Lecture 1b: EM Theory. Quiz 2 -> Up to & incl- boosting (proof) Recall: GMM -> easy ex. of EM alg. > there are many application $P(x \in B) = \sum_{j=1}^{n} P(j) N(x \in Alj \setminus S_{j})$ $\sum_{j=1}^{n} P(j) = 1$ $\sum_{j=1}^{n} P(j) = 1$ $\sum_{j=1}^{n} P(j) = 1$ $\sum_{j=1}^{n} P(j) = 1$ Optimization Method -> EM Alg. Q E → posterior prob. ⇒ Pα, (j|t) = P(j| Xt, Q()) Note: $p^{(\ell)}(j|t) = \mathbb{E}[S(j|t)|X, \underline{B}^{(\ell)}]$ illustrate why we can substitue $S(j|+) \rightarrow p^{(e)}(j|+)$ 2) $M: \widehat{\Omega}(j) = \sum_{t=1}^{n} p^{(t)}(j|t) \Rightarrow \text{effective } \# \text{ of samples } \widehat{n} \neq \text{th component}.$ $\sum_{i=1}^{m} \widehat{n}(i) = n$ Mixture proportions: $p^{(e+1)}(j) = \frac{\hat{n}(j)}{n}$ Component means: $\mu_{j}^{(eti)} = \frac{1}{\widehat{n}(j)} \sum_{t=1}^{N} p^{(e)}(j|t) \times t$ Component Covariance matrix: $\sum_{j}^{(\ell+1)} = \frac{1}{\widehat{p(j)}} \sum_{t=1}^{n} p^{(\ell)} (j|t) (x_{t} - \underline{M}_{j}^{(\ell)}) (X_{t} - \underline{M}_{j}^{(\ell)})^{T}$ Posterier prob. of It belonging to component j: $p^{(e)}(j|t) = \mathbb{E}[S(j|t)|X, \underline{0}^{(e)}] \rightarrow \text{Question}$

=
$$\Pr$$
 (sample Xt belongs to component $\hat{j} \mid \underline{\theta}^{(\ell)}$)

$$= \frac{P(j) P(Xt|j, \underline{O}^{(e)})}{\sum_{j'=1}^{m} P(j') P(Xt|j', \underline{D}^{(e)})} \longrightarrow \text{Bayes Rule}$$

Theorem to guarantee.

Let
$$f(Q) = \sum_{t} \log_{t} P(Xt; Q)$$
 be the log-likelihood on $X = \{X_1, ..., X_n\}$.

Theorem: [EM Thm]

Unless
$$\underline{\theta}^{(e)}$$
 is a stationary point of max $f(\underline{\theta})$, we have $f(\underline{\theta}^{(e_1)}) > f(\underline{\theta}^{(e)})$ $\forall \ \ell \in \mathbb{N}$ in any case!

Idea: Majorization - Minimization (MM) Alg-

Consider the opt- (minimum) prob.
$$\min_{x \in \mathbb{R}^d} f(x) \Rightarrow \underline{Bad}$$
 function

We assume that this is difficult. (directly)

Method Iterative

Successively minimize an auxiliary function
$$u(x, x^{(\ell)})$$

We hope
$$\{\chi^{(e)}\}_{e=1}^{\infty}$$
 converges to a stationary point!

to strong

we want to guarantee $f(x^{(\ell+1)}) \leq f(x^{(\ell)})$

 $\frac{\text{Defn}: \text{ An auxiliary function of } f: \mathbb{R}^d \to \mathbb{R} \text{ is a bivariate func.}}{u: \mathbb{R}^d \times \mathbb{R}^d \to \mathbb{R} \text{ s.t.}}$

 $\mathbb{O} \quad \mathsf{u}(\mathsf{x},\mathsf{x}) = \mathsf{f}(\mathsf{x})$

② $u(x, x^{(e)}) \ge f(x)$ $\forall (x, x^{(e)}) \in \mathbb{R}^{d_x} \mathbb{R}^{d}$

says that u(·, x(e)) is a Majorizor of for!

Say me have an auxiliary f^{\triangle} $u(x, x^{(e)})$ of f(x)

MM procedure is

$$\chi^{(l+1)} = \underset{\chi}{\operatorname{argmin}} \mathcal{N}(\chi, \chi^{(e)})$$

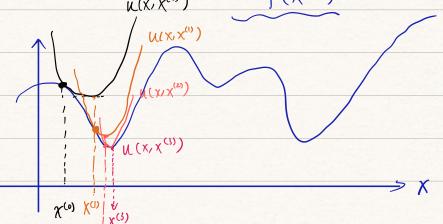
assume this is easier than prime prob. (substaintially)

[Props]: $f(x^{(e+1)}) \leq f(x^{(e)}) \Rightarrow monotonically decrease!$

Pf: f(x(e+1)) ≤ u(x(e+1), x(e)) = min u(x, x(e))

$$\begin{cases} (x^{(e)}, x^{(e)}) \end{cases}$$

 $\mu(x,x_{(0)}) = f(x_{(6)})$



Rmk: Equality holds iff
$$f(x^{(\ell^+)}) = u(x^{(\ell^+)}, x^{(\ell)})$$

 $u(x^{(\ell^+)}, x^{(\ell)}) = u(x^{(\ell)}, x^{(\ell)})$

Rmk: would like
$$x^{(e)} \longrightarrow \text{Stationary point of } \boxed{-(x)}$$
Under additional condition \rightarrow continuity & differentiability of u , we can show (x)

> observation

Complete Dataset (X, Z)

incomplete

Goal: Estimate of from incomplete X

$$MLE \rightarrow \frac{\hat{\theta}}{MLE} = \operatorname{argmin} - \log p(X|\underline{\theta})$$

Intractable since
$$P(X|\theta) = \sum_{z} P(X,Z|\theta)$$

 $N(\theta, \theta) = -\mathbb{E}[\log P(X, Z|\theta) | X, \theta] - H(P(\theta|X, \theta))$

$$= -\sum_{\underline{z}} [\log P(X, \underline{z} | \theta)] P(\underline{z} | X, \theta) + \sum_{\underline{z}} P(\underline{z} | X, \theta) (\log P(\underline{z} | X, \theta))$$

$$= -\sum_{\underline{z}} P(\underline{z} | X, \theta) \log P(X | \theta)$$

$$= -\log P(X | \theta)$$

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[Jenson Inequality] Let $g \to \text{convex} \Rightarrow g(\underline{E} Y) \leq \underline{E} [g(Y)]$

$$\text{Verity Part } \textcircled{2}$$

$$\text{P(A)} \Rightarrow P(X | \underline{z}, \theta) P(\underline{z} | \theta)$$

$$= -\log \sum_{\underline{z}} P(X | \underline{z}, \theta) P(\underline{z} | \theta) P(\underline{z} | \theta)$$

$$= -\log \sum_{\underline{z}} P(X | \underline{z}, \theta) P(\underline{z} | \theta) P(X | \underline{z}, \theta)$$

$$= -\log \sum_{\underline{z}} P(\underline{z} | X, \theta^{(0)}) P(\underline{z} | X, \theta^{(0)})$$

$$= -\log \sum_{\underline{z}} P(\underline{z} | X, \theta^{(0)}) \log P(\underline{z} | X, \theta^{(0)})$$

$$= -\sum_{\underline{z}} P(\underline{z} | X, \theta^{(0)}) \log P(X, \underline{z} | \theta) - H(P(\underline{z} | X, \theta^{(0)})$$

$$= -\sum_{\underline{z}} P(\underline{z} | X, \theta^{(0)}) \log P(X, \underline{z} | \theta) - H(P(\underline{z} | X, \theta^{(0)})$$

$$= - \mathbb{E} \left[\left| \log P(X, 2|\theta) \right| X, \theta^{(1)} \right] - H \left(p(\cdot|X, \theta^{(1)}) \right]$$

$$= - \mathbb{E} \left[\left| \log P(X, 2|\theta) \right| X, \theta^{(1)} \right] - H \left(p(\cdot|X, \theta^{(1)}) \right]$$

$$= \mathcal{U}(\theta, \theta^{(0)})$$

Hence, by the properties of Auxiliany function $u(0,0^{(0)})$ of function f.

$$\Leftrightarrow f(0^{(1+1)}) \leqslant f(0^{(1)})$$

Application of Jensen's Ineq. explained.

$$= -\log \sum_{z} P(2|X, \underline{\theta}^{(1)}) \frac{P(X|Z, \theta) P(2|\theta)}{P(2|X, \theta^{(1)})}$$

$$= -\log \mathbb{E} \left[\frac{P(X|Z,\theta) P(Z|\theta)}{P(Z|X,\theta^{(1)})} \right]$$

$$= -\log \mathbb{E} \left[\frac{P(X|Z,\theta) P(Z|\theta)}{P(Z|X,\theta^{(1)})} \right]$$

$$=$$
 g $\mathbb{E}(\Upsilon)$

$$= \sum_{z} P(2|X, \theta^{(1)}) - \log \frac{P(X|Z, \theta) P(2|\theta)}{P(2|X, \theta^{(1)})}$$

$$= - \sum_{z} P(z|X, \theta^{(1)}) \log P(x|z, \theta) P(z|\theta)$$

$$= f(0)$$

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Other Application of MM: Ranking
            m=20 teams
            Each team has a skill level Die [0,1]
BTL Model
                 Pr(i \text{ beats } j) = \frac{\theta i}{\theta i + \theta j}

\mathfrak{D} = \{ b \hat{j} : i, j \in [m] \} \qquad \underline{\theta} = \{ \theta_1, \dots, \theta_m \}

        bij = # of times i beat j in one season
   \mathcal{L}(\mathcal{D}, \underline{\theta}) = \prod_{\hat{i}, \hat{i}} \left( \frac{\theta_{\hat{i}}}{\theta_{\hat{i}} + \theta_{\hat{j}}} \right)^{b_{\hat{i}} \hat{j}} \implies \text{Here}, i \neq \hat{j} \neq \emptyset
       MLE \rightarrow L(\mathcal{D}; \mathcal{Q}) = \sum_{ij} bij [log \theta i - log (\theta i + \theta j)]
       Minimize -l(\mathcal{D}; \mathcal{L}) = f(\mathcal{D}) = -\sum_{(i)} bij [log \theta i - log (\theta i + \theta j)]
 0 Gradient \Rightarrow Q \downarrow \Rightarrow \theta i \leftrightarrow \theta j [related] in second term \log (\theta i + \theta j)
 \frac{\partial f}{\partial \theta_R} = -\sum_{j} b_{kj} \left( \frac{1}{\theta_R} - \frac{1}{\theta_R f \theta_j} \right)
                                             Couple
(ink jok + 5 bik · Di+ OR
 2) MM framework - find Auxiliany ul.,.) of for
            For a concave for h(.),
                         h(y) \leqslant h(x) + h'(x) (y - x)
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upper Bound -> property [2]

we want to apply this to h(y) = log (y)

$$h(y) \leq \log x + \frac{y-x}{x}$$

A majorizer of f is:

$$\frac{1}{1} \frac{\partial u}{\partial t} = -\frac{1}{2} \frac{\partial u}{\partial t} \left[\log \theta - \log \left(\theta \right)^{(1)} + \theta \right]^{(1)} - \frac{\partial u}{\partial t} + \frac{\partial u}{\partial t} \right] \\
= -\frac{1}{1} \frac{\partial u}{\partial t} \log \theta + \frac{1}{1} \frac{\partial u}{\partial t} + \frac{1}{1} \frac{\partial u}{\partial t} + \frac{\partial u}{\partial u} + \frac{\partial u}{\partial t} + \frac{\partial u}{\partial u} + \frac{\partial u}{\partial t} + \frac{\partial u}{\partial$$

Update Way:
$$Q_{\hat{l}}^{(H)} = \operatorname{argmin} u(\underline{\theta}, \underline{\theta}^{(l)})$$

$$= \frac{\sum_{j \neq i} b_{j}}{\sum_{j \neq i} b_{j} + b_{j} \cdot b_{j}$$