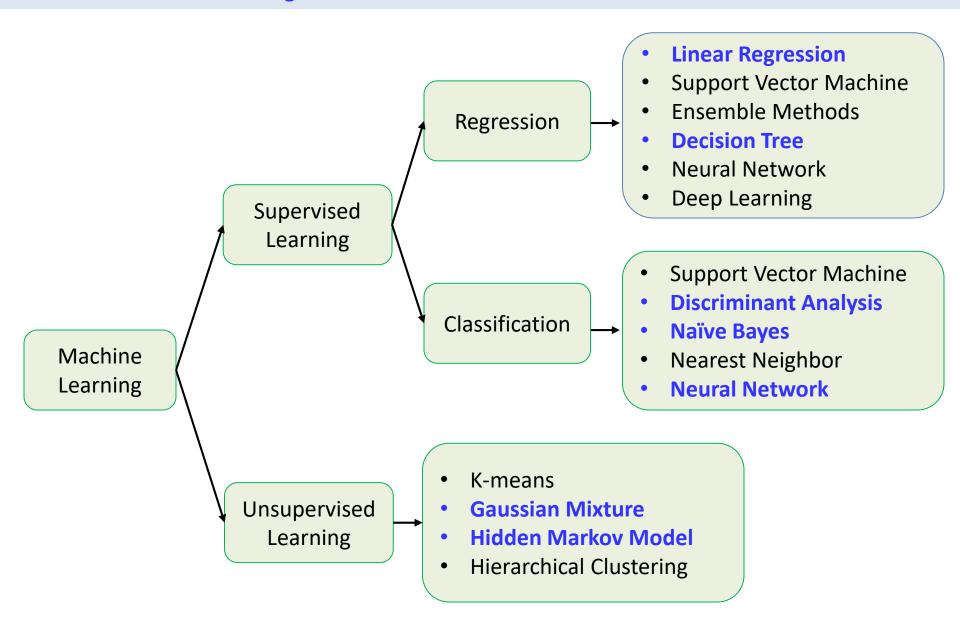
L8. Machine Learning (Regression)

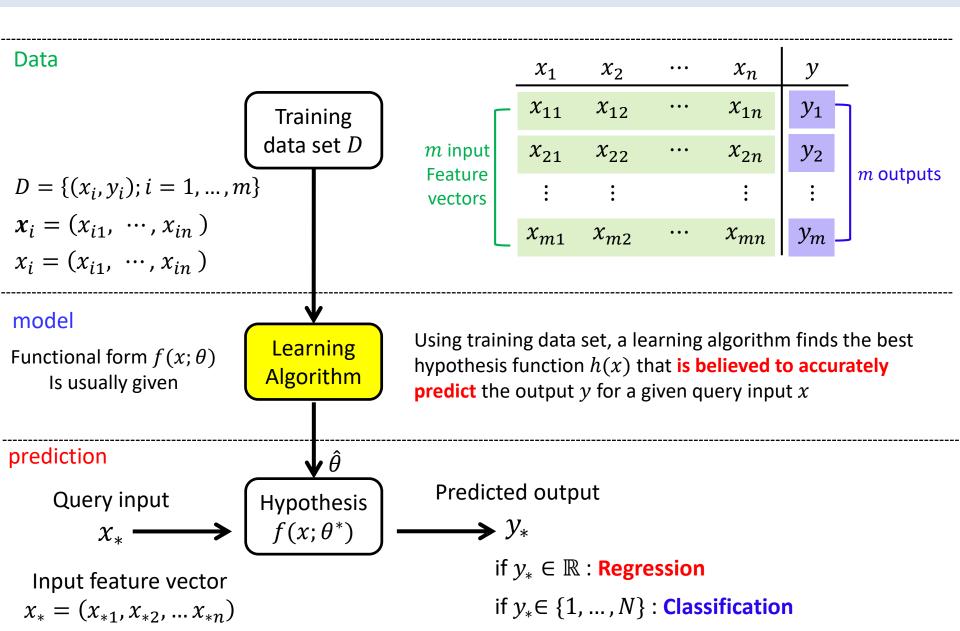
What is Machine Learning?

- Data: observations, experience,...
- Model: a form of prior knowledge, assumptions, belief
 - ✓ Functional model
 - ✓ Probabilistic model
- Prediction: the new knowledge obtained by combining the data and model
 - ✓ Regression
 - ✓ Classification
 - ✓ Clustering

What is Machine Learning?



Supervised learning



Two different learning approaches

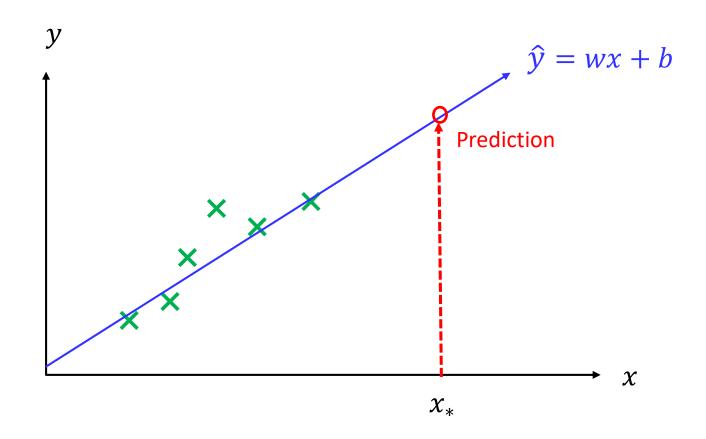
- Machine Learning as Optimization
 - ✓ Relate variables through a basis function (parametric function)
 - ✓ Formulate learning problem as an optimization problem.
 - ✓ Employ optimization algorithm to solve the formulated problem
- Machine Learning as Probabilistic Modeling (not necessarily Bayesian)
 - ✓ Relate variables through probability distributions
 - ✓ Formulate learning problem as inference
 - ✓ If Bayesian, treat parameters with probability distributions
 - ✓ Requires inference methods (integral or sampling) to solve the formulated problem

Let's explore different views on Machine Learning by taking a linear regression as an example

Load Map

- 1. Optimization Approach (Normal Equation)
- 2. Maximum Likelihood Estimation (MLE) Approach
- 3. Maximum A Posteriori Estimation (MAP) Approach
- 4. Full Bayesian Approach
 - ✓ Analytical approach
 - ✓ Sampling approach
- 5. Regularization regression (Ridge and Lasso)
 - ✓ Optimization view
 - ✓ Bayesian View

1D Linear Regression

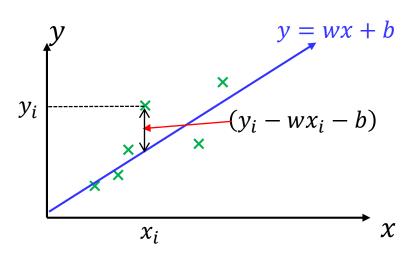


- Data : $(x_1, y_1), ..., (x_m, y_m)$
- Model: Linear model $\hat{y} = wx + b$ $(y = w^Tx + b \text{ for multidimensional})$
- Learning: What are w and b?
- **Prediction**: What is $\hat{y}_* = wx_* + b$

Learning as optimization

Define an objective (cost) function

$$J(w,b) = \sum_{i=1}^{m} (y_i - wx_i - b)^2$$



Minimize the error function with respect to w and b

$$\frac{dJ(w,b)}{dw} = -2\sum_{i=1}^{m} x_i(y_i - wx_i - b) = 0 \rightarrow w^* = \frac{\sum_{i=1}^{m} (y_i - b^*)x_i}{\sum_{i=1}^{n} x_i^2} = \frac{\sum_{i=1}^{m} (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^{m} (x_i - \bar{x})^2}$$

$$\frac{dJ(w,b)}{db} = -2\sum_{i=1}^{m} (y_i - wx_i - b) = 0 \to b^* = \frac{\sum_{i=1}^{m} (y_i - w^*x_i)}{n}$$

Learning as optimization

Notation for general cases $x_i \in \mathbb{R}^n$

A linear regression model

$$\hat{y}_i = w_0 + w_1 x_{i1} + \dots + w_n x_{in}$$
 with $w = (w_0, w_1, \dots, w_n)^T$ and $x_i = (x_{i1}, \dots, x_{in})^T$

• If we introduce $x_{i0} = 1$,

$$\hat{y}_i = w^T x_i$$

with
$$w = (w_0, w_1, ..., w_n)^T$$
 and $x_i = (x_{i0}, x_{i1}, ..., x_{in})^T$

In a Matrix form

$$\begin{pmatrix} \hat{y}_1 \\ \vdots \\ \hat{y}_m \end{pmatrix} = \begin{pmatrix} -x_1^T - \\ \vdots \\ -x_m^T - \end{pmatrix} \begin{pmatrix} w_0 \\ \vdots \\ w_n \end{pmatrix} = \begin{pmatrix} x_1^T w \\ \vdots \\ x_m^T w \end{pmatrix} \longrightarrow \hat{y} = Xw$$

m: # of data points

with
$$\hat{y} = (\hat{y}_1, \dots, \hat{y}_m)^T$$
, $X = \begin{pmatrix} -x_1^T - \\ \vdots \\ -x_m^T - \end{pmatrix}$

Learning as optimization (Normal Equation)

The cost function for the optimization can be defined as:

$$J(w) = \frac{1}{2} \sum_{i=1}^{m} (x_i^T w - y_i)^2 = \frac{1}{2} \|y - wX\|_2^2 = \frac{1}{2} (Xw - y)^T (Xw - y)$$

$$\sqrt{\sum_{i=1}^{m} z_i^2} = \|z\|_2 = \sqrt{z^T z}$$

• The optimum parameters \widehat{w} can be computed by minizing the cost function :

$$\widehat{w} = \arg\min_{w} J(w) = \arg\min_{w} \frac{1}{2} ||y - wX||_{2}^{2}$$

• For reference, other vector norms are summarized here:

$$||z||_1 = \sum_{i=1}^n |z_i|, \qquad ||z||_p = \left(\sum_{i=1}^n |z_i|^p\right)^{\frac{1}{p}} (p \ge 1), \qquad ||z||_{\infty} = \max_i |z_i|$$

Learning as optimization (Normal Equation)

Linear algebra approach for finding the optimum parameters:

$$\widehat{w} = \arg\min_{w} J(w) = \arg\min_{w} \frac{1}{2} ||y - wX||_{2}^{2}$$

Since, J(w) is differentiable in w, the optimality condition : $\nabla_w J(w) = 0$ at $w = \widehat{w}$

$$\nabla_{w}J(w) = \nabla_{w}\frac{1}{2}(Xw - y)^{T}(Xw - y)$$

$$= \frac{1}{2}\nabla_{w}(w^{T}X^{T}Xw - w^{T}X^{T}y - y^{T}Xw + y^{T}y)$$

$$= \frac{1}{2}\nabla_{w}\operatorname{tr}(w^{T}X^{T}Xw - w^{T}X^{T}y - y^{T}Xw + y^{T}y)$$

$$= \frac{1}{2}\nabla_{w}\left(\operatorname{tr}(w^{T}X^{T}Xw) - 2\operatorname{tr}(y^{T}Xw)\right)$$

$$= \frac{1}{2}(X^{T}Xw + X^{T}Xw - 2X^{T}y)$$

$$= X^{T}Xw - X^{T}y$$

$$\nabla_{w}J(\widehat{w}) = X^{T}X\widehat{w} - X^{T}y = 0$$

$$\rightarrow X^{T}X\widehat{w} = X^{T}y$$

$$\rightarrow \widehat{w} = (X^{T}X)^{-1}X^{T}y \text{ (when } X \text{ is full column rank)}$$

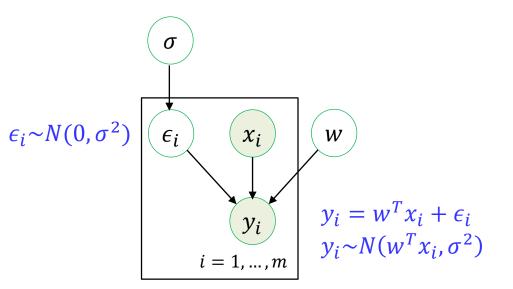
Probabilistic view on linear regression

Assume there is uncertainty in the predicted value :

$$y_i = w^T x_i + \epsilon_i$$
 with $\epsilon_i \sim N(0, \sigma^2)$

• Then the probabilistic model on output y_i can be represented as

$$y_i \sim N(w^T x_i, \sigma^2)$$
 or $p(y_i | w^T x_i, \sigma) = N(y_i | w^T x_i, \sigma^2)$



An error ϵ_i is independently identically distributed (i.i.d assumption)

The likelihood of the data is defined as

$$p(y|X, w, \sigma) = \prod_{i=1}^{m} N(y_i|w^T x_i, \sigma^2) = \prod_{i=1}^{m} \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(y_i - w^T x_i)^2}{2\sigma^2}\right)$$

Learning as probabilistic model (MLE Approach)

The log likelihood is

$$L(w,\sigma) = \log p(y|X, w, \sigma)$$

$$= \log \prod_{i=1}^{m} \frac{1}{\sqrt{2\pi\sigma^{2}}} \exp\left(-\frac{(y_{i} - w^{T}x_{i})^{2}}{2\sigma^{2}}\right)$$

$$= \log\left(\frac{1}{\sqrt{2\pi\sigma^{2}}}\right)^{m} \exp\left(-\sum_{i=1}^{m} \frac{(y_{i} - w^{T}x_{i})^{2}}{2\sigma^{2}}\right)$$

$$= m \log \frac{1}{\sqrt{2\pi\sigma^{2}}} - \frac{1}{\sigma^{2}} \sum_{i=1}^{m} (y_{i} - w^{T}x_{i})^{2}$$

$$= m \log \frac{1}{\sqrt{2\pi\sigma^{2}}} - \frac{1}{\sigma^{2}} J(w)$$

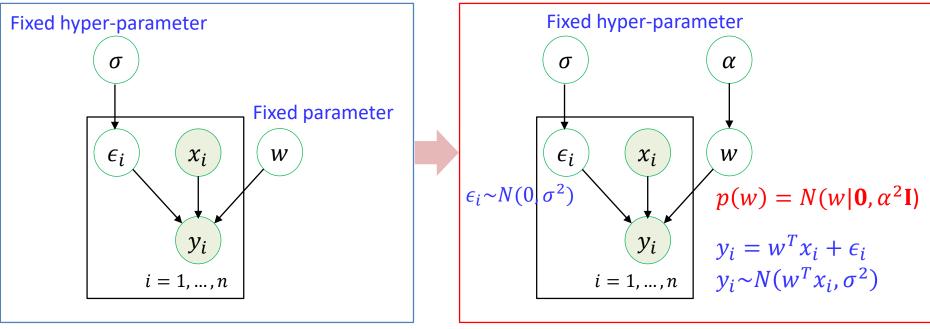
The optimum parameters is determined by maximizing log likelihood

$$(w^*, \sigma) = \max_{(w, \sigma)} L(w, \sigma) = \max_{(w, \sigma)} \log p(y|X, w, \sigma)$$

Minimizing the square error sum J(w) =

MLE approach (point estimation)

Bayesian approach



- Consider the parameter w as stochastic variables (represented as a distribution)
- (Assume σ is known for simple derivation)
- Find the distribution on parameter w

$$p(w|y,X) = \frac{p(y|X,w)p(w|X)}{\int_{w} p(y|X,w)p(w|X)dw} \rightarrow p(w|y) = \frac{p(y|w)p(w)}{\int_{w} p(y|w)p(w)dw}$$

We will assume X is fixed for the data y

Multivariate regression likelihood is

$$p(y|w) = \prod_{i=1}^{m} p(y_i|x_i, w)$$

$$= \frac{1}{(2\pi\sigma^2)^{m/2}} \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^{m} (y_i - w^T x_i)^2\right) \qquad m = \text{\# of data points}$$

Multivariate Gaussian prior on parameter w

$$p(w) = N(w|\mathbf{0}, \alpha^2 \mathbf{I})$$

$$p(w) = \frac{1}{(2\pi\alpha^2)^{n/2}} \exp\left(-\frac{1}{2\alpha^2} w^T w\right)$$

$$n = \text{Dimension of } w$$

We want to find the posterior

$$p(w|X,y) \propto p(y|X,w)p(w)$$

$$= \frac{1}{(2\pi\sigma^2)^{m/2}} \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^{m} (y_i - w^T x_i)^2\right) \frac{1}{(2\pi\alpha^2)^{n/2}} \exp\left(-\frac{1}{2\alpha^2} w^T w\right)$$

Take log:

$$\begin{split} \log p(w|X,y) &= -\frac{1}{2\sigma^2} \sum_{i=1}^m (y_i - w^T x_i)^2 - \frac{1}{2\alpha^2} w^T w + const \\ &= -\frac{1}{2\sigma^2} \sum_{i=1}^m y_i^2 + \frac{1}{\sigma^2} \sum_{i=1}^m y_i x_i^T w - \frac{1}{2\sigma^2} \sum_{i=1}^m w^T x_i x_i^T w - \frac{1}{2\alpha^2} w^T w + const \\ &= -\frac{1}{2\sigma^2} y^T y + \frac{1}{\sigma^2} y^T X w - \frac{1}{2\sigma^2} w^T X^T X w - \frac{1}{2\alpha^2} w^T w + const \\ &= -\frac{1}{2\sigma^2} y^T y + \frac{1}{\sigma^2} y^T X w - \frac{1}{2} w^T \left[\frac{1}{\sigma^2} X^T X + \frac{1}{\alpha^2} I \right] w + const \end{split}$$

Posterior distribution is

$$p(w|X,y) = N(w|\mu_w, \Sigma_w)$$

$$\mu_w = \Sigma_w \left(\frac{1}{\sigma^2} X^T y\right) \qquad \Sigma_w = \left[\frac{1}{\sigma^2} X^T X + \frac{1}{\alpha^2} I\right]^{-1}$$

Predictive distribution

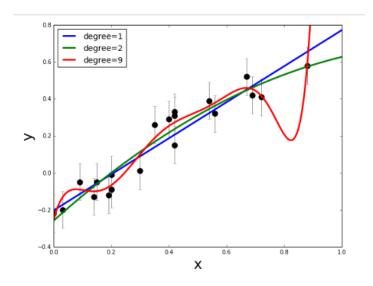
$$p(y_*|x_*, X, y) = \int_{w} p(y_*|x_*, w) p(w|X, y) dw$$

Jupyter Demo Simulation
Bayesian Regression Analytical

Predictive distribution

$$p(y_*|x_*, X, y) = \int_{w} p(y_*|x_*, w) p(w|X, y) dw$$

Jupyter Demo Simulation
Bayesian Regression (Sampling using PyMC)



What is a good regression function?

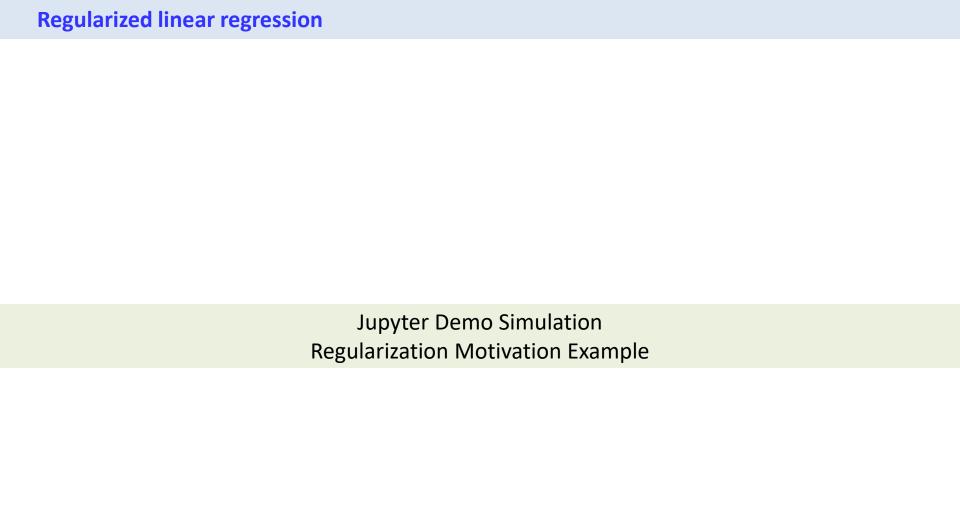
• The goal of regression is to come up with some good prediction function:

$$\hat{f}(x) = x^T \widehat{w}$$

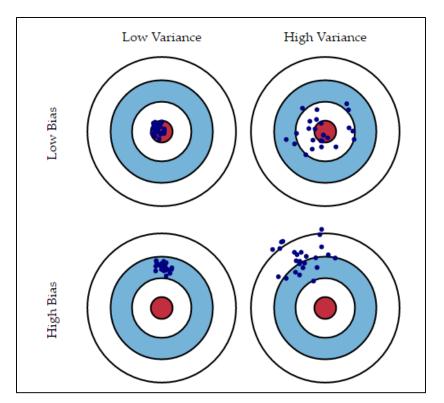
• So far, we have found \widehat{w} by finding (Ordinary Least Square Estimation)

$$\widehat{w} = \arg\min_{w} J(w) = \arg\min_{w} \frac{1}{2} ||y - wX||_{2}^{2}$$

- To see whether $\hat{f}(x)$ is a good candidate, we need to check
 - ✓ Is \widehat{w} close to the true w?
 - ✓ Will $\hat{f}(x)$ fit future observation well? (Generalization)



The Bias and Variance Trade off



Each hit represents an individual realization of our model, given the chance variability in the training data we gather.

- Imagine you could repeat the whole model building process more than once: each time you gather new data and run a new analysis creating a new model.
- Due to randomness in the underlying data sets, the resulting models will have a range of predictions.
 - Bias measures how far off in general these models' predictions are from the correct value
 - Variance measures the variability of a model prediction for a given data point

The Bias and Variance Trade off

Estimation : $\hat{f}(x) = x^T \hat{w}$

True : f(x)

Observation : $y = f(x) + \epsilon, \epsilon \sim N(0, \sigma^2)$

• The expected prediction error of a regression fit $\hat{f}(x_0)$, using square-loss error :

$$\begin{aligned} \mathsf{EPP}(x_0) &= \mathsf{E} \left[\left(y - \hat{f}(x_0) \right)^2 \middle| x_0 \right] \\ &= \mathsf{E} \left[y^2 + \hat{f}(x_0)^2 - 2y \hat{f}(x_0) \middle| x_0 \right] \\ &= \mathsf{E} \left[y^2 \middle| x_0 \right] + \mathsf{E} \left[\hat{f}(x_0)^2 \right] - \mathsf{E} \left[2y \hat{f}(x_0) \middle| x_0 \right] \\ &= \mathsf{E} \left[y^2 \middle| x_0 \right] + \mathsf{E} \left[\hat{f}(x_0)^2 \right] - 2f(x_0) \mathsf{E} \left[\hat{f}(x_0) \right] \end{aligned}$$

$$E[2y\widehat{f}(x_0)|x_0] = E[2(f(x) + \epsilon)\widehat{f}(x_0)|x_0]$$

$$= 2E[f(x)\widehat{f}(x_0)|x_0] + 2E[\epsilon\widehat{f}(x_0)|x_0]$$

$$= 2f(x_0)E[\widehat{f}(x_0)|x_0] + 2E[\epsilon\widehat{f}(x_0)|x_0] \quad (\because f(x_0) \text{ is constant})$$

$$= 2f(x_0)E[\widehat{f}(x_0)] \qquad (\because \epsilon \perp \widehat{f}(x_0))$$

The Bias and Variance Trade off

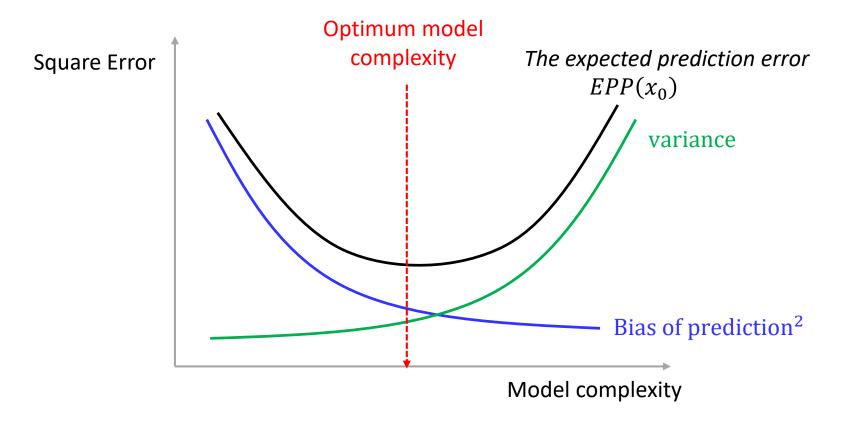
Estimation : $\hat{f}(x) = x^T \hat{w}$

True : f(x)

Observation : $y = f(x) + \epsilon, \epsilon \sim N(0, \sigma^2)$

The expected prediction error of a regression fit $\hat{f}(x_0)$, using square-loss error :

$$\begin{split} \text{EPP}(x_0) &= \mathbb{E}\left[\left(y - \hat{f}(x_0)\right)^2 \middle| x_0\right] \\ &= \mathbb{E}[y^2 + \hat{f}(x_0)^2 - 2y\hat{f}(x_0) \middle| x_0\right] \\ &= \mathbb{E}[y^2 \middle| x_0\right] + \mathbb{E}[\hat{f}(x_0)^2] - \mathbb{E}[2y\hat{f}(x_0) \middle| x_0\right] \\ &= \mathbb{E}[y^2 \middle| x_0\right] + \mathbb{E}[\hat{f}(x_0)^2] - 2f(x_0)\mathbb{E}[\hat{f}(x_0)] \\ &= \text{var}[y \middle| x_0\right] + \mathbb{E}[y \middle| x_0\right]^2 + \text{var}[\hat{f}(x_0)] + \mathbb{E}[\hat{f}(x_0)]^2 - 2f(x_0)\mathbb{E}[\hat{f}(x_0)] \\ &= \text{var}[y \middle| x_0\right] + \text{var}[\hat{f}(x_0)] + \mathbb{E}[y \middle| x_0\right]^2 + \mathbb{E}[\hat{f}(x_0)]^2 - 2f(x_0)\mathbb{E}[\hat{f}(x_0)] \\ &= \text{var}[y \middle| x_0\right] + \text{var}[\hat{f}(x_0)] + f(x_0)^2 + \mathbb{E}[\hat{f}(x_0)]^2 - 2f(x_0)\mathbb{E}[\hat{f}(x_0)] \\ &= \text{var}[y \middle| x_0\right] + \mathbb{E}\left[(\hat{f}(x_0) - \mathbb{E}[\hat{f}(x_0)])^2\right] + (f(x_0) - \mathbb{E}[\hat{f}(x_0)])^2 \\ &= \sigma^2 + \text{variance of prediction} + \text{Bias of prediction}^2 \\ &\text{Irreducible} \end{split}$$

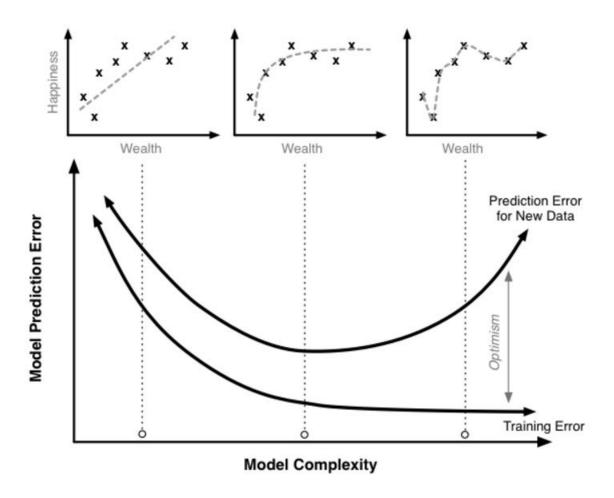


Model complexity is related with

- The number of model parameters
- The size of model parameters

• ...

How to measure prediction error



How to measure (estimate) model prediction error?

- Statistical measure (i.e., R² value)
- Information Theoretic Approaches (i.e., BIC, AIC measure)
- Holdout set or Cross Validation (Training data vs Test data set)

Ordinary Least Square Estimation

 OLS estimates find the parameter that minimize the bias between the predicted and true values:

$$\widehat{w} = \underset{w}{\operatorname{argmin}} \|y - wX\|_{2}^{2}$$

- OLS estimates often have low bias but large variance
- → Poor generalization toward unseen test data set
- All features have a weight
- → Smaller subset with strong effects is more interpretable
- w_i' s are unconstrained
- → They can explore and hence are susceptible to very high variance

We need some shrinkage (or regulation) to constraint \widehat{w}

Ridge regression

Ridge regression introduces a regularization with the L-2 norm:

$$\widehat{w} = \underset{w}{\operatorname{argmin}} \|y - wX\|_{2}^{2} + \lambda_{2} \|w\|_{2}^{2} \qquad \|w\|_{2} = \sqrt{\sum_{i=1}^{k} w_{i}^{2}},$$

- Sacrifice a little of bias to reduce the variance of predicted values
- → More stable and generalize better
- Keep all the repressors in the model
- → Not easily interpretable model

Lasso (Least Absolute Shrinkage and Selection Operator)

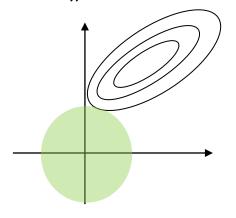
Lasso regression introduces a regularization with the L-1 norm:

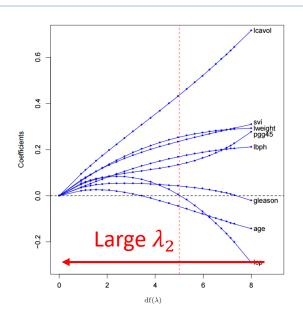
$$\widehat{w} = \underset{w}{\operatorname{argmin}} \|y - wX\|_{2}^{2} + \lambda_{1} \|w\|_{1} \qquad \|w\|_{1} = \sum_{i=1}^{k} |w_{i}|$$

- Only a small subset of features with $\widehat{w}_i \neq 0$ are selected
- → Increases the interpretability
- More difficult to implement than Ridge Regression

Ridge regression

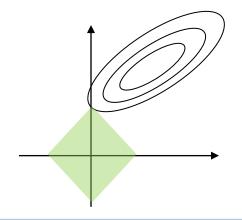
$$\widehat{w} = \underset{w}{\operatorname{argmin}} \|y - wX\|_{2}^{2} + \lambda_{2} \|w\|_{2}^{2}$$

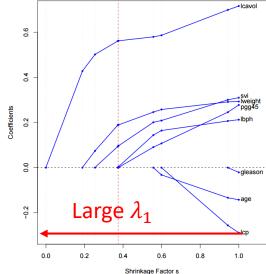




Lasso (Least Absolute Shrinkage and Selection Operator)

$$\widehat{w} = \underset{w}{\operatorname{argmin}} \|y - wX\|_{2}^{2} + \lambda_{1} \|w\|_{1}$$



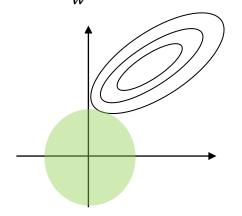


Bayesian view on Ridge regression

Ridge regression

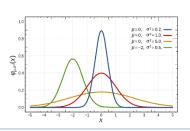
$$\widehat{w} = \underset{w}{\operatorname{argmin}} \|y - wX\|_{2}^{2} + \lambda_{2} \|w\|_{2}^{2}$$

$$\widehat{w} = \underset{w}{\operatorname{argmin}} \|y - wX\|_{2}^{2} + \lambda_{2} \|w\|_{2}^{2} \qquad \widehat{w} = \underset{w}{\operatorname{argmax}} \log p(w|X, y) = \log p(y|X, w) p(w)$$



MAP estimation view

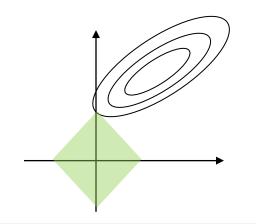
$$p(w) = \frac{1}{(2\pi\alpha^2)^{k/2}} \exp\left(-\frac{1}{2\alpha^2} w^T w\right)$$



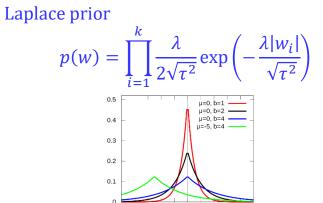
Lasso (Least Absolute Shrinkage and Selection Operator)

$$\widehat{w} = \operatorname{argmin} \|y - wX\|_2^2 + \lambda_1 \|w\|_1$$

$$\widehat{w} = \underset{w}{\operatorname{argmin}} \|y - wX\|_{2}^{2} + \lambda_{1} \|w\|_{1} \qquad \widehat{w} = \underset{w}{\operatorname{argmax}} \log p(w|X, y) = \log p(y|X, w) p(w)$$



MAP estimation view



Bayesian view on Ridge regression

•
$$p(y|X,w) = \prod_{i=1}^{m} \frac{1}{(2\pi\sigma^2)^{1/2}} \exp\left(-\frac{(y_i - w^T x_i)^2}{2\sigma^2}\right)$$

m : number of data points

n: dimension of w

•
$$p(w) = N(w|\mathbf{0}, \tau^2 \mathbf{I}) = \prod_{i=1}^{n} \frac{1}{(2\pi\tau^2)^{1/2}} \exp\left(-\frac{(w_i - 0)^2}{2\tau^2}\right) = \frac{1}{(2\pi\tau^2)^{n/2}} \exp\left(-\frac{1}{2\tau^2} w^T w\right)$$

• $p(w|X,y) \propto p(y|X,w)p(w)$

$$= \prod_{i=1}^{m} \frac{1}{(2\pi\sigma^{2})^{1/2}} \exp\left(-\frac{(y_{i} - w^{T}x_{i})^{2}}{2\sigma^{2}}\right) \frac{1}{(2\pi\tau^{2})^{n/2}} \exp\left(-\frac{1}{2\tau^{2}}w^{T}w\right)$$

$$= \left(\frac{1}{(2\pi\sigma^{2})^{1/2}}\right)^{m} \frac{1}{(2\pi\tau^{2})^{n/2}} \exp\left(-\sum_{i=1}^{m} \frac{(y_{i} - w^{T}x_{i})^{2}}{2\sigma^{2}} - \frac{1}{2\tau^{2}}w^{T}w\right)$$

•
$$\log p(w|X,y) = m \log \frac{1}{(2\pi\sigma^2)^{1/2}} + \log \frac{1}{(2\pi\tau^2)^{n/2}} - \frac{1}{2\sigma^2} \left(\sum_{i=1}^m (y_i - w^T x_i)^2 + \frac{\sigma^2}{\tau^2} w^T w \right)$$

Maximum A Posteriori (MAP) estimation with Gaussian prior = Ridge Regression

$$(w^*) = \underset{w}{\operatorname{argmax}} \log p(w|X, y) = \underset{w}{\operatorname{argmin}} ||y - wX||_2^2 + \lambda_2 ||w||_2^2$$

Bayesian view on Lasso regression

•
$$p(y|X,w) = \prod_{i=1}^{m} \frac{1}{(2\pi\sigma^2)^{1/2}} \exp\left(-\frac{(y_i - w^T x_i)^2}{2\sigma^2}\right)$$

m: number of data points

n: dimension of w

•
$$p(w) = \text{Lap}(w|\lambda, \tau) = \prod_{i=1}^{n} \frac{\lambda}{2\sqrt{\tau^2}} \exp\left(-\frac{\lambda|w_i|}{\sqrt{\tau^2}}\right)$$

• $p(w|X,y) \propto p(y|X,w)p(w)$

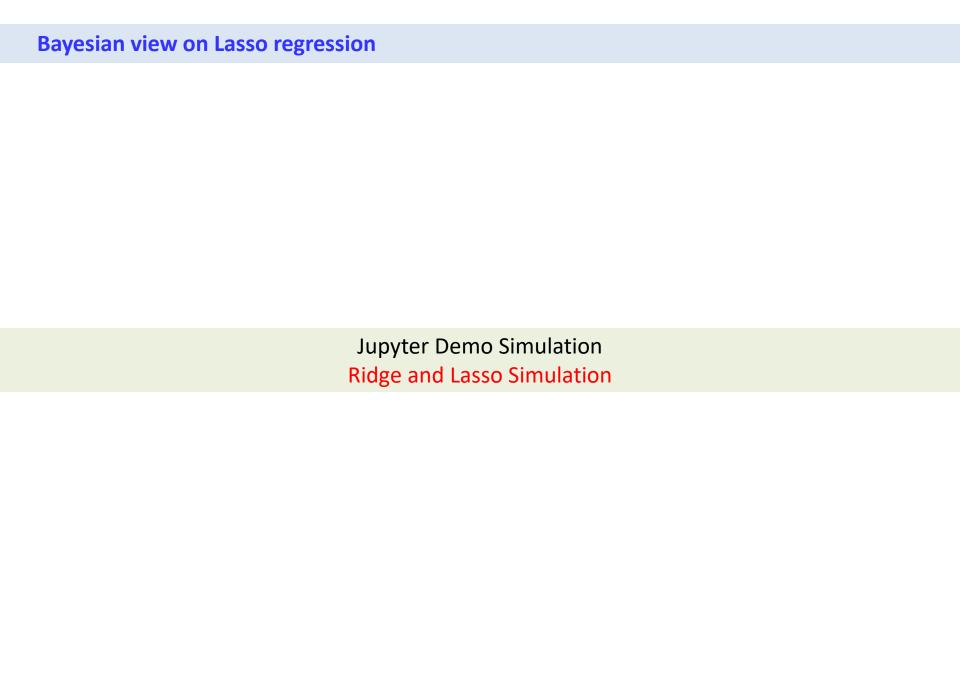
$$= \prod_{i=1}^{m} \frac{1}{(2\pi\sigma^{2})^{1/2}} \exp\left(-\frac{(y_{i} - w^{T}x_{i})^{2}}{2\sigma^{2}}\right) \left(\frac{\lambda}{2\sqrt{\tau^{2}}}\right)^{n} \exp\left(-\frac{\lambda}{\sqrt{\tau^{2}}} \sum_{i=1}^{n} |w_{i}|\right)$$

$$= \left(\frac{1}{(2\pi\sigma^{2})^{1/2}}\right)^{m} \left(\frac{\lambda}{2\sqrt{\tau^{2}}}\right)^{n} \exp\left(-\sum_{i=1}^{m} \frac{(y_{i} - w^{T}x_{i})^{2}}{2\sigma^{2}} - \frac{\lambda}{\sqrt{\tau^{2}}} \sum_{i=1}^{n} |w_{i}|\right)$$

•
$$\log p(w|X,y) = m \log \frac{1}{(2\pi\sigma^2)^{1/2}} + n \log \frac{\lambda}{2\sqrt{\tau^2}} - \frac{1}{2\sigma^2} \left(\sum_{i=1}^m (y_i - w^T x_i)^2 + \frac{2\sigma^2 \lambda}{\sqrt{\tau^2}} \sum_{i=1}^n |w_i| \right)$$

Maximum A Posteriori estimation with Laplacian prior = Lasso regression

$$(w^*) = \underset{w}{\operatorname{argmax}} \log p(w|X, y) = \underset{w}{\operatorname{argmin}} ||y - wX||_2^2 + \lambda_1 ||w||_1$$



Bayesian Model Selection

Bayesian Approach for Model Selection

- Model fitting proceeds by assuming a particular model is true, and tuning the model so it provides the best possible fit to the data
- **Model selection**, on the other hand, asks the larger question of whether the assumptions of the model are compatible with the data.

Linear model:

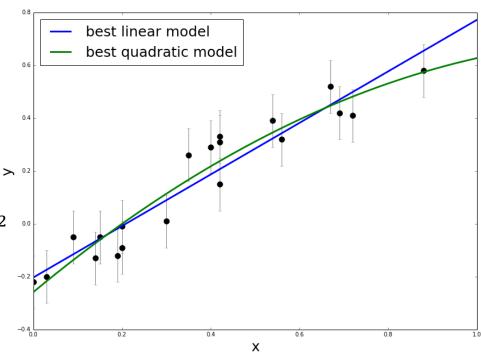
$$y_{M1} = f_{M1}(x; w) = w_0 + w_1 x$$

 $y \sim N(y_{M1}, \sigma_y^2)$

Quadratic model:

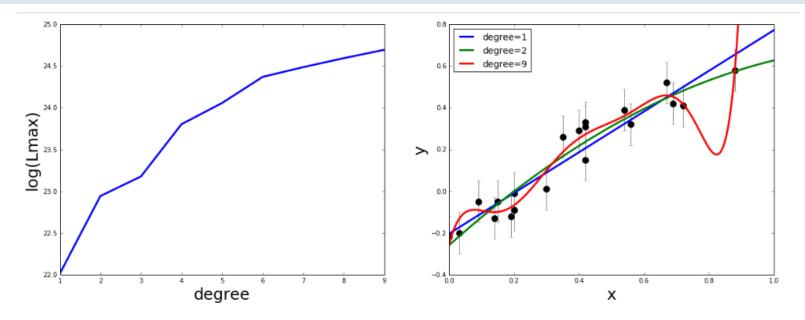
$$y_{M2} = f_{M2}(x; w) = w_0 + w_1 x + w_2 x^2$$

$$y \sim N(y_{M2}, \sigma_y^2)$$



Which model is better?

Model Complexity and Generality



- Comparing maximum likelihood $p(D|w,M_1)$ and $p(D|w,M_2)$ is not a good idea $\lim_w p(D|w,M_1) \quad v.\,s. \quad \lim_w p(D|w,M_2)$
- As more complex model is used, model better fits the data, however, this model cannot predict well on unseen test data
- Balancing between model fitting and generalization is a fundamental question in ML
- In frequentist approach, a complex model is penalized by additional regularization term

The Bayesian approach addresses this by integrating over the model parameter space, which in effect acts to automatically penalize overly-complex models.

Bayesian Approach for Model Selection

• The parameter posterior given the model *M* is expressed

$$p(w|D,M) = \frac{p(D|w,M)p(w|M)}{p(D|M)}$$

The model posterior can be expressed

$$p(M|D) = \frac{p(D|M)p(M)}{p(D)}$$

-p(M) is model prior representing preference on a certain model

$$-p(D|M) = \int_{W} p(D, w|M) dw = \int_{\Omega} p(D|w, M) p(w|M) d\omega$$

(Integration over the entire parameter space $w \in \Omega$)

• The odd ratio between two models, M_1 and M_2 , can be expressed

$$O_{21} = \frac{p(M_2|D)}{p(M_1|D)} = \frac{p(D|M_2)}{p(D|M_1)} \frac{p(M_2)}{p(M_1)}$$

 0_{21} > threshold Choose M_2

$$rac{p(D|M_2)}{p(D|M_1)}$$
 : Bayes factors $rac{p(M_2)}{p(M_1)}$: Prior odd ratio