Regression analysis and resampling methods

Johan Mylius Kroken^{1,2}, and Nanna Bryne^{1,2}

² Center for Computational Science, University of Oslo, Norway

September 15, 2022

ABSTRACT

The \dots

1. Introduction

2. Theory

* DECIDE ON z vs. y

$$MSE(\mathbf{y}, \tilde{\mathbf{y}}) = \frac{1}{n} (\mathbf{y} - \tilde{\mathbf{y}})^2 = \frac{1}{n} \sum_{i=0}^{n-1} (y_i - \tilde{y}_i)^2$$
(1)

$$R^{2}(\mathbf{y}, \tilde{\mathbf{y}}) = 1 - \frac{(\mathbf{y} - \tilde{\mathbf{y}})^{2}}{(\mathbf{y}(1 - \bar{y}))^{2}} = 1 - \frac{\sum_{i=0}^{n-1} (y_{i} - \tilde{y}_{i})^{2}}{\sum_{i=0}^{n-1} (y_{i} - \bar{y})^{2}}$$
(2)

$$\bar{y} = \frac{1}{n} \sum_{i=0}^{n-1} y_i$$

2.1. Regression

We assume the vector $\mathbf{y} \in \mathbb{R}^n$ consisting of observed values y_i to take the form $\mathbf{y} = f(\mathbf{x}) + \boldsymbol{\varepsilon}$ where $f(\mathbf{x}) \in \mathbb{R}^n$ is a continous function and $\boldsymbol{\varepsilon} \in \mathbb{R}^n$ is a normally distributed noise. We approximate f by $\tilde{\mathbf{y}} = X\boldsymbol{\beta}$, where $X \in \mathbb{R}^{n \times p}$ is the design matrix of n row vectors $\mathbf{x}_i \in \mathbb{R}^p$, and $\boldsymbol{\beta} \in \mathbb{R}^p$ are the unknown parameters to be determined. The integers n and p represent the number of data points and features, respectively.

For an observed value y_i we have $y_i = \mathbf{x}_i^{\mathsf{T}} \boldsymbol{\beta} + \varepsilon_i = (X\boldsymbol{\beta})_i + \varepsilon_i^{\mathsf{T}}$. The inner product $\mathbf{x}_i^{\mathsf{T}} \boldsymbol{\beta}$ is non-stochastic, hence

$$\mathbb{E}\big[(X\boldsymbol{\beta})_i\big] = (X\boldsymbol{\beta})_i$$

and since

$$\mathbb{E}\big[\varepsilon_i\big] \stackrel{\text{per def.}}{=} 0,$$

we have the expectation value

$$\mathbb{E}[y_i] = \mathbb{E}[(X\boldsymbol{\beta})_i + \varepsilon_i]$$
$$= \mathbb{E}[(X\boldsymbol{\beta})_i] + \mathbb{E}[\varepsilon_i]$$
$$= (X\boldsymbol{\beta})_i.$$

To find the variance, we need the expetation value of the outer product $\mathbf{y}\mathbf{y}^{\mathsf{T}}$,

$$\mathbb{E}[\mathbf{y}\mathbf{y}^{\mathsf{T}}] = \mathbb{E}[(X\boldsymbol{\beta} + \boldsymbol{\varepsilon})(X\boldsymbol{\beta} + \boldsymbol{\varepsilon})^{\mathsf{T}}]$$

$$= \mathbb{E}[X\boldsymbol{\beta}\boldsymbol{\beta}^{\mathsf{T}}X^{\mathsf{T}} + X\boldsymbol{\beta}\boldsymbol{\varepsilon}^{\mathsf{T}} + \boldsymbol{\varepsilon}\boldsymbol{\beta}^{\mathsf{T}}X^{\mathsf{T}} + \boldsymbol{\varepsilon}\boldsymbol{\varepsilon}^{\mathsf{T}}]$$

$$= X\boldsymbol{\beta}\boldsymbol{\beta}^{\mathsf{T}}X^{\mathsf{T}} + \mathbb{I}\sigma^{2}.$$
(3)

The variance becomes

$$\operatorname{Var}[y_i] = \mathbb{E}[(\mathbf{y}\mathbf{y}^{\mathsf{T}})_{ii}] - (\mathbb{E}[y_i])^2$$
$$= (X\boldsymbol{\beta})_i(X\boldsymbol{\beta})_i + \sigma^2 - (X\boldsymbol{\beta})_i(X\boldsymbol{\beta})_i$$
$$= \sigma^2.$$

Taking $\hat{\boldsymbol{\beta}}$ as the Ordinary Least Square (OLS) expression for the optimal parameter, i.e.

$$\hat{\boldsymbol{\beta}} = (X^{\mathsf{T}}X)^{-1}X^{\mathsf{T}}\mathbf{y},$$

we get the expected value

$$\mathbb{E}[\hat{\boldsymbol{\beta}}] = \mathbb{E}[(X^{\mathsf{T}}X)^{-1}X^{\mathsf{T}}\mathbf{y}]$$

$$= (X^{\mathsf{T}}X)^{-1}X^{\mathsf{T}}\mathbb{E}[\mathbf{y}]$$

$$= (X^{\mathsf{T}}X)^{-1}X^{\mathsf{T}}X\boldsymbol{\beta}$$

$$= \boldsymbol{\beta}.$$

The variance is then

$$\begin{aligned} \operatorname{Var} \left[\hat{\boldsymbol{\beta}} \right] &= \mathbb{E} \left[\hat{\boldsymbol{\beta}} \hat{\boldsymbol{\beta}}^{\mathsf{T}} \right] - \mathbb{E} \left[\hat{\boldsymbol{\beta}} \right] \mathbb{E} \left[\hat{\boldsymbol{\beta}}^{\mathsf{T}} \right] \\ &= \mathbb{E} \left[(X^{\mathsf{T}} X)^{-1} X^{\mathsf{T}} \mathbf{y} \mathbf{y}^{\mathsf{T}} X ((X^{\mathsf{T}} X)^{-1})^{\mathsf{T}} \right] - \boldsymbol{\beta} \boldsymbol{\beta}^{\mathsf{T}} \\ &= (X^{\mathsf{T}} X)^{-1} X^{\mathsf{T}} \mathbb{E} \left[\mathbf{y} \mathbf{y}^{\mathsf{T}} \right] X (X^{\mathsf{T}} X)^{-1} - \boldsymbol{\beta} \boldsymbol{\beta}^{\mathsf{T}} \\ &\stackrel{(3)}{=} (X^{\mathsf{T}} X)^{-1} X^{\mathsf{T}} (X \boldsymbol{\beta} \boldsymbol{\beta}^{\mathsf{T}} X^{\mathsf{T}} + \mathbb{I} \sigma^2) X (X^{\mathsf{T}} X)^{-1} \\ &= \boldsymbol{\beta} \boldsymbol{\beta}^{\mathsf{T}} + (X^{\mathsf{T}} X)^{-1} X^{\mathsf{T}} \sigma^2 X (X^{\mathsf{T}} X)^{-1} - \boldsymbol{\beta} \boldsymbol{\beta}^{\mathsf{T}} \\ &= \sigma^2 (X^{\mathsf{T}} X)^{-1}. \end{aligned}$$

¹ Institute of Theoretical Astrophysics, University of Oslo, P.O. Box 1029 Blindern, N-0315 Oslo, Norway

 $^{^{1}}$ Can also write $X_{ij}\beta_{j},$ but we try to be consistent with our notation.

- 2.1.1. Ordinary Least Squares (OLS)
- 2.1.2. Ridge regression
- 2.1.3. Lasso regression
- 2.2. Resampling
- 2.2.1. Bootstrap method
- 2.2.2. Cross-validation
- 3. Analysis
- 4. Conclusion