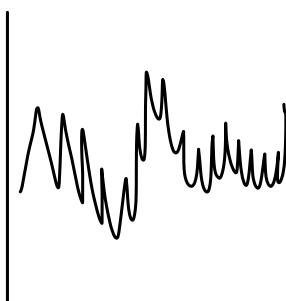


MAT 3379

Time Series Analysis

Study Guide

Winter 2024



$$X_t = Y_t + m_t + S_t$$

Examples

- White noise: $\{Z_t\}$ sequence of independent random variables with mean 0 and variance 1
- Random Walk: $\{Z_t\}$ sequence of iid random variables with mean 0 and variance σ_z^2 . $X_t = \sum_{i=1}^t Z_i$, $t=1,2,\dots$
- Model with trend: Sometimes a trend is present in time series. $X_t = 1+2t + Z_t$, $t=1,2,\dots$
Trend is $m_t = 1+2t$
- Economics Trend: $X_t = P_t e^{rt}$
 P_t is real price, r is interest rate

Time Series Structure

$$X_t = m_t + Y_t + S_t$$

- m_t is a trend
- S_t is a seasonal part
- Y_t is a stationary part

Eliminate Trend

- Differencing: Compute $\nabla X_t = X_t - X_{t-1}$, $t=2,\dots,n$
- Polynomial Fitting: Assume $m_t = a + bt$
Estimate a and b by minimizing $\sum_{t=1}^n (X_t - a - bt)^2$, $\hat{m}_t = \hat{a} + \hat{b}t$.
Detrended time series: $\hat{Y}_t = X_t - \hat{m}_t$
- Exponential Smoothing: $\alpha \in (0,1)$
Trend: $\hat{m}_1 = X_1$, $\hat{m}_t = \alpha X_t + (1-\alpha)\hat{m}_{t-1}$, $t=2,\dots,n$
De-trended time series: $\hat{Y}_t = X_t - \hat{m}_t$
- Moving Average Smoothing: $q \in \mathbb{Z}^+$
 $\hat{m}_t = (2q+1)^{-1} \sum_{j=-q}^q X_{t+j}$, $q+1 \leq t \leq n-q$
Detrended time series: $\hat{Y}_t = X_t - \hat{m}_t$

Mean Function

- $\mu_X(t) = E[X_t]$

Covariance Function

- $\gamma_X(t,s) = \text{Cov}(X_t, X_s) = E[X_t X_s] - E[X_t]E[X_s]$
- Note: $\gamma_X(t,t) = \text{Var}(X_t)$

Properties of Covariance

- For $a \in \mathbb{R}$, $\text{Cov}(X, a) = 0$
- For $a, b \in \mathbb{R}$, $\text{Cov}(X, aU + bV) = a \text{Cov}(X, U) + b \text{Cov}(X, V)$
- $\text{Cov}(X, Y)^2 \leq \text{Var}(X) \text{Var}(Y)$

Stationary Time Series

- $\mu_X(t)$ does not depend on t
- $\gamma_X(t,s)$ depends only on $h = t-s$
- Covariance function is non-negative definite

Some useful Properties

- $\text{Cov}(A, B) = E[AB] - E[A]E[B]$
- If A, B are independent, then $\text{Cov}(A, B) = 0$
- $E[aA + bB] = aE[A] + bE[B]$
- $\text{Cov}(A, A) = \text{Var}(A)$
- $\text{Var}(A + bB) = \text{Var}(A) + b^2 \text{Var}(B)$ if A, B are independent

MA(1) Model

- $\{Z_t\}$ white noise
- $\theta \in \mathbb{R}$ ($\theta \neq 0$)
- $X_t = Z_t + \theta Z_{t-1}$
- $\mu_X(t) = 0$
- $\gamma_X(t, t+h) = \begin{cases} (1+\theta^2)\sigma_z^2 & h=0 \\ \theta\sigma_z^2 & h=1 \\ 0 & h \geq 2 \end{cases}$

Autocorrelation

$$\rho_X(h) = \frac{\gamma_X(h)}{\gamma_X(0)} = \frac{\text{Cov}(X_0, X_h)}{\text{Var}(X_0)}$$

Partial Autocorrelation

- Correlation between X_t and X_{t+h} after conditioning out the "in-between" variables $X_{t+1}, \dots, X_{t+h-1}$

PACF between X_1 and X_3 when conditioning out X_2 :

$$\rho_{13.2} = \frac{\text{Corr}(X_1, X_3) - \text{Corr}(X_1, X_2)\text{Corr}(X_2, X_3)}{\sqrt{1 - \text{Corr}^2(X_1, X_2)} \sqrt{1 - \text{Corr}^2(X_2, X_3)}}$$

Sample Mean, Sample Autocovariance, Sample Autocorrelation

- Sample Mean: $\hat{\mu} = \bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$
- Sample Variance: $\hat{\sigma}_x^2 = \hat{\gamma}_X(0) = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$
- Sample Autocovariance: $\hat{\gamma}_X(h) = \frac{1}{n-1} \sum_{t=1}^{n-h} (X_t - \bar{X})(X_{t+h} - \bar{X})$
- Sample Autocorrelation: $\hat{\rho}_X(h) = \frac{\hat{\gamma}_X(h)}{\hat{\gamma}_X(0)}$
- Sample PACF at lag 2: $\hat{\alpha}(2) = \frac{\hat{\rho}_X(2) - \hat{\rho}_X(1)^2}{1 - \hat{\rho}_X^2(1)}$

Linear Processes

Let $\{Z_t\}$ be a white noise
Let $\{\psi_j\}$ be a sequence of constants

$$X_t = \sum_{j=0}^{\infty} \psi_j Z_{t-j}, \quad t=1,2,\dots$$

The model is called:

- linear model
- (infinite order) moving average
- (causal) moving average
- (one-sided) moving average

The linear process is well-defined if $\sum_{j=0}^{\infty} |\psi_j| < \infty$

The linear process is also stationary with

- $E[X_t] = 0$
- $\gamma_X(h) = \sigma_z^2 \sum_{j=0}^{\infty} \psi_j \psi_{j+h}$

Difference and Backward operators

- Difference operator ∇ defined as $\nabla X_t = X_t - X_{t-1}$
- Backward operator B defined as $BX_t = X_{t-1}$

ARMA models

Let $\{Z_t\}$ be a white noise. Then ARMA(p, q) model is a stationary solution to

$$\phi(B)X_t = \theta(B)Z_t \quad (*)$$

where

$$\phi(z) = 1 - \phi_1 z - \phi_2 z^2 - \dots - \phi_p z^p$$

$$\theta(z) = 1 + \theta_1 z + \theta_2 z^2 + \dots + \theta_q z^q$$

- $\phi(z)$ autoregressive polynomial
- $\theta(z)$ moving average polynomial
- $(*)$ does not guarantee a stationary solution or a causal solution

Stationarity and Causality

The existence of causal and stationary solution means that ARMA model can be written as

$$X_t = \sum_{j=0}^{\infty} \psi_j Z_{t-j}$$

How to check if there is a stationary and causal solution?

- Take AR polynomial $\phi(z)$
- Find roots of ϕ : $\phi(z) = 0$
- If all roots are such that $|z| > 1$, then the model is stationary
- If all roots are such that $|z| > 1$, then the model is causal

Linear Representations

Given an ARMA model. If it is stationary and causal, then we can write it as

$$X_t = \sum_{j=1}^{\infty} \psi_j Z_{t-j}$$

The linear representation of the AR(1) model is

$$X_t = \sum_{j=1}^{\infty} \phi^j Z_{t-j} \quad |\phi| < 1$$

The linear representation of the ARMA(1,1) model is

$$X_t = \sum_{j=0}^{\infty} \psi_j Z_{t-j}$$

where $\psi_0 = 1$ and $\psi_j = \phi^{j-1}(\theta + \phi)$

Recursive approach to ACF

To find a recursive definition for the covariance of an AR(p) model, first multiply every term by X_{t-h} in the model. Then apply expected value to all terms and use that to find the recursive definition and initial cases for the covariance.

For AR(1), the ACF is

$$\gamma_X(0) = \sigma_z^2 \frac{1}{1-\phi^2}, \quad \gamma_X(h) = \phi^{|h-1|} \quad h \geq 1$$