

Monte Carlo Optimization

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Introduction

This section will cover topics to optimization problems, and solutions Monte Carlo methods provide. Topics covered include

1. Stochastic Search and Simulated Annealing
2. EM Algorithm and MC EM

Light bulb example

This exercise is taken from Flury and Zoppe, 2000, see Exercises in EM.

Below is the setup for the first exercise.

The First Exercise

Suppose there are two light bulb survival experiments. In the first, there are N bulbs whose exact lifetimes y_i for $i \in \{1, \dots, N\}$ are recorded. The lifetimes have an exponential distribution, such that $y_i \sim \text{Exp}(\theta)$. In the second experiment, there are M bulbs. After some time $t > 0$, a researcher walks into the room and only records how many lightbulbs are still burning out of M bulbs. Depending on whether the lightbulbs are still burning or out, the results from the second experiment are right- or -left-censored. There are indicators E_1, \dots, E_M for each of the bulbs in the second experiment. If the bulb is still burning, $E_i = 1$, else $E_i = 0$.

Given this information, our task is to solve for an MLE estimator for θ .

Our first step in solving this is finding the joint likelihood for the observed and unobserved data (i.e. complete-data likelihood).

Let X_1, \dots, X_M be the (unobserved) lifetimes for the second experiment, and let $Z = \sum_{i=1}^M E_i$ be the number of light bulbs still burning. Thus, the observed data from both the experiments combined is $\mathcal{Y} = (Y_1, \dots, Y_N, E_1, \dots, E_M)$ and the unobserved data is $\mathcal{X} = (X_1, \dots, X_M)$.

The complete data log-likelihood is obtained by

$$\begin{aligned} L(\theta|X, Y) &= \prod_{i=1}^N \frac{1}{\theta} e^{-y_i/\theta} \times \prod_{i=1}^M \frac{1}{\theta} e^{-x_i/\theta} \\ &= \theta^{-N} e^{-N\bar{y}/\theta} \times \theta^{-M} e^{-\sum_{i=1}^M x_i/\theta} \end{aligned}$$

And log-likelihood is obtained by

$$\begin{aligned} \log(L(\theta)) &= -N \times \log(\theta) - N\bar{y}/\theta - M \times \log(\theta) + \sum_{i=1}^M x_i/\theta \\ &= -N(\log(\theta) + \bar{y}/\theta) - M \times \log(\theta) + \sum_{i=1}^M x_i/\theta \end{aligned}$$

Or as written by Flury and Zoppe,

$$\log^c(L(\theta|\mathcal{Y}, \mathcal{X})) = -N(\log(\theta) + \bar{Y}/\theta) - \sum_{i=1}^M (\log(\theta) + X_i/\theta)$$

The next step, is to take the expectation of $\log(L(\theta))$ with respect to observed data.

$$\begin{aligned} E[\log(L(\theta))|\mathcal{Y}, \mathcal{X}] &= E[-N(\log(\theta) + \bar{Y}/\theta) - \sum_{i=1}^M (\log(\theta) + X_i/\theta)|\mathcal{Y}, \mathcal{X}] \\ &= -N(\log(\theta) + \bar{Y}/\theta) - E[\sum_{i=1}^M (\log(\theta) + X_i/\theta)|\mathcal{Y}, \mathcal{X}] \\ &= -N(\log(\theta) + \bar{Y}/\theta) - M \times \log(\theta) + E[\frac{1}{\theta} \sum_{i=1}^M X_i|\mathcal{Y}, \mathcal{X}] \\ &= -N(\log(\theta) + \bar{Y}/\theta) - M \times \log(\theta) + \frac{1}{\theta} \sum_{i=1}^M E[X_i|\mathcal{Y}, \mathcal{X}] \\ &= -N(\log(\theta) + \bar{Y}/\theta) - M \times \log(\theta) + \frac{1}{\theta} \sum_{i=1}^M E[X_i|E_i] \end{aligned}$$

which is linear for unobserved X_i . But

(2)

$$E[X_i|\mathcal{Y}] = E[X_i|E_i] = \begin{cases} t + \theta & \text{if } E_i = 1 \\ \theta - t \frac{e^{-t/\theta}}{1 - e^{-t/\theta}} & \text{if } E_i = 0 \end{cases}$$

For the first case, $E_i = 1$, so

$$\begin{aligned} E[x_i|x_i > t] &= E[x_i + t] \\ &= t + E[x_i] \\ &= t + \theta \end{aligned}$$

For the second case, $E_i = 0$, then

$$\int_0^t P(X_i > x|X_i < t) dx = \int_0^t \frac{P(x < X_i < t)}{P(X_i < t)} dx$$

For the denominator, we get

$$\begin{aligned} P(X_i < t) &= \int_0^t \frac{1}{\theta} e^{-x_i/\theta} dx \\ &= \frac{1}{\theta} (-\theta e^{-x_i/\theta}) \Big|_0^t \\ &= 1 - e^{-t/\theta} \end{aligned}$$

and for the numerator we obtain

$$\begin{aligned} P(x < X_i < t) &= \int_x^t \frac{1}{\theta} e^{-x_i/\theta} dx \\ &= \frac{1}{\theta} (-\theta e^{-x_i/\theta}) \Big|_x^t \\ &= e^{-x/\theta} - e^{-t/\theta} \end{aligned}$$

Altogether, we obtain

$$\begin{aligned}
\int_0^t P(X_i > x | X_i < t) dx &= \int_0^t \frac{P(x < X_i < t)}{P(X_i < t)} dx \\
&= \int_0^t \frac{e^{-x/\theta} - e^{-t/\theta}}{(1 - e^{-t/\theta})} dx \\
&= \frac{1}{(1 - e^{-t/\theta})} \int_0^t (e^{-x/\theta} - e^{-t/\theta}) dx \\
&= \frac{1}{(1 - e^{-t/\theta})} \left(\int_0^t e^{-x/\theta} dx - \int_0^t e^{-t/\theta} dx \right) \\
&= \frac{1}{(1 - e^{-t/\theta})} (\theta(1 - e^{-t/\theta}) - x \times e^{-t/\theta} \Big|_0^t) \\
&= \theta - t \times \frac{e^{-t/\theta}}{1 - e^{-t/\theta}}
\end{aligned}$$

If we plug in the conditional expected values

$$E[X_i | \mathcal{Y}] = E[X_i | E_i] = \begin{cases} t + \theta & \text{if } E_i = 1 \\ \theta - t \frac{e^{-t/\theta}}{1 - e^{-t/\theta}} & \text{if } E_i = 0 \end{cases}$$

into the log-likelihood, we will obtain

$$\begin{aligned}
\log(L(\theta)) &= -N(\log(\theta) + \bar{y}/\theta) - M \times \log(\theta) + \sum_{i=1}^M x_i/\theta \\
&= -N \times \log(\theta) - N\bar{y}/\theta - M \times \log(\theta) + \sum_{i=1}^M x_i/\theta \\
&= -(N + M) \times \log(\theta) - N\bar{y}/\theta + \sum_{i=1}^M x_i/\theta \\
&= -(N + M) \times \log(\theta) - \frac{1}{\theta} (N\bar{y} + \sum_{i=1}^M x_i) \\
&= -(N + M) \log(\theta) - \frac{1}{\theta} [N\bar{Y} + Z(t + \theta) + (M - Z)(\theta - t \times \frac{e^{-t/\theta}}{1 - e^{-t/\theta}})]
\end{aligned}$$

and therefore the j th step consists of replacing X_i in (1) by its expected value (2), using the current numerical parameter value $\theta^{(j-1)}$. The result is

$$(3) \quad \log(L(\theta)) = -(N + M) \log(\theta) - \frac{1}{\theta} [N\bar{Y} + Z(t + \theta^{(j-1)}) + (M - Z)(\theta^{(j-1)} - t p^{(j-1)})]$$

where

$$p^{(j)} = \frac{e^{-t/\theta^{(j)}}}{1 - e^{-t/\theta^{(j)}}}$$

```
set.seed(5678)
theta = 5
rate = 1/theta

t = 5
```

```

N = 100
M = 50
y = rexp(n = N, rate = rate)
x = rexp(n = M, rate = rate)
x = sort(x)
E = as.integer(x > t)

N.ybar = sum(y)
Z = sum(E)
t = 5

theta.j = 0.1
theta.jp1 = 0.5
for(i in 1:10){
  theta.j = theta.jp1
  p = (exp(-t/theta.j)/(1-exp(-t/theta.j)))
  theta.jp1 = (N.ybar + Z*( t + theta.j) + (M-Z)*(theta.j - t*p) ) / (N+M)
  print(theta.jp1)
}

## [1] 4.624345
## [1] 5.366158
## [1] 5.445061
## [1] 5.45323
## [1] 5.454073
## [1] 5.45416
## [1] 5.454169
## [1] 5.45417
## [1] 5.45417
## [1] 5.45417

## compare against MLE from observed data
mean(y)

## [1] 6.036602

```