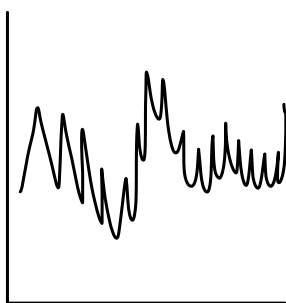


# MAT 3379

## Time Series Analysis

### Study Guide

Winter 2024



$$X_t = Y_t + m_t + S_t$$

## Examples

- White noise:  $\{Z_t\}$  sequence of independent random variables with mean 0 and variance 1
- Random Walk:  $\{Z_t\}$  sequence of iid random variables with mean 0 and variance  $\sigma_z^2$ .  $X_t = \sum_{i=1}^t Z_i$ ,  $t=1,2,\dots$
- Model with trend: Sometimes a trend is present in time series.  $X_t = 1+2t + Z_t$ ,  $t=1,2,\dots$   
Trend is  $m_t = 1+2t$
- Economics Trend:  $X_t = P_t e^{rt}$   
 $P_t$  is real price,  $r$  is interest rate

## Time Series Structure

$$X_t = m_t + Y_t + S_t$$

- $m_t$  is a trend
- $S_t$  is a seasonal part
- $Y_t$  is a stationary part

## Eliminate Trend

- Differencing: Compute  $\nabla X_t = X_t - X_{t-1}$ ,  $t=2,\dots,n$
- Polynomial Fitting: Assume  $m_t = a + bt$   
Estimate  $a$  and  $b$  by minimizing  $\sum_{t=1}^n (X_t - a - bt)^2$ ,  $\hat{m}_t = \hat{a} + \hat{b}t$ .  
Detrended time series:  $\hat{Y}_t = X_t - \hat{m}_t$
- Exponential Smoothing:  $\alpha \in (0,1)$   
Trend:  $\hat{m}_1 = X_1$ ,  $\hat{m}_t = \alpha X_t + (1-\alpha)\hat{m}_{t-1}$ ,  $t=2,\dots,n$   
De-trended time series:  $\hat{Y}_t = X_t - \hat{m}_t$
- Moving Average Smoothing:  $q \in \mathbb{Z}^+$   
 $\hat{m}_t = (2q+1)^{-1} \sum_{j=-q}^q X_{t+j}$ ,  $q+1 \leq t \leq n-q$   
Detrended time series:  $\hat{Y}_t = X_t - \hat{m}_t$

## Mean Function

- $\mu_X(t) = E[X_t]$

## Covariance Function

- $\gamma_X(t,s) = \text{Cov}(X_t, X_s) = E[X_t X_s] - E[X_t]E[X_s]$
- Note:  $\gamma_X(t,t) = \text{Var}(X_t)$

## Properties of Covariance

- For  $a \in \mathbb{R}$ ,  $\text{Cov}(X, a) = 0$
- For  $a, b \in \mathbb{R}$ ,  $\text{Cov}(X, aU + bV) = a \text{Cov}(X, U) + b \text{Cov}(X, V)$
- $\text{Cov}(X, Y)^2 \leq \text{Var}(X) \text{Var}(Y)$

## Stationary Time Series

- $\mu_X(t)$  does not depend on  $t$
- $\gamma_X(t,s)$  depends only on  $h = t-s$
- Covariance function is non-negative definite

## Some useful Properties

- $\text{Cov}(A, B) = E[AB] - E[A]E[B]$
- If  $A, B$  are independent, then  $\text{Cov}(A, B) = 0$
- $E[aA + bB] = aE[A] + bE[B]$
- $\text{Cov}(A, A) = \text{Var}(A)$
- $\text{Var}(A + bB) = \text{Var}(A) + b^2 \text{Var}(B)$  if  $A, B$  are independent

## MA(1) Model

- $\{Z_t\}$  white noise
- $\theta \in \mathbb{R}$  ( $\theta \neq 0$ )
- $X_t = Z_t + \theta Z_{t-1}$
- $\mu_X(t) = 0$
- $\gamma_X(t, t+h) = \begin{cases} (1+\theta^2)\sigma_z^2 & h=0 \\ \theta\sigma_z^2 & h=1 \\ 0 & h \geq 2 \end{cases}$

## Autocorrelation

$$\rho_X(h) = \frac{\gamma_X(h)}{\gamma_X(0)} = \frac{\text{Cov}(X_0, X_h)}{\text{Var}(X_0)}$$

## Partial Autocorrelation

- Correlation between  $X_t$  and  $X_{t+h}$  after conditioning out the "in-between" variables  $X_{t+1}, \dots, X_{t+h-1}$

PACF between  $X_1$  and  $X_3$  when conditioning out  $X_2$ :

$$\rho_{13.2} = \frac{\text{Corr}(X_1, X_3) - \text{Corr}(X_1, X_2)\text{Corr}(X_2, X_3)}{\sqrt{1 - \text{Corr}^2(X_1, X_2)} \sqrt{1 - \text{Corr}^2(X_2, X_3)}}$$

## Sample Mean, Sample Autocovariance, Sample Autocorrelation

- Sample Mean:  $\hat{\mu} = \bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$
- Sample Variance:  $\hat{\sigma}_x^2 = \hat{\gamma}_X(0) = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$
- Sample Autocovariance:  $\hat{\gamma}_X(h) = \frac{1}{n-1} \sum_{t=1}^{n-h} (X_t - \bar{X})(X_{t+h} - \bar{X})$
- Sample Autocorrelation:  $\hat{\rho}_X(h) = \frac{\hat{\gamma}_X(h)}{\hat{\gamma}_X(0)}$
- Sample PACF at lag 2:  $\hat{\alpha}(2) = \frac{\hat{\rho}_X(2) - \hat{\rho}_X(1)^2}{1 - \hat{\rho}_X^2(1)}$

## Linear Processes

Let  $\{Z_t\}$  be a white noise  
Let  $\{\psi_j\}$  be a sequence of constants

$$X_t = \sum_{j=0}^{\infty} \psi_j Z_{t-j}, \quad t=1,2,\dots$$

The model is called:

- linear model
- (infinite order) moving average
- (causal) moving average
- (one-sided) moving average

The linear process is well-defined if  $\sum_{j=0}^{\infty} |\psi_j| < \infty$

The linear process is also stationary with

$$\begin{aligned} E[X_t] &= 0 \\ \gamma_X(h) &= \sigma_z^2 \sum_{j=0}^{\infty} \psi_j \psi_{j+h} \end{aligned}$$

## Difference and Backward operators

- Difference operator  $\nabla$  defined as  $\nabla X_t = X_t - X_{t-1}$
- Backward Operator  $B$  defined as  $BX_t = X_{t-1}$

## ARMA models

Let  $\{Z_t\}$  be a white noise. Then ARMA( $p, q$ ) model is a stationary solution to

$$\phi(B)X_t = \theta(B)Z_t \quad (*)$$

where

$$\phi(z) = 1 - \phi_1 z - \phi_2 z^2 - \dots - \phi_p z^p$$

$$\theta(z) = 1 + \theta_1 z + \theta_2 z^2 + \dots + \theta_q z^q$$

- $\phi(z)$  autoregressive polynomial
- $\theta(z)$  moving average polynomial
- $(*)$  does not guarantee a stationary solution or a causal solution

## Stationarity and Causality

The existence of causal and stationary solution means that ARMA model can be written as

$$X_t = \sum_{j=0}^{\infty} \psi_j Z_{t-j}$$

How to check if there is a stationary and causal solution?

- Take AR polynomial  $\phi(z)$
- Find roots of  $\phi$ :  $\phi(z) = 0$
- If all roots are such that  $|z| > 1$ , then the model is stationary
- If all roots are such that  $|z| > 1$ , then the model is causal

## Linear Representations

Given an ARMA model. If it is stationary and causal, then we can write it as

$$X_t = \sum_{j=1}^{\infty} \psi_j Z_{t-j}$$

The linear representation of the AR(1) model is

$$X_t = \sum_{j=1}^{\infty} \phi^j Z_{t-j} \quad |\phi| < 1$$

The linear representation of the ARMA(1,1) model is

$$X_t = \sum_{j=0}^{\infty} \psi_j Z_{t-j}$$

where  $\psi_0 = 1$  and  $\psi_j = \phi^{j-1}(\theta + \phi)$

## Recursive approach to ACF

To find a recursive definition for the covariance of an AR( $p$ ) model, first multiply every term by  $X_{t-h}$  in the model. Then apply expected value to all terms and use that to find the recursive definition and initial cases for the covariance.

For AR(1), the ACF is

$$\gamma_X(0) = \sigma_z^2 \frac{1}{1 - \phi^2}, \quad \gamma_X(h) = \phi(h-1) \quad h \geq 1$$