

Feel free to work with other students, but make sure you write up the homework and code on your own (no copying homework or code; no pair programming). Feel free to ask students or instructors for help debugging code or whatever else, though.

1 (Murphy 2.16) Suppose $\theta \sim \text{Beta}(a, b)$ such that

$$\mathbb{P}(\theta; a, b) = \frac{1}{B(a, b)} \theta^{a-1} (1-\theta)^{b-1} = \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \theta^{a-1} (1-\theta)^{b-1}$$

where $B(a, b) = \Gamma(a)\Gamma(b)/\Gamma(a+b)$ is the Beta function and $\Gamma(x)$ is the Gamma function. Derive the mean, mode, and variance of θ .

Mean:

For a continuous rv, $\mathbb{E}[X] = \int_{\mathcal{X}} x p(x) dx$. In addition, Beta is in $[0, 1]$.

Thus, $\mathbb{E}[\theta] = \int_0^1 \theta p(\theta) d\theta$

$$= \int_0^1 \theta \left(\frac{1}{B(a, b)} \theta^{a-1} (1-\theta)^{b-1} \right) d\theta$$

$$= \int_0^1 \frac{\theta^a (1-\theta)^{b-1}}{B(a, b)} d\theta$$

$$= \frac{1}{B(a, b)} \int_0^1 \theta^a (1-\theta)^{b-1} d\theta$$

note that $B(x, y) = \int_0^1 \theta^{x-1} (1-\theta)^{y-1} d\theta$

$$= \frac{B(a+1, b)}{B(a, b)}$$

$$= \frac{\Gamma(a+1)\Gamma(b)}{\Gamma(a+b)} \cdot \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)}$$

since $\Gamma(x+1) = x\Gamma(x)$

$$= \frac{a\Gamma(a) \cdot \Gamma(a+b)}{(a+b)\Gamma(a+b) \cdot \Gamma(a)}$$

$$= \boxed{\frac{a}{a+b}}$$

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variance

$$\begin{aligned}
 \sigma^2 &= E[\theta^2] - \mu^2 \\
 &= \int_0^1 \theta^2 p(\theta) d\theta - \mu^2 \\
 &= \frac{1}{B(a,b)} \int_0^1 \theta^{a+1} (1-\theta)^b d\theta - \mu^2 \quad \text{substitute } B(a+2,b) \\
 &= \frac{B(a+2,b)}{B(a,b)} - \mu^2 \quad \Gamma(a+2) = (a+1)a\Gamma(a) \\
 &= \frac{(a+1)a\Gamma(a)\Gamma(b)}{\Gamma(a+2+b)\Gamma(a+b)} - \mu^2 \quad \Gamma(a+b+2) = (a+b+1)(a+b)\Gamma(a+b) \\
 &= \frac{(a+1)a}{(a+b+1)(a+b)} - \frac{a^2}{(a+b)^2} \\
 &= \frac{(a+1)a(a+b) - a^2(a+b+1)}{(a+b+1)(a+b)^2} \\
 &= \frac{(a^2+a)(a+b) - a^3 - a^2b - a^2}{(a+b+1)(a+b)^2} \\
 &= \frac{a^2 + a^2b + a^2 + ab - a^3 - a^2b - a^2}{(a+b+1)(a+b)^2} \\
 &= \boxed{\frac{ab}{(a+b)^2(a+b+1)}}
 \end{aligned}$$

mode

Since at the mode, the probability is at maximum, we can calculate it by setting the derivative to 0.

$$0 = \frac{dp(\theta)}{d\theta}$$

$$0 = \frac{1}{B(a,b)} \left[\frac{d}{d\theta} (\theta^{a-1} (1-\theta)^{b-1}) \right]$$

$$0 = \frac{1}{B(a,b)} \left[-\theta^{a-1}(b-1)(1-\theta)^{b-2} + (a-1)\theta^{a-2}(1-\theta)^{b-1} \right]$$

$$\theta^{a-1}(b-1)(1-\theta)^{b-2} = (a-1)\theta^{a-2}(1-\theta)^{b-1}$$

$$\theta = \frac{a-1}{b-1} (1-\theta) \Rightarrow \theta \left(1 + \frac{a-1}{b-1} \right) = \frac{a-1}{b-1} \Rightarrow \theta = \boxed{\frac{a-1}{a+b-2}}$$

2 (Murphy 9) Show that the multinomial distribution

$$\text{Cat}(\mathbf{x}|\boldsymbol{\mu}) = \prod_{i=1}^K \mu_i^{x_i}$$

is in the exponential family and show that the generalized linear model corresponding to this distribution is the same as multinomial logistic regression (softmax regression).

Show in the exponential family:

$$\begin{aligned}\text{Cat}(\mathbf{x}|\boldsymbol{\mu}) &= \exp\left(\log \prod_{i=1}^K \mu_i^{x_i}\right) \\ &= \exp\left(\sum_{i=1}^K x_i \log \mu_i\right) \\ &= \exp\left(\sum_{i=1}^K x_i \log \mu_i\right)\end{aligned}$$

In order to have $a(\boldsymbol{\eta})$, we pull out the k^{th} term

$$\begin{aligned}\text{Cat}(\mathbf{x}|\boldsymbol{\mu}) &= \exp\left(\sum_{i=1}^{K-1} x_i \log \mu_i + x_K \log \mu_K\right) \quad \text{since } \sum_{i=1}^K x_i = 1, \\ &= \exp\left[\sum_{i=1}^{K-1} x_i \log \mu_i + \left(1 - \sum_{i=1}^{K-1} x_i\right) \log \mu_K\right] \quad x_K = 1 - \sum_{i=1}^{K-1} x_i \\ &= \exp\left(\sum_{i=1}^{K-1} x_i \log \mu_i + \log \mu_K - \sum_{i=1}^{K-1} x_i \log \mu_K\right) \\ &= \exp\left(\sum_{i=1}^{K-1} x_i (\log \mu_i - \log \mu_K) + \log \mu_K\right) \\ &= \exp\left(\sum_{i=1}^{K-1} x_i \log\left(\frac{\mu_i}{\mu_K}\right) + \log \mu_K\right) \\ &= \exp\left(\begin{bmatrix} \log\left(\frac{\mu_1}{\mu_K}\right) \\ \log\left(\frac{\mu_2}{\mu_K}\right) \\ \vdots \\ \log\left(\frac{\mu_{K-1}}{\mu_K}\right) \end{bmatrix}^T \begin{bmatrix} x_1 \\ x_2 \\ \vdots \\ x_{K-1} \end{bmatrix} + \log \mu_K\right) \quad \blacksquare\end{aligned}$$

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Since the exponential family is in the form

$p(x; \eta) = b(x) \exp(\eta^T T(x) - a(\eta))$, if we let

$$\vec{\eta} = \begin{bmatrix} \log(\frac{\mu_1}{\mu_k}) \\ \vdots \\ \log(\frac{\mu_{k-1}}{\mu_k}) \end{bmatrix}, \quad \vec{x} = \begin{bmatrix} x_1 \\ \vdots \\ x_{k-1} \end{bmatrix}, \quad a(\vec{\eta}) = -\log \mu_k, \quad b(y) = 1,$$

then $\text{Cat}(x; \mu)$ is in the exponential family.

Show GLM = softmax regression.

write μ_i in terms of μ_k :

$$\eta_i = \log\left(\frac{\mu_i}{\mu_k}\right)$$

$$\mu_i = e^{\eta_i} \cdot \mu_k \quad (1)$$

Find μ_k in terms of η_i :

$$\mu_k = 1 - \sum_{i=1}^{k-1} \mu_i$$

$$\mu_k = 1 - \sum_{i=1}^{k-1} e^{\eta_i} \mu_k$$

$$\mu_k = \frac{1}{1 - \sum_{i=1}^{k-1} e^{\eta_i}} \quad (2)$$

write μ_i in terms of η_i :

$$\mu_i = e^{\eta_i} \mu_k \quad (\text{from (1)})$$

$$= \frac{e^{\eta_i}}{1 - \sum_{i=1}^{k-1} e^{\eta_i}}$$

$$= \text{softmax function } S(\eta)_i$$

