

# ***STATISTICAL INFERENCES (2cr)***

## Chapter 8 Sampling Distributions & Data Descriptions

Zhong Guan

Math, IUSB

# Outline

- 1 8.3 Sampling Distributions
- 2 8.4 Sampling Distribution of Means
  - Central Limit Theorem
  - Inferences on the Population Mean
  - Sampling Distribution of Difference between Two Means

# Definition

**Sampling distribution** is the probability distribution of a **statistic**.

For example, the probability distribution of the sampling mean  $\bar{X}$  is called the sampling distribution of the mean.

## Definition

**Sampling distribution** is the probability distribution of a **statistic**.

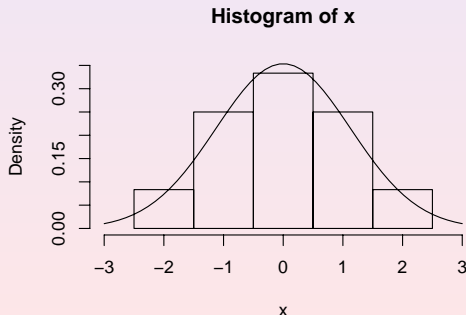
For example, the probability distribution of the sampling mean  $\bar{X}$  is called the sampling distribution of the mean.

# Examples

**Example 1.** Suppose we have a population containing the following values:

-2, -1, -1, -1, 0, 0, 0, 0, 1, 1, 1, 2

Distribution of the data is close to:  $N(0, 7/6)$ .

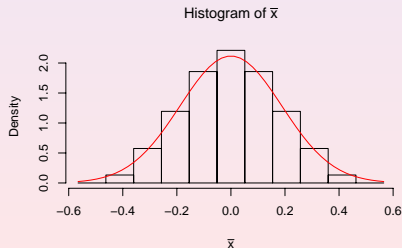


# Distribution of All Sample Means

Let  $X_1, X_2, \dots, X_n$  be a sample of size  $n = 9$ .

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$$

There are  $\binom{12}{9} = 220$  sample means. The distribution of all the sample means is also normal  $N(0, 0.0354)$ .

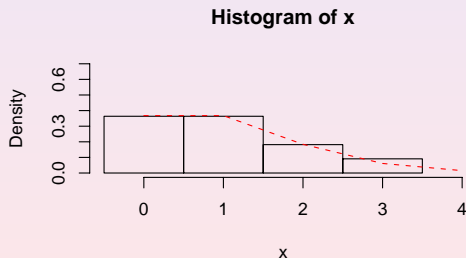


# Examples

**Example 2.** Suppose we have a skewed population containing the following values:

0, 0, 0, 0, 1, 1, 1, 1, 2, 2, 3

Distribution of the data is close to Poisson:  $P(1)$ .

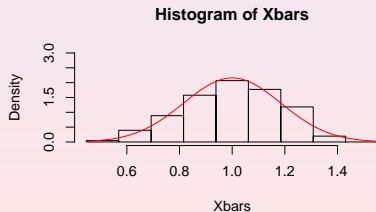


# Distribution of All Sample Means

Let  $X_1, X_2, \dots, X_n$  be a sample of size  $n = 8$ .

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$$

There are  $\binom{11}{8} = 165$  sample means. The distribution of all the sample means is also close to normal  $N(1, 0.034)$ .





# Outline

## 1 8.3 Sampling Distributions

## 2 8.4 Sampling Distribution of Means

- Central Limit Theorem
- Inferences on the Population Mean
- Sampling Distribution of Difference between Two Means

# Exact distribution of sample mean

Let  $X_1, X_2, \dots, X_n$  be a sample of size  $n$ . That is,  $X_1, X_2, \dots, X_n$  are independent random variables having the same distribution with mean  $\mu$  and variance  $\sigma^2$ .

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i = \frac{1}{n} (X_1 + X_2 + \dots + X_n)$$

$$E(\bar{X}) = \mu, \quad \text{Var}(\bar{X}) = \frac{\sigma^2}{n}$$

If the sample is from  $N(\mu, \sigma^2)$ , then exactly

$$\bar{X} \sim N\left(\mu, \frac{\sigma^2}{n}\right) \quad \text{and} \quad \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \sim N(0, 1).$$

What if the sample is not from  $N(\mu, \sigma^2)$ ?

# Exact distribution of sample mean

Let  $X_1, X_2, \dots, X_n$  be a sample of size  $n$ . That is,  $X_1, X_2, \dots, X_n$  are independent random variables having the same distribution with mean  $\mu$  and variance  $\sigma^2$ .

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i = \frac{1}{n} (X_1 + X_2 + \dots + X_n)$$

$$E(\bar{X}) = \mu, \quad \text{Var}(\bar{X}) = \frac{\sigma^2}{n}$$

If the sample is from  $N(\mu, \sigma^2)$ , then exactly

$$\bar{X} \sim N\left(\mu, \frac{\sigma^2}{n}\right) \quad \text{and} \quad \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \sim N(0, 1).$$

What if the sample is not from  $N(\mu, \sigma^2)$ ?

# Exact distribution of sample mean

Let  $X_1, X_2, \dots, X_n$  be a sample of size  $n$ . That is,  $X_1, X_2, \dots, X_n$  are independent random variables having the same distribution with mean  $\mu$  and variance  $\sigma^2$ .

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i = \frac{1}{n} (X_1 + X_2 + \dots + X_n)$$

$$E(\bar{X}) = \mu, \quad \text{Var}(\bar{X}) = \frac{\sigma^2}{n}$$

If the sample is from  $N(\mu, \sigma^2)$ , then exactly

$$\bar{X} \sim N\left(\mu, \frac{\sigma^2}{n}\right) \quad \text{and} \quad \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \sim N(0, 1).$$

What if the sample is not from  $N(\mu, \sigma^2)$ ?

# Exact distribution of sample mean

Let  $X_1, X_2, \dots, X_n$  be a sample of size  $n$ . That is,  $X_1, X_2, \dots, X_n$  are independent random variables having the same distribution with mean  $\mu$  and variance  $\sigma^2$ .

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i = \frac{1}{n}(X_1 + X_2 + \dots + X_n)$$

$$E(\bar{X}) = \mu, \quad \text{Var}(\bar{X}) = \frac{\sigma^2}{n}$$

If the sample is from  $N(\mu, \sigma^2)$ , then **exactly**

$$\bar{X} \sim N\left(\mu, \frac{\sigma^2}{n}\right) \quad \text{and} \quad \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \sim N(0, 1).$$

What if the sample is not from  $N(\mu, \sigma^2)$ ?

# Exact distribution of sample mean

Let  $X_1, X_2, \dots, X_n$  be a sample of size  $n$ . That is,  $X_1, X_2, \dots, X_n$  are independent random variables having the same distribution with mean  $\mu$  and variance  $\sigma^2$ .

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i = \frac{1}{n}(X_1 + X_2 + \dots + X_n)$$

$$E(\bar{X}) = \mu, \quad \text{Var}(\bar{X}) = \frac{\sigma^2}{n}$$

If the sample is from  $N(\mu, \sigma^2)$ , then **exactly**

$$\bar{X} \sim N\left(\mu, \frac{\sigma^2}{n}\right) \quad \text{and} \quad \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \sim N(0, 1).$$

What if the sample is not from  $N(\mu, \sigma^2)$ ?

# Central Limit Theorem

**Central Limit Theorem** If  $\bar{X}$  is the sample mean of a random sample  $X_1, X_2, \dots, X_n$  of size  $n$  from a distribution, discrete or continuous, with mean  $\mu$  and variance  $\sigma^2$ , then the distribution of

$$Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}}$$

approaches the standard normal distribution  $N(0, 1)$  as  $n \rightarrow \infty$ , i.e.,

$$\lim_{n \rightarrow \infty} P(Z \leq z) = \Phi(z).$$

So if  $n \geq 30$  (or smaller for symmetric distribution), then the distribution of  $\bar{X}$  is approximately  $N(\mu, \sigma^2/n)$  and

$$P(a < Z \leq b) \approx \Phi(b) - \Phi(a), \quad a < b.$$

# Central Limit Theorem

**Central Limit Theorem** If  $\bar{X}$  is the sample mean of a random sample  $X_1, X_2, \dots, X_n$  of size  $n$  from a distribution, discrete or continuous, with mean  $\mu$  and variance  $\sigma^2$ , then the distribution of

$$Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}}$$

approaches the standard normal distribution  $N(0, 1)$  as  $n \rightarrow \infty$ , i.e.,

$$\lim_{n \rightarrow \infty} P(Z \leq z) = \Phi(z).$$

So if  $n \geq 30$  (or smaller for symmetric distribution), then the distribution of  $\bar{X}$  is approximately  $N(\mu, \sigma^2/n)$  and

$$P(a < Z \leq b) \approx \Phi(b) - \Phi(a), \quad a < b.$$



# Central Limit Theorem

**Central Limit Theorem** If  $\bar{X}$  is the sample mean of a random sample  $X_1, X_2, \dots, X_n$  of size  $n$  from a distribution, discrete or continuous, with mean  $\mu$  and variance  $\sigma^2$ , then the distribution of

$$Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}}$$

approaches the standard normal distribution  $N(0, 1)$  as  $n \rightarrow \infty$ , i.e.,

$$\lim_{n \rightarrow \infty} P(Z \leq z) = \Phi(z).$$

So if  $n \geq 30$  (or smaller for symmetric distribution), then the distribution of  $\bar{X}$  is **approximately**  $N(\mu, \sigma^2/n)$  and

$$P(a < Z \leq b) \approx \Phi(b) - \Phi(a), \quad a < b.$$

# Examples

**Example 1.** Let  $\bar{X}$  be the sample mean of a random sample  $X_1, X_2, \dots, X_n$  of size  $n$  from  $U[0, 1]$ , the uniform distribution on  $[0, 1]$ .

$$\mu = E(X_1) = \frac{1}{2}, \quad \sigma^2 = \text{Var}(X_1) = \frac{1}{12}.$$

$$Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} = \sqrt{12n}(\bar{X} - \frac{1}{2})$$

# Examples

**Example 1.** Let  $\bar{X}$  be the sample mean of a random sample  $X_1, X_2, \dots, X_n$  of size  $n$  from  $U[0, 1]$ , the uniform distribution on  $[0, 1]$ .

$$\mu = E(X_1) = \frac{1}{2}, \quad \sigma^2 = \text{Var}(X_1) = \frac{1}{12}.$$

$$Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} = \sqrt{12n}(\bar{X} - \frac{1}{2})$$

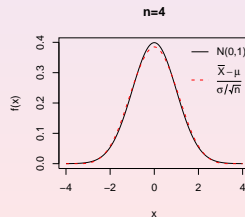
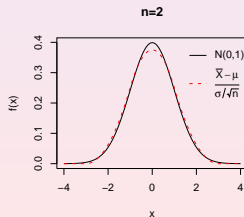
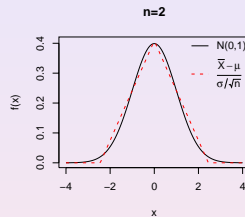
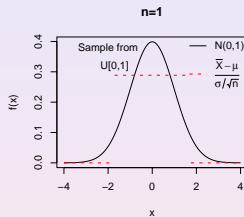
# Examples

**Example 1.** Let  $\bar{X}$  be the sample mean of a random sample  $X_1, X_2, \dots, X_n$  of size  $n$  from  $U[0, 1]$ , the uniform distribution on  $[0, 1]$ .

$$\mu = E(X_1) = \frac{1}{2}, \quad \sigma^2 = \text{Var}(X_1) = \frac{1}{12}.$$

$$Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} = \sqrt{12n}(\bar{X} - \frac{1}{2})$$

# Examples



# Examples

**Example 2.** Let  $\bar{X}$  be the sample mean of a random sample  $X_1, X_2, \dots, X_n$  of size  $n$  from  $EXP(\theta)$ , the exponential distribution with

$$\mu = E(X_1) = \theta, \quad \sigma^2 = \text{Var}(X_1) = \theta^2.$$

$$Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} = \frac{\bar{X} - \theta}{\theta/\sqrt{n}}$$

# Examples

**Example 2.** Let  $\bar{X}$  be the sample mean of a random sample  $X_1, X_2, \dots, X_n$  of size  $n$  from  $EXP(\theta)$ , the exponential distribution with

$$\mu = E(X_1) = \theta, \quad \sigma^2 = \text{Var}(X_1) = \theta^2.$$

$$Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} = \frac{\bar{X} - \theta}{\theta/\sqrt{n}}$$

# Examples

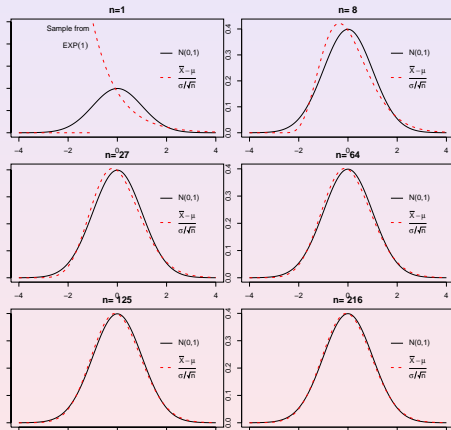
**Example 2.** Let  $\bar{X}$  be the sample mean of a random sample  $X_1, X_2, \dots, X_n$  of size  $n$  from  $EXP(\theta)$ , the exponential distribution with

$$\mu = E(X_1) = \theta, \quad \sigma^2 = \text{Var}(X_1) = \theta^2.$$

$$Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} = \frac{\bar{X} - \theta}{\theta/\sqrt{n}}$$



# Examples



# Examples

**Example 3.** Let  $\bar{X}$  be the sample mean of a random sample  $X_1, X_2, \dots, X_n$  of size  $n$  from  $\chi^2(k)$ , the chi-square distribution with

$$\mu = E(X_1) = k, \quad \sigma^2 = \text{Var}(X_1) = 2k.$$

$$Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} = \frac{\bar{X} - k}{2k/\sqrt{n}}$$

# Examples

**Example 3.** Let  $\bar{X}$  be the sample mean of a random sample  $X_1, X_2, \dots, X_n$  of size  $n$  from  $\chi^2(k)$ , the chi-square distribution with

$$\mu = E(X_1) = k, \quad \sigma^2 = \text{Var}(X_1) = 2k.$$

$$Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} = \frac{\bar{X} - k}{2k/\sqrt{n}}$$

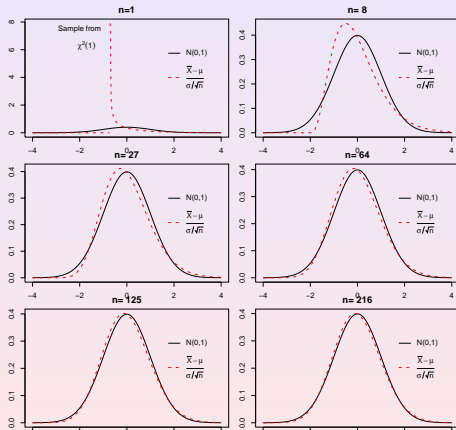
# Examples

**Example 3.** Let  $\bar{X}$  be the sample mean of a random sample  $X_1, X_2, \dots, X_n$  of size  $n$  from  $\chi^2(k)$ , the chi-square distribution with

$$\mu = E(X_1) = k, \quad \sigma^2 = \text{Var}(X_1) = 2k.$$

$$Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} = \frac{\bar{X} - k}{2k/\sqrt{n}}$$

# Examples



# Examples

The central limit theorem is also valid for discrete distribution.

**Example 4.** Let  $\bar{X}$  be the sample mean of a random sample  $X_1, X_2, \dots, X_n$  of size  $n$  from  $b(1, p)$  with

$$\mu = p, \quad \sigma^2 = p(1 - p).$$

$$Y = \sum_{i=1}^n X_i \sim b(n, p).$$

By the CLT,

$$\bar{X} \sim N(\mu, \sigma^2/n), \text{ approximately.}$$

$$Y = n\bar{X} \sim N(n\mu, n^2\sigma^2/n) = N[np, np(1-p)], \text{ approximately.}$$

# Examples

The central limit theorem is also valid for discrete distribution.

**Example 4.** Let  $\bar{X}$  be the sample mean of a random sample  $X_1, X_2, \dots, X_n$  of size  $n$  from  $b(1, p)$  with

$$\mu = p, \quad \sigma^2 = p(1 - p).$$

$$Y = \sum_{i=1}^n X_i \sim b(n, p).$$

By the CLT,

$$\bar{X} \sim N(\mu, \sigma^2/n), \text{ approximately.}$$

$$Y = n\bar{X} \sim N(n\mu, n^2\sigma^2/n) = N[np, np(1-p)], \text{ approximately.}$$

# Examples

The central limit theorem is also valid for discrete distribution.

**Example 4.** Let  $\bar{X}$  be the sample mean of a random sample  $X_1, X_2, \dots, X_n$  of size  $n$  from  $b(1, p)$  with

$$\mu = p, \quad \sigma^2 = p(1 - p).$$

$$Y = \sum_{i=1}^n X_i \sim b(n, p).$$

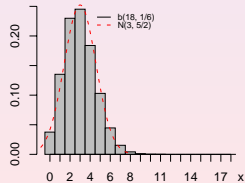
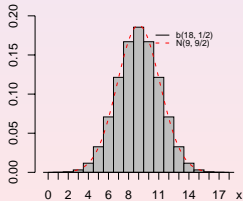
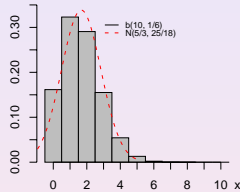
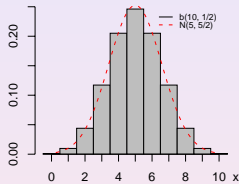
By the CLT,

$$\bar{X} \sim N(\mu, \sigma^2/n), \text{ approximately.}$$

$$Y = n\bar{X} \sim N(n\mu, n^2\sigma^2/n) = N[np, np(1-p)], \text{ approximately.}$$



# Examples



## Examples

**Example 5.** If a certain machine makes resistors have a mean resistance of 40 ohms and standard deviation of 2 ohms, what is the probability that a random sample of 36 of these resistors will have an average resistance of more than 40.5 ohms.

# Examples

**Solution of Example 5.**  $n = 36$ ,  $\mu = 40$ ,  $\sigma = 2$ . By the CLT,  $\bar{X}$  has normal distribution with mean  $\mu_{\bar{X}} = \mu = 40$  and standard deviation  $\sigma_{\bar{X}} = \sigma/\sqrt{n} = 2/\sqrt{36} = 2/6 = 1/3$ . The  $z$  value of  $x = 40.5$  is

$$z = \frac{x - \mu_{\bar{X}}}{\sigma_{\bar{X}}} = \frac{40.5 - 40}{1/3} = 3(0.5) = 1.5$$

Using Table A.3

$$P(\bar{X} > 40.5) = P(Z > 1.5) = 1 - P(Z \leq 1.5) = 1 - 0.9332 = 0.0668$$

# Examples

**Solution of Example 5.**  $n = 36$ ,  $\mu = 40$ ,  $\sigma = 2$ . By the CLT,  $\bar{X}$  has normal distribution with mean  $\mu_{\bar{X}} = \mu = 40$  and standard deviation  $\sigma_{\bar{X}} = \sigma/\sqrt{n} = 2/\sqrt{36} = 2/6 = 1/3$ . The  $z$  value of  $x = 40.5$  is

$$z = \frac{x - \mu_{\bar{X}}}{\sigma_{\bar{X}}} = \frac{40.5 - 40}{1/3} = 3(0.5) = 1.5$$

Using Table A.3

$$P(\bar{X} > 40.5) = P(Z > 1.5) = 1 - P(Z \leq 1.5) = 1 - 0.9332 = 0.0668$$

# Outline

## 1 8.3 Sampling Distributions

## 2 8.4 Sampling Distribution of Means

- Central Limit Theorem
- Inferences on the Population Mean
- Sampling Distribution of Difference between Two Means

# Examples

**Example 6.[Example 5. (Cont.)]** An observed sample of 36 of these resistors indicates a sample average of 39.1 ohms. Does this sample information appear to support or refute the conjecture that  $\mu = 40$  ohms?

**Solution of Example 6.** If the conjecture  $\mu = 40$  is true, then by the CLT  $\bar{X}$  with  $n = 36$  is approximately normal with mean  $\mu = 40$  and standard deviation  $\sigma/\sqrt{n} = 2/\sqrt{36} = 1/3$ . How likely can the value of  $\bar{X}$  be as far away from the center  $\mu = 40$  as the observed  $\bar{x} = 39.1$ ?

That is, if  $\mu = 40$ ,

$$P(|\bar{X} - 40| \geq |39.1 - 40|) = P(|\bar{X} - 40| \geq 0.9) = ?$$

# Examples

**Example 6.[Example 5. (Cont.)]** An observed sample of 36 of these resistors indicates a sample average of 39.1 ohms. Does this sample information appear to support or refute the conjecture that  $\mu = 40$  ohms?

**Solution of Example 6.** If the conjecture  $\mu = 40$  is true, then by the CLT  $\bar{X}$  with  $n = 36$  is approximately normal with mean  $\mu = 40$  and standard deviation  $\sigma/\sqrt{n} = 2/\sqrt{36} = 1/3$ .

How likely can the value of  $\bar{X}$  be as far away from the center  $\mu = 40$  as the observed  $\bar{x} = 39.1$ ?

That is, if  $\mu = 40$ ,

$$P(|\bar{X} - 40| \geq |39.1 - 40|) = P(|\bar{X} - 40| \geq 0.9) = ?$$

# Examples

**Example 6.[Example 5. (Cont.)]** An observed sample of 36 of these resistors indicates a sample average of 39.1 ohms. Does this sample information appear to support or refute the conjecture that  $\mu = 40$  ohms?

**Solution of Example 6.** If the conjecture  $\mu = 40$  is true, then by the CLT  $\bar{X}$  with  $n = 36$  is approximately normal with mean  $\mu = 40$  and standard deviation  $\sigma/\sqrt{n} = 2/\sqrt{36} = 1/3$ . How likely can the value of  $\bar{X}$  be as far away from the center  $\mu = 40$  as the observed  $\bar{x} = 39.1$ ?

That is, if  $\mu = 40$ ,

$$P(|\bar{X} - 40| \geq |39.1 - 40|) = P(|\bar{X} - 40| \geq 0.9) = ?$$



# Examples

**Example 6.[Example 5. (Cont.)]** An observed sample of 36 of these resistors indicates a sample average of 39.1 ohms. Does this sample information appear to support or refute the conjecture that  $\mu = 40$  ohms?

**Solution of Example 6.** If the conjecture  $\mu = 40$  is true, then by the CLT  $\bar{X}$  with  $n = 36$  is approximately normal with mean  $\mu = 40$  and standard deviation  $\sigma/\sqrt{n} = 2/\sqrt{36} = 1/3$ . How likely can the value of  $\bar{X}$  be as far away from the center  $\mu = 40$  as the observed  $\bar{x} = 39.1$ ?

That is, if  $\mu = 40$ ,

$$P(|\bar{X} - 40| \geq |39.1 - 40|) = P(|\bar{X} - 40| \geq 0.9) = ?$$

## Solution of Example 6.

By the CLT, if  $\mu = 40$ ,  $Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} = \frac{\bar{X} - 40}{2/\sqrt{36}} = 3(\bar{X} - 40)$ .

$$P(|\bar{X} - 40| \geq 0.9) = P(|Z| \geq 3(0.9)) = P(|Z| \geq 2.7)$$

$$= P(Z \geq 2.7) + P(Z \leq -2.7)$$

$$= P(Z \geq 2.7) + P(-Z \geq 2.7) = 2P(Z \geq 2.7)$$

$$= 2[1 - P(Z < 2.7)] = 2(1 - 0.996533) \approx 0.007$$

One would experience by chance that an  $\bar{x}$  is 0.9 ohms from the mean in only 7 in 1000 samples of size 36. This sample is an evidence against the conjecture  $\mu = 40$  ohms.

## Solution of Example 6.

By the CLT, if  $\mu = 40$ ,  $Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} = \frac{\bar{X} - 40}{2/\sqrt{36}} = 3(\bar{X} - 40)$ .

$$P(|\bar{X} - 40| \geq 0.9) = P(|Z| \geq 3(0.9)) = P(|Z| \geq 2.7)$$

$$= P(Z \geq 2.7) + P(Z \leq -2.7)$$

$$= P(Z \geq 2.7) + P(-Z \geq 2.7) = 2P(Z \geq 2.7)$$

$$= 2[1 - P(Z < 2.7)] = 2(1 - 0.996533) \approx 0.007$$

One would experience by chance that an  $\bar{x}$  is 0.9 ohms from the mean in only 7 in 1000 samples of size 36. This sample is an evidence against the conjecture  $\mu = 40$  ohms.

## Solution of Example 6.

By the CLT, if  $\mu = 40$ ,  $Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} = \frac{\bar{X} - 40}{2/\sqrt{36}} = 3(\bar{X} - 40)$ .

$$P(|\bar{X} - 40| \geq 0.9) = P(|Z| \geq 3(0.9)) = P(|Z| \geq 2.7)$$

$$= P(Z \geq 2.7) + P(Z \leq -2.7)$$

$$= P(Z \geq 2.7) + P(-Z \geq 2.7) = 2P(Z \geq 2.7)$$

$$= 2[1 - P(Z < 2.7)] = 2(1 - 0.996533) \approx 0.007$$

One would experience by chance that an  $\bar{x}$  is 0.9 ohms from the mean in only 7 in 1000 samples of size 36. This sample is an evidence against the conjecture  $\mu = 40$  ohms.

## Solution of Example 6.

By the CLT, if  $\mu = 40$ ,  $Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} = \frac{\bar{X} - 40}{2/\sqrt{36}} = 3(\bar{X} - 40)$ .

$$P(|\bar{X} - 40| \geq 0.9) = P(|Z| \geq 3(0.9)) = P(|Z| \geq 2.7)$$

$$= P(Z \geq 2.7) + P(Z \leq -2.7)$$

$$= P(Z \geq 2.7) + P(-Z \geq 2.7) = 2P(Z \geq 2.7)$$

$$= 2[1 - P(Z < 2.7)] = 2(1 - 0.996533) \approx 0.007$$

One would experience by chance that an  $\bar{x}$  is 0.9 ohms from the mean in only 7 in 1000 samples of size 36. This sample is an evidence against the conjecture  $\mu = 40$  ohms.

## Solution of Example 6.

By the CLT, if  $\mu = 40$ ,  $Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} = \frac{\bar{X} - 40}{2/\sqrt{36}} = 3(\bar{X} - 40)$ .

$$P(|\bar{X} - 40| \geq 0.9) = P(|Z| \geq 3(0.9)) = P(|Z| \geq 2.7)$$

$$= P(Z \geq 2.7) + P(Z \leq -2.7)$$

$$= P(Z \geq 2.7) + P(-Z \geq 2.7) = 2P(Z \geq 2.7)$$

$$= 2[1 - P(Z < 2.7)] = 2(1 - 0.996533) \approx 0.007$$

One would experience by chance that an  $\bar{x}$  is 0.9 ohms from the mean in only 7 in 1000 samples of size 36. This sample is an evidence against the conjecture  $\mu = 40$  ohms.

## Solution of Example 6.

By the CLT, if  $\mu = 40$ ,  $Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} = \frac{\bar{X} - 40}{2/\sqrt{36}} = 3(\bar{X} - 40)$ .

$$P(|\bar{X} - 40| \geq 0.9) = P(|Z| \geq 3(0.9)) = P(|Z| \geq 2.7)$$

$$= P(Z \geq 2.7) + P(Z \leq -2.7)$$

$$= P(Z \geq 2.7) + P(-Z \geq 2.7) = 2P(Z \geq 2.7)$$

$$= 2[1 - P(Z < 2.7)] = 2(1 - 0.996533) \approx 0.007$$

One would expect by chance that an  $\bar{x}$  is 0.9 ohms from the mean in only 7 in 1000 samples of size 36. This sample is an evidence against the conjecture  $\mu = 40$  ohms.

## Solution of Example 6.

- The probability  $P(|\bar{X} - 40| \geq 0.9 | \mu = 40)$  is called the  $p$ -value of the sample mean  $\bar{x} = 39.1$ .
- Under the condition that the conjecture or hypothesis  $\mu = 40$  is true,  $p$ -value is the probability that we can observed an  $\bar{X}$  as extreme as  $\bar{x} = 39.1$ .
- $p$ -value is **NOT** the probability that the conjecture or hypothesis  $\mu = 40$  is true.



## Solution of Example 6.

- The probability  $P(|\bar{X} - 40| \geq 0.9 | \mu = 40)$  is called the  $p$ -value of the sample mean  $\bar{x} = 39.1$ .
- Under the condition that the conjecture or hypothesis  $\mu = 40$  is true,  $p$ -value is the probability that we can observed an  $\bar{X}$  as extreme as  $\bar{x} = 39.1$ .
- $p$ -value is NOT the probability that the conjecture or hypothesis  $\mu = 40$  is true.

## Solution of Example 6.

- The probability  $P(|\bar{X} - 40| \geq 0.9 | \mu = 40)$  is called the  $p$ -value of the sample mean  $\bar{x} = 39.1$ .
- Under the condition that the conjecture or hypothesis  $\mu = 40$  is true,  $p$ -value is the probability that we can observed an  $\bar{X}$  as extreme as  $\bar{x} = 39.1$ .
- $p$ -value is **NOT** the probability that the conjecture or hypothesis  $\mu = 40$  is true.

# Outline

## 1 8.3 Sampling Distributions

## 2 8.4 Sampling Distribution of Means

- Central Limit Theorem
- Inferences on the Population Mean
- Sampling Distribution of Difference between Two Means

# Exact Distribution of Difference between Two Means

If  $\bar{X}_1$  and  $\bar{X}_2$  are the sample means of independent random samples of size  $n_1$  and  $n_2$  from two **normal** distributions with means  $\mu_1$  and  $\mu_2$  and variances  $\sigma_1^2$  and  $\sigma_2^2$ , respectively, then the distribution of  $\bar{X}_1 - \bar{X}_2$  is **exactly** normal with mean  $\mu_1 - \mu_2$  and variance  $(\sigma_1^2/n_1) + (\sigma_2^2/n_2)$ . So the distribution of

$$Z = \frac{(\bar{X}_1 - \bar{X}_2) - (\mu_1 - \mu_2)}{\sqrt{(\sigma_1^2/n_1) + (\sigma_2^2/n_2)}}$$

**exactly** standard normal  $N(0, 1)$ .

# CLT for Difference between Two Means

## Central Limit Theorem for Difference between Two Means

If  $\bar{X}_1$  and  $\bar{X}_2$  are the sample means of independent random samples of size  $n_1$  and  $n_2$  from two **nonnormal** distributions, **discrete or continuous**, with means  $\mu_1$  and  $\mu_2$  and variances  $\sigma_1^2$  and  $\sigma_2^2$ , respectively, then the distribution of

$$Z = \frac{(\bar{X}_1 - \bar{X}_2) - (\mu_1 - \mu_2)}{\sqrt{(\sigma_1^2/n_1) + (\sigma_2^2/n_2)}}$$

**approximately** standard normal distribution  $N(0, 1)$  if  $n_1, n_2 \geq 30$ .

# Examples

**Example 7.** Two independent samples of size 18 are selected from two types of paints,  $A$  and  $B$ . The average drying time, in hours, is recorded for each sample. Assume that **the populations are normal** with the same means and the population standard deviations are both known to be 1.0. Find the probability that the difference between the two means is greater than 1.0.

**Solution of Example 7.**  $\mu_A = \mu_B$ ,  $\sigma_A = \sigma_B = 1$ , and  $n_A = n_B = 18$ . Because **the populations are normal**

$$Z = \frac{(\bar{X}_A - \bar{X}_B) - (\mu_A - \mu_B)}{\sqrt{(\sigma_A^2/n_A) + (\sigma_B^2/n_B)}} = \frac{(\bar{X}_1 - \bar{X}_2) - 0}{\sqrt{(1/18) + (1/18)}} = 3(\bar{X}_1 - \bar{X}_2)$$

**is exactly standard normal.**

# Examples

**Example 7.** Two independent samples of size 18 are selected from two types of paints,  $A$  and  $B$ . The average drying time, in hours, is recorded for each sample. Assume that **the populations are normal** with the same means and the population standard deviations are both known to be 1.0. Find the probability that the difference between the two means is greater than 1.0.

**Solution of Example 7.**  $\mu_A = \mu_B$ ,  $\sigma_A = \sigma_B = 1$ , and  $n_A = n_B = 18$ . Because **the populations are normal**

$$Z = \frac{(\bar{X}_A - \bar{X}_B) - (\mu_A - \mu_B)}{\sqrt{(\sigma_A^2/n_A) + (\sigma_B^2/n_B)}} = \frac{(\bar{X}_1 - \bar{X}_2) - 0}{\sqrt{(1/18) + (1/18)}} = 3(\bar{X}_1 - \bar{X}_2)$$

is **exactly** standard normal.

## Solution of Example 7.

So

$$\begin{aligned}P(|\bar{X}_A - \bar{X}_B| \geq 1) &= P(|Z| \geq 3) = 2(1 - P(Z < 3)) = 2(1 - 0.9987) \\&= 2(0.0013) = 0.0026\end{aligned}$$

One would experience by chance that a difference between the two means is bigger than 1 in only 2.6 in 1000 pairs of samples of size 18.



## Solution of Example 7.

So

$$\begin{aligned}P(|\bar{X}_A - \bar{X}_B| \geq 1) &= P(|Z| \geq 3) = 2(1 - P(Z < 3)) = 2(1 - 0.9987) \\&= 2(0.0013) = 0.0026\end{aligned}$$

One would expect by chance that a difference between the two means is bigger than 1 in only 2.6 in 1000 pairs of samples of size 18.

## Examples

**Example 7.(Cont.)** Suppose a difference of 1.0 in means was observed in real samples.

(a) Does this seem to be a reasonable results if the two population mean drying times truly are equal?

No.

(b) If someone selected 10,000 pairs of samples of size 18 under the condition that  $\mu_A = \mu_B$ , in how many of these 10,000 experiments would there be a difference  $\bar{x}_A - \bar{x}_B$  is as large as 1.0?

26.

## Examples

**Example 7.(Cont.)** Suppose a difference of 1.0 in means was observed in real samples.

(a) Does this seem to be a reasonable results if the two population mean drying times truly are equal?

No.

(b) If someone selected 10,000 pairs of samples of size 18 under the condition that  $\mu_A = \mu_B$ , in how many of these 10,000 experiments would there be a difference  $\bar{x}_A - \bar{x}_B$  is as large as 1.0?

26.

## Examples

**Example 7.(Cont.)** Suppose a difference of 1.0 in means was observed in real samples.

(a) Does this seem to be a reasonable results if the two population mean drying times truly are equal?

No.

(b) If someone selected 10,000 pairs of samples of size 18 under the condition that  $\mu_A = \mu_B$ , in how many of these 10,000 experiments would there be a difference  $\bar{x}_A - \bar{x}_B$  is as large as 1.0?

26.

## Examples

**Example 7.(Cont.)** Suppose a difference of 1.0 in means was observed in real samples.

(a) Does this seem to be a reasonable results if the two population mean drying times truly are equal?

No.

(b) If someone selected 10,000 pairs of samples of size 18 under the condition that  $\mu_A = \mu_B$ , in how many of these 10,000 experiments would there be a difference  $\bar{x}_A - \bar{x}_B$  is as large as 1.0?

26.

## Examples

**Example 7.(Cont.)** Suppose a difference of 1.0 in means was observed in real samples.

(a) Does this seem to be a reasonable results if the two population mean drying times truly are equal?

No.

(b) If someone selected 10,000 pairs of samples of size 18 under the condition that  $\mu_A = \mu_B$ , in how many of these 10,000 experiments would there be a difference  $\bar{x}_A - \bar{x}_B$  is as large as 1.0?

26.