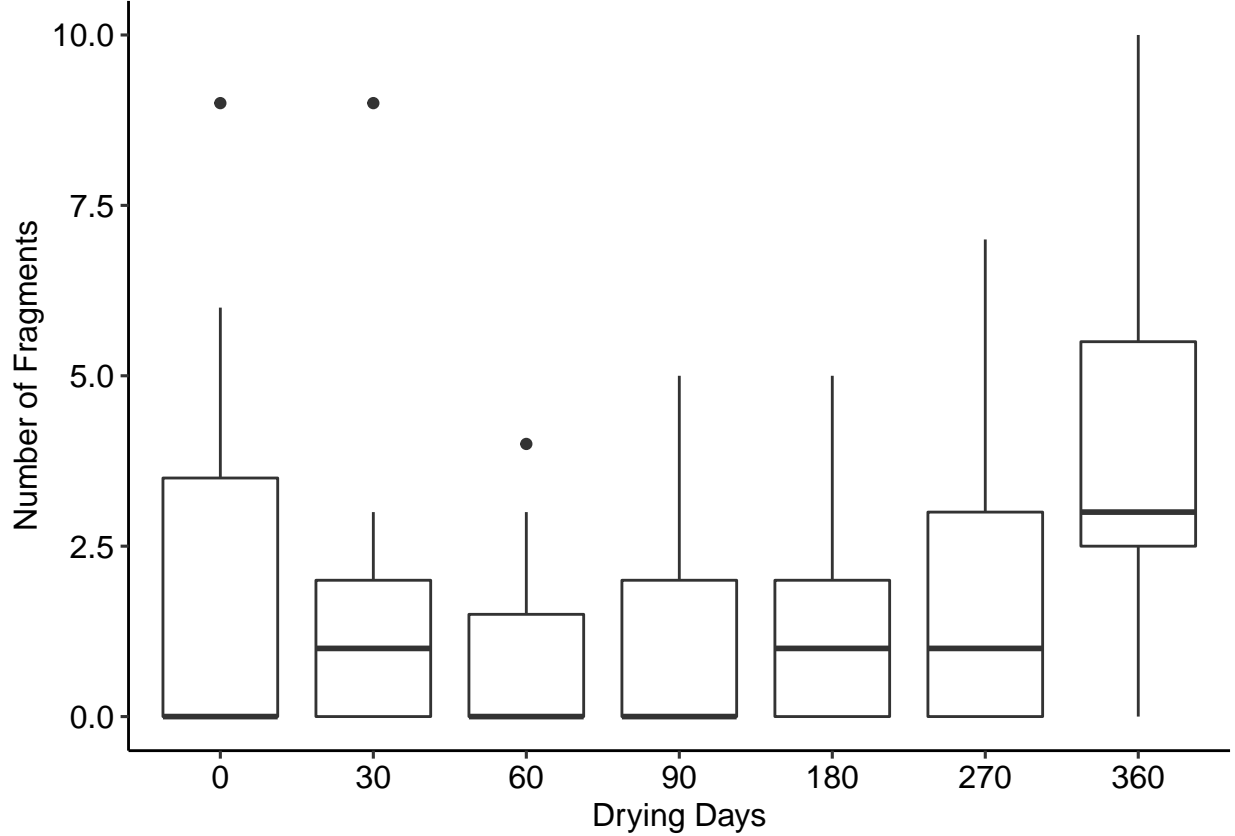
There are some observations have multiple features in Fractographic Features variable. Thus, we create 6 dummy variables to indicate which feature appears for a given observation.

EDA

We conduct a basic exploratory data analysis (EDA) to visualize the data and find some patterns between variables. We focus on the two binary responses and number of fragments.

Table 1: Bisecting Fracture and Fracture-like percentage for different drying-day groups

	0	30	60	90	180	270	360
Bisecting percentage	0.47	0.6	0.67	0.47	0.67	0.60	0.87
Fracture-like percentage	0.47	0.8	0.73	0.53	0.73	0.93	0.93



In the table, we give a bisecting percentage and fracture-like percentage for each groups of bones in different drying days. We can see a pattern that with the drying days increasing, it's more likely to have a bisecting or fracture-like happened during the standardized fracturing apparatus.

In the boxplot, we can see several outliers in the first three groups and large amount of zero value. It's hard to figure a clear pattern in the first four groups, but later on, number of fragments are increased as the drying days increased.

Modeling & Interpretation

We used logistic regression, poisson regression, negative binomial, and multinomial logit model based on different types of response variables in our dataset.

Fracture Type

COMMINUTED relative to BUTTERFULY:

Intercept: This is the multinomial logit estimate for COMMINUTED relative to BUTTERFULY when the predictor variables in the model are evaluated at zero. For day equals to zero, the logit for preferring

COMMINUTED relative to BUTTERFULY is 0.8269632. Note that evaluating days equals to zero indicated that the day of starting the experiment.

Days: This is the multinomial logit estimate for a one unit increase in Days for COMMINUTED relative to BUTTERFULY. If a subject was to increase Days by one point, the multinomial log-odds of preferring COMMINUTED relative to BUTTERFULY would be expected to increase by 10.00216.

T-statistics for Days in $\text{Log}(\text{Prob}(\text{COMMINUTED})/\text{Prob}(\text{BUTTERFULY}))$ equals to 0.8158939, which is lower than 2. Then, the Days does not make any difference for log-odds of preferring COMMINUTED relative to BUTTERFULY

non applicable relative to BUTTERFULY

Intercept: This is the multinomial logit estimate for non applicable relative to BUTTERFULY when the predictor variables in the model are evaluated at zero. For day equals to zero, the logit for preferring non applicable relative to BUTTERFULY is 1.5166589. Note that evaluating days equals to zero indicated that the day of starting the experiment

Days: This is the multinomial logit estimate for a one unit increase in Days for non applicable relative to BUTTERFULY. If a subject was to increase Days by one point, the multinomial log-odds of preferring non applicable relative to BUTTERFULY would be expected to decrease by 0.0024548581.

T-statistics for days in $\text{Log}(\text{Prob}(\text{non applicable})/\text{Prob}(\text{BUTTERFULY}))$ equals to $|-0.0024548581/0.002647403| = 0.8948441$, which is lower than 2. Then, the Days does not make any difference for log-odds of preferring non applicable relative to BUTTERFULY

$$\text{Log}(\text{Prob}(\text{OBLIQUE})/\text{Prob}(\text{BUTTERFULY})) = -0.1913727 - 0.0027752170 * \text{Days}$$

OBLIQUE relative to BUTTERFULY:

Intercept: This is the multinomial logit estimate for OBLIQUE relative to BUTTERFULY when the predictor variables in the model are evaluated at zero. For day equals to zero, the logit for preferring OBLIQUE relative to BUTTERFULY is -0.1913727. Note that evaluating days equals to zero indicated that the day of starting the experiment Days: This is the multinomial logit estimate for a one unit increase in Days for OBLIQUE relative to BUTTERFULY. If a subject were to increase Days by one point, the multinomial log-odds of preferring OBLIQUE relative to BUTTERFULY would be expected to decrease by 0.0027752170.

T-statistics for days in $\text{Log}(\text{Prob}(\text{OBLIQUE})/\text{Prob}(\text{BUTTERFULY}))$ equals to $|-0.0027752170 / 0.004192975| = 0.661873$, which is lower than 2. Then, the Days does not make any difference for log-odds of preferring OBLIQUE relative to BUTTERFULY

$$\text{Log}(\text{Prob}(\text{SEGMENTAL})/\text{Prob}(\text{BUTTERFULY})) = -0.5994698 - 0.0006693986 * \text{Days}$$

SEGMENTAL relative to BUTTERFULY:

Intercept: This is the multinomial logit estimate for SEGMENTAL relative to BUTTERFULY when the predictor variables in the model are evaluated at zero. For day equals to zero, the logit for preferring SEGMENTAL relative to BUTTERFULY is -0.5994698. Note that evaluating days equals to zero indicated that the day of starting the experiment Days: This is the multinomial logit estimate for a one unit increase in Days for SEGMENTAL relative to BUTTERFULY. If a subject were to increase Days by one point, the multinomial log-odds of preferring SEGMENTAL relative to BUTTERFULY would be expected to decrease by -0.0006693986.

T-statistics for days in $\text{Log}(\text{Prob}(\text{SEGMENTAL})/\text{Prob}(\text{BUTTERFULY}))$ equals to $|-0.0006693986/0.004110398| = 0.1628549$, which is lower than 2. Then, the Days does not make any difference for log-odds of preferring SEGMENTAL relative to BUTTERFULY.

$$\text{Log}(\text{Prob}(\text{TRANSVERSE})/\text{Prob}(\text{BUTTERFULY})) = -1.9804151 - 0.004371236 * \text{Days}$$

TRANSVERSE relative to BUTTERFULY:

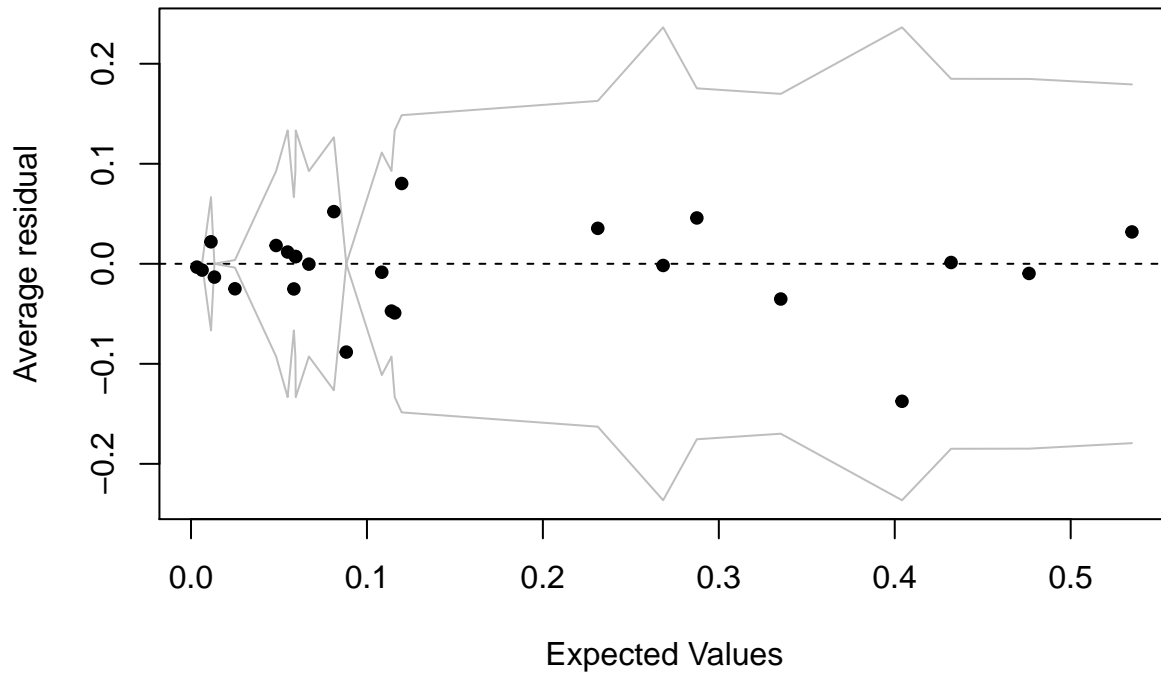
Intercept : This is the multinomial logit estimate for TRANSVERSE relative to BUTTERFULY when the predictor variables in the model are evaluated at zero. For day equals to zero, the logit for preferring TRANSVERSE relative to BUTTERFULY is -1.9804151. Note that evaluating days equals to zero indicated that the day of starting the experiment Days: This is the multinomial logit estimate for a one unit increase in Days for TRANSVERSE relative to BUTTERFULY. If a subject were to increase Days by one point, the multinomial log-odds of preferring TRANSVERSE relative to BUTTERFULY would be expected to decrease by 0.004371236.

T-statistics for days in $\text{Log}(\text{Prob}(\text{TRANSVERSE})/\text{Prob}(\text{BUTTERFULY}))$ equals to $|-0.0043712363/0.010436283| = 0.4188499$, which is lower than 2. Then, the Days does not make any difference for log-odds of preferring TRANSVERSE relative to BUTTERFULY

According to the Binned residual plot, there is not clearly pattern.

```
## Call:
## multinom(formula = bone$`Fracture Type` ~ bone$days, trace = FALSE)
##
## Coefficients:
##              (Intercept) bone$days
## COMMINUTED           0.83   0.00217
## non applicable        1.52  -0.00245
## OBLIQUE              -0.19  -0.00278
## SEGMENTAL            -0.60  -0.00067
## TRANSVERSE          -1.98  -0.00437
##
## Std. Errors:
##              (Intercept) bone$days
## COMMINUTED           0.52   0.0026
## non applicable        0.50   0.0027
## OBLIQUE              0.69   0.0042
## SEGMENTAL            0.76   0.0041
## TRANSVERSE          1.43   0.0104
##
## Residual Deviance: 281
## AIC: 301
```

Binned residual plot



##	(Intercept)	bone\$days
## COMMINUTED	1.58	0.82
## non applicable	3.04	-0.89
## OBLIQUE	-0.28	-0.66
## SEGMENTAL	-0.79	-0.16
## TRANSVERSE	-1.39	-0.42

Fracture Surface

non applicable relative to INTERMEDIATE:

$\text{Log}(\text{Prob}(\text{non applicable})/\text{Prob}(\text{INTERMEDIATE})) = 2.185304 - 0.008706689 \cdot \text{Days}$

Intercept: This is the multinomial logit estimate for non applicable relative to INTERMEDIATE when the predictor variables in the model are evaluated at zero. For day equals to zero, the logit for preferring non applicable relative to INTERMEDIATE is 2.185304. Note that evaluating days equals to zero indicated that the day of starting the experiment Days: This is the multinomial logit estimate for a one unit increase in Days for non applicable relative to INTERMEDIATE. If a subject were to increase Days by one point, the multinomial log-odds of preferring non applicable relative to INTERMEDIATE would be expected to decrease by 0.008706689.

T-statistics for days in $\text{Log}(\text{Prob}(\text{non applicable})/\text{Prob}(\text{INTERMEDIATE}))$ equals to $|0.008706689/0.002536340| = 3.432777$, which is greater than 2. Then, the Days make a significant difference for log-odds of preferring non applicable relative to INTERMEDIATE

ROUGH relative to INTERMEDIATE:

$\text{Log}(\text{Prob}(\text{ROUGH})/\text{Prob}(\text{INTERMEDIATE})) = -23.616556 + 0.060230466 \cdot \text{Days}$

Intercept: This is the multinomial logit estimate for ROUGH relative to INTERMEDIATE when the predictor variables in the model are evaluated at zero. For day equals to zero, the logit for preferring ROUGH relative to INTERMEDIATE is -23.616556. Note that evaluating days equals to zero indicated that the day of starting the experiment Days: This is the multinomial logit estimate for a one unit increase in Days for ROUGH relative to INTERMEDIATE. If a subject were to increase Days by one point, the multinomial log-odds of preferring ROUGH relative to INTERMEDIATE would be expected to increase by 0.060230466.

T-statistics for days in $\text{Log}(\text{Prob}(\text{ROUGH})/\text{Prob}(\text{INTERMEDIATE}))$ equals to $|0.060230466/0.002940204| = 20.48513$, which is greater than 2. Then, the Days make a significant difference for log-odds of preferring ROUGH relative to INTERMEDIATE

SMOOTH relative to INTERMEDIATE:

$\text{Log}(\text{Prob}(\text{SMOOTH})/\text{Prob}(\text{INTERMEDIATE})) = 2.255591 - 0.007492372 * \text{Days}$

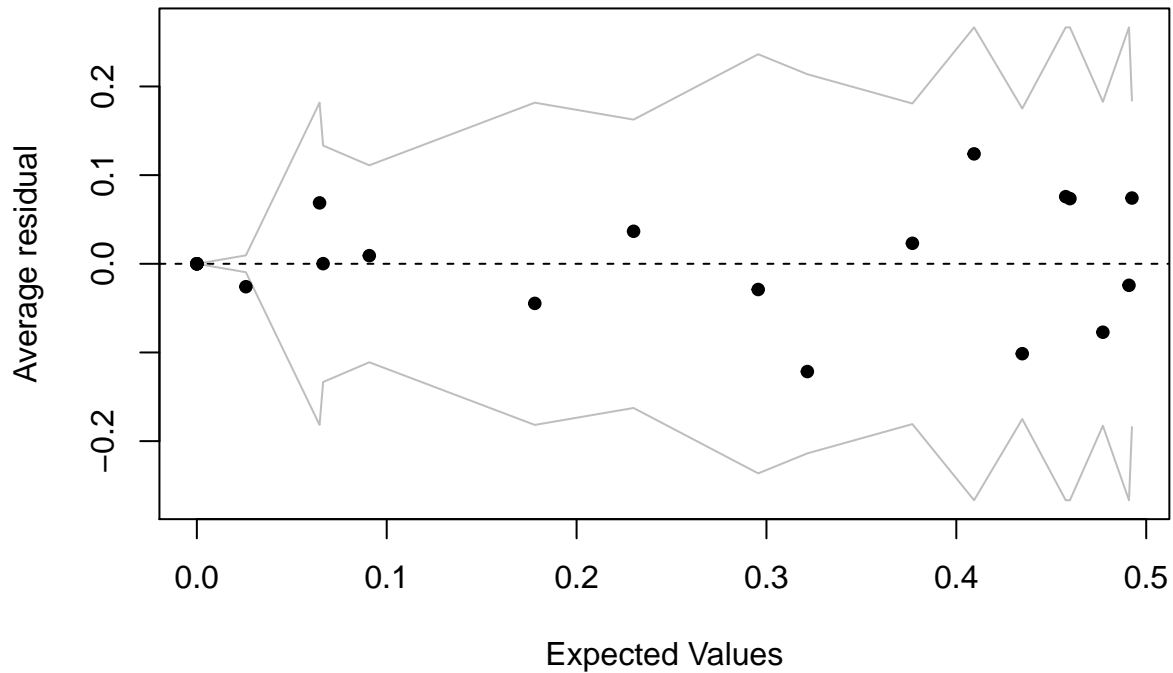
Intercept: This is the multinomial logit estimate for SMOOTH relative to INTERMEDIATE when the predictor variables in the model are evaluated at zero. For day equals to zero, the logit for preferring SMOOTH relative to INTERMEDIATE is 2.255591. Note that evaluating days equals to zero indicated that the day of starting the experiment Days: This is the multinomial logit estimate for a one unit increase in Days for SMOOTH relative to INTERMEDIATE. If a subject were to increase Days by one point, the multinomial log-odds of preferring SMOOTH relative to INTERMEDIATE would be expected to decrease by 0.007492372.

T-statistics for days in $\text{Log}(\text{Prob}(\text{SMOOTH})/\text{Prob}(\text{INTERMEDIATE}))$ equals to $|-0.007492372/0.002378579| = 3.149936$, which is greater than 2. Then, the Days make a significant difference for log-odds of preferring SMOOTH relative to INTERMEDIATE

According to the Binned residual plot, there is not clearly pattern.

```
## Call:
## multinom(formula = bone$`Fracture Surface` ~ bone$days, trace = FALSE)
##
## Coefficients:
##              (Intercept) bone$days
## non applicable          2.2   -0.0087
## ROUGH                 -23.6    0.0602
## SMOOTH                 2.3   -0.0075
##
## Std. Errors:
##              (Intercept) bone$days
## non applicable    5.7e-01    0.0025
## ROUGH             3.6e-05    0.0029
## SMOOTH             5.6e-01    0.0024
##
## Residual Deviance: 208
## AIC: 220
```

Binned residual plot



```
##               (Intercept) bone$days
## non applicable    3.8e+00    -3.4
## ROUGH            -6.6e+05    20.5
## SMOOTH           4.0e+00    -3.1
```

Bisect Fracture

Since response variable “bisect fracture” is a binary outcomes(0 or 1), we are using logistic regression to figure out relationship between the probability of “bisect fracture” and value of “days”.

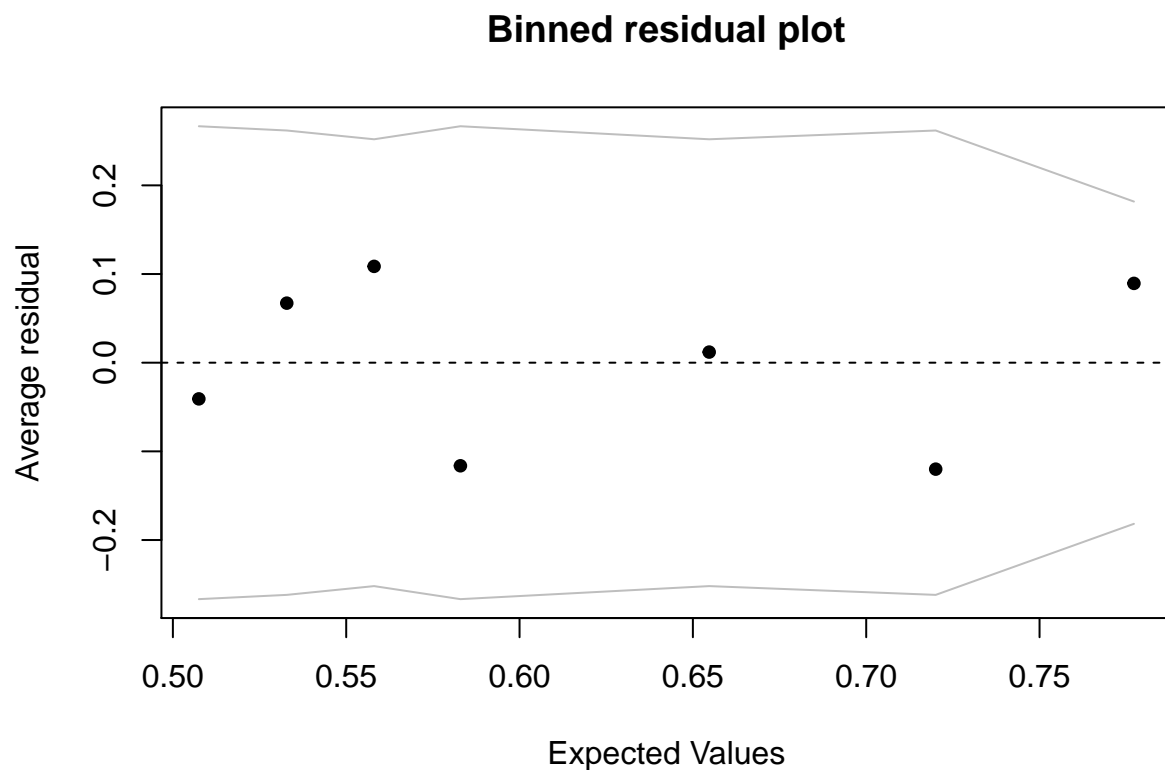
In addition, when “bisect fracture” is 0, it means the bone did not break by the standard fracturing machine, otherwise, it means the bone broke by the machine.

```
##
## Call:
## glm(formula = bone$`Bisecting Fracture` ~ bone$days, family = binomial())
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -1.73    -1.23     0.71     1.04     1.16
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)  0.02996   0.30127   0.10   0.921
## bone$days    0.00339   0.00174   1.94   0.052 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
```

```
## (Dispersion parameter for binomial family taken to be 1)
##
## Null deviance: 139.55 on 104 degrees of freedom
## Residual deviance: 135.53 on 103 degrees of freedom
## AIC: 139.5
##
## Number of Fisher Scoring iterations: 4
```

From the output, the estimated intercept coefficient 0.029 correspond to the log odds of broke the bone with average days, in other words, on average days, expected probability of broke the bone is $\text{expit}(0.029) = .507 = 50.7\%$. And one-unit increase in the predictor “days” increases the log odds of broke the bone (versus not broke the bone) by 0.0034.

Further, p-value of days is 0.052 indicates that there is no strong evident but moderate evidence in support of the differences of the bone broke by the machine or not affected by drying time.



Since the discrete residual plot does not tell us much about the model fit, we are using Binned residual plot for this model. Due to every points fall within the bounds, our model fits well in this case.

Fracture like

Since “Fracture-like” is a binary outcomes (0 or 1) as well, therefore, as idea with above part, we used logistic regression to figure out their relationship.

In addition, when “Fracture-like” is 0, it means the bone has no surface feature after cut by the standard fracturing machine, otherwise, it means fracturing machine made a mark on the bone.

```
##
## Call:
## glm(formula = bone$`Fracture-like` ~ bone$days, family = binomial())
```



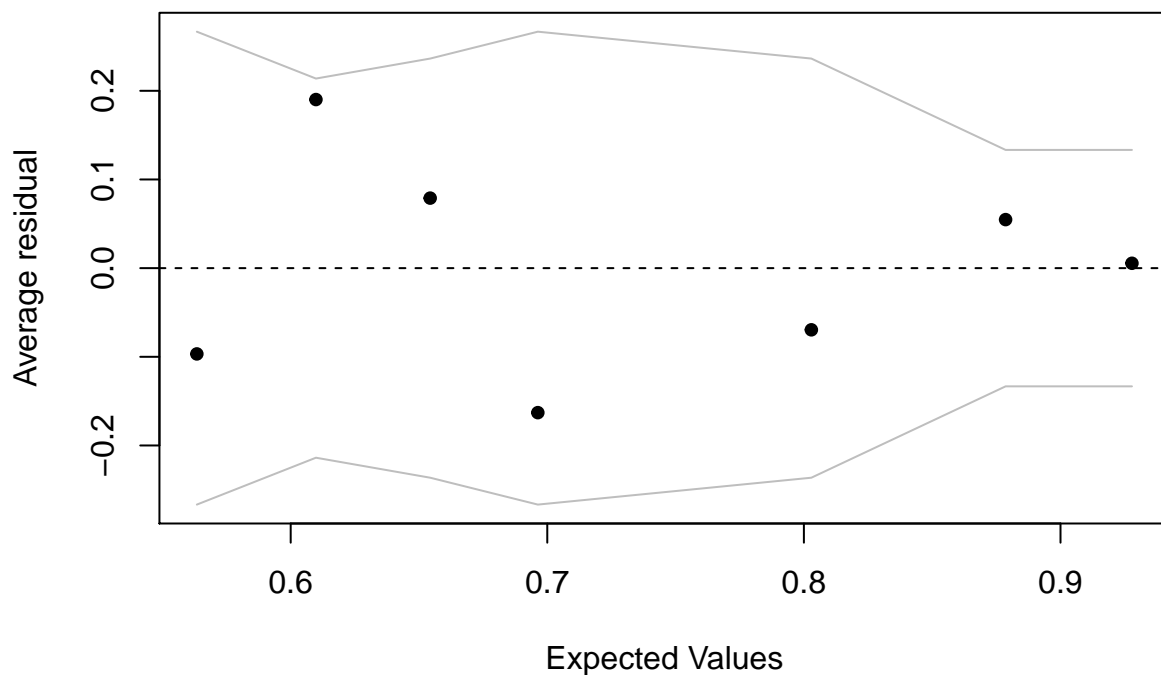
```
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -2.293  -1.288   0.509   0.921   1.071
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)  0.25510    0.32022   0.80  0.4257
## bone$days    0.00639    0.00230   2.78  0.0055 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 121.78  on 104  degrees of freedom
## Residual deviance: 112.13  on 103  degrees of freedom
## AIC: 116.1
##
## Number of Fisher Scoring iterations: 4
```

From the output, the estimated intercept coefficient 0.2551 correspond to the log odds of made a mark on bone with average days, in other words, on average days, expected probability of made a mark on bone is $\text{expit}(0.2551) = 0.563 = 56.3\%$.

And one-unit increase in the predictor “days” increases the log odds of made a mark on bone (versus no surface feature on bone) by 0.0064.

Also, p-value of days is 0.0055 indicates that there is strong evidence in determining whether the bone has any surface feature after cut by the machine affected by drying time.

Binned residual plot



Same logic with above part, our model fit well in this case.

Fracture Angle

As “Fracture Angle” as no natural-order variable(nominal), since we cannot compare the order of angle of fresh bone and dry bone, and also N/A(non applicable) in this variable, so we used multinomial logit model in this case, that the outcome variable in this model has more than 2 levels.

```
## Call:
## multinom(formula = bone$`Fracture Angle` ~ bone$days, trace = FALSE)
##
## Coefficients:
##              (Intercept) bone$days
## non applicable          2.2   -0.0057
## OBLIQUE                2.2   -0.0030
## RIGHT ANGLES          -1.2   -0.0077
##
## Std. Errors:
##              (Intercept) bone$days
## non applicable          0.63    0.0029
## OBLIQUE                0.62    0.0027
## RIGHT ANGLES           1.48    0.0104
##
## Residual Deviance: 200
## AIC: 212
```

From the output, the baseline is Fracture Angle = Intermediate.

A one-unit increase in the variable “days” is associated with the decrease in the log odds of being in Intermediate fracture angle vs. non applicable fracture angle in the amount of 0.0057.

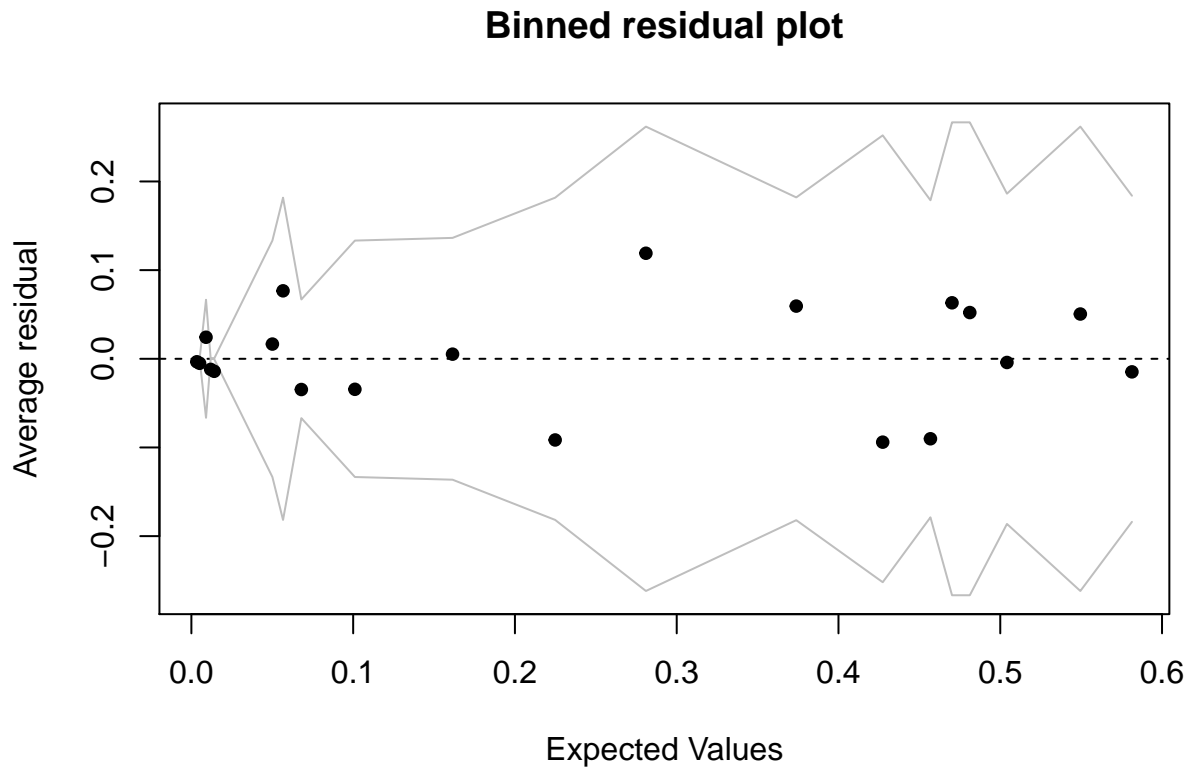
A one-unit increase in the variable “days” is associated with the decrease in the log odds of being in Intermediate fracture angle vs. oblique fracture angle in the amount of 0.003.

In other words, if one-unit increase in “days”, your chances of staying in the intermediate fracture angle category are higher compared to staying in non applicable and oblique.

Since summary of multinomial logit model does not provide p-values, we can get significance of the coefficients by calculating z score and input it into the formula: $z = (x_i - \bar{x})/s$ where coefficients of summary represents $x_i - \bar{x}$, and Std.Errors of summary represents s .

```
##              (Intercept) bone$days
## non applicable    0.00037    0.047
## OBLIQUE          0.00034    0.266
## RIGHT ANGLES     0.41049    0.458
```

From output, it represents p-value of coefficients. Only non applicable has strong evidence in determining whether affected by drying time because of its p-value is $0.047 < 5\%$.



It looks like our model fit kind of well in this case, but there might not good enough when expected values small.

Number of Fragments

Considering “Number of Fragments” as a numerical variable, we use the poisson regression at first.

```
fit.Number_of_Fragments1 <- glm(bone$`Number of Fragments` ~ bone$days, family="poisson")
summary(fit.Number_of_Fragments1) #estimated coefficients
```

```
##
## Call:
## glm(formula = bone$`Number of Fragments` ~ bone$days, family = "poisson")
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -2.480  -1.685  -0.741   0.649   4.522
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)  0.196437   0.123824   1.59    0.11
## bone$days    0.002574   0.000549   4.69 2.8e-06 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for poisson family taken to be 1)
##
##      Null deviance: 296.20  on 104  degrees of freedom
```

```
## Residual deviance: 274.71 on 103 degrees of freedom
## AIC: 452
##
## Number of Fisher Scoring iterations: 6
```

```
pchisq(fit.Number_of_Fragments1$deviance,df=fit.Number_of_Fragments1$df.residual, lower.tail=FALSE)
```

```
## [1] 1.5e-17
```

```
regout <- glm(bone$`Number of Fragments`~bone$days, family=quasipoisson)
summary.glm(regout)$dispersion
```

```
## [1] 2.9
```

Here our null hypothesis is that our model is correctly specified. The low p-value shows that we can reject this hypothesis, thus the poisson regression model fits poorly. Next, we check the dispersion of the data. We can see the variance of random component is roughly 3 times the size of the mean. There may be a over-dispersion problem. So we consider to use a negative binomial model.

```
fit.Number_of_Fragments2 <- glm.nb(bone$`Number of Fragments` ~ bone$days)
summary(fit.Number_of_Fragments2)
```

```
##
## Call:
## glm.nb(formula = bone$`Number of Fragments` ~ bone$days, init.theta = 0.833740805,
## link = log)
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -1.579  -1.301  -0.433   0.350   2.204
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)  0.24931    0.20348   1.23   0.220
## bone$days    0.00225    0.00103   2.18   0.029 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for Negative Binomial(0.83) family taken to be 1)
##
##      Null deviance: 117.06 on 104 degrees of freedom
## Residual deviance: 111.39 on 103 degrees of freedom
## AIC: 386.7
##
## Number of Fisher Scoring iterations: 1
##
##
##              Theta:  0.834
##             Std. Err.: 0.212
##
## 2 x log-likelihood: -380.656
```

```
pchisq(fit.Number_of_Fragments2$deviance,df=fit.Number_of_Fragments2$df.residual, lower.tail=FALSE)
```

```
## [1] 0.27
```

Using the negative binomial model, we have a better fit. The only variable “days” has a coefficient of 0.002247, which is statistically significant. This means that for each one-unit increase in “days”, the expected log count

of the number of fragments increases by 0.002247. It can be also interpreted as for each 100 unit increase in “days”, the expected count of the number of fragments will be increased by 1.25 times.

Fracture Outline

We assume “Fracture Outline” as no natural-order(nominal) variable, same reason as “Fracture Angle” part, thus we used multinomial logit model.

```
##                (Intercept) bone$days
## INTERMEDIATE      -1.71    -0.0155
## non applicable      0.13    -0.0043
## TRANSVERSE       -3.38    -0.0059
```

From the output, the baseline is Fracture Angle = Cruved/Jagged.

A one-unit increase in the variable “days” is associated with the decrease in the log odds of being in curved fracture outline vs. intermediate fracture outline in the amount of 0.0155.

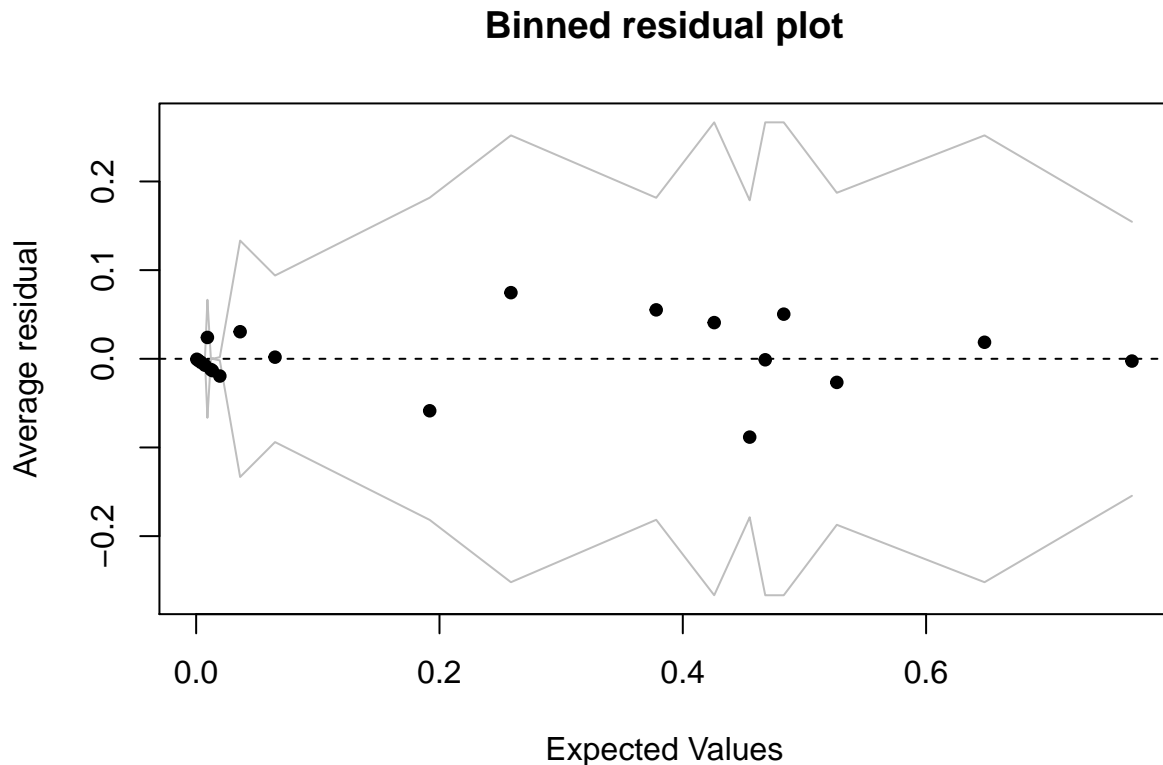
A one-unit increase in the variable “days” is associated with the decrease in the log odds of being in curved fracture outline vs. non applicable fracture outline in the amount of 0.0043.

In other words, if one-unit increase in “days”, your chances of staying in the curved/jagged fracture outline category are higher compared to staying in non applicable, intermediate and transverse.

Get significance of the coefficients by calculating z score and input it into the formula:

```
##                (Intercept) bone$days
## INTERMEDIATE      0.027     0.184
## non applicable      0.686     0.017
## TRANSVERSE        0.014     0.565
```

From output, it represents p-value of coefficients. Only the non applicable has strong evidence to determining whether it affected by drying time because of its p value is $0.017 < 5\%$.



Looks like our model fit kind of well in this case, except there might not good enough when expected values really small.

Fractographic Features

We have 6 different features, since some of features appear at the same time, therefore we create 6 dummy variables for different features, and build a logistic regression model for each feature to figuring out their relationship. Then, we will use chi-square test to evaluate whether 2 features are correlated.

Chi-square test for Independence

Null hypothesis of chi-square: 2 variables/features are independent with each other.

```
tb.1 = table(bone$`WAKE HACKLE`, bone$`BONE HACKLE`)
chisq.test(tb.1)
```

```
## Warning in chisq.test(tb.1): Chi-squared approximation may be incorrect
```

```
##
## Pearson's Chi-squared test with Yates' continuity correction
##
## data:  tb.1
## X-squared = 12, df = 1, p-value = 5e-04
```

Since $p\text{-value } 0.0005 < 0.01$ significant level, therefore, we have strong evidence to reject the null hypothesis. Thus, “WAKE HACKLE” related to “BONE HACKLE”.

```
tb.2 = table(bone$`WAKE HACKLE`, bone$`BONE MIRROR`)
chisq.test(tb.2)
```

```
## Warning in chisq.test(tb.2): Chi-squared approximation may be incorrect
```

```
##  
## Pearson's Chi-squared test with Yates' continuity correction  
##  
## data: tb.2  
## X-squared = 5e-30, df = 1, p-value = 1
```

Since p-value $1 > 0.05$ significant level, therefore, we have strong evidence to retain the null hypothesis. Thus, “WAKE HACKLE” is independent of “BONE MIRROR”.

```
tb.3 = table(bone$`WAKE HACKLE`, bone$`CANTILEVER CURL`)  
chisq.test(tb.3)
```

```
## Warning in chisq.test(tb.3): Chi-squared approximation may be incorrect
```

```
##  
## Pearson's Chi-squared test with Yates' continuity correction  
##  
## data: tb.3  
## X-squared = 0.1, df = 1, p-value = 0.7
```

Since p-value $0.7 > 0.05$ significant level, therefore, we have strong evidence to retain the null hypothesis. Thus, “WAKE HACKLE” is independent of “COMPRESSIONAL CURL”.

```
tb.4 = table(bone$`WAKE HACKLE`, bone$`ARREST RIDGE`)  
chisq.test(tb.4)
```

```
## Warning in chisq.test(tb.4): Chi-squared approximation may be incorrect
```

```
##  
## Pearson's Chi-squared test with Yates' continuity correction  
##  
## data: tb.4  
## X-squared = 5, df = 1, p-value = 0.03
```

Since p-value $0.03 < 0.05$ significant level, therefore, we have somewhat strong evidence to reject the null hypothesis. Thus, “WAKE HACKLE” is related to “ARREST RIDGE”.

```
tb.5 = table(bone$`WAKE HACKLE`, bone$`WALLNER LINES`)  
chisq.test(tb.5)
```

```
## Warning in chisq.test(tb.5): Chi-squared approximation may be incorrect
```

```
##  
## Pearson's Chi-squared test with Yates' continuity correction  
##  
## data: tb.5  
## X-squared = 4, df = 1, p-value = 0.05
```

Since p-value $0.05 = 0.05$ significant level, therefore, we have somewhat strong evidence to retain the null hypothesis. Thus, “WAKE HACKLE” is independent of “WALLNER LINES”.

```
tb.6 = table(bone$`BONE HACKLE`, bone$`BONE MIRROR`)  
chisq.test(tb.6)
```

```
## Warning in chisq.test(tb.6): Chi-squared approximation may be incorrect
```

```
##  
## Pearson's Chi-squared test with Yates' continuity correction  
##
```

```
## data:  tb.6
## X-squared = 7e-31, df = 1, p-value = 1
```

Since p-value $1 > 0.05$ significant level, therefore, we have strong evidence to retain the null hypothesis. Thus, “BONE HACKLE” is independent of “BONE MIRROR”.

```
tb.7 = table(bone$`BONE HACKLE`, bone$`CANTILEVER CURL`)
chisq.test(tb.7)
```

```
## Warning in chisq.test(tb.7): Chi-squared approximation may be incorrect
```

```
##
## Pearson's Chi-squared test with Yates' continuity correction
##
## data:  tb.7
## X-squared = 3e-30, df = 1, p-value = 1
```

Since p-value $1 > 0.05$ significant level, therefore, we have strong evidence to retain the null hypothesis. Thus, “BONE HACKLE” is independent of “CANTILEVER CURL”.

```
tb.8 = table(bone$`BONE HACKLE`, bone$`ARREST RIDGE`)
chisq.test(tb.8)
```

```
## Warning in chisq.test(tb.8): Chi-squared approximation may be incorrect
```

```
##
## Pearson's Chi-squared test with Yates' continuity correction
##
## data:  tb.8
## X-squared = 8, df = 1, p-value = 0.005
```

Since p-value $0.005 < 0.01$ significant level, therefore, we have strong evidence to reject the null hypothesis. Thus, “BONE HACKLE” is related to “ARREST RIDGE”.

```
tb.9 = table(bone$`BONE HACKLE`, bone$`WALLNER LINES`)
chisq.test(tb.9)
```

```
## Warning in chisq.test(tb.9): Chi-squared approximation may be incorrect
```

```
##
## Pearson's Chi-squared test with Yates' continuity correction
##
## data:  tb.9
## X-squared = 2e-30, df = 1, p-value = 1
```

Since p-value $1 > 0.05$ significant level, therefore, we have strong evidence to retain the null hypothesis. Thus, “BONE HACKLE” is independent of “WALLER LINES”.

```
tb.10 = table(bone$`BONE MIRROR`, bone$`CANTILEVER CURL`)
chisq.test(tb.10)
```

```
## Warning in chisq.test(tb.10): Chi-squared approximation may be incorrect
```

```
##
## Pearson's Chi-squared test with Yates' continuity correction
##
## data:  tb.10
## X-squared = 2e-30, df = 1, p-value = 1
```

```
tb.11 = table(bone$`BONE MIRROR`, bone$`ARREST RIDGE`)
chisq.test(tb.11)
```



```
## Warning in chisq.test(tb.11): Chi-squared approximation may be incorrect
```

```
##  
## Pearson's Chi-squared test with Yates' continuity correction  
##  
## data:  tb.11  
## X-squared = 2e-30, df = 1, p-value = 1
```

```
tb.12 = table(bone$`BONE MIRROR`, bone$`WALLNER LINES`)  
chisq.test(tb.12)
```

```
## Warning in chisq.test(tb.12): Chi-squared approximation may be incorrect
```

```
##  
## Pearson's Chi-squared test with Yates' continuity correction  
##  
## data:  tb.12  
## X-squared = 1e-30, df = 1, p-value = 1
```

Since triple p-value $1 > 0.05$ significant level, therefore, we have strong evidence to retain the null hypothesis. Thus, “BONE MIRROR” is independent of “CANTILEVER CURL”, “ARREST RIDGE” and “WALLNER LINES”.

```
tb.13 = table(bone$`CANTILEVER CURL`, bone$`ARREST RIDGE`)  
chisq.test(tb.13)
```

```
## Warning in chisq.test(tb.13): Chi-squared approximation may be incorrect
```

```
##  
## Pearson's Chi-squared test with Yates' continuity correction  
##  
## data:  tb.13  
## X-squared = 2e-30, df = 1, p-value = 1
```

```
tb.14 = table(bone$`CANTILEVER CURL`, bone$`WALLNER LINES`)  
chisq.test(tb.14)
```

```
## Warning in chisq.test(tb.14): Chi-squared approximation may be incorrect
```

```
##  
## Pearson's Chi-squared test with Yates' continuity correction  
##  
## data:  tb.14  
## X-squared = 1e-30, df = 1, p-value = 1
```

Since both p-value $1 > 0.05$ significant level, therefore, we have strong evidence to retain the null hypothesis. Thus, “CANTILEVER CURL” is independent of “ARREST RIDGE” and “WALLNER LINES”.

```
tb.15 = table(bone$`ARREST RIDGE`, bone$`WALLNER LINES`)  
chisq.test(tb.15)
```

```
## Warning in chisq.test(tb.15): Chi-squared approximation may be incorrect
```

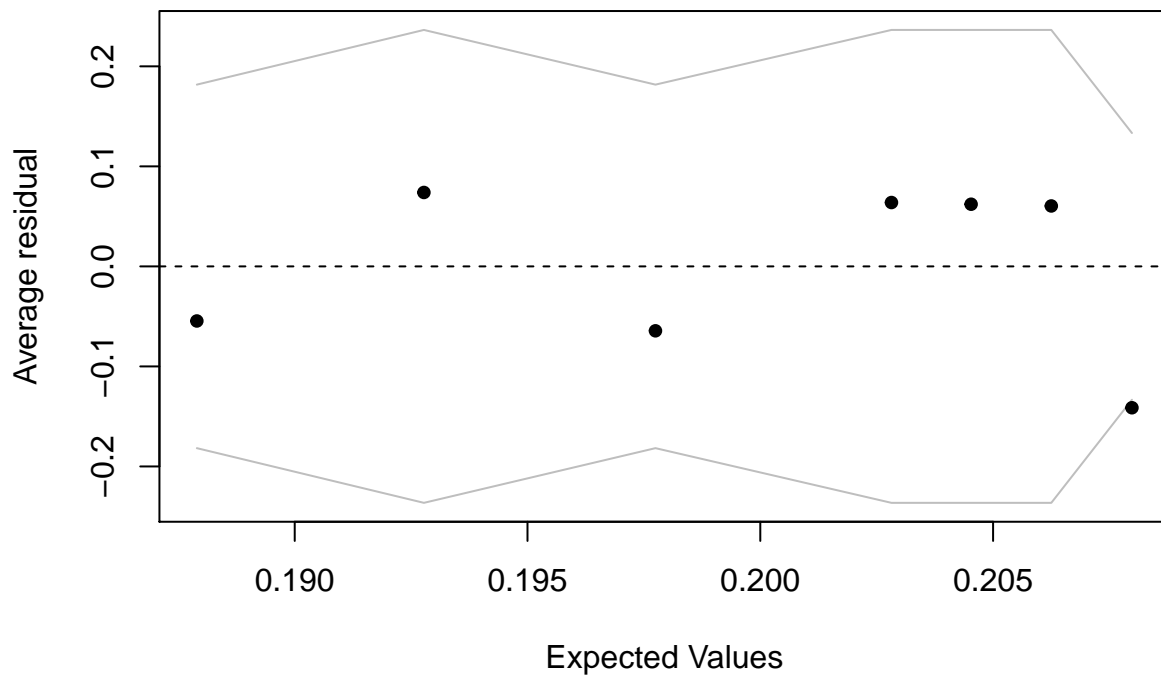
```
##  
## Pearson's Chi-squared test with Yates' continuity correction  
##  
## data:  tb.15  
## X-squared = 1e-30, df = 1, p-value = 1
```

Since p-value $1 > 0.05$ significant level, therefore, we have strong evidence to retain the null hypothesis. Thus, “ARREST RIDGE” is independent of “WALLNER LINES”.

Logistic Regression for six feature v.s. days

```
##
## Call:
## glm(formula = bone$`WAKE HACKLE` ~ bone$days, family = binomial())
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -0.683  -0.676  -0.664  -0.645   1.829
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept) -1.337128   0.367130  -3.64  0.00027 ***
## bone$days   -0.000352   0.001985  -0.18  0.85938
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 105.08  on 104  degrees of freedom
## Residual deviance: 105.05  on 103  degrees of freedom
## AIC: 109.1
##
## Number of Fisher Scoring iterations: 4
```

Binned residual plot



From the output, the estimated intercept coefficient -1.337 correspond to the log odds of having a wake hackle feature on bone with average days, in other words, on average days, expected probability of having a wake

hackle feature on bone is $\text{expit}(-1.337) = 0.208 = 20.8\%$.

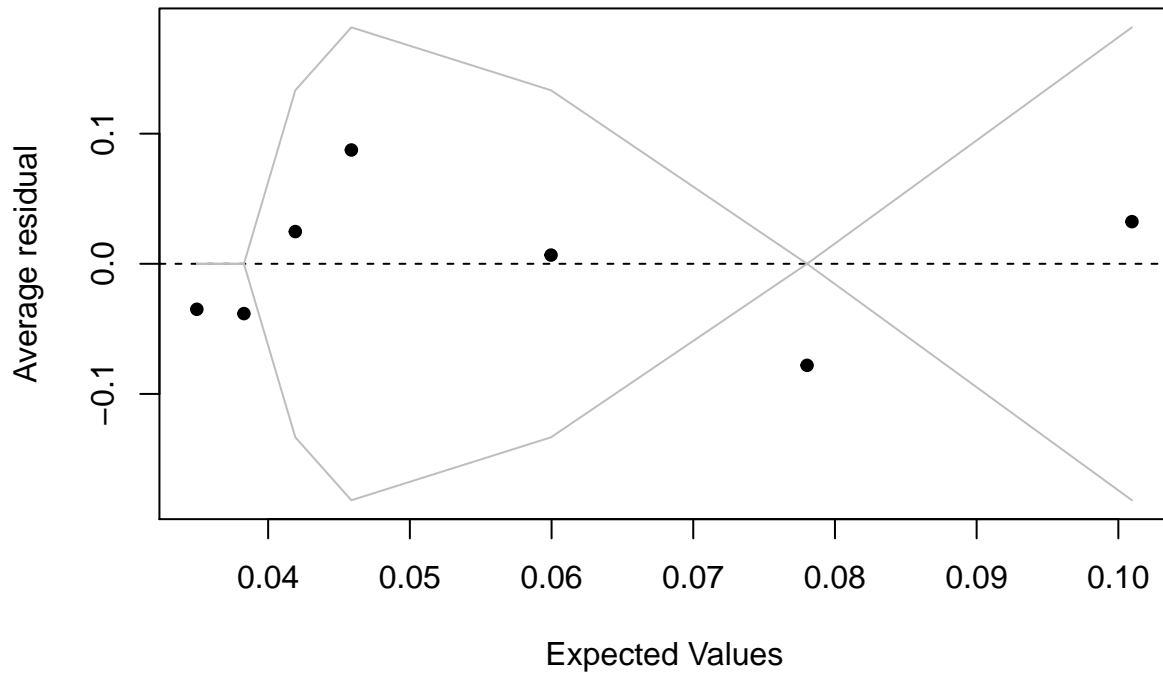
And one-unit increase in the predictor “days” increases the log odds of having a wake hackle feature on bone (versus no wake hackle feature on bone) by -0.00035.

We can see the influence of drying days are relatively small and the 0.86 p-value shows no evidence in support of the differences of the wake hackle feature when drying time changes.

Binned residual plot shows our model fits well in this case.

```
##
## Call:
## glm(formula = bone$`BONE HACKLE` ~ bone$days, family = binomial())
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -0.461  -0.403  -0.306  -0.279   2.519
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept) -3.31778    0.73705   -4.50  6.7e-06 ***
## bone$days    0.00314    0.00324    0.97    0.33
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 45.997  on 104  degrees of freedom
## Residual deviance: 45.067  on 103  degrees of freedom
## AIC: 49.07
##
## Number of Fisher Scoring iterations: 6
```

Binned residual plot

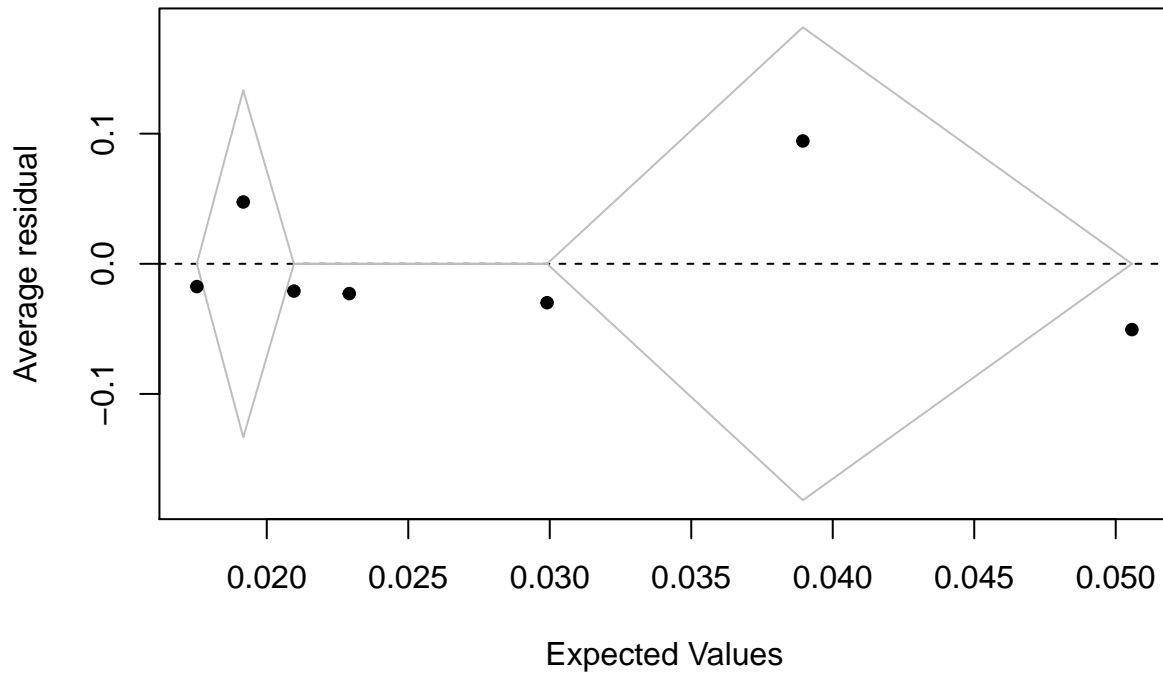


Again, the low coefficient and p value of the variable days, shows no evidence in support of the differences of the bone hackle feature when drying time changes.

Binned residual plot with only a few points outside the bound shows our model fits quite well in this case.

```
##
## Call:
## glm(formula = bone$`BONE MIRROR` ~ bone$days, family = binomial())
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -0.322  -0.282  -0.215  -0.197   2.812
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept) -4.02622    1.02982  -3.91 9.2e-05 ***
## bone$days    0.00304    0.00449   0.68  0.5
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 27.246  on 104  degrees of freedom
## Residual deviance: 26.795  on 103  degrees of freedom
## AIC: 30.8
##
## Number of Fisher Scoring iterations: 6
```

Binned residual plot



Again, the low coefficient and p value of the variable days, shows no evidence in support of the differences of the bone mirror feature when drying time changes.

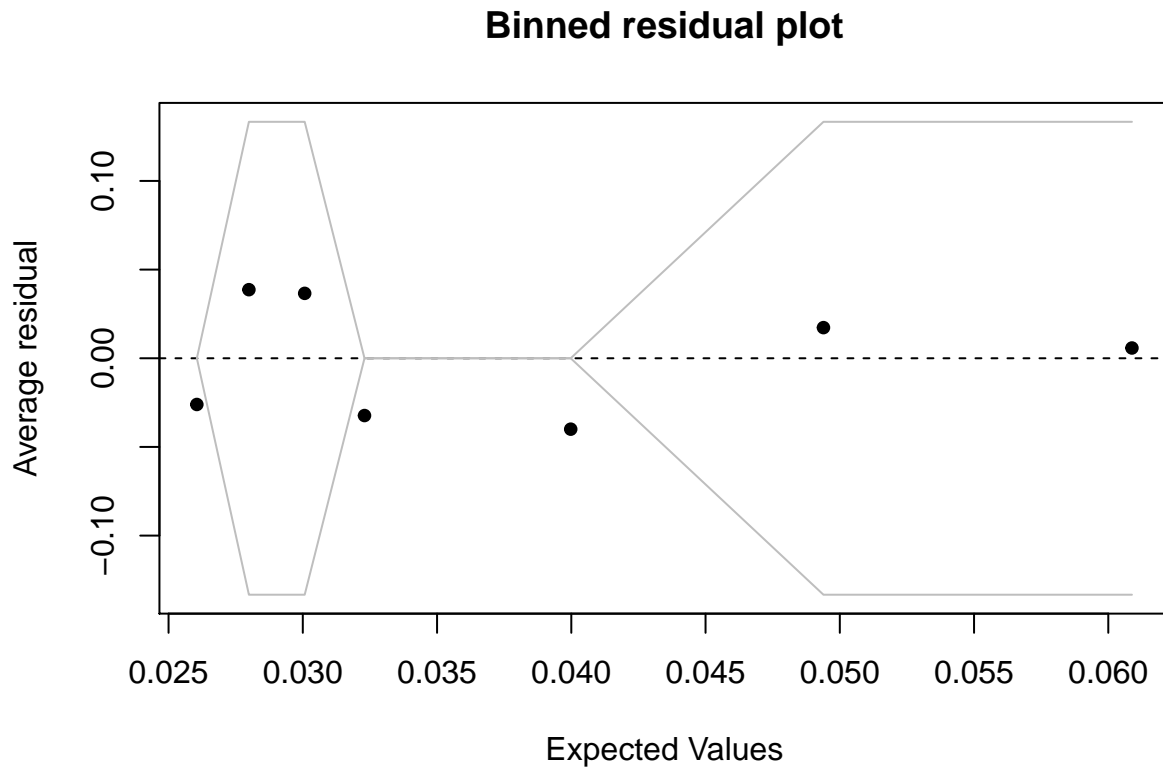
Binned residual plot with most points outside the bound shows our model fits bad in this case.

```
## # A tibble: 2 x 2
##   `BONE MIRROR`      n
##   <fct>          <int>
## 1 0             102
## 2 1              3
```

We only have 3 bones have the bone mirror feature, that may be the reason for the poor fitting.

```
##
## Call:
## glm(formula = bone$`CANTILEVER CURL` ~ bone$days, family = binomial())
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -0.354  -0.318  -0.256  -0.238   2.674
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept) -3.62120    0.86550  -4.18  2.9e-05 ***
## bone$days    0.00246    0.00392   0.63    0.53
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
```

```
##
## Null deviance: 33.987 on 104 degrees of freedom
## Residual deviance: 33.601 on 103 degrees of freedom
## AIC: 37.6
##
## Number of Fisher Scoring iterations: 6
```

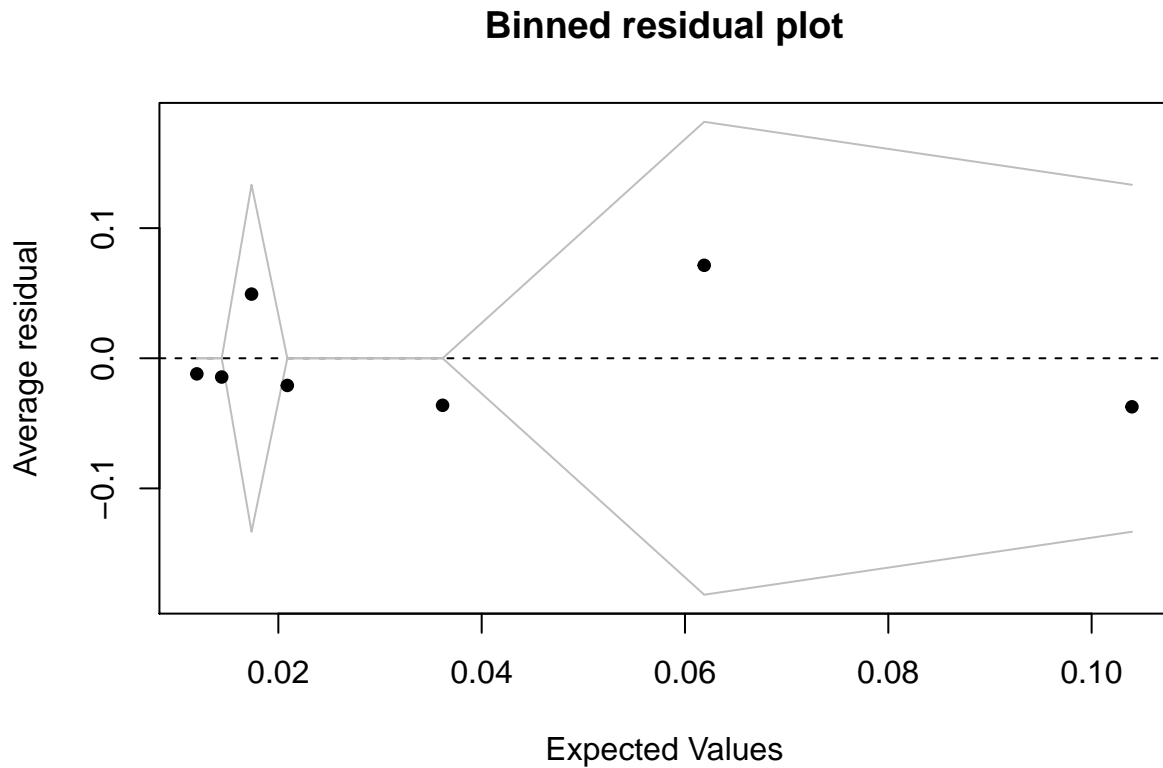


Again, the low coefficient and p value of the varibale days, shows no evidence in support of the differences of the cantilever curl feature when drying time changes.

Binned residual plot with most points inside the bound shows our model fits quite good in this case.

```
##
## Call:
## glm(formula = bone$`ARREST RIDGE` ~ bone$days, family = binomial())
##
## Deviance Residuals:
##    Min       1Q   Median       3Q      Max
## -0.469 -0.357 -0.205 -0.171  2.847
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept) -4.41229    1.10916  -3.98  6.9e-05 ***
## bone$days    0.00627    0.00417   1.50    0.13
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
```

```
##
## Null deviance: 33.987 on 104 degrees of freedom
## Residual deviance: 31.525 on 103 degrees of freedom
## AIC: 35.52
##
## Number of Fisher Scoring iterations: 6
```



```
## # A tibble: 2 x 2
##   `ARREST RIDGE`      n
##   <fct>          <int>
## 1 0             101
## 2 1              4
```

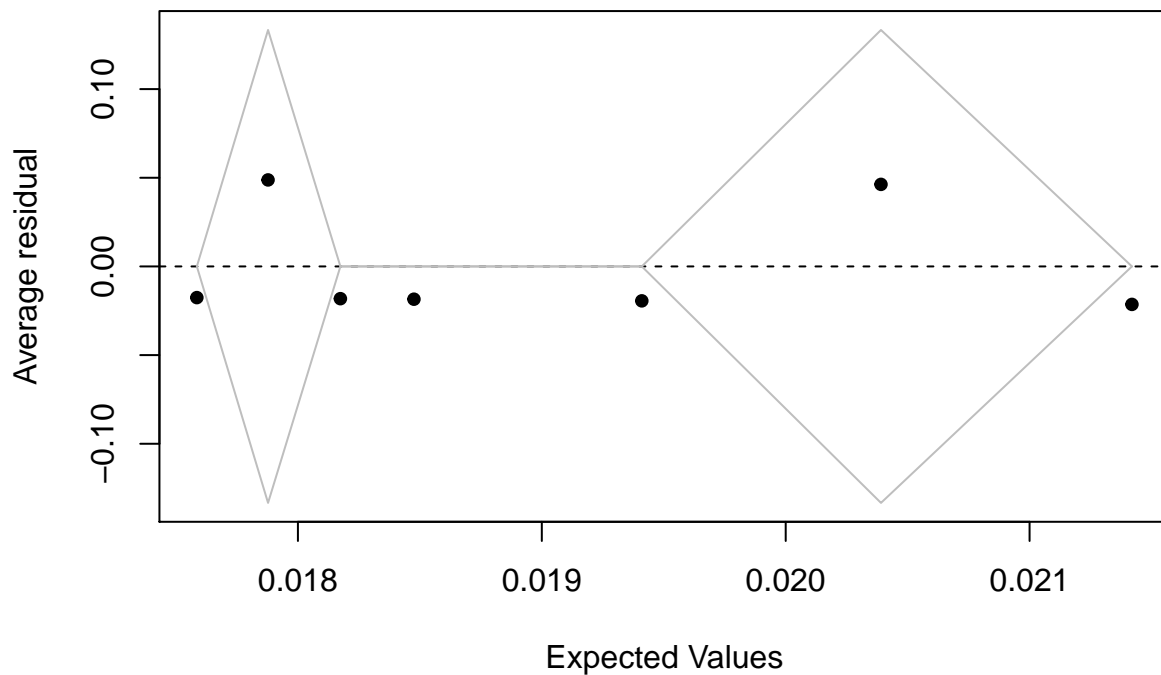
Again, the low coefficient and p value of the varibale days, shows no evidence in support of the differences of the arrest ridge feature when drying time changes.

Binned residual plot with most points outside the bound shows our model fits poor in this case. The possible reason is that we only have 4 bones have the arrest ridge feature

```
##
## Call:
## glm(formula = bone$`WALLNER LINES` ~ bone$days, family = binomial())
##
## Deviance Residuals:
##   Min       1Q   Median       3Q      Max
## -0.208  -0.203  -0.193  -0.190   2.837
##
## Coefficients:
```

```
##           Estimate Std. Error z value Pr(>|z|)
## (Intercept) -4.022938   1.108503  -3.63  0.00028 ***
## bone$days    0.000559   0.005660   0.10  0.92136
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
## Null deviance: 19.805  on 104  degrees of freedom
## Residual deviance: 19.795  on 103  degrees of freedom
## AIC: 23.8
##
## Number of Fisher Scoring iterations: 6
```

Binned residual plot



```
## # A tibble: 2 x 2
##   `WALLNER LINES`    n
##   <fct>           <int>
## 1 0             103
## 2 1              2
```

Again, the low coefficient and p value of the varibale days, shows no evidence in support of the differences of the wallner lines feature when drying time changes.

Binned residual plot with most points outside the bound shows our model fits poor in this case. The possible reason is that we only have 2 bones have the wallner lines feature

Conclusion