

FINAL EXAM FORMULA SHEET

Probability	&	Counting
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Addition Rule $(E \cap F \neq \emptyset)$: P(E or F) = P(E) + P(F) - P(F)P(E and F)

Addition Rule for mutually exclusive events where $(E \cap F = \emptyset)$:

$$P(E \ or \ F) = P(E) + P(F)$$

Multiplication Rule for dependent events:

$$P(E \text{ and } F) = P(E) \cdot P(F|E) = P(F) \cdot P(E|F)$$

Multiplication Rule for independent events:

 $P(E \text{ and } F) = P(E) \cdot P(F)$

Combination (order doesn't matter & repeats not allowed)

$$_{n}C_{k} = \frac{n!}{k! (n-k)!}$$

Expected Value

 $E[x] = x_1p_1 + x_2p_2 + \cdots + x_np_n$

Permutation (order matters & repeats not allowed)

$$_{n}P_{k} = \frac{n!}{(n-k)!}$$

Conditional Probability:

$$P(F|E) = \frac{P(E \text{ and } F)}{P(E)}$$

Location: $L = n \cdot \frac{P}{100}$

Pth %ile: $P = \frac{L}{n} \cdot 100$

Sample Mean: $\bar{x} = \frac{\sum x}{n}$

Sample Standard Deviation:

$$S = \sqrt{\frac{\sum x^2 - \frac{1}{n}(\sum x)^2}{n - 1}}$$

Discrete Random Variables:

$$\mu = E(X) = \sum_{i=1}^{n} x_i p(x_i)$$

$$\mu = E(X) = \sum_{i=1}^{n} x_i \, p(x_i) \qquad \qquad \sigma^2 = E(X^2) - \mu^2 = \sum_{i=1}^{n} x_i^2 \, p(x_i) - \mu^2$$

Binomial Random Variables:

$$P(X = x) = \frac{n!}{x!(n-x)!} p^{x} (1-p)^{n-x}$$

$$\mu = E(X) = np$$
 $\sigma_X = \sqrt{np(1-p)}$

Poisson Random Variables:

$$P(X = x; \mu) = \frac{e^{-\mu}\mu^{x}}{x!} \qquad E(X) = \mu \qquad \sigma_{X} = \sqrt{\mu}$$

Continuous Random Variables:

$$P(a \le X \le b) = \int_a^b f(x)dx$$

$$\mu = E(X) = \int_{-\infty}^{\infty} f(x)dx \qquad Var(X) = E(X^2) - \mu^2 = \int_{-\infty}^{\infty} x^2 f(x)dx - \mu^2$$

Uniform Distribution: U[a, b]

$$\mu = E(X) = \frac{a+b}{2}$$
 $Var(X) = \frac{(b-a)^2}{12}$

Normal Distribution:

z-score:
$$z = \frac{x-\mu}{\sigma}$$

Jointly Distributed Random Variables: Discrete Case

1) Joint probability mass function (pmf)

$$p(x,y) = P(X = x \text{ and } Y = y)$$

$$p(x,y) \ge 0$$

$$\sum_{x} \sum_{y} p(x,y) = 1$$

$$P[(X,Y) \in A] = \sum_{(x,y)\in A} \sum_{y} p(x,y)$$

2) Marginal probability mass function of X

$$p_X(x) = \sum_{y: p(x,y) > 0} p(x,y)$$

Expected value of X

$$\mu_X = E(X) = \sum_X [x \cdot p_X(x)]$$

Variance of X

$$Var(X) = \sigma_X^2 = \sum_{x} [x^2 \cdot p_X(x)] - \mu_X^2$$

Standard deviation of X

$$\sigma_X = \sqrt{Var(X)} = \sqrt{\sum_X [x^2 \cdot p_X(x)] - \mu_X^2}$$

3) Marginal probability mass function of Y

$$p_Y(y) = \sum_{x: p(x,y) > 0} p(x,y)$$

Expected value of Y

$$\mu_Y = E(Y) = \sum_{y} [y \cdot p_Y(y)]$$

Variance of Y

$$Var(Y) = \sigma_Y^2 = \sum_{y} [y^2 \cdot p_Y(y)] - \mu_Y^2$$

Standard deviation of Y

$$\sigma_Y = \sqrt{Var(Y)} = \sqrt{\sum_y [y^2 \cdot p_Y(y)] - \mu_Y^2}$$

- 4) Independent Random Variables
 - X and Y are independent if

$$p(x,y) = p_X(x) \cdot p_Y(y)$$
 for all values of x and y

5) Expected value of a function h(X,Y)

$$E[h(X,Y)] = \sum_{x} \sum_{y} [h(x,y) \cdot p(x,y)]$$

Jointly Distributed Random Variables: Continuous Case

6) Joint probability density function (pdf)

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) dx dy = 1$$

$$P[(X, Y) \in A] = \int_{A} \int_{A} f(x, y) dx dy$$

$$P(a \le X \le b, c \le Y \le d) = \int_{a}^{b} \int_{c}^{d} f(x, y) dy dx$$

7) Marginal probability density function of X

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy$$

Expected value of X

$$\mu_X = E(X) = \int_{-\infty}^{\infty} x \cdot f_X(x) \ dx$$

Variance of X

$$Var(X) = \sigma_X^2 = \int_{-\infty}^{\infty} x^2 \cdot f_X(x) \ dx - \mu_X^2$$

Standard deviation of X

$$\sigma_X = \sqrt{Var(X)} = \sqrt{\int_{-\infty}^{\infty} x^2 \cdot f_X(x) \ dx - \mu_X^2}$$

8) Marginal probability density function of Y

$$f_Y(y) = \int_{-\infty}^{\infty} f(x, y) dx$$

Expected value of Y

$$\mu_Y = E(Y) = \int_{-\infty}^{\infty} y \cdot f_Y(y) dy$$

Variance of Y

$$Var(Y) = \sigma_Y^2 = \int_{-\infty}^{\infty} y^2 \cdot f_Y(y) \, dy - \mu_Y^2$$

Standard deviation of Y

$$\sigma_Y = \sqrt{Var(Y)} = \sqrt{\int_{-\infty}^{\infty} y^2 \cdot f_Y(y) \ dy - \mu_Y^2}$$

- 9) Independent Random Variables
 - X and Y are independent if

$$f(x,y) = f_X(x) \cdot f_Y(y)$$

10) Expected value of a function h(X, Y)

$$E[h(X,Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} h(x,y) \cdot f(x,y) dxdy$$

Sampling Distribution of Sample Means: Central Limit Theorem

For any given population with mean μ and standard deviation σ , a sampling distribution of sample means will have the following three characteristics if <u>either</u> the sample size, n, is at least 30 or the population is normally distributed.

1. The mean of a sampling distribution of sample means equals

$$\mu_{\bar{X}} = \mu$$

2. The standard deviation of a sampling distribution of sample means equals

$$\sigma_{\bar{X}} = \frac{\sigma}{\sqrt{n}}$$

3. Formula for z-score:

$$z = \frac{\bar{x} - \mu_{\bar{X}}}{\sigma_{\bar{X}}} = \frac{\bar{x} - \mu}{\frac{\sigma}{\sqrt{n}}}$$

Sampling Distribution of Sample Proportions: Central Limit Theorem

1) For large samples, the sample proportion is approximately normally distributed, with the mean and standard deviation given by the following formula:

$$\mu_{\hat{p}} = p$$

$$\sigma_{\hat{p}} = \sqrt{\frac{p(1-p)}{n}}$$

where p is the population proportion and n is the sample size.

- A sample is large if the interval $[p 3\sigma_{\hat{p}}, p + 3\sigma_{\hat{p}}]$ lies wholly within the interval [0, 1].
- 2) Formula for z-score

$$z = \frac{\hat{p} - \mu_{\hat{p}}}{\sigma_{\hat{p}}} = \frac{\hat{p} - p}{\sqrt{\frac{p(1-p)}{n}}}$$

Types of Errors in		Reality		
Hypothesis Testing		H_0 is true H_0 is false		
Dagislan	Reject H_0	Type 1 Error	Correct Decision	
Decision	Fail to reject H_0	Correct Decision	Type 2 Error	

С	$\alpha = 1 - c$	\mathbf{Z}_{α}	± Z 0.72
0.90	0.10	1.28	±1.645
0.95	0.05	1.645	±1.96
0.98	0.02	2.05	±2.33
0.99	0.01	2.33	±2.575

Inference about a Population Mean				
Inference	Condition	Confidence Interval	Test Statistic	df
n ≥ 30	σ known	$\bar{x} \pm z_{\alpha/2} \left(\frac{\sigma}{\sqrt{n}} \right)$	$z = \frac{\bar{x} - \mu}{\sigma / \sqrt{n}}$	N/A
	σ unknown	$\bar{x} \pm z_{\alpha/2} \left(\frac{s}{\sqrt{n}} \right)$	$z = \frac{\bar{x} - \mu}{s / \sqrt{n}}$	N/A
n < 30 and normal population	σ known	$\bar{x} \pm z_{\alpha/2} \left(\frac{\sigma}{\sqrt{n}} \right)$	$z = \frac{\bar{x} - \mu}{\sigma / \sqrt{n}}$	N/A
	σ unknown	$\bar{x} \pm t_{\alpha/2} \left(\frac{s}{\sqrt{n}} \right)$	$t = \frac{\bar{x} - \mu}{s / \sqrt{n}}$	n – 1

Inference about a Population Proportion

$$\hat{p} \pm 3\sqrt{\frac{\hat{p}(1-\hat{p})}{n}} \in [0,1] \qquad \hat{p} \pm z_{\alpha/2} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}} \qquad z = \frac{\hat{p}-p}{\sqrt{\frac{p}{n}(1-p)}} \qquad \text{N/A}$$

Calculate p-values

- For a left-tailed test, p-value = P(Z ≤ z).
- For a right-tailed test, p-value = P(Z ≥ z).
- For a two-tailed test, p-value = P(|Z| ≥ |z|).

Conclusions Using p-Values

- If p-value $\leq \alpha$, then reject the null hypothesis.
- If p-value > α , then fail to reject the null hypothesis.

Inference about Two Population Means				
Inference	Condition	Confidence Interval	Test Statistic	df
	$n_1 \ge 30$ $n_2 \ge 30$	$(\bar{x}_1 - \bar{x}_2) \pm z_{\alpha/2} \sqrt{\frac{s_1^2}{n_1} + \frac{s_1^2}{n_2}}$	$z = \frac{(\bar{x}_1 - \bar{x}_2) - (\mu_1 - \mu_2)}{\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}}$	N/A
Independent samples	$n < 30$ σ_1 and σ_2 unknown but assumed	$(\bar{x}_1 - \bar{x}_2) \pm t_{\alpha/2} \sqrt{s_p^2 \left(\frac{1}{n_1} + \frac{1}{n_2}\right)}$	$t = \frac{(\bar{x}_1 - \bar{x}_2) - (\mu_1 - \mu_2)}{\sqrt{s_p^2 \left(\frac{1}{n_1} + \frac{1}{n_2}\right)}}$	$n_1 + n_2 - 2$
po	equal, normal populations	$s_p^2 = \frac{(n_1 - 1)s_1^2 + (n_2 - 1)s_2^2}{n_1 + n_2 - 2}$		
Paired sample	·	$\bar{d} \pm t_{\alpha/2} \frac{s_d}{\sqrt{n}}$	$t = \frac{\bar{d} - \mu_d}{s_d / \sqrt{n}}$	n-1

Rejection Regions (in the case of z- test statistic)	 For left-tailed test: z ≤ -z_α or (-∞, -z_α) For right-tailed test: z ≥ z_α or (z_α, ∞) For two-tailed test: z ≥ z_{α/2} or (-∞, -z_{α/2}) ∪ (z_{α/2}, ∞)
Rejection Regions (in the case of <i>t</i> - test statistic)	 For left-tailed test: t ≤ -t_α or (-∞, -t_α) For right-tailed test: t ≥ t_α or (t_α, ∞) For two-tailed test: t ≥ t_{α/2} or (-∞, -t_{α/2}) ∪ (t_{α/2}, ∞)