STAT 391

Homework 5

Out Tuesday May 1, 2018 Due Tuesday May 8, 2018 ©Marina Meilă mmp@stat.washington.edu

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Problem 1 – Bias and variance for the Poisson distribution

This is a thought experiment of a kind statisticians often do. Imagine that... n data points $x_{1:n}$ are sampled i.i.d. from a Poisson distribution with parameter λ (and because it's a thought experiment, we assume we know λ .) Recall also that $\lambda^{ML} = \sum_{i=1}^{n} x_i/n$.

- **a.** Calculate $E[\lambda^{ML}]$ as a function of λ . All expectations are under the true distribution of the data.
- **b.** Calculate $Var(\lambda^{ML})$ as a function of λ .
- **c.** Assume that n is large enough for the Central Limit Theorem to apply. Express the probability that $\lambda^{ML} \geq \lambda + 1$ as a function of n, λ and Φ the CDF of the standard normal.
- **d.** Numerical answer for $\lambda = 10$, n = 100. Use a table/calculator/computer, don't submit code.

Problem 2 – Confidence Intervals and Boostrap

Use the data set \mathcal{D} from hw5-data1.dat, generated by an unknown density f, to answer the following questions. Formula and numerical result OK, no proofs needed.

- **a.** Estimate μ^{ML} the mean of f.
- **b.** Estimate $\sigma^{2,ML}$ the Maximum Likelihood variance of f. Then calculate $\sigma^{2,C}$ the unbiased estimator of $Var\ f$.
- c. Estimate the variance and standard deviation of μ^{ML} , pretending that $\sigma^{2,C}$ is the true variance Var f.
- **d.** Use the CLT approximation to obtain the Confidence Interval (CI) for confidence level 1δ , for $\delta = 0.01$.
- **e.** Now estimate the variance of μ^{ML} by Bootstrap; denote this by $\sigma^{2,B}$. Take B=1000 bootstrap samples, and calculate from them the numerical value of $\sigma^{2,B}$.
- **f.** Use the CLT approximation again to obtain the Confidence Interval (CI) for confidence level 1δ , for $\delta = 0.01$, from the new variance estimator $\sigma^{2,B}$. You will obtain slightly different values in $\mathbf{c} + \mathbf{d}$, vs $\mathbf{e} + \mathbf{f}$. How different are they? The natural unit of measure in probability is the standard deviation of the quantity measured. Take $SD_C(\mu^{ML})$, the standard deviation of μ^{ML} (NOT SQUARED) computed in (c) as the unit. Then, calculate

$$e_r = \frac{|SD_C(\mu^{ML}) - \sigma^B|}{SD_C(\mu^{ML})}.$$

If this is "small" (that is, much smaller than 1), then the two CI's are "close". You can also measure the overlap of the intervals, relative to the length of one of them. Note that this will be exactly equal to $1 - e_r$.

Problem 3 – Median of Means (MOM)

A recent method for estimating the mean of a distribution is MOM¹.

The MOM estimator of the mean is computed as follows: 1. Divide the data set into K equal subsets of equal size m (assume that n = mK exactly). 2. Compute μ_k the mean of subset k, k = 1 : K. 3. Compute the median of the μ_k 's.

$$\mu^{MOM} = \text{median}(\mu_1, \dots \mu_K), \text{ where } \mu_k = \frac{1}{m} \sum_{i=(k-1)m+1}^{km} x_i, \text{ and } m = \frac{n}{K}.$$
 (1)

¹See e.g. http://www.ub.edu/focm2017/slides/Lugosi.pdf.

Why the trouble? It can be proved that μ^{MOM} is **robust**, that is, it is less influenced by outliers than μ^{ML} the mean of the data. Verify this on the data set hw5-data2.dat.

- **a.** Compute μ^{ML} the mean of the data, and μ^{MOM} for K=56 (n=2800 for this data set).
- **b.** Extract B = 1000 bootstrap samples, and compute $\mu^{MOM,b}$ and $\mu^{ML,b}$ for b = 1:B. Then estimate the variance of μ^{MOM} , μ^{ML} by bootstrap. Does the experiment agree with the theory?
- [c. Extra credit] Repeat a, b for the data from the previous problem. Do you observe any difference? Some theory The choice of K depends on the desired confidence level δ . If $K = 8 \log_2 \frac{1}{\delta}$ then

$$|\mu^{MOM} - \mu| \le 2 \frac{\sigma}{\sqrt{m}}, \text{ with probability } \ge 1 - \delta,$$
 (2)

where m = n/K, and μ, σ are the unknown mean and standard deviation.

So, K = 56 corresponds to $\delta = \frac{1}{128}$.