Time Series

John Green

April 2024

 John Green
 Time Series
 April 2024
 1 / 21

Outline

- Notes on macroeconomic data
- Applications
- Lags
- 4 Autocovariance and autocorrelation
- Forecasting

 John Green
 Time Series
 April 2024
 2 / 21

Macroeconomic data

- Volatility clustering
- Seasonal adjustments
- Data revisions and sources
- Unemployment rate
- Labor force participation rate

 John Green
 Time Series
 April 2024
 3 / 21

Uses of time series econometrics

Time series analysis is a highly valuable skill in industry:

- Forecasting (eg interest rates, returns to a stock, GDP growth)
- Dynamic treatment effects: impact of a policy after 1 month, 3 months, 2 years...
- Dynamic system analysis

But, we will have to deal with problems like correlation over time.

4/21

Time series data

- Generally similar to panel data, but with only one unit
- Takes the form $\{Y_t\}_{1 < t < T}$ with covariates $\{X_t\}_{1 < t < T}$
- For now, we assume timing is evenly spaced and there are no missing periods

 John Green
 Time Series
 April 2024
 5/21

Lags

- Lags are very important in time series analysis
- X_{t-1} is often a very good predictor of X_t
- More generally, the j^{th} lag of Y_t is Y_{t-j}
- ullet We might take a difference: $\Delta_j Y_t = Y_t Y_{t-j}$

6/21

Log variables

- Oftentimes we will use log variables instead of the raw values
- Two benefits:
 - Makes exponential growth linear
 - Differences in logs are (approximately) equal to the percentage change:

$$\log(1+\alpha) \approx \alpha$$

So long as α is "small"

• So, $\log X_t - \log X_{t-1} = \log \frac{X_t}{X_{t-1}}$ is roughly the percentage change in X from t-1 to t

<ロト 4回ト 4 直ト 4 直ト - 直 - 釣り()

7/21

Autocorrelation

- Just like with panel data, we will often have correlation between lagged values
- First autocovariance of Y_t : $cov(Y_t, Y_{t-1})$
- First autocorrelation of Y_t : $corr(Y_t, Y_{t-1}) \equiv \rho_1$
- ullet Can calculate for any lag j in the same way and get the j^{th} serial correlation coefficient

 John Green
 Time Series
 April 2024
 8 / 21

Forecasting and stationarity

- We will often be interested in making an out-of-sample (OOS) forecast of our variable
- This is similar to what we did before trying to target the oracle prediction
- We no longer care about causal interpretation of coefficients, OVB, etc.; we just want to get an accurate OOS prediction
- This requires that our data be *stationary* (or jointly stationary)
 - Intuitively, our out-of-sample data needs to be similar to our sample Data
 - Technically, we need the joint distribution across time to be independent of the time period we are looking at

 John Green
 Time Series
 April 2024
 9 / 21

Forecasting error

• An s periods ahead forecast, using estimated coefficients based on t = 1, ..., T, is:

$$\hat{Y}_{T+s|T}$$

• The forecast error is:

$$Y_{T+s} - \hat{Y}_{T+s|T}$$

• And we might evaluate our model based on the mean squared forecasting error:

$$extit{MSFE} = \mathbb{E}\left[\left(Y_{T+s} - \hat{Y}_{T+s|T}\right)^2
ight]$$

 John Green
 Time Series
 April 2024
 10 / 21

Autoregression

- Regress Y_t on its past (lagged values)
- p^{th} order autoregression is Y_t on Y_{t-1}, \ldots, Y_{t-p} , denoted AR(p):

$$Y_t = \beta_0 + \beta_1 Y_{t-1} + \dots + \beta_p Y_{t-p} + u_t$$

- Not causal!
- ullet Can easily test of value in period t-s is useful with t-test or multiple periods with F-test
- For the AR(1) model $Y_t = \beta Y_{t-1} + u_t$, where $|\beta| < 1$ and u = 0, covariance and correlation between t and t s are easy:

$$corr(Y_t, Y_{t-s}) = \beta^s$$



11 / 21

Moving averages

Moving average model:

$$Y_t = \epsilon_t + \theta \epsilon_{t-1}$$

Make substitutions:

$$Y_t = \epsilon_t + \theta Y_{t-1} + \theta^2 Y_{t-2} + \dots$$

- Some nice properties:
 - $Var(Y_t) = \sigma^2(1 + \theta^2)$
 - $cov(Y_t, Y_{t-1}) = \theta \sigma^2$
 - $cov(Y_t, Y_{t-j}) = 0 \ \forall j > 0$
- unfortunately, we can't estimate with OLS; have to use NLS
- Box-pierce test for serial correlation: H_0 is the series is white noise



John GreenTime SeriesApril 202412/21

Covariates

 We can add in other variables (and their lags) to form an autoregressive distributed lag (ADL) model:

$$Y_t = \beta_0 + \beta_1 Y_{t-1} + \dots + \beta_p Y_{t-p} + \beta_1 X_{t-1} + \dots + \beta_p X_{t-r} + u_t$$

with p lags of Y and r lags of X, denoted ADL(p, r).



 John Green
 Time Series
 April 2024
 13 / 21

Forecast error

 Sometimes we look at the root mean squared forecast error, which gives us a "typical" error:

$$\textit{RMSFE} = \sqrt{\textit{E}\left[\left(Y_{T+1} - \hat{Y}_{T+1|T}\right)^{2}\right]}$$

Again, MSFE has two components:

$$MSFE = E(u_{T+1})^2 + E\left[\left(\hat{\beta}_0 - \beta_0\right) + \left(\hat{\beta}_1 - \beta_1\right)Y_T + \left(\hat{\beta}_2 - \beta_2\right)X_T\right]^2$$

- Unavoidable oracle uncertainty (first term)
- Uncertainty from estimating coefficients (second term)



John Green Time Series April 2024 14 / 21

Forecast error

Three methods for estimating the MSFE:

- If number of regressors is small, use the in-sample regression error, $\widehat{MSFE}_{SER} = s_{\hat{u}}^2 = \frac{SSR}{T-p-1}$ which ignores the uncertainty from $\hat{\beta}$
- ② If errors are homoscedastic, we can use this approximation for the full MSFE:

$$\widehat{\mathit{MSFE}}_{\mathit{FPE}} = \left(\frac{T+p+1}{T} \right) s_{\hat{u}}^2 = \left(\frac{T+p+1}{T-p-1} \right) \left(\frac{\mathit{SSR}}{T} \right)$$

1 Method with fewest assumptions: pseudo out of sample. Create a rolling estimate through t=s, calculate error for $t=s+1,\ldots,T$, repeat; $\widehat{MSFE}_{POOS}=\frac{1}{P}\sum_{s=T-P+1}^{T}\widetilde{u}_{s}^{2}$

15 / 21

Number of lags

- How can we choose number of lags p?
- We can use a bunch of t or F tests... but model will usually be too large
- Information criterion helps us select by allowing more bias but lower variance (kind of like with the shrinkage estimators)

$$\mathsf{AIC}(p) = \mathsf{In}\left(\frac{\mathsf{SSR}(p)}{T}\right) + (p+1)\frac{2}{T}$$
 $\mathsf{BIC}(p) = \mathsf{In}\left(\frac{\mathsf{SSR}(p)}{T}\right) + (p+1)\frac{\mathsf{In}\ T}{T}$

- BIC is consistent; AIC is not but may be good if you think more lags are more important
- Can generalize to ADL model

(ロト 4回 ト 4 重 ト 4 重 ト) Q (O

John Green Time Series April 2024 16 / 21

Misc. forecast topics

- Forecast evaluation
- Combining forecasts
- Vector autoregression (VAR):

$$Y_{t} = \beta_{01} + \beta_{11}Y_{t-1} + \dots + \beta_{1p}Y_{t-p} + \gamma_{11}X_{t-1} + \dots + \gamma_{1p}X_{t-p} + u_{1t}$$

$$X_{t} = \beta_{02} + \beta_{21}Y_{t-1} + \dots + \beta_{2p}Y_{t-p} + \gamma_{21}X_{t-1} + \dots + \gamma_{2p}X_{t-p} + u_{2t}$$

- Granger causality
- Direct vs. iterative prediction



 John Green
 Time Series
 April 2024
 17 / 21

Asset returns

- Two features:
 - Volatility clustering
 - Pat tails
- ARCH model:

$$Y_t = \mu + u_t$$

$$u_t \sim N(0, \sigma_t^2)$$

$$\sigma_t^2 = \alpha_0 + \alpha_1 u_{t-1}^2$$

Generalized ARCH (GARCH):

$$\sigma_t^2 = \alpha_0 + \alpha_1 u_{t-1}^2 + \phi_1 \sigma_{t-1}^2$$

• These capture both of our desired features



18 / 21

Stationarity

- All these models are assuming stationarity
- First threat is a *trend*, a persistent long-term movement which might be deterministic or stochastic
 - $Y_t = \beta_0 + Y_{t-1} + u_t$ where u_t is serially correlated and β_0 gives the drift
- If Y_t is a random walk, then ΔY_t is stationary
- Trends may cause:
 - Ocefficients biased to 0; bad forecast
 - 2 t-test doesn't work
 - Spurious correlation between 2 random walks



 John Green
 Time Series
 April 2024
 19 / 21

Dickey-Fuller Test

- Dickey-Fuller test has H_0 of random walk vs. H_A of stationarity
- Use DF table for critical value
 - Depends if model has intercept only or also has a time trend
- If DF fails to reject and has a unit root, use first differences model

 John Green
 Time Series
 April 2024
 20 / 21

Breaks

- Second threat are structural breaks where the model itself (coefficients) changes over time
- If we suspect a break at a specific point we can test this using dummies for before and after the break

$$Y_{t} = \beta_{0} + \beta_{1} Y_{t-1} + \delta_{1} X_{t-1} + \gamma_{0} D_{t}(\tau) + \gamma_{1} [D_{t}(\tau) \times Y_{t-1}] + \gamma_{2} [D_{t}(\tau) \times X_{t-1}] + u_{t}$$

- ullet Use a Chow test statistic on γ
- If we don't know when the break happens we can test dates on $\underline{t} <= t <= \overline{t}$ and take the largest Chow statistic (QLR statistic)
- The more QLR stats we compute, the larger the critical value



 John Green
 Time Series
 April 2024
 21 / 21