# STA414 Assignment 2

### Q1 Implementing the model

(a) Implementing log\_prior.

```
using Revise # lets you change A2funcs without restarting julia!
include("A2_src.jl")
using Plots
using Statistics: mean
using Zygote
using Test
using Logging
using .A2funcs: log1pexp # log(1 + exp(x)) stable
using .A2funcs: factorized_gaussian_log_density
using .A2funcs: skillcontour!
using .A2funcs: plot_line_equal_skill!

function log_prior(zs)
   return factorized_gaussian_log_density(0,0,zs)
end
```

(b) Implementing logp\_a\_beats\_b.

```
function logp_a_beats_b(za,zb)
  return -log1pexp(zb-za)
end
```

(c) Implementing all\_games\_log\_likelihood.

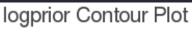
```
function all_games_log_likelihood(zs,games)
  zs_a = zs[games[:,1],:]
  zs_b = zs[games[:,2],:]
  likelihood = logp_a_beats_b.(zs_a, zs_b)
  return sum(likelihood, dims = 1)
end
```

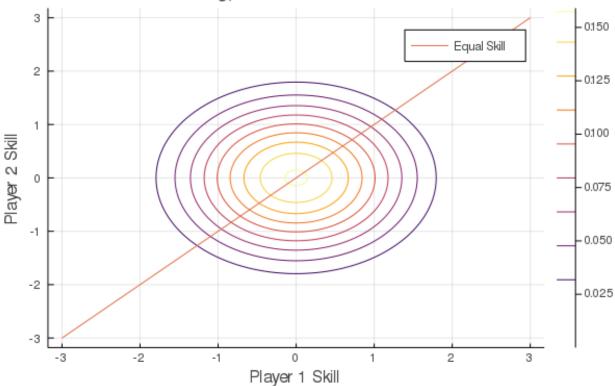
(d) Implementing joint\_log\_density.

```
function joint_log_density(zs,games)
  return sum(log_prior(zs) .+ all_games_log_likelihood(zs, games), dims=1)
end
```

## Q2 Examining the posterior for only two players and toy data

(a) plot the log prior graph.

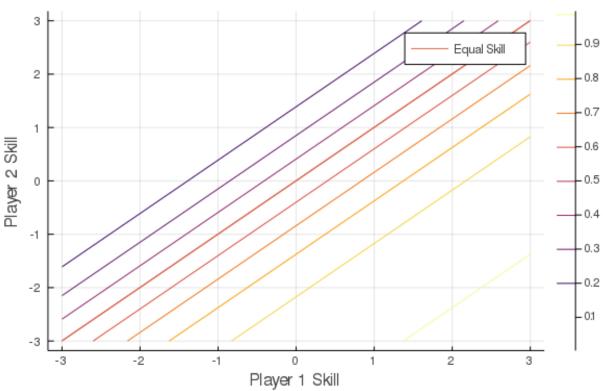




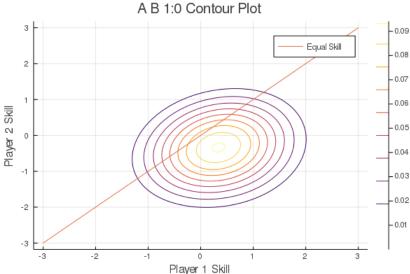
#### (b) Plot the log likelihood Contour Plot.

```
plot(title="likelihood Contour Plot",
    xlabel = "Player 1 Skill",
    ylabel = "Player 2 Skill"
    )
skillcontour!(zs -> exp.(logp_a_beats_b(zs[1], zs[2])))
plot_line_equal_skill!()
savefig(joinpath("plots","likelihood.png"))
```

# likelihood Contour Plot

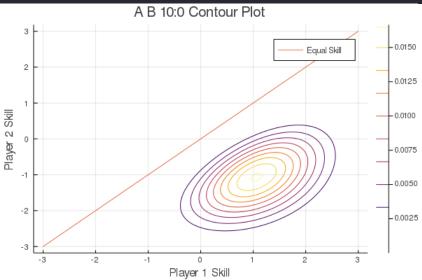


#### (c) Plot the Contour of A B 1:0.

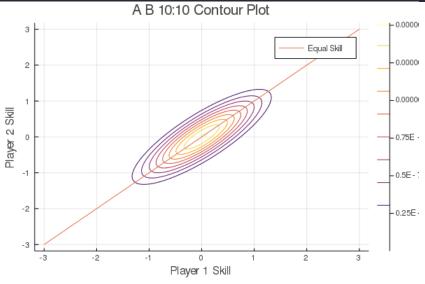


#### (d) Plot Contour of A B 10:0.

```
plot(title="A B 10:0 Contour Plot",
     xlabel = "Player 1 Skill",
     ylabel = "Player 2 Skill"
     )
skillcontour!(zs -> exp.(joint_log_density(zs, two_player_toy_games(10,0))))
plot_line_equal_skill!()
savefig(joinpath("plots","toy_svi_a10b0.png"))
```



#### (e) Plot Contour of A B 10:10.



## Q3 Stochastic Variational Inference on Two Players and Toy Data

(a) Implement elbo.

```
function elbo(params,logp,num_samples)
  samples = exp.(params[2]) .*randn(size(params[1])[1],num_samples) .+params[1]
  logp_estimate = logp(samples)
  logq_estimate = factorized_gaussian_log_density(params[1],params[2],samples)
  return sum(logp_estimate-logq_estimate)/num_samples
end
```

(b) Implement neg\_toy\_elbo.

```
function neg_toy_elbo(params; games = two_player_toy_games(1,0), num_samples = 100)
  # TODO: Write a function that takes parameters for q,
  # evidence as an array of game outcomes,
  # and returns the -elbo estimate with num_samples many samples from q
  logp(zs) = joint_log_density(zs,games)
  return -elbo(params,logp,num_samples)
end
```

(c) Implement fit\_toy\_variational\_dist.

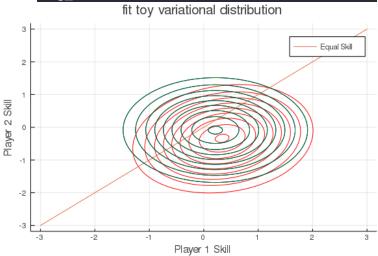
```
function fit_toy_variational_dist(init_params, toy_evidence;    num_itrs=200,
 lr= 1e-2, num_q_samples = 10)
 params_cur = init_params
 for i in 1:num_itrs
   grad_params = gradient(params -> neg_toy_elbo(params; games = toy_evidence,
   num_samples = num_q_samples), params_cur)[1]
   mu_new = params_cur[1] - lr * grad_params[1]
   ls_new = params_cur[2] - lr * grad_params[2]
   params_cur = (mu_new, ls_new)
   plot(title="fit toy variational distribution", xlabel = "Player 1 Skill",
   ylabel = "Player 2 Skill");
   target_post(zs) = exp.(joint_log_density(zs, toy_evidence))
   skillcontour!(target_post,colour=:red)
   plot_line_equal_skill!()
   mu = params_cur[1]
   logsig = params_cur[2]
   var_log_prior(zs) = factorized_gaussian_log_density(mu, logsig, zs)
   var_post(zs) = exp.(var_log_prior(zs))
   display(skillcontour!(var_post, colour=:blue))
 return params_cur
```

(d) Optimized variational approximation contours Plot of A B 1:0.

```
#01(d) 1:0
plot(title="A B 1:0 Toy SVI",
    xlabel = "Player 1 Skill",
    ylabel = "Player 2 Skill"
    )
# Plot the original graph
target_post(zs) = exp.(joint_log_density(zs, two_player_toy_games(1,0)))
skillcontour!(target_post,colour=:red)
plot_line_equal_skill!()

# Plot SVI
num_q_samples = 10
data = two_player_toy_games(1,0)
opt_params = fit_toy_variational_dist(toy_params_init, two_player_toy_games(1,0), num_itrs=200, lr= 1e-2, num_q_samples = 10)
println("neg_elbo (final loss): $(neg_toy_elbo(opt_params; games = two_player_toy_games(1,0), num_samples = num_q_samples))")
mu = opt_params[1]
logsig = opt_params[2]
var_log_prior(zs) = factorized_gaussian_log_density(mu, logsig, zs)
var_post(zs) = exp.(var_log_prior(zs))
display(skillcontour!(var_post, colour=:green))
savefig(joinpath("plots","toy_svi_alb0.png"))
```

## neg\_elbo (final loss): 0.7439268493244543



(e) Optimized variational approximation contours Plot of A B 10:0.

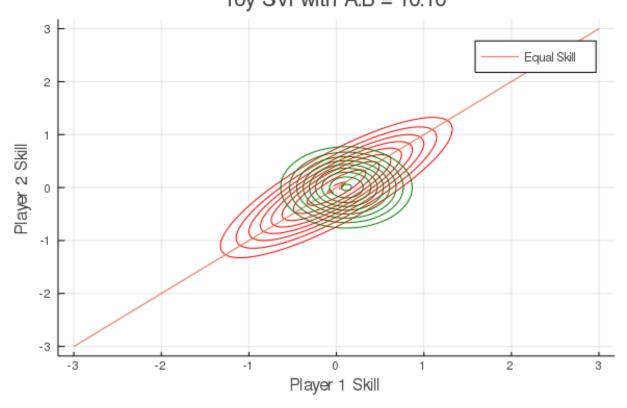
```
plot(title="A B 10:0 Toy SVI",
    xlabel = "Player 1 Skill",
    ylabel = "Player 2 Skill"
    )
    target_post(zs) = exp.(joint_log_density(zs, two_player_toy_games(10,0)))
    skillcontour!(target_post,colour=:red)
    plot_line_equal_skill!()

# Plot the new graph
    opt_params = fit_toy_variational_dist(toy_params_init, two_player_toy_games(10,0); num_itrs=200, lr= 1e-2, num_q_samples = 10)
    num_q_samples = 10
    println("neg_elbo (final loss): $(neg_toy_elbo(opt_params; games = two_player_toy_games(10,0), num_samples = num_q_samples))")
    mu = opt_params[1]
    logsig = opt_params[2]
    var_log_prior(zs) = factorized_gaussian_log_density(mu, logsig, zs)
    var_post(zs) = exp.(var_log_prior(zs))
    display(skillcontour!(var_post, colour=:green))
    savefig(joinpath("plots","toy_svi_a10b0.png"))
```

neg\_elbo (final loss): 2.9131816235518913

#### (f) Optimized variational approximation contours Plot of A B 10:10

## neg elbo (final loss): 15.746682651790952 Toy SVI with A:B = 10:10

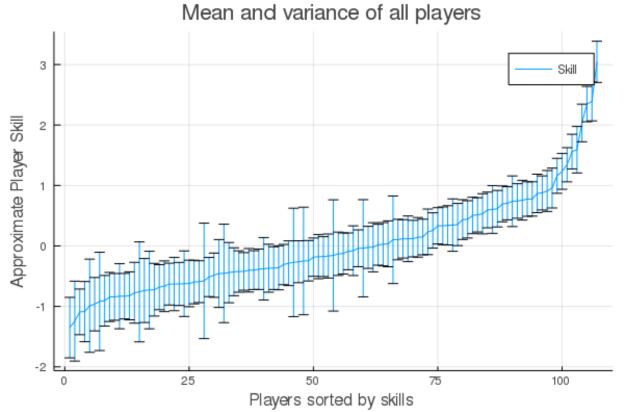


## Q4 Approximate inference conditioned on real data

(a) Yes.

(b) loss: 1185.287000374451

```
perm = sortperm(means)
plot(title="Mean and variance of all players",
    xlabel = "Players sorted by skills",
    ylabel = "Approximate Player Skill"
    )
plot!(means[perm], yerror=exp.(logstd[perm]), label="Skill")
savefig(joinpath("plots","player_mean_var.png"))
```



(d)

```
desc_perm = sortperm(means, rev=true)
 print("Top 10 players are: ")
for i in 1:10
    print(player_names[desc_perm][i],"\n")
                                Top 10 players are: Novak-Djokovic
                                Roger-Federer
                                Rafael-Nadal
                                Andy-Murray
                                Robin-Soderling
                                David-Ferrer
                                Jo-Wilfried-Tsonga
                                Tomas-Berdych
                                Juan-Martin-Del-Potro
                                Richard-Gasquet
                                                           T / / / 4 3
(f)
                            Solve for A, A= (1 7)
                           \widehat{\text{SMOR}} \quad \binom{\mathbb{R}_{\Delta}}{\mathbb{R}_{b}} \sim \mathcal{N} \left( \begin{pmatrix} \mu_{\Delta} \\ \mu_{b} \end{pmatrix}, \begin{pmatrix} \sigma_{i} & \sigma \\ \sigma & \sigma_{32a} \end{pmatrix} \right) = \mathcal{N} \left( \mu_{1} \Sigma \right)
                                     \Rightarrow \begin{bmatrix} \gamma_{A} \\ \gamma_{b} \end{bmatrix} \sim \mathcal{N} \begin{bmatrix} A\mu, A\Sigma A \end{bmatrix} = \mathcal{N} \begin{bmatrix} \mu_{A} \mu_{B} \\ \mu_{B} \end{bmatrix}, \begin{bmatrix} \sigma_{01}\sigma_{02} & -\sigma_{02} \\ -\sigma_{02} & \sigma_{03} \end{bmatrix} \end{bmatrix}
                                      => P(Ya/X)~ N(ha-Mb, 611+62)
                                         => let Ma= N(MA-Mb, OTHOS)
(g)
                                             P(ya>0)= 1-colf(ya,0)
exact_rf_better = 1 - cdf(Normal(0,1), (mu_RN - mu_RF)/sqrt(var_RF + var_RN))
MC size = 10000
samples_RF = randn(MC_size) * exp(-1.1412578223971224) + mu_RF
samples_RN = randn(MC_size) * exp(-1.2163985016732244) + mu_RN
MC_rf_better = count(x->x==1, samples_RF .> samples_RN) / MC_size
The result of SM of RF is bette than RN is: 0.5755
```

(h)

```
# Q4(h)
lowest_idx = desc_perm[num_players]
mu_lowest = means[lowest_idx]
var_lowest = exp(logstd[lowest_idx])^2
# Exact prob
exact_rf_better_than_worst = 1 - cdf(Normal(0,1), (mu_lowest - mu_RF)/sqrt(var_RF + var_lowest))
# SM apporach
samples_lowest = randn(MC_size) * exp(logstd[lowest_idx]) .+ mu_lowest
MC_h = count(x->x==1, samples_RF .> samples_lowest) / MC_size
The result of SM of RF is better the worst is: 1.0
```

(i)

(b), (c), and (e)