## Question 1: $M^X/M/1/4$ Queue

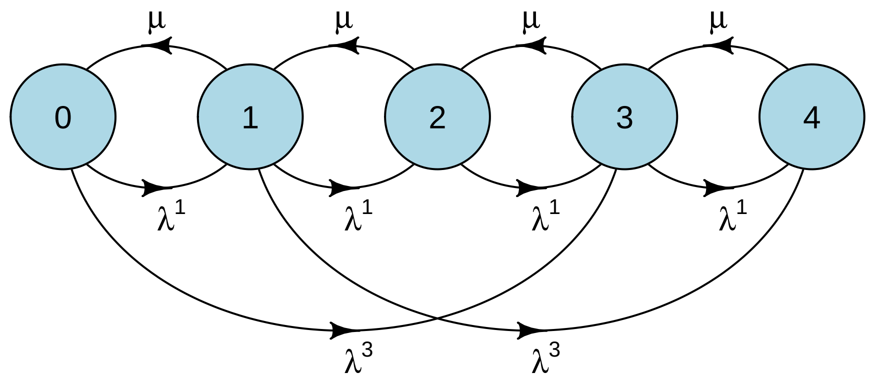


Figure 1: $M^X/M/1/4$, with $\lambda^1$ and $\lambda^3$ arrivals. See Annex A for R Code.

a) We can illustrate the steady state system, with $\lambda^1$ denoting the arrival of a batch with 1 customer, and $\lambda^3$ denoting a batch with 3 customers. *These are not indicies, but a limit of the software.* Now, from state space diagram, in Figure 1, we can form the steady state rate equations:

$$\mu P_1 = (\lambda_1 + \lambda_3)P_0, \qquad \mu P_2 = (\lambda_1 + \lambda_3)P_1, \qquad \mu P_3 = \lambda_1 P_2, \qquad \mu P_4 = \lambda_1 P_3$$

These rate questions lead us to an expression for each state:

$$P_1 = \left[\frac{\lambda_1 + \lambda_3}{\mu}\right] P_0, \qquad P_2 = \left[\frac{\lambda_1 + \lambda_3}{\mu}\right] P_1, \qquad P_3 = \frac{\lambda_1}{\mu}, \qquad P_4 = \frac{\lambda_1}{\mu} P_3$$

$$\lambda_1, \lambda_3 = 1, \quad \mu = 1 \implies P_1 = 2P_0, \quad P_2 = 2P_1, \quad P_3 = P_2, \quad P_4 = P_3$$

Now, since $\sum_{n=0}^{4} P_n = 1$, then

$$\begin{aligned} \sum_{n=0}^{4} P_n &= P_0 + 2P_0 + 4P_0 + 4P_0 + 4P_0 = 1 \\ \therefore P_0 &= \frac{1}{15} \end{aligned}$$

Hence, we obtain: $P_0 = \frac{1}{15}$, $P_1 = \frac{2}{15}$, $P_2 = \frac{4}{15}$, $P_3 = \frac{4}{15}$ and $P_4 = \frac{4}{15}$.

b) From the steady state probabilities, we can deduce the expected length of the system $L_2$:

$$\begin{aligned} L_s = E(N_s) &= \sum_{n=1}^{4} nP_n = \frac{1}{15}(0 + 1(2) + 2(4) + 3(4) + 4(4)) \\ &= \frac{38}{15} \quad \approx 2.53 \quad (3s.f.) \end{aligned}$$

c) For expected waiting time $W_s$ in the system, we notice that the waiting time in the queue is $W_q = E(T)E(N)$, where $E(T) = 1/\mu$ denotes the expected time for a set of individuals to be served. Hence $W_q = L_s/\mu$. Now, we also have that the mean waiting time is the mean time spent in the queue plus the mean service time: $W_s = W_q + 1/\mu$.
Hence, $W_s = \frac{38}{15} + 1 = \frac{53}{15} \implies$ 3 mins 32 seconds, expected total waiting time.

## Question 2: $M^X/M/1/\infty$ Queue

Queue Process $M/M/1$ with infinite capacity and variable arrival rate. An arrival rate of $\lambda_n \leq 1$, and service rate $\mu = 1$. Arrival and Service are independent Poisson processes.

a) $\lambda_n = \frac{n^4}{(n+1)^4}$, with $\lambda_0 = 1$
In a steady state, this system has: $\lambda_0 P_0 = P_1$, $\lambda_1 P_1 = P_2 \ldots$
This leads to $P_n = P_0 \prod_{m=0}^{n} \lambda_m$. Hence, if a steady state does exist then $\sum_{n\geq 0} P_n = 1$. This is only possible if $\sum_{n\geq 0} \prod_{m=0}^{n} \lambda_m$ is finite – converges. Importantly, we reduce the product to:

$$\begin{aligned}\prod_{m=0}^{n} \lambda_m &= 1 \cdot \frac{1^4}{2^4} \cdot \frac{2^4}{3^4} \cdot \ldots \cdot \frac{n^4}{(n+1)^4} = \left[\frac{n!}{(n+1)!}\right]^4 \\ &= \frac{1}{(n+1)^4}\end{aligned}$$

Hence, we obtain

$$\sum_{n\geq 0} \frac{1}{(n+1)^4} = \sum_{n\geq 1} \frac{1}{n^4}$$

Now, by observing the similar general problem, $\sum_{k=1}^{\infty} \frac{1}{k^\alpha} < \infty$ if $\alpha > 1$, we see that $\sum_{n\geq 0} \prod_{m=0}^{n} \lambda_m$ converges for this $\lambda_n$, where $\alpha = 4 > 1$. Hence, it is clear that a form for $P_0$ exists ($P_0 = \frac{90}{\pi^4}$), and a steady state probability distribution <u>does exist</u>.
Now, for the expected system size, $L_s = E(N_s)$

$$\begin{aligned}E(N_s) &= \sum_{n\geq 0} nP_n \\ &= P_0 \sum_{n\geq 0} n \cdot \frac{1}{(n+1)^4} = P_0 \sum_{n\geq 1} \frac{n-1}{n^4} \\ &= P_0 \sum_{n\geq 1} \left[\frac{1}{n^3} - \frac{1}{n^4}\right]\end{aligned}$$

Clearly, this converges and, since $\frac{1}{n^4} < \frac{1}{n^3}$ for all $n$, $L_s$ is <u>finite</u> and positive.

b) $\lambda_n = \frac{\sqrt{n}}{\sqrt{n+1}}$, with $\lambda_0 = 1$
Similarly, as before, we have $P_n = P_0 \prod_{m=0}^{n} \lambda_m$, where the product decomposes to:

$$\begin{aligned}\prod_{m=0}^{n} \lambda_m &= 1 \cdot \frac{\sqrt{1}}{\sqrt{2}} \cdot \frac{\sqrt{2}}{\sqrt{3}} \cdot \ldots \cdot \frac{\sqrt{n}}{\sqrt{n+1}} = \sqrt{\frac{n!}{(n+1)!}} \\ &= \frac{1}{\sqrt{n+1}}\end{aligned}$$

This leads us to:

$$\sum_{n\geq 0} P_n = P_0 \sum_{n\geq 1} \frac{1}{\sqrt{n}}$$

Now, studying the same general form $\sum_{k=1}^{\infty} \frac{1}{k^\alpha} = \infty$ if $\alpha \leq 1$, we see that $\alpha = \frac{1}{2} \leq 1$. Hence, this does not converge and there is no viable form for $P_n$ for a steady state. Therefore there is no steady state probability distribution.

c) $\lambda_n = \frac{n^{\frac{3}{2}}}{(n+1)^{\frac{3}{2}}}$, with $\lambda_0 = 1$.
Similar, as above, we have $\lambda_0 P_0 = P_1$, $\lambda_1 P_1 = P_2 \dots$
Now we have a $P_n$ with a product term of the form:

$$\begin{aligned}\prod_{m=0}^{n} \lambda_m &= 1 \cdot \frac{1^{\frac{3}{2}}}{2^{\frac{3}{2}}} \cdot \frac{2^{\frac{3}{2}}}{3^{\frac{3}{2}}} \cdot \ldots \cdot \frac{n^{\frac{3}{2}}}{(n+1)^{\frac{3}{2}}} = \left[\frac{n!}{(n+1)!}\right]^{\frac{3}{2}} \\ &= \frac{1}{(n+1)^{\frac{3}{2}}}\end{aligned}$$

This leads to $\sum_{n\geq 0} P_n$ of the form:

$$P_0 \sum_{n\geq 1} \frac{1}{n^{\frac{3}{2}}}$$

This clearly has $\alpha = 1.5 > 1$, hence this will converge and deliver a density function for $P_n$, so a steady state distribution does exist.
Now, for the form of the expected system size $L_s = E[N_s]$:

$$\begin{aligned}E[N_s] &= \sum_{n\geq 0} n P_n \\ &= P_0 \sum_{n\geq 0} n \cdot \frac{1}{(n+1)^{\frac{3}{2}}} \\ &= P_0 \sum_{n\geq 1} \left[\frac{1}{n^{\frac{1}{2}}} - \frac{1}{n^{\frac{3}{2}}}\right]\end{aligned}$$

Notice that we have one sum that converges and one that doenst. Hence the overall term does not converge, and there is no finite form for $L_s$. We can expect the system to have infinite size.

### Question 3: Markov Chain

a) With the given transition matrix, we can illustrate the markov chain, with ‘a’ denoting $\alpha$:

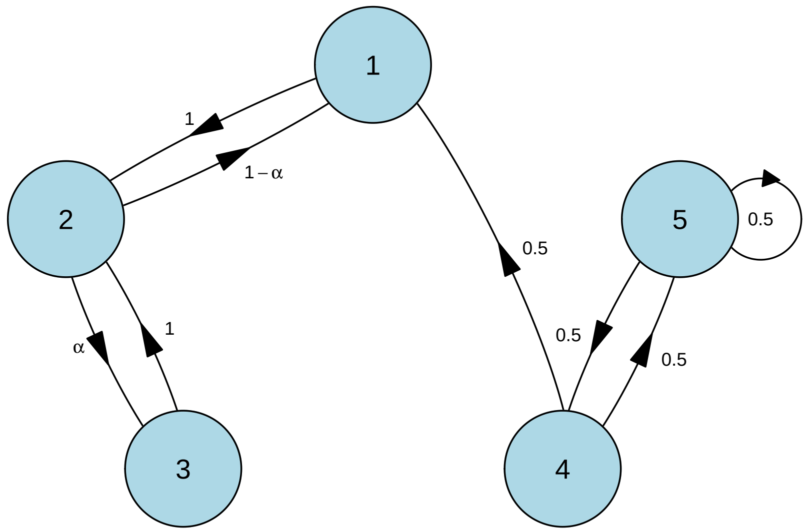


Figure 2: 5 State Markov Chain - See Annex A for R Code

### Annex A: Queue System R COde

```
require('heemod')
require('shape')
require('diagram')
Tr2 <- define_transition( state_names = c('4', '3', '2', '1', '0'),
  0, mu, 0, 0, 0,
  lambda^1, 0, mu, 0, 0,
  0, lambda^1, 0, mu, 0,
  lambda^3, 0, lambda^1, 0, mu,
  0, lambda^3 , 0, lambda^1, 0)

curves <- matrix(nrow = 5, ncol = 5, 0.55)
plot(Tr2, pos=c(5),curve=curves, endhead=FALSE,
     arr.length=0.3, latex=FALSE,
     lwd = 1, box.lwd = 1, box.col = "lightblue",
     cex.txt = 1, box.size = 0.07)
```

### Annex B: Markov Chain Plot in R

```
library('heemod')
library('shape')
library('diagram')

Tr <- define_transition(
  state_names = c('1', '2', '3', '4', '5'),
  0, 1, 0, 0, 0,
  1-alpha, 0, alpha, 0, 0,
  0, 1, 0, 0, 0,
  .5, 0, 0, 0, .5,
  0, 0, 0, .5, .5);

curves <- matrix(nrow = 5, ncol = 5, 0.065)

plot(Tr,
     curve=curves,
     self.shiftx = 0.125,
     self.shifty = 0,
     self.arrpos = 1.4,
     arr.type= "triangle",
     self.cex=0.7,
     latex = FALSE,
     lwd = 1,
     box.lwd = 1,
     box.col = "lightblue",
     cex.txt = 0.7,
     box.size = 0.09)
```