Building a Bayesian model Posterior distribution Prediction Bayesian intervals

Chapter 3

Inference from posterior distribution

Assumption in order:

- (1) Likelihood,
- (2) Parameters,
- (3) Prior.

The assumption results in a Posterior.

Likelihood

 The probability of the data given the parameters. (Assumption)

Likelihood

- The probability of the data given the parameters.
 (Assumption)
- Which parameters? The parameters you need to define a probability distribution.

Likelihood

- The probability of the data given the parameters. (Assumption)
- Which parameters? The parameters you need to define a probability distribution.
- The likelihood is not a distribution for the parameters.

For the data:

$$X = \begin{cases} 1 & \text{if the child is girl,} \\ 0 & \text{if the child is boy.} \end{cases}$$

What parameter do we need for defining the distribution of the data ? $(\mathbb{P}(X|?))$

For the data:

$$X = \begin{cases} 1 & \text{if the child is girl,} \\ 0 & \text{if the child is boy.} \end{cases}$$

What parameter do we need for defining the distribution of the data ? $(\mathbb{P}(X|?))$

• The parameter is the probability of a child being a girl, p.

• If you observe n_F girls of n births, the distribution given the data, is binomial:

$$\mathbb{P}(n_{F}|n,p) = \frac{n!}{n_{F}!(n-n_{F})!} p^{n_{F}} (1-p)^{n-n_{F}}$$

Priors

• Prior incorporates prior information, before the data.

Priors

- Prior incorporates prior information, before the data.
- Often no information is available:
 - Vague prior small impact on the posterior.
 - regularizing prior forces the posterior in certain direction (for more complex model).

• If you observe n_F girls of n births, the distribution given the data, is binomial:

$$\mathbb{P}(n_{F}|n,p) = \frac{n!}{n_{F}!(n-n_{F})!} p^{n_{F}} (1-p)^{n-n_{F}}$$

Posterior

From the prior, parameters, and likelihood, one gets posterior:

By Bayes theorem

$$p(p|n_F,n) = \frac{p(n_F|n,p)p(p)}{p(n_F)}$$

Posterior

From the prior, parameters, and likelihood, one gets posterior:

By Bayes theorem

$$p(p|n_F,n) = \frac{p(n_F|n,p)p(p)}{p(n_F)}$$

Or in words:

$$Posterior = \frac{Likelihood \cdot Prior}{Average \ Likelihood}$$

Inference from posterior distribution

The posterior distribution

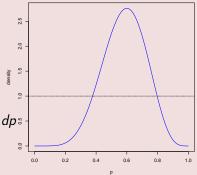
$$p(p|n,y) \propto p^{y}(1-p)^{n-y}\mathbb{I}_{[0,1]}(p)$$

.

 What is the probability that p is less then 0.5:

$$\mathbb{P}(p < 0.5 | n, y) = \int_0^{0.5} p(p|n, y) dp$$
 :

(here we would need the actual density)



```
p_grid <- seq(from = 0, to = 1, length.out = 10)
print(p_grid)</pre>
```

- [1] 0.0000000 0.1111111 0.2222222 0.3333333
 - [5] 0.4444444 0.5555556 0.6666667 0.7777778
 - [9] 0.8888889 1.0000000

• Grid approximation of the posterior:

```
p_grid \leftarrow seq(from = 0, to = 1, length.out = 10)
```

• Grid approximation of the posterior:

```
p_grid <- seq(from = 0, to = 1, length.out = 10)
```

.

```
prior <- rep(1, 10)
print(prior)</pre>
```

[1] 1 1 1 1 1 1 1 1 1 1

• Grid approximation of the posterior:

```
p_grid <\!\!- seq(from = 0, to = 1, length.out = 10) prior <\!\!- rep(1, 10)
```

Grid approximation of the posterior:

```
p\_grid \leftarrow seq(from = 0, to = 1, length.out = 10)
prior \leftarrow rep(1, 10)
```

• .

```
nf = 6
n = 10
likelihood <- dbinom(x = nf, size = n, prob = p_grid)
print(likelihood)</pre>
```

```
[1] 0.0000000000 0.0002466915 0.0092547850
```

```
[4] 0.0569018950 0.1541821742 0.2409096472
```

```
[7] 0.2276075801 0.1133711163 0.0157882546
```

[10] 0.0000000000

• Grid approximation of the posterior:

Grid approximation of the posterior:

```
\begin{array}{lll} p\_grid & <& seq(from = 0,\ to = 1,\ length.out = 10)\\ prior & <& rep(1,\ 10)\\ nf = 6\\ n = 10\\ likelihood & <& dbinom(x = nf,\ size = n,\ prob = p\_grid) \end{array}
```

.

```
posterior <- likelihood * prior
print(posterior) # sum(posterior) = 0.8182621</pre>
```

```
[4] 0.0569018950 0.1541821742 0.2409096472 [7] 0.2276075801 0.1133711163 0.0157882546 [10] 0.0000000000
```

• Grid approximation of the posterior:

• Grid approximation of the posterior:

• .

```
posterior <- posterior / sum(posterior)
print(posterior) # sum(posterior) = 1</pre>
```

```
[4] 0.0695399334 0.1884263806 0.2944162197 [7] 0.2781597338 0.1385510953 0.0192948614 [10] 0.0000000000
```

• Grid approximation of the posterior:

```
p_grid <- seq(from = 0, to = 1, length.out = 1000)
prior <- rep(1, 1000)

nf = 6
n = 10
likelihood <- dbinom(x = nf, size = n, prob = p_grid)
posterior <- likelihood * prior
posterior <- posterior / sum(posterior)</pre>
```

Grid approximation of the posterior:

Sampling:

Posterior distribution

What is the probabilility that p is less then 0.5?

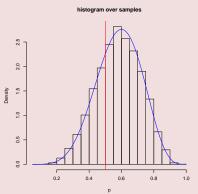


Figure : Histogram of draws from the posterior. Blue line is the posterior distribution

Calculating $\mathbb{P}(p < 0.5)$

•
$$\mathbb{P}(p|n=10, n_f=6) = \int_0^{0.5} p(p|n=10, n_f=6) dp = 0.274$$

Calculating $\mathbb{P}(p < 0.5)$

- $\mathbb{P}(p|n=10, n_f=6) = \int_0^{0.5} p(p|n=10, n_f=6) dp = 0.274$
- Approximation by sampling $(p_1, p_2, ..., p_s)$:

$$P_{est} = \frac{1}{10^4} \sum_{i=1}^{10^4} \mathbb{I}_{[0,0.5)}(p_i) = 0.281$$

In code:

```
samples <- sample(p_grid ,
prob = posterior ,
size = s,
replace = T)
P_est <- mean(samples < 0.5)</pre>
```

• What is the effect of the code s (the number of samples) on the approximation?

```
samples <- sample(p_grid ,
prob = posterior ,
size = s,
replace = T)
P_est <- mean(samples < 0.5)</pre>
```

- What is the effect of the code s (the number of samples) on the approximation?
- s = 2 three runs of the code

$$P_{est}^1 = 0.5, P_{est}^2 = 0, P_{est}^3 = 0.$$

```
samples <- sample(p_grid ,
prob = posterior ,
size = s,
replace = T)
P_est <- mean(samples < 0.5)</pre>
```

- What is the effect of the code s (the number of samples) on the approximation?
- s = 2 three runs of the code

$$P_{est}^1 = 0.5, P_{est}^2 = 0, P_{est}^3 = 0.$$

• s = 10 three runs of the code

$$P_{est}^1 = 0, P_{est}^2 = 0.3, P_{est}^3 = 0.2.$$

```
samples <- sample(p_grid ,
prob = posterior ,
size = s,
replace = T)
P_est <- mean(samples < 0.5)</pre>
```

- What is the effect of the code s (the number of samples) on the approximation?
- s = 2 three runs of the code

$$P_{est}^1 = 0.5, P_{est}^2 = 0, P_{est}^3 = 0.$$

• s = 10 three runs of the code

$$P_{est}^1 = 0, P_{est}^2 = 0.3, P_{est}^3 = 0.2.$$

• $s = 10^5$ three runs of the code

$$P_{\text{est}}^1 = 0.272, P_{\text{est}}^2 = 0.274, P_{\text{est}}^3 = 0.274.$$

Predictions

- I will denote future/unobserved data with a star.
- n_f^* number tails, n^* number of coin tosses.

Predictions

- I will denote future/unobserved data with a star.
- n_f^* number tails, n^* number of coin tosses.
- For a given \hat{p} , the distribution of n_f^* is the likelihood

$$p(n_f^*|n^*,\hat{p}) = \frac{n^*!}{n^*!(n^*-n_f^*)!}\hat{p}^{n_f^*}(1-\hat{p})^n$$

Predictions

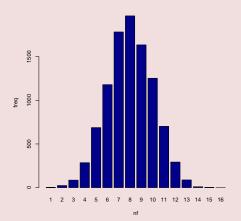
- I will denote future/unobserved data with a star.
- n_f^* number tails, n^* number of coin tosses.
- For a given \hat{p} , the distribution of n_f^* is the likelihood

$$p(n_f^*|n^*,\hat{p}) = \frac{n^*!}{n^*!(n^*-n_f^*)!}\hat{p}^{n_f^*}(1-\hat{p})^n$$

In R:

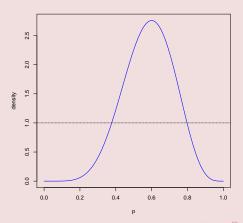
samples n_f^*

Distribution of $p(n_f^*|n=16,\hat{p}=0.5)$



which \hat{p}

Which p̂?



Predictions

• The answer no single \hat{p} !

Predictions

- The answer no single \hat{p} !
- The predictive distribution prior to observing the data (y, n)

$$p(n_f^*|n^*) = \int p(n_f^*|n^*,p)p(p)dp.$$

Predictions

- The answer no single \hat{p} !
- The predictive distribution prior to observing the data (y, n)

$$p(n_f^*|n^*) = \int p(n_f^*|n^*,p)p(p)dp.$$

• The predictive distribution after observing the data (y, n)

$$p(n_f^*|n^*, n, n_f) = \int p(n_f^*|n^*, p)p(p|n, n_f)dp.$$

Sampling from the posterior

Prior predictive distribution

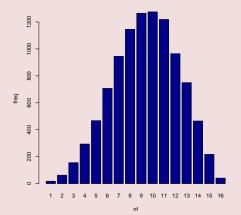
Sampling from the posterior

Prior predictive distribution

Posterior predictive distribution

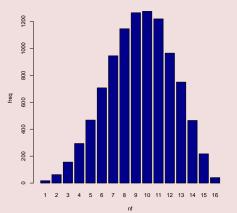
samples n_f^*

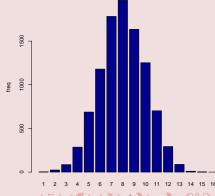
Distribution of
$$p(n_f^*|n^*=16, n=10, n_f=6)$$



samples n_f^*

Distribution of $p(n_f^*|n^*=16, n=10, n_f=6)$





• a P% Credibility interval (CIP) is not unique.

- a P% Credibility interval (CIP) is not unique.
- CI_P for a random variable X is any interval [a, b] such that

$$\mathbb{P}(X \in [a,b]) = P.$$

- a P% Credibility interval (CI_P) is not unique.
- CI_P for a random variable X is any interval [a, b] such that

$$\mathbb{P}(X \in [a,b]) = P.$$

 An approximation of the percentile interval (PI) is obtained by taking quantiles from samples.

- a P% Credibility interval (CIP) is not unique.
- CI_P for a random variable X is any interval [a, b] such that

$$\mathbb{P}(X \in [a,b]) = P.$$

- An approximation of the percentile interval (PI) is obtained by taking quantiles from samples.
- Highest posterior interval (HPDI) is the smallest possible interval

