

Comp 767 Assignment 2

Jonathan Pearce 260672004

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Problem 1.B.i We start by providing the explicit form of $V^\pi(s)$, the true value function at state s

$$V^\pi(s) = \mathbb{E}[r_0 + \gamma r_1 + \dots + \gamma^{k-1} r_{k-1} + \gamma^k V^\pi(s_k)]$$

Where r_i is the i th reward received along the trajectory under π . $V^\pi(s_k)$ is the true value function at the k th state in the trajectory. γ is the discount factor. It follows from the linearity of expectation,

$$\begin{aligned} &= \mathbb{E}[r_0] + \mathbb{E}[\gamma r_1] + \dots + \mathbb{E}[\gamma^{k-1} r_{k-1}] + \mathbb{E}[\gamma^k V^\pi(s_k)] \\ &= \mathbb{E}[r_0] + \gamma \mathbb{E}[r_1] + \dots + \gamma^{k-1} \mathbb{E}[r_{k-1}] + \gamma^k \mathbb{E}[V^\pi(s_k)] \end{aligned}$$

We now expand and provide an upper bound for Δ_t using the formula for $\bar{V}_{\pi,t}^k(s)$ from the question and $V_\pi(s)$ from above:

$$\begin{aligned} \Delta_t &= \max_s | \bar{V}_{\pi,t}^k(s) - V_\pi(s) | \\ &= \max_s \left| \frac{1}{n} \sum_{i=1}^n \left(\sum_{j=1}^k \gamma^{j-1} r_{j-1}^{(i)} \right) + \gamma^k \bar{V}_{\pi,t-1}^k(s_k^{(i)}) - V_\pi(s) \right| \\ &= \max_s \left| \frac{1}{n} \sum_{i=1}^n \left(\sum_{j=1}^k \gamma^{j-1} r_{j-1}^{(i)} \right) + \gamma^k \bar{V}_{\pi,t-1}^k(s_k^{(i)}) - \left(\mathbb{E}[r_0] + \gamma \mathbb{E}[r_1] + \dots + \gamma^{k-1} \mathbb{E}[r_{k-1}] + \gamma^k \mathbb{E}[V^\pi(s_k)] \right) \right| \\ &= \max_s \left| \left(\frac{1}{n} \sum_{i=1}^n r_0^{(i)} - \mathbb{E}[r_0] \right) + \dots + \left(\frac{1}{n} \sum_{i=1}^n \gamma^{k-1} r_{k-1}^{(i)} - \gamma^{k-1} \mathbb{E}[r_{k-1}] \right) + \left(\frac{1}{n} \sum_{i=1}^n \gamma^k \bar{V}_{\pi,t-1}^k(s_k^{(i)}) - \gamma^k \mathbb{E}[V^\pi(s_k)] \right) \right| \\ &= \left(\frac{1}{n} \sum_{i=1}^n r_0^{(i)} - \mathbb{E}[r_0] \right) + \dots + \left(\frac{1}{n} \sum_{i=1}^n \gamma^{k-1} r_{k-1}^{(i)} - \gamma^{k-1} \mathbb{E}[r_{k-1}] \right) + \max_s \left| \left(\frac{1}{n} \sum_{i=1}^n \gamma^k \bar{V}_{\pi,t-1}^k(s_k^{(i)}) - \gamma^k \mathbb{E}[V^\pi(s_k)] \right) \right| \\ &= \left(\frac{1}{n} \sum_{i=1}^n r_0^{(i)} - \mathbb{E}[r_0] \right) + \dots + \gamma^{k-1} \left(\frac{1}{n} \sum_{i=1}^n r_{k-1}^{(i)} - \mathbb{E}[r_{k-1}] \right) + \gamma^k \max_s \left| \left(\frac{1}{n} \sum_{i=1}^n \bar{V}_{\pi,t-1}^k(s_k^{(i)}) - \mathbb{E}[V^\pi(s_k)] \right) \right| \end{aligned}$$

$\left| \bar{V}_{\pi,t-1}^k(s_k^{(i)}) - \mathbb{E}[V^\pi(s_k)] \right| \leq \Delta_{t-1}$ by assumption. Therefore we obtain,

$$\leq \left(\frac{1}{n} \sum_{i=1}^n r_0^{(i)} - \mathbb{E}[r_0] \right) + \dots + \gamma^{k-1} \left(\frac{1}{n} \sum_{i=1}^n r_{k-1}^{(i)} - \mathbb{E}[r_{k-1}] \right) + \gamma^k \Delta_{t-1}$$

Using Hoeffding's inequality, we can perform a deviation analysis on the j th term in the equation above,

$$\mathbb{P} \left[\left| \frac{1}{n} \sum_{i=1}^n r_j^{(i)} - \mathbb{E}[r_j] \right| \geq \epsilon \right] \leq 2 \exp\left(\frac{-2n^2\epsilon^2}{n(1 - (-1))^2}\right) = 2 \exp\left(\frac{-n\epsilon^2}{2}\right)$$

However, we want a deviation analysis that holds for all k terms (of this form) simultaneously. To do this we use the union bound

$$\mathbb{P} \left[\bigcup_{j=1}^k \left| \frac{1}{n} \sum_{i=1}^n r_j^{(i)} - \mathbb{E}[r_j] \right| \geq \epsilon \right] \leq \sum_{j=1}^k \mathbb{P} \left[\left| \frac{1}{n} \sum_{i=1}^n r_j^{(i)} - \mathbb{E}[r_j] \right| \geq \epsilon \right] \leq 2k \exp\left(\frac{-n\epsilon^2}{2}\right)$$

Equating to δ and solving for ϵ :

$$\begin{aligned} \delta &= 2ke^{-\frac{n\epsilon^2}{2}} \\ \frac{\delta}{2k} &= e^{-\frac{n\epsilon^2}{2}} \\ \ln\left(\frac{2k}{\delta}\right) &= \frac{n\epsilon^2}{2} \end{aligned}$$

$$\sqrt{\frac{2}{n} \ln\left(\frac{2k}{\delta}\right)} = \epsilon$$

Therefore with probability at least $1 - \delta$, $\left|\frac{1}{n} \sum_{i=1}^n r_j^{(i)} - \mathbb{E}[r_j]\right| \leq \epsilon = \sqrt{\frac{2}{n} \ln\left(\frac{2k}{\delta}\right)}$ holds for all k terms simultaneously. It follows,

$$\begin{aligned} \Delta_t &\leq \left(\frac{1}{n} \sum_{i=1}^n r_0^{(i)} - \mathbb{E}[r_0]\right) + \dots + \gamma^{k-1} \left(\frac{1}{n} \sum_{i=1}^n r_{k-1}^{(i)} - \mathbb{E}[r_{k-1}]\right) + \gamma^k \Delta_{t-1} \\ &\leq \epsilon + \gamma\epsilon + \dots + \gamma^{k-1}\epsilon + \gamma^k \Delta_{t-1} \\ &= \epsilon \left(\frac{1 - \gamma^k}{1 - \gamma}\right) + \gamma^k \Delta_{t-1} \end{aligned}$$

We conclude that if $\epsilon = \sqrt{\frac{2}{n} \ln\left(\frac{2k}{\delta}\right)}$ then, $\Delta_t \leq \epsilon \left(\frac{1 - \gamma^k}{1 - \gamma}\right) + \gamma^k \Delta_{t-1}$ holds with probability $1 - \delta$.

Problem 1.B.ii We start by providing the explicit form of $V^\pi(s)$, the true value function at state s

$$V^\pi(s) = \mathbb{E} \left[(1 - \lambda) \sum_{k=1}^{\infty} \lambda^{k-1} \left(\sum_{j=1}^k \gamma^{j-1} r_{j-1} + \gamma^k V^\pi(s_k) \right) \right]$$

Where r_i is the i th reward received along the trajectory under π . $V^\pi(s_k)$ is the true value function at the k th state in the trajectory. γ is the discount factor. λ is the decay parameter. It follows from the linearity of expectation,

$$\begin{aligned} &= (1 - \lambda) \sum_{k=1}^{\infty} \lambda^{k-1} \left(\sum_{j=1}^k \mathbb{E}[\gamma^{j-1} r_{j-1}] + \mathbb{E}[\gamma^k V^\pi(s_k)] \right) \\ &= (1 - \lambda) \sum_{k=1}^{\infty} \lambda^{k-1} \left(\sum_{j=1}^k \gamma^{j-1} \mathbb{E}[r_{j-1}] + \gamma^k \mathbb{E}[V^\pi(s_k)] \right) \\ &= (1 - \lambda) \left((\mathbb{E}[r_0] + \gamma \mathbb{E}[V^\pi(s_1)]) + \lambda (\mathbb{E}[r_0] + \gamma \mathbb{E}[r_1] + \gamma^2 \mathbb{E}[V^\pi(s_2)]) + \dots \right) \\ &= \left(\mathbb{E}[r_0] + \gamma \lambda \mathbb{E}[r_1] + (\gamma \lambda)^2 \mathbb{E}[r_2] + \dots \right) + (1 - \lambda) \left(\gamma \mathbb{E}[V^\pi(s_1)] + \lambda \gamma^2 \mathbb{E}[V^\pi(s_2)] + \lambda^2 \gamma^3 \mathbb{E}[V^\pi(s_3)] + \dots \right) \\ &= \sum_{j=0}^{\infty} (\gamma \lambda)^j \mathbb{E}[r_j] + (1 - \lambda) \sum_{j=0}^{\infty} \gamma (\gamma \lambda)^j \mathbb{E}[V^\pi(s_{j+1})] \end{aligned}$$

We can now expand and simplify $\bar{V}_{\pi,t}^\lambda(s)$:

$$\bar{V}_{\pi,t}^\lambda(s) = \frac{1}{n} \sum_{i=1}^n (1 - \lambda) \sum_{k=1}^{\infty} \lambda^{k-1} \left(\sum_{j=1}^{k-1} \gamma^{j-1} r_{j+1}^{(i)} + \gamma^k \bar{V}_{\pi,t-1}^\lambda(s_{k+1}^{(i)}) \right)$$

Using the rewriting of $V^\pi(s)$ above, we obtain:

$$\begin{aligned} &= \frac{1}{n} \sum_{i=1}^n \left(\sum_{j=0}^{\infty} (\gamma \lambda)^j r_j^{(i)} + (1 - \lambda) \sum_{j=0}^{\infty} \gamma (\gamma \lambda)^j \bar{V}_{\pi,t-1}^\lambda(s_{j+1}^{(i)}) \right) \\ &= \sum_{j=0}^{\infty} (\gamma \lambda)^j \frac{1}{n} \sum_{i=1}^n r_j^{(i)} + (1 - \lambda) \sum_{j=0}^{\infty} \gamma (\gamma \lambda)^j \frac{1}{n} \sum_{i=1}^n \bar{V}_{\pi,t-1}^\lambda(s_{j+1}^{(i)}) \end{aligned}$$

We now expand and provide an upper bound for Δ_t using the formula for $\bar{V}_{\pi,t}^\lambda(s)$ above and $V_\pi(s)$ also above:

$$\begin{aligned} \Delta_t &= \max_s \left| \bar{V}_{\pi,t}^\lambda(s) - V_\pi(s) \right| \\ &= \max_s \left| \sum_{j=0}^{\infty} (\gamma \lambda)^j \frac{1}{n} \sum_{i=1}^n r_j^{(i)} + (1 - \lambda) \sum_{j=0}^{\infty} \gamma (\gamma \lambda)^j \frac{1}{n} \sum_{i=1}^n \bar{V}_{\pi,t-1}^\lambda(s_{j+1}^{(i)}) - V_\pi(s) \right| \\ &= \max_s \left| \sum_{j=0}^{\infty} (\gamma \lambda)^j \frac{1}{n} \sum_{i=1}^n r_j^{(i)} + (1 - \lambda) \sum_{j=0}^{\infty} \gamma (\gamma \lambda)^j \frac{1}{n} \sum_{i=1}^n \bar{V}_{\pi,t-1}^\lambda(s_{j+1}^{(i)}) - \sum_{j=0}^{\infty} (\gamma \lambda)^j \mathbb{E}[r_j] - (1 - \lambda) \sum_{j=0}^{\infty} \gamma (\gamma \lambda)^j \mathbb{E}[V^\pi(s_{j+1})] \right| \\ &= \max_s \left| \sum_{j=0}^{\infty} (\gamma \lambda)^j \frac{1}{n} \sum_{i=1}^n r_j^{(i)} - \sum_{j=0}^{\infty} (\gamma \lambda)^j \mathbb{E}[r_j] + (1 - \lambda) \sum_{j=0}^{\infty} \gamma (\gamma \lambda)^j \frac{1}{n} \sum_{i=1}^n \bar{V}_{\pi,t-1}^\lambda(s_{j+1}^{(i)}) - (1 - \lambda) \sum_{j=0}^{\infty} \gamma (\gamma \lambda)^j \mathbb{E}[V^\pi(s_{j+1})] \right| \end{aligned}$$

$$\begin{aligned}
&= \max_s \left| \sum_{j=0}^{\infty} (\gamma\lambda)^j \left(\frac{1}{n} \sum_{i=1}^n r_j^{(i)} - \mathbb{E}[r_j] \right) + (1-\lambda) \sum_{j=0}^{\infty} \gamma(\gamma\lambda)^j \left(\frac{1}{n} \sum_{i=1}^n \bar{V}_{\pi,t-1}^{\lambda}(s_{j+1}^{(i)}) - \mathbb{E}[V^{\pi}(s_{j+1})] \right) \right| \\
&= \sum_{j=0}^{\infty} (\gamma\lambda)^j \left(\frac{1}{n} \sum_{i=1}^n r_j^{(i)} - \mathbb{E}[r_j] \right) + (1-\lambda) \sum_{j=0}^{\infty} \gamma(\gamma\lambda)^j \max_s \left| \left(\frac{1}{n} \sum_{i=1}^n \bar{V}_{\pi,t-1}^{\lambda}(s_{j+1}^{(i)}) - \mathbb{E}[V^{\pi}(s_{j+1})] \right) \right| \\
&\left| \bar{V}_{\pi,t-1}^k(s_k^{(i)}) - \mathbb{E}[V^{\pi}(s_k)] \right| \leq \Delta_{t-1} \text{ by assumption. Therefore we obtain,}
\end{aligned}$$

$$\begin{aligned}
&\leq \sum_{j=0}^{\infty} (\gamma\lambda)^j \left(\frac{1}{n} \sum_{i=1}^n r_j^{(i)} - \mathbb{E}[r_j] \right) + (1-\lambda) \sum_{j=0}^{\infty} \gamma(\gamma\lambda)^j \Delta_{t-1} \\
&= \sum_{j=0}^{k-1} (\gamma\lambda)^j \left(\frac{1}{n} \sum_{i=1}^n r_j^{(i)} - \mathbb{E}[r_j] \right) + \sum_{j=k}^{\infty} (\gamma\lambda)^j \left(\frac{1}{n} \sum_{i=1}^n r_j^{(i)} - \mathbb{E}[r_j] \right) + (1-\lambda) \sum_{j=0}^{\infty} \gamma(\gamma\lambda)^j \Delta_{t-1}
\end{aligned}$$

Here we bound the terms in the summation from $j = 0 \dots (k-1)$ using Hoeffding's inequality and the union bound in the exact same way as in part i. Therefore with probability at least $1 - \delta$, $\left| \frac{1}{n} \sum_{i=1}^n r_j^{(i)} - \mathbb{E}[r_j] \right| \leq \epsilon = \sqrt{\frac{2}{n} \ln(\frac{2k}{\delta})}$ holds for all k terms simultaneously. For the terms in the summation from $j = k, \dots \infty$ we will assume maximum variance. Because $r_i \in [-1, 1] \forall i$, $\left| \frac{1}{n} \sum_{i=1}^n r_j^{(i)} - \mathbb{E}[r_j] \right| \leq 1 - (-1) = 2$. In order to make our bound on Δ_t as tight as possible we select the value k that minimizes the sum of these two summations; $j = 0, \dots (k-1)$ and $j = k, \dots \infty$.

$$\begin{aligned}
\Delta_t &\leq \min_k \left(\sum_{j=0}^{k-1} (\gamma\lambda)^j \epsilon + \sum_{j=k}^{\infty} (\gamma\lambda)^j 2 \right) + (1-\lambda) \sum_{j=0}^{\infty} \gamma(\gamma\lambda)^j \Delta_{t-1} \\
&= \min_k \left(\epsilon \sum_{j=0}^{k-1} (\gamma\lambda)^j + 2 \sum_{j=k}^{\infty} (\gamma\lambda)^j \right) + (1-\lambda) \gamma \Delta_{t-1} \sum_{j=0}^{\infty} (\gamma\lambda)^j \\
&= \min_k \left(\epsilon \sum_{j=0}^{k-1} (\gamma\lambda)^j + 2(\gamma\lambda)^k \sum_{j=0}^{\infty} (\gamma\lambda)^j \right) + (1-\lambda) \gamma \Delta_{t-1} \sum_{j=0}^{\infty} (\gamma\lambda)^j \\
&= \min_k \left(\epsilon \frac{1 - (\gamma\lambda)^k}{1 - \gamma\lambda} + 2(\gamma\lambda)^k \frac{1}{1 - \gamma\lambda} \right) + (1-\lambda) \gamma \Delta_{t-1} \frac{1}{1 - \gamma\lambda} \\
&= \min_k \left(\frac{1 - (\gamma\lambda)^k}{1 - \gamma\lambda} \epsilon + 2 \frac{(\gamma\lambda)^k}{1 - \gamma\lambda} \right) + \frac{(1-\lambda)\gamma}{1 - \gamma\lambda} \Delta_{t-1}
\end{aligned}$$

We conclude that if $\epsilon = \sqrt{\frac{2}{n} \ln(\frac{2k}{\delta})}$ then, $\Delta_t \leq \min_k \left(\frac{1 - (\gamma\lambda)^k}{1 - \gamma\lambda} \epsilon + 2 \frac{(\gamma\lambda)^k}{1 - \gamma\lambda} \right) + \frac{(1-\lambda)\gamma}{1 - \gamma\lambda} \Delta_{t-1}$ holds with probability $1 - \delta$.

Note for part ii, the second term in the minimum expression has a factor of 2 that does not appear in the original paper. The paper did not appear to specify any information on how the rewards are distributed in $[-1, 1]$. Therefore that factor of 2 assumes a worst case of an empirical average being -1 and the expectation being 1 or vice versa. However if the rewards are sampled uniformly from $[-1, 1]$ then the expectation of any reward would be 0, therefore the maximum deviation would be 1 and my result would be the same as the result in the paper.

Problem 2.B.i Before performing policy evaluation or control, certainty-equivalence has to process all of the trajectories in \mathcal{D} in order to build (or update) the model of the environment. In contrast methods that estimate the q-value function are capable of running online, which means they can refine their estimates of the q-value function immediately after receiving more data, instead of collecting data and processing at the end like certainty-equivalence. Certainty-equivalence estimates a model of the environment (transitions and rewards). Therefore certainty-equivalence has space complexity $O(|\mathcal{S}|^2|\mathcal{A}|)$. Methods that estimate the q-value functions are more space efficient and have only $O(|\mathcal{S}||\mathcal{A}|)$ space complexity. Therefore it appears methods that estimate the q-value function are more space and time efficient than certainty-equivalence. However, the main benefit of certainty-equivalence (and model based RL algorithms in general) is that it is more sample efficient than model free methods that estimate the q-value functions directly [1].

Problem 2.B.ii Assuming $\hat{R}(s, a) \in [0, R_{\max}]$. By Hoeffding's inequality we have,

$$\mathbb{P}\left[\left|\hat{R}(s, a) - R(s, a)\right| \geq \epsilon\right] \leq 2 \exp\left(\frac{-2n^2\epsilon^2}{n(R_{\max} - 0)^2}\right) = 2 \exp\left(\frac{-2n\epsilon^2}{R_{\max}^2}\right)$$

However we want a deviation analysis that holds for all $|\mathcal{S} \times \mathcal{A}|$ state-action pairs simultaneously. To do this we use the union bound,

$$\mathbb{P}\left[\bigcup_{(s,a) \in |\mathcal{S} \times \mathcal{A}|} \left|\hat{R}(s, a) - R(s, a)\right| \geq \epsilon\right] \leq \sum_{(s,a) \in |\mathcal{S} \times \mathcal{A}|} \mathbb{P}\left[\left|\hat{R}(s, a) - R(s, a)\right| \geq \epsilon\right] \leq 2|\mathcal{S} \times \mathcal{A}| \exp\left(\frac{-2n\epsilon^2}{R_{\max}^2}\right)$$

Equating to δ and solving for ϵ :

$$\begin{aligned} \delta &= 2|\mathcal{S} \times \mathcal{A}| \exp\left(\frac{-2n\epsilon^2}{R_{\max}^2}\right) \\ \frac{\delta}{2|\mathcal{S} \times \mathcal{A}|} &= \exp\left(\frac{-2n\epsilon^2}{R_{\max}^2}\right) \\ \ln \frac{2|\mathcal{S} \times \mathcal{A}|}{\delta} &= \frac{2n\epsilon^2}{R_{\max}^2} \\ \sqrt{\frac{R_{\max}^2 \ln \frac{2|\mathcal{S} \times \mathcal{A}|}{\delta}}{2n}} &= \epsilon \\ R_{\max} \sqrt{\frac{1}{2n} \ln \frac{2|\mathcal{S} \times \mathcal{A}|}{\delta}} &= \epsilon \end{aligned}$$

Therefore with probability at least $1 - \delta$, $\left|\hat{R}(s, a) - R(s, a)\right| \leq \epsilon = R_{\max} \sqrt{\frac{1}{2n} \ln \frac{2|\mathcal{S} \times \mathcal{A}|}{\delta}}$ holds for all state-action pairs simultaneously. Equivalently, with probability $1 - \delta$,

$$\max_{s,a} \left|\hat{R}(s, a) - R(s, a)\right| \leq R_{\max} \sqrt{\frac{1}{2n} \ln \frac{2|\mathcal{S} \times \mathcal{A}|}{\delta}} \quad (1)$$

Now, assuming $\hat{P}(s'|s, a) \in [0, 1]$. By Hoeffding's inequality we have,

$$\mathbb{P}\left[\left|\hat{P}(s'|s, a) - P(s'|s, a)\right| \geq \epsilon\right] \leq 2 \exp\left(\frac{-2n^2\epsilon^2}{n(1-0)^2}\right) = 2 \exp(-2n\epsilon^2)$$

However we want a deviation analysis that holds for all $|\mathcal{S} \times \mathcal{A} \times \mathcal{S}|$ state-action-state triples simultaneously. To do this we use the union bound

$$\mathbb{P}\left[\bigcup_{(s,a,s') \in |\mathcal{S} \times \mathcal{A} \times \mathcal{S}|} \left|\hat{P}(s'|s, a) - P(s'|s, a)\right| \geq \epsilon\right] \leq \sum_{(s,a,s') \in |\mathcal{S} \times \mathcal{A} \times \mathcal{S}|} \mathbb{P}\left[\left|\hat{P}(s'|s, a) - P(s'|s, a)\right| \geq \epsilon\right] \leq 2|\mathcal{S} \times \mathcal{A} \times \mathcal{S}| \exp(-2n\epsilon^2)$$

Equating to δ and solving for ϵ :

$$\begin{aligned} \delta &= 2|\mathcal{S} \times \mathcal{A} \times \mathcal{S}| \exp(-2n\epsilon^2) \\ \frac{\delta}{2|\mathcal{S} \times \mathcal{A} \times \mathcal{S}|} &= \exp(-2n\epsilon^2) \\ \ln \frac{2|\mathcal{S} \times \mathcal{A} \times \mathcal{S}|}{\delta} &= 2n\epsilon^2 \\ \sqrt{\frac{1}{2n} \ln \frac{2|\mathcal{S} \times \mathcal{A} \times \mathcal{S}|}{\delta}} &= \epsilon \end{aligned}$$

Therefore with probability at least $1 - \delta$, $\left| \hat{P}(s'|s, a) - P(s'|s, a) \right| \leq \epsilon = \sqrt{\frac{1}{2n} \ln \frac{2|\mathcal{S} \times \mathcal{A} \times \mathcal{S}|}{\delta}}$ holds for all state-action-state triples simultaneously. Equivalently, with probability $1 - \delta$,

$$\max_{s,a,s'} \left| \hat{P}(s'|s, a) - P(s'|s, a) \right| \leq \sqrt{\frac{1}{2n} \ln \frac{2|\mathcal{S} \times \mathcal{A} \times \mathcal{S}|}{\delta}} \quad (2)$$

Substituting $\delta = \frac{\epsilon}{2}$ in (1) and (2), we obtain the following two inequalities that hold simultaneously with probability $1 - \delta$,

$$\begin{aligned} \max_{s,a} \left| \hat{R}(s, a) - R(s, a) \right| &\leq R_{\max} \sqrt{\frac{1}{2n} \ln \frac{4|\mathcal{S} \times \mathcal{A}|}{\delta}} \\ \max_{s,a,s'} \left| \hat{P}(s'|s, a) - P(s'|s, a) \right| &\leq \sqrt{\frac{1}{2n} \ln \frac{4|\mathcal{S} \times \mathcal{A} \times \mathcal{S}|}{\delta}} \end{aligned}$$

We now provide a proof of the simulation lemma (Kearns & Singh 1998; Near-Optimal Reinforcement Learning in Polynomial Time).

We begin by making this observation,

$$\begin{aligned} \mathbb{E}[V_M^\pi(s)] &= \mathbb{E}[r_0 + \gamma r_1 + \gamma^2 r_2 + \dots] \\ &= \mathbb{E}[r_0] + \mathbb{E}[\gamma r_1] + \mathbb{E}[\gamma^2 r_2] + \dots \\ &= \mathbb{E}[r_0] + \gamma \mathbb{E}[r_1] + \gamma^2 \mathbb{E}[r_2] + \dots \end{aligned}$$

Assuming the reward distribution is uniform from 0 to R_{\max} we obtain,

$$\begin{aligned} &= \frac{R_{\max}}{2} + \gamma \frac{R_{\max}}{2} + \gamma^2 \frac{R_{\max}}{2} + \dots \\ &= \frac{R_{\max}}{2} (1 + \gamma + \gamma^2 + \dots) \\ &= \frac{R_{\max}}{2} \frac{1}{1 - \gamma} \\ &= \frac{R_{\max}}{2(1 - \gamma)} \end{aligned}$$

Next we develop a bound for $\left\| \hat{P}^\pi(s, \cdot) - P^\pi(s, \cdot) \right\|_1$ using some properties from [2],

$$\begin{aligned} \max_{s,a} \left\| \hat{P}^\pi(s, a) - P^\pi(s, a) \right\|_1 &\leq \max_{s,a} |\mathcal{S}| \left\| \hat{P}^\pi(s, a) - P^\pi(s, a) \right\|_\infty \\ &= |\mathcal{S}| \max_{s,a,s'} \left| \hat{P}(s'|s, a) - P(s'|s, a) \right| \\ &\leq |\mathcal{S}| \sqrt{\frac{1}{2n} \ln \frac{4|\mathcal{S} \times \mathcal{A} \times \mathcal{S}|}{\delta}} \\ &= \epsilon_P \end{aligned}$$

For all states $s \in \mathcal{S}$

$$\begin{aligned} |V_M^\pi(s) - V_M^\pi(s)| &= \left| \left(\hat{R}^\pi(s, \cdot) + \gamma \hat{P}^\pi(s, \cdot) V_M^\pi \right) - \left(R^\pi(s, \cdot) + \gamma P^\pi(s, \cdot) V_M^\pi \right) \right| \\ &= \left| \left(\hat{R}^\pi(s, \cdot) - R^\pi(s, \cdot) \right) + \left(\gamma \hat{P}^\pi(s, \cdot) V_M^\pi - \gamma P^\pi(s, \cdot) V_M^\pi \right) \right| \\ &\leq \left| \hat{R}^\pi(s, \cdot) - R^\pi(s, \cdot) \right| + \left| \gamma \hat{P}^\pi(s, \cdot) V_M^\pi - \gamma P^\pi(s, \cdot) V_M^\pi \right| \\ &\leq \epsilon_R + \left| \gamma \hat{P}^\pi(s, \cdot) V_M^\pi - \gamma P^\pi(s, \cdot) V_M^\pi \right| \\ &= \epsilon_R + \gamma \left| \hat{P}^\pi(s, \cdot) V_M^\pi - P^\pi(s, \cdot) V_M^\pi \right| \\ &= \epsilon_R + \gamma \left| \hat{P}^\pi(s, \cdot) V_M^\pi - P^\pi(s, \cdot) V_M^\pi + P^\pi(s, \cdot) V_M^\pi - P^\pi(s, \cdot) V_M^\pi \right| \\ &= \epsilon_R + \gamma \left| \left(\hat{P}^\pi(s, \cdot) - P^\pi(s, \cdot) \right) V_M^\pi + P^\pi(s, \cdot) (V_M^\pi - V_M^\pi) \right| \\ &\leq \epsilon_R + \gamma \left| \left(\hat{P}^\pi(s, \cdot) - P^\pi(s, \cdot) \right) V_M^\pi \right| + \gamma \left| P^\pi(s, \cdot) (V_M^\pi - V_M^\pi) \right| \\ &\leq \epsilon_R + \gamma \left| \left(\hat{P}^\pi(s, \cdot) - P^\pi(s, \cdot) \right) V_M^\pi \right| + \gamma \|V_M^\pi - V_M^\pi\|_\infty \end{aligned}$$

In our setting we have n large and $\hat{R} \approx R$, therefore $V_M^\pi \approx \mathbb{E}[V_M^\pi(s)] \cdot \mathbf{1}$. Since $\hat{P}^\pi(s, \cdot)$ and $P^\pi(s, \cdot)$ are valid probability distributions both summing to 1, their difference is a distribution that sums up to 0. Therefore we can shift V_M^π down by $\mathbb{E}[V_M^\pi(s)] \cdot \mathbf{1}$ without changing the value of our current bound.

$$\begin{aligned}
&= \epsilon_R + \gamma \left| \left(\hat{P}^\pi(s, \cdot) - P^\pi(s, \cdot) \right) \left(V_M^\pi - \mathbb{E}[V_M^\pi] \cdot \mathbf{1} \right) \right| + \gamma \|V_M^\pi - V_M^\pi\|_\infty \\
&= \epsilon_R + \gamma \left| \left(\hat{P}^\pi(s, \cdot) - P^\pi(s, \cdot) \right) \left(V_M^\pi - \frac{R_{\max}}{2(1-\gamma)} \cdot \mathbf{1} \right) \right| + \gamma \|V_M^\pi - V_M^\pi\|_\infty \\
&\leq \epsilon_R + \gamma \left\| \hat{P}^\pi(s, \cdot) - P^\pi(s, \cdot) \right\|_1 \left\| V_M^\pi - \frac{R_{\max}}{2(1-\gamma)} \right\|_\infty + \gamma \|V_M^\pi - V_M^\pi\|_\infty \\
&\leq \epsilon_R + \gamma \epsilon_P \left\| V_M^\pi - \frac{R_{\max}}{2(1-\gamma)} \right\|_\infty + \gamma \|V_M^\pi - V_M^\pi\|_\infty \\
&= \epsilon_R + \frac{\gamma \epsilon_P R_{\max}}{2(1-\gamma)} + \gamma \|V_M^\pi - V_M^\pi\|_\infty
\end{aligned}$$

Because the above holds for all $s \in \mathcal{S}$ we therefore have,

$$\|V_M^\pi - V_M^\pi\|_\infty \leq \epsilon_R + \frac{\gamma \epsilon_P R_{\max}}{2(1-\gamma)} + \gamma \|V_M^\pi - V_M^\pi\|_\infty$$

We can now solve for $\|V_M^\pi - V_M^\pi\|_\infty$

$$\begin{aligned}
\|V_M^\pi - V_M^\pi\|_\infty - \gamma \|V_M^\pi - V_M^\pi\|_\infty &\leq \epsilon_R + \frac{\gamma \epsilon_P R_{\max}}{2(1-\gamma)} \\
\|V_M^\pi - V_M^\pi\|_\infty (1-\gamma) &\leq \epsilon_R + \frac{\gamma \epsilon_P R_{\max}}{2(1-\gamma)} \\
\|V_M^\pi - V_M^\pi\|_\infty &\leq \frac{\epsilon_R}{1-\gamma} + \frac{\gamma \epsilon_P R_{\max}}{2(1-\gamma)^2}
\end{aligned}$$

(3)

Problem 2.B.iii

$$V_M^*(s) - V_M^{\pi_M^*}(s) = V_M^{\pi_M^*}(s) - V_M^{\pi_M^*}(s) + V_M^{\pi_M^*}(s) - V_M^{\pi_M^*}(s)$$

Because $V_M^{\pi_M^*}(s) \leq V_M^{\pi_M^*}(s)$, we obtain

$$\begin{aligned}
&\leq V_M^{\pi_M^*}(s) - V_M^{\pi_M^*}(s) + V_M^{\pi_M^*}(s) - V_M^{\pi_M^*}(s) \\
&\leq \|V_M^{\pi_M^*} - V_M^{\pi_M^*}\|_\infty + \|V_M^{\pi_M^*} - V_M^{\pi_M^*}\|_\infty \\
&\leq 2 \sup_{\pi: S \rightarrow A} \|V_M^\pi - V_M^\pi\|_\infty
\end{aligned}$$

Problem 2.B.iv

$$\begin{aligned}
V_M^*(s) - V_M^{\pi_M^*}(s) &\leq 2 \sup_{\pi: S \rightarrow A} \|V_M^\pi(s) - V_M^\pi(s)\|_\infty \\
&\leq 2 \left(\frac{\epsilon_R}{1-\gamma} + \frac{\gamma \epsilon_P R_{\max}}{2(1-\gamma)^2} \right) \\
&= 2 \left(\frac{R_{\max} \sqrt{\ln \frac{4|S \times A|}{\delta}}}{\sqrt{2n}(1-\gamma)} + \frac{\gamma R_{\max} |S| \sqrt{\ln \frac{4|S \times A \times S|}{\delta}}}{2\sqrt{2n}(1-\gamma)^2} \right)
\end{aligned}$$

Because $\frac{\gamma}{2} \in [0, \frac{1}{2}]$

$$\leq 2 \left(\frac{R_{\max} \sqrt{\ln \frac{4|S \times A|}{\delta}}}{\sqrt{2n}(1-\gamma)} + \frac{R_{\max} |S| \sqrt{\ln \frac{4|S \times A \times S|}{\delta}}}{\sqrt{2n}(1-\gamma)^2} \right)$$

Because, $\sqrt{\ln \frac{4|S \times A|}{\delta}} \leq \sqrt{\ln \frac{4|S \times A \times S|}{\delta}}$, $(1-\gamma) \in [0, 1]$ and $|S| \geq 1$.

$$\leq 4 \left(\frac{R_{\max} |S| \sqrt{\ln \frac{4|S \times A \times S|}{\delta}}}{\sqrt{2n}(1-\gamma)^2} \right)$$

It follows,

$$V_M^*(s) - V_M^{\pi_M^*}(s) = O\left(\frac{R_{\max}|S|\sqrt{\ln\frac{4|S|\times A\times|S|}{\delta}}}{\sqrt{n}(1-\gamma)^2}\right)$$

If we suppress R_{\max} and the logarithm, which grows slowly relative to $|S|$, we obtain the desired result

$$V_M^*(s) - V_M^{\pi_M^*}(s) = O\left(\frac{|S|}{\sqrt{n}(1-\gamma)^2}\right)$$

References

1. https://people.eecs.berkeley.edu/~cbfinn/_files/mbrl_bootcamp.pdf
2. <https://www.cs.ubc.ca/~schmidtm/Courses/540-F14/norms.pdf>