

UNIVERSITY OF CANTERBURY

Exam

Prescription Number: **EMTH211-18S2**

Time allowed: 180 minutes.

Write your answers in the spaces provided.

There is a *total* of 90 points.

Use black or blue ink. Do not use pencil.

Only UC approved calculators are allowed.

There is no formula sheet for this test.

Show all working. Write neatly. Marks can be lost for poorly presented answers.

| | |
|--------------|--|
| Family name: | |
| Given names: | |
| Student ID: | |

| MARKS | |
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| Office Use Only | |
| Q1 | |
| Q2 | |
| Q3 | |
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| Q5 | |
| Q6 | |
| Total | |

Question 1

[15 points]

The matrix $A = \begin{bmatrix} 0.8 & 0.4 \\ 0.2 & 0.6 \end{bmatrix}$ has dominant eigenvalue $\lambda_1 = 1$.

(a) Calculate the trace of A .

$$\text{Tr}(A) = 0.8 + 0.6 = 1.4$$

(b) Use (a) to find the second eigenvalue λ_2 of A .

$$\lambda_1 + \lambda_2 = 1.4, \text{ so } \lambda_2 = 0.4$$

(c) Find the LU decomposition of A .

$$\begin{bmatrix} 0.8 & 0.4 \\ 0.2 & 0.6 \end{bmatrix} \rightarrow \begin{bmatrix} 0.8 & 0.4 \\ 0 & 0.5 \end{bmatrix}$$

where we have used the row operation $R_2 - 1/4 * R_1$. Hence, we have

$$L = \begin{bmatrix} 1 & 0 \\ 0.25 & 1 \end{bmatrix}$$

and

$$U = \begin{bmatrix} 0.8 & 0.4 \\ 0 & 0.5 \end{bmatrix}.$$

TURN OVER

- (d) Use (c) to carry out one iteration of the inverse power method with initial vector $\begin{bmatrix} 1 \\ 0 \end{bmatrix}$ to approximate an eigenvector for λ_2 . Use the ∞ -norm to normalise your vector. *You do not need to approximate the eigenvalue λ_2 .*

We need to calculate

$$\mathbf{z} = A^{-1} \begin{bmatrix} 1 \\ 0 \end{bmatrix},$$

by solving

$$A\mathbf{z} = \begin{bmatrix} 1 \\ 0 \end{bmatrix}.$$

Now $A\mathbf{z} = LU\mathbf{z}$. We first solve

$$L\mathbf{y} = \begin{bmatrix} 1 \\ 0 \end{bmatrix}.$$

This yields $y_1 = 1, y_2 = -0.25$. We then solve

$$U\mathbf{z} = \begin{bmatrix} 1 \\ -0.25 \end{bmatrix}$$

yielding

$$\mathbf{z} = \begin{bmatrix} 3/2 \\ -1/2 \end{bmatrix}.$$

Normalising \mathbf{z} with respect to the ∞ -norm gives us the first approximation: $\begin{bmatrix} 3 \\ -1 \end{bmatrix}$.

Question 2

[15 points]

A population of female *internet sharks* is distributed into three age classes:

- *baby sharks* are between 0 and 1 year old
- *mummy sharks* are between 1 and 2 year old, and
- *grandma sharks* are between 2 and 3 year old.

Baby sharks are not able to give birth. No internet shark gets older than 3 years. The average number of babies born to a mummy shark is 9.93 and the average number of babies born to a grandma shark is 0.1. Only 1 out of 5 baby sharks survives their first year, but 7 out of 10 mummy sharks become grandma sharks.

(a) Write down the Leslie matrix L describing this population.

$$L = \begin{bmatrix} 0 & 9.93 & 0.1 \\ 0.2 & 0 & 0 \\ 0 & 0.7 & 0 \end{bmatrix}.$$

(b) Suppose that in a certain year y_0 , there are equally many baby, mummy and grandma sharks. What is the fraction of mummy sharks within the total population in year $y_0 + 1$?

$$L \begin{bmatrix} 1 \\ 1 \\ 1 \end{bmatrix} = \begin{bmatrix} 10.03 \\ 0.2 \\ 0.7 \end{bmatrix}.$$

So the fraction of mummy sharks is

$$\frac{0.2}{10.93},$$

which is about 1.8%.

TURN OVER

(c) When using the matrix L from (a) in MatLab, we obtain:

```
eig(L)
ans=
    1.4127667
   -1.4057171
   -0.0070495
```

What is the destiny of the population of internet sharks? Explain your reasoning.

The dominant eigenvalue is 1.413, which is strictly larger than 1. Hence, the population will keep on expanding.

(d) Suppose we are harvesting a fraction $h < 1$ of baby sharks at the end of each growth year. We are not harvesting any mummy or grandma sharks. Write down the matrix H such that, if \mathbf{x} is the population in year z_0 , $(I - H)L\mathbf{x}$ is the population after harvesting. (That is, the population at the start of year $z_0 + 1$).

$$H = \begin{bmatrix} h & 0 & 0 \\ 0 & 0 & 0 \\ 0 & 0 & 0 \end{bmatrix}$$

- (e) Let h be as in (d), namely, the fraction of baby sharks that are harvested. Determine the smallest value h_0 such that if $h > h_0$, we are sure that the population of internet sharks eventually goes extinct. *Hint: h_0 is the fraction for which the population remains stable; express finding h_0 as an eigenvalue problem.*

The population is stable if

$$(I - H)L\mathbf{x} = \mathbf{x}.$$

Hence, $(I - H)L$ should have eigenvalue 1. Now

$$\begin{bmatrix} 1-h & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{bmatrix} \begin{bmatrix} 0 & 9.93 & 0.1 \\ 0.2 & 0 & 0 \\ 0 & 0.7 & 0 \end{bmatrix} = \begin{bmatrix} 0 & (1-h)0.93 & (1-h)0.1 \\ 0.2 & 0 & 0 \\ 0 & 0.7 & 0 \end{bmatrix}.$$

$\lambda = 1$ needs to be a solution of the characteristic equation. This happens if and only if

$$\det \begin{bmatrix} -1 & (1-h)0.93 & (1-h)0.1 \\ 0.2 & -1 & 0 \\ 0 & 0.7 & -1 \end{bmatrix} = 0.$$

This equation gives us

$$\begin{aligned} -1 - 0.2((h-1)0.93 + (h-1)(0.07)) &= 0 \\ -1 - 2(h-1) &= 0 \\ 1 &= 2h \\ h &= 1/2 \end{aligned}$$

We conclude that $h_0 = 1/2$. That is, if we harvest more than half of the baby sharks each year, eventually, the population will go extinct.

TURN OVER

Question 3

[15 points]

(a) Let

$$\mathbf{x}_1 = \begin{bmatrix} 1 \\ 1 \\ 0 \\ 1 \end{bmatrix}, \quad \mathbf{x}_2 = \begin{bmatrix} 1 \\ 1 \\ 1 \\ 1 \end{bmatrix} \quad \text{and} \quad \mathbf{x}_3 = \begin{bmatrix} 1 \\ 0 \\ 0 \\ 0 \end{bmatrix}.$$

with $W = \text{span}(\mathbf{x}_1, \mathbf{x}_2, \mathbf{x}_3)$. Find an orthogonal basis for W . Furthermore, find the projection matrix that projects onto W^\perp .

SOLUTION:

$$\mathbf{v}_1 = \mathbf{x}_3 = \begin{bmatrix} 1 \\ 0 \\ 0 \\ 0 \end{bmatrix}$$

$$\mathbf{v}_2 = \mathbf{x}_1 - \frac{\mathbf{x}_1 \cdot \mathbf{v}_1}{\mathbf{v}_1 \cdot \mathbf{v}_1} \mathbf{v}_1 = \begin{bmatrix} 1 \\ 1 \\ 0 \\ 1 \end{bmatrix} - \frac{1}{1} \begin{bmatrix} 1 \\ 0 \\ 0 \\ 0 \end{bmatrix} = \begin{bmatrix} 0 \\ 1 \\ 0 \\ 1 \end{bmatrix}$$

$$\mathbf{v}_3 = \mathbf{x}_2 - \frac{\mathbf{x}_2 \cdot \mathbf{v}_1}{\mathbf{v}_1 \cdot \mathbf{v}_1} \mathbf{v}_1 - \frac{\mathbf{x}_2 \cdot \mathbf{v}_2}{\mathbf{v}_2 \cdot \mathbf{v}_2} \mathbf{v}_2 = \begin{bmatrix} 1 \\ 1 \\ 1 \\ 1 \end{bmatrix} - \frac{1}{1} \begin{bmatrix} 1 \\ 0 \\ 0 \\ 0 \end{bmatrix} - \frac{2}{2} \begin{bmatrix} 0 \\ 1 \\ 0 \\ 1 \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \\ 1 \\ 0 \end{bmatrix}$$

The projection onto W is given by

$$\mathbf{P} = \begin{bmatrix} 1 & 0 & 0 \\ 0 & \frac{\sqrt{2}}{2} & 0 \\ 0 & 0 & 1 \\ 0 & \frac{\sqrt{2}}{2} & 0 \end{bmatrix} \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & \frac{\sqrt{2}}{2} & 0 & \frac{\sqrt{2}}{2} \\ 0 & 0 & 1 & 0 \end{bmatrix} = \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & \frac{1}{2} & 0 & \frac{1}{2} \\ 0 & 0 & 1 & 0 \\ 0 & \frac{1}{2} & 0 & \frac{1}{2} \end{bmatrix}$$

and so the projection onto W^\perp will be

$$\mathbf{I} - \mathbf{P} = \begin{bmatrix} 0 & 0 & 0 & 0 \\ 0 & \frac{1}{2} & 0 & -\frac{1}{2} \\ 0 & 0 & 0 & 0 \\ 0 & -\frac{1}{2} & 0 & \frac{1}{2} \end{bmatrix}.$$

(b) Let $A = P R$ with

$$P = \begin{bmatrix} 1 & 2 \\ -1 & 1 \\ 0 & 0 \\ 1 & -1 \end{bmatrix} \quad \text{and} \quad R = \begin{bmatrix} 1 & 1 \\ 0 & 1/2 \end{bmatrix}.$$

Find the least squares solution to

$$A\mathbf{x} = \begin{bmatrix} 0 \\ 1 \\ 1 \\ 0 \end{bmatrix}.$$

SOLUTION:

The columns of P are orthogonal but are not normalized. Thus we write

$$P = \begin{bmatrix} \frac{1}{3}\sqrt{3} & \frac{1}{3}\sqrt{6} \\ -\frac{1}{3}\sqrt{3} & \frac{1}{6}\sqrt{6} \\ 0 & 0 \\ \frac{1}{3}\sqrt{3} & -\frac{1}{6}\sqrt{6} \end{bmatrix} \begin{bmatrix} \sqrt{3} & 0 \\ 0 & \sqrt{6} \end{bmatrix} = Q D$$

say. The least square solution will be the solution of

$$D R \mathbf{x} = Q^T \mathbf{b};$$

that is

$$\begin{bmatrix} \sqrt{3} & \sqrt{3} \\ 0 & \frac{1}{2}\sqrt{6} \end{bmatrix} \mathbf{x} = \begin{bmatrix} -\frac{1}{3}\sqrt{3} \\ \frac{1}{6}\sqrt{6} \end{bmatrix}.$$

Therefore the least square solution is

$$\mathbf{x} = \begin{bmatrix} -\frac{2}{3} \\ \frac{1}{3} \end{bmatrix}.$$

TURN OVER

Question 4

[15 points]

(a) Orthogonally diagonalize

$$A = \begin{bmatrix} 3 & 0 & 0 & 1 \\ 0 & 2 & 0 & 0 \\ 0 & 0 & 2 & 0 \\ 1 & 0 & 0 & 3 \end{bmatrix}.$$

SOLUTION:

The eigenvalues of A are 2, 2, 2 and 4. Now

$$[A - 2I] = \begin{bmatrix} 1 & 0 & 0 & 1 \\ 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 \\ 1 & 0 & 0 & 1 \end{bmatrix} \rightarrow \begin{bmatrix} 1 & 0 & 0 & 1 \\ 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 \end{bmatrix}$$

and so the geometric multiplicity of 2 is 3 as expected. The eigenspace E_2 is given by

$$s_1 \begin{bmatrix} 1 \\ 0 \\ 0 \\ -1 \end{bmatrix} + s_2 \begin{bmatrix} 0 \\ 1 \\ 0 \\ 0 \end{bmatrix} + s_3 \begin{bmatrix} 0 \\ 0 \\ 1 \\ 0 \end{bmatrix}.$$

We need to choose an orthonormal basis (or use Gram-Schmidt). An orthonormal basis is given by

$$\frac{\sqrt{2}}{2} \begin{bmatrix} 1 \\ 0 \\ 0 \\ -1 \end{bmatrix}, \quad \begin{bmatrix} 0 \\ 1 \\ 0 \\ 0 \end{bmatrix}, \quad \begin{bmatrix} 0 \\ 0 \\ 1 \\ 0 \end{bmatrix}.$$

Since A is symmetric, the remaining eigenvector must be orthogonal to these three eigenvectors. Therefore we can choose

$$\frac{\sqrt{2}}{2} \begin{bmatrix} 1 \\ 0 \\ 0 \\ 1 \end{bmatrix}$$

to obtain an orthonormal basis. Therefore

$$Q = \begin{bmatrix} 0 & 0 & \frac{1}{2}\sqrt{2} & \frac{1}{2}\sqrt{2} \\ 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & -\frac{1}{2}\sqrt{2} & \frac{1}{2}\sqrt{2} \end{bmatrix}$$

with

$$D = \begin{bmatrix} 2 & 0 & 0 & 0 \\ 0 & 2 & 0 & 0 \\ 0 & 0 & 2 & 0 \\ 0 & 0 & 0 & 4 \end{bmatrix}.$$

(b) Find a SVD for

$$A = \begin{bmatrix} 1 & 1 \\ 0 & 0 \end{bmatrix}.$$

SOLUTION:

We begin by computing the singular values.

$$B = A^T A = \begin{bmatrix} 1 & 1 \\ 1 & 1 \end{bmatrix}.$$

The eigenvalues are 0 (the columns are linearly dependent) and 2 (trace formula). Thus the singular values of A are

$$\sigma_1 = \sqrt{2} \quad \text{and} \quad \sigma_2 = 0$$

and

$$\Sigma = \begin{bmatrix} \sqrt{2} & 0 \\ 0 & 0 \end{bmatrix}.$$

The eigenvalues are distinct and so the eigenvectors of B will be orthogonal. Thus \mathbf{v}_1 will be the unit eigenvector of B associated with 2. Now

$$\begin{bmatrix} -1 & 1 \\ 1 & -1 \end{bmatrix} \rightarrow \begin{bmatrix} -1 & 1 \\ 0 & 0 \end{bmatrix}$$

and so

$$\mathbf{v}_1 = \frac{1}{\sqrt{2}} \begin{bmatrix} 1 \\ 1 \end{bmatrix}.$$

Similarly

$$\begin{bmatrix} 1 & 1 \\ 1 & 1 \end{bmatrix} \rightarrow \begin{bmatrix} 1 & 1 \\ 0 & 0 \end{bmatrix}$$

and so

$$\mathbf{v}_2 = \frac{1}{\sqrt{2}} \begin{bmatrix} -1 \\ 1 \end{bmatrix}.$$

Therefore

$$V = [\mathbf{v}_1 \quad \mathbf{v}_2] = \frac{1}{\sqrt{2}} \begin{bmatrix} 1 & -1 \\ 1 & 1 \end{bmatrix}.$$

Since $\sigma_1 \neq 0$, we have

$$\mathbf{u}_1 = \frac{1}{\sigma_1} A \mathbf{v}_1 = \frac{1}{2} \begin{bmatrix} 2 \\ 0 \end{bmatrix} = \begin{bmatrix} 1 \\ 0 \end{bmatrix}.$$

Since $\sigma_2 = 0$, \mathbf{u}_2 is any vector such that $\mathbf{u}_1, \mathbf{u}_2$ forms an orthonormal basis for \mathbb{R}^2 . Therefore we choose

$$\mathbf{u}_2 = \begin{bmatrix} 0 \\ 1 \end{bmatrix}$$

and so

$$U = [\mathbf{u}_1 \quad \mathbf{u}_2] = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$

which completes the SVD.

TURN OVER

Question 5

[10 points]

Suppose we have two data vectors $\mathbf{x} = (x_1, x_2, \dots, x_n)^\top$ and $\mathbf{y} = (y_1, y_2, \dots, y_n)^\top$ of length n . We denote by \bar{x} and \bar{y} the sample means and by $\tilde{\mathbf{x}}$ and $\tilde{\mathbf{y}}$ the centered vectors. The column vector of n -ones is denoted by $\mathbf{1}$.

- (a) Give the definitions of the sample mean \bar{x} and the centered vector $\tilde{\mathbf{x}}$. Show that $\tilde{\mathbf{x}}$ and $\mathbf{1}$ are orthogonal.

The (sample) mean is defined as

$$\bar{x} = \frac{1}{n} \sum_{i=1}^n x_i$$

and the centered data as $\tilde{\mathbf{x}} = \mathbf{x} - \mathbf{1}\bar{x}$. We need to show for orthogonality that $\tilde{\mathbf{x}} \cdot \mathbf{1} = 0$.

$$\tilde{\mathbf{x}} \cdot \mathbf{1} = (\mathbf{x} - \mathbf{1}\bar{x}) \cdot \mathbf{1} = \sum_{i=1}^n (x_i - \bar{x}) = -n\bar{x} + \sum_{i=1}^n x_i = -n\bar{x} + n\bar{x} = 0.$$

- (b) Give the definition of the sample covariance $cov(\mathbf{x}, \mathbf{y})$ and show that centering does not change the sample covariance, i.e. $cov(\mathbf{x}, \mathbf{y}) = cov(\tilde{\mathbf{x}}, \tilde{\mathbf{y}})$.

The (sample) covariance is defined as $cov(\mathbf{x}, \mathbf{y}) = \frac{1}{n-1} \tilde{\mathbf{x}} \cdot \tilde{\mathbf{y}}$.

Option 1

The covariance is location invariant, i.e. for all $a, b \in \mathbb{R}$ we have $cov(\mathbf{x} + \mathbf{1}a, \mathbf{y} + \mathbf{1}b) = cov(\mathbf{x}, \mathbf{y})$. Hence,

$$cov(\tilde{\mathbf{x}}, \tilde{\mathbf{y}}) = cov(\mathbf{x} - \mathbf{1}\bar{x}, \mathbf{y} - \mathbf{1}\bar{y}) = cov(\mathbf{x}, \mathbf{y}).$$

Option 2

For any vector \mathbf{x} we have $\tilde{\tilde{\mathbf{x}}} = \mathbf{0}$ and therefor $\tilde{\tilde{\mathbf{x}}} = \tilde{\mathbf{x}}$. Hence,

$$cov(\tilde{\mathbf{x}}, \tilde{\mathbf{y}}) = \frac{1}{n-1} \tilde{\tilde{\mathbf{x}}} \cdot \tilde{\tilde{\mathbf{y}}} = \frac{1}{n-1} \tilde{\mathbf{x}} \cdot \tilde{\mathbf{y}} = cov(\mathbf{x}, \mathbf{y}).$$

- (c) Assume that x and y report temperatures in degree Celsius. How will the covariance $cov(x, y)$ **and** the correlation $corr(x, y)$ change, when x and y are transformed into degree Fahrenheit? The transformation from Celsius to Fahrenheit is $9/5x - 32$.

We use the scale invariance of the covariance

$$cov(9/5x - 32, 9/5y - 32) = cov(9/5x, 9/5y) = 9/5 cov(x, 9/5y) = \frac{81}{25} cov(x, y).$$

The correlation does not change. It is invariant to adding a constant or multiplying by a positive number.

Optional

$$corr(9/5x - 32, 9/5y - 32) = corr(9/5x, 9/5y) = sign(9/5)^2 corr(x, y) = corr(x, y).$$

- (d) Assume that x and y report the same temperature, x in Celsius and y in Fahrenheit. What is the value of $corr(x, y)$? Give a reason for your answer.

The correlation will be 1 since Celsius and Fahrenheit have a perfect linear relationship and with positive slope.

Question 6

[20 points]

1. Let $\mathbf{x} = (1, 2, 4, 5)^\top$ and $\mathbf{y} = (1, 2, 3, 10)^\top$. We consider the regression

$$\mathbf{y} = b_0 + b_1\mathbf{x} + \mathbf{e}.$$

(a) Compute the slope \hat{b}_1 of the least squares regression line.

$$\hat{b}_1 = \frac{\tilde{\mathbf{x}} \cdot \tilde{\mathbf{y}}}{\|\tilde{\mathbf{x}}\|^2}$$

$$\bar{x} = 3 \text{ and } \bar{y} = 4,$$

$$\tilde{\mathbf{x}} = (-2, -1, 1, 2),$$

$$\tilde{\mathbf{y}} = (-3, -2, -1, 6),$$

$$\tilde{\mathbf{x}} \cdot \tilde{\mathbf{y}} = 6 + 2 - 1 + 12 = 19, \text{ and}$$

$$\|\tilde{\mathbf{x}}\|^2 = 4 + 1 + 1 + 4 = 10.$$

$$\text{Hence, } \hat{b}_1 = 19/10 = 1.9.$$

(b) Compute the intercept \hat{b}_0 of the regression line.

$$\hat{b}_0 = \bar{y} - \hat{b}_1\bar{x} = 4 - 1.9 \times 3 = -1.7.$$

- (c) Compute a 90% confidence interval for b_1 from \hat{b}_1 . A table with the critical values of the t-distribution is given at the end of this booklet.

$$\begin{aligned} SSR &= (1 - 1.9 + 1.7)^2 + (2 - 1.9 \times 2 + 1.7)^2 + (3 - 1.9 \times 4 + 1.7)^2 + (10 - 1.9 \times 5 + 1.7)^2 \\ &= 0.64 + 0.01 + 8.41 + 4.84 \\ &= 13.9 \end{aligned}$$

$$se(\hat{b}_1) = \sqrt{\frac{1}{n-2} \frac{SSR}{\|\tilde{\mathbf{x}}\|^2}} = \sqrt{\frac{1}{2} \frac{13.9}{10}} \approx 0.834$$

Hence, $b_1 = \hat{b}_1 \pm t_2(0.95)se(\hat{b}_1) \approx 1.9 \pm 2.92 \times 0.834 \approx 1.9 \pm 2.435$
 Or equivalently, $b_1 \in [-0.535, 4.335]$.

- (d) Compute the goodness-of-fit measure R^2 for the regression.

$$R^2 = 1 - \frac{SSR}{SST}$$

We already know that $SSR = 13.9$ from the last question. $SST = \|\tilde{\mathbf{y}}\|^2 = 9+4+1+36 = 50$. Hence,

$$R^2 = 1 - \frac{13.9}{50} = 0.722.$$

2. For a multiple linear regression model

$$\mathbf{y} = b_0 + b_1\mathbf{x}_1 + b_2\mathbf{x}_2 + \cdots + b_p\mathbf{x}_p + \mathbf{e}$$

we denote the vector of least squares estimators by $\hat{\mathbf{b}} = (\hat{b}_0, \hat{b}_1, \hat{b}_2, \dots, \hat{b}_p)^\top$.

- (a) Give the definition of the model matrix and the normal equation for the vector of estimators $\hat{\mathbf{b}}$.

The first column of the model matrix is the vector $\mathbf{1}$. The second to $(p+1)$ -th columns are given by the data vectors $\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_p$

$$X = [\mathbf{1}, \mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_p].$$

The normal equation is

$$X^\top X \hat{\mathbf{b}} = X^\top \mathbf{y} \quad \text{or} \quad \hat{\mathbf{b}} = (X^\top X)^{-1} X^\top \mathbf{y}$$

- (b) Some researchers want to investigate the energy loss in power lines of same diameter and different lengths. They regress the lost energy on voltage, and current, as well as on resistance, and length of the power line. But when they repeat their measurements under very similar conditions the OLS estimator produces quite different results.

Explain what might have gone wrong in this regression?

Resistance is a linear function of the length of a power line, when diameter (and temperature) are constant. We will have multicollinearity in the data. This causes instability for the solution of the least squares problem which results in instable estimators.

- (c) Try to find evidence that nothing is wrong with the regression by using a statistical test. Use the level $\alpha = 0.05$ for the test. In the regression above $p = 4$. Assume that $n = 20$ and $\det(X^\top X) = 0.5$. Interpret your result in one or two sentences. We can apply a Haitovsky test. The test statistic is

$$H = \left(1 + \frac{2p+5}{6} - n\right) \ln(1 - \det(X^\top X)) = \left(-16\frac{5}{6}\right) \ln(0.5) \approx 11.668.$$

We need to compare the value of the test statistic with the critical value of the Chi-square distribution with $p(p-1)/2$ degrees of freedom $\chi^2_{0.05}(6) = 12.592$.

The test fails to reject the hypothesis that we do not have multicollinearity in the regressors. This is not evidence that there is multicollinearity, we just cannot rule it out.

- (d) Suggest a way how this can be fixed and give a reason why this would help.

Option 1

We need to drop either resistance or length from the regression since these two are the reason for the multicollinearity.

Option 2

We can collect more data. This will give more precise information about the regressors which can compensate the instabilities caused by multicollinearity.

TURN OVER

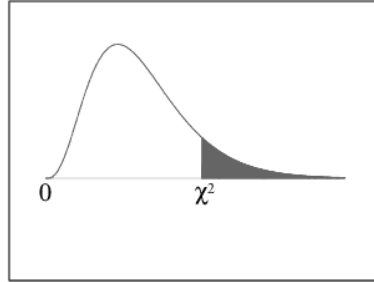
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| cum. prob | <i>t</i> .50 | <i>t</i> .75 | <i>t</i> .80 | <i>t</i> .85 | <i>t</i> .90 | <i>t</i> .95 | <i>t</i> .975 | <i>t</i> .99 | <i>t</i> .995 | <i>t</i> .999 | <i>t</i> .9995 |
|-----------|------------------|--------------|--------------|--------------|--------------|--------------|---------------|--------------|---------------|---------------|----------------|
| one-tail | 0.50 | 0.25 | 0.20 | 0.15 | 0.10 | 0.05 | 0.025 | 0.01 | 0.005 | 0.001 | 0.0005 |
| two-tails | 1.00 | 0.50 | 0.40 | 0.30 | 0.20 | 0.10 | 0.05 | 0.02 | 0.01 | 0.002 | 0.001 |
| df | | | | | | | | | | | |
| 1 | 0.000 | 1.000 | 1.376 | 1.963 | 3.078 | 6.314 | 12.71 | 31.82 | 63.66 | 318.31 | 636.62 |
| 2 | 0.000 | 0.816 | 1.061 | 1.386 | 1.886 | 2.920 | 4.303 | 6.965 | 9.925 | 22.327 | 31.599 |
| 3 | 0.000 | 0.765 | 0.978 | 1.250 | 1.638 | 2.353 | 3.182 | 4.541 | 5.841 | 10.215 | 12.924 |
| 4 | 0.000 | 0.741 | 0.941 | 1.190 | 1.533 | 2.132 | 2.776 | 3.747 | 4.604 | 7.173 | 8.610 |
| 5 | 0.000 | 0.727 | 0.920 | 1.156 | 1.476 | 2.015 | 2.571 | 3.365 | 4.032 | 5.893 | 6.869 |
| 6 | 0.000 | 0.718 | 0.906 | 1.134 | 1.440 | 1.943 | 2.447 | 3.143 | 3.707 | 5.208 | 5.959 |
| 7 | 0.000 | 0.711 | 0.896 | 1.119 | 1.415 | 1.895 | 2.365 | 2.998 | 3.499 | 4.785 | 5.408 |
| 8 | 0.000 | 0.706 | 0.889 | 1.108 | 1.397 | 1.860 | 2.306 | 2.896 | 3.355 | 4.501 | 5.041 |
| 9 | 0.000 | 0.703 | 0.883 | 1.100 | 1.383 | 1.833 | 2.262 | 2.821 | 3.250 | 4.297 | 4.781 |
| 10 | 0.000 | 0.700 | 0.879 | 1.093 | 1.372 | 1.812 | 2.228 | 2.764 | 3.169 | 4.144 | 4.587 |
| 11 | 0.000 | 0.697 | 0.876 | 1.088 | 1.363 | 1.796 | 2.201 | 2.718 | 3.106 | 4.025 | 4.437 |
| 12 | 0.000 | 0.695 | 0.873 | 1.083 | 1.356 | 1.782 | 2.179 | 2.681 | 3.055 | 3.930 | 4.318 |
| 13 | 0.000 | 0.694 | 0.870 | 1.079 | 1.350 | 1.771 | 2.160 | 2.650 | 3.012 | 3.852 | 4.221 |
| 14 | 0.000 | 0.692 | 0.868 | 1.076 | 1.345 | 1.761 | 2.145 | 2.624 | 2.977 | 3.787 | 4.140 |
| 15 | 0.000 | 0.691 | 0.866 | 1.074 | 1.341 | 1.753 | 2.131 | 2.602 | 2.947 | 3.733 | 4.073 |
| 16 | 0.000 | 0.690 | 0.865 | 1.071 | 1.337 | 1.746 | 2.120 | 2.583 | 2.921 | 3.686 | 4.015 |
| 17 | 0.000 | 0.689 | 0.863 | 1.069 | 1.333 | 1.740 | 2.110 | 2.567 | 2.898 | 3.646 | 3.965 |
| 18 | 0.000 | 0.688 | 0.862 | 1.067 | 1.330 | 1.734 | 2.101 | 2.552 | 2.878 | 3.610 | 3.922 |
| 19 | 0.000 | 0.688 | 0.861 | 1.066 | 1.328 | 1.729 | 2.093 | 2.539 | 2.861 | 3.579 | 3.883 |
| 20 | 0.000 | 0.687 | 0.860 | 1.064 | 1.325 | 1.725 | 2.086 | 2.528 | 2.845 | 3.552 | 3.850 |
| 21 | 0.000 | 0.686 | 0.859 | 1.063 | 1.323 | 1.721 | 2.080 | 2.518 | 2.831 | 3.527 | 3.819 |
| 22 | 0.000 | 0.686 | 0.858 | 1.061 | 1.321 | 1.717 | 2.074 | 2.508 | 2.819 | 3.505 | 3.792 |
| 23 | 0.000 | 0.685 | 0.858 | 1.060 | 1.319 | 1.714 | 2.069 | 2.500 | 2.807 | 3.485 | 3.768 |
| 24 | 0.000 | 0.685 | 0.857 | 1.059 | 1.318 | 1.711 | 2.064 | 2.492 | 2.797 | 3.467 | 3.745 |
| 25 | 0.000 | 0.684 | 0.856 | 1.058 | 1.316 | 1.708 | 2.060 | 2.485 | 2.787 | 3.450 | 3.725 |
| 26 | 0.000 | 0.684 | 0.856 | 1.058 | 1.315 | 1.706 | 2.056 | 2.479 | 2.779 | 3.435 | 3.707 |
| 27 | 0.000 | 0.684 | 0.855 | 1.057 | 1.314 | 1.703 | 2.052 | 2.473 | 2.771 | 3.421 | 3.690 |
| 28 | 0.000 | 0.683 | 0.855 | 1.056 | 1.313 | 1.701 | 2.048 | 2.467 | 2.763 | 3.408 | 3.674 |
| 29 | 0.000 | 0.683 | 0.854 | 1.055 | 1.311 | 1.699 | 2.045 | 2.462 | 2.756 | 3.396 | 3.659 |
| 30 | 0.000 | 0.683 | 0.854 | 1.055 | 1.310 | 1.697 | 2.042 | 2.457 | 2.750 | 3.385 | 3.646 |
| 40 | 0.000 | 0.681 | 0.851 | 1.050 | 1.303 | 1.684 | 2.021 | 2.423 | 2.704 | 3.307 | 3.551 |
| 60 | 0.000 | 0.679 | 0.848 | 1.045 | 1.296 | 1.671 | 2.000 | 2.390 | 2.660 | 3.232 | 3.460 |
| 80 | 0.000 | 0.678 | 0.846 | 1.043 | 1.292 | 1.664 | 1.990 | 2.374 | 2.639 | 3.195 | 3.416 |
| 100 | 0.000 | 0.677 | 0.845 | 1.042 | 1.290 | 1.660 | 1.984 | 2.364 | 2.626 | 3.174 | 3.390 |
| 1000 | 0.000 | 0.675 | 0.842 | 1.037 | 1.282 | 1.646 | 1.962 | 2.330 | 2.581 | 3.098 | 3.300 |
| Z | 0.000 | 0.674 | 0.842 | 1.036 | 1.282 | 1.645 | 1.960 | 2.326 | 2.576 | 3.090 | 3.291 |
| | 0% | 50% | 60% | 70% | 80% | 90% | 95% | 98% | 99% | 99.8% | 99.9% |
| | Confidence Level | | | | | | | | | | |

Chi-Square Distribution Table



The shaded area is equal to α for $\chi^2 = \chi^2_{\alpha}$.

| df | $\chi^2_{.995}$ | $\chi^2_{.990}$ | $\chi^2_{.975}$ | $\chi^2_{.950}$ | $\chi^2_{.900}$ | $\chi^2_{.100}$ | $\chi^2_{.050}$ | $\chi^2_{.025}$ | $\chi^2_{.010}$ | $\chi^2_{.005}$ |
|------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|
| 1 | 0.000 | 0.000 | 0.001 | 0.004 | 0.016 | 2.706 | 3.841 | 5.024 | 6.635 | 7.879 |
| 2 | 0.010 | 0.020 | 0.051 | 0.103 | 0.211 | 4.605 | 5.991 | 7.378 | 9.210 | 10.597 |
| 3 | 0.072 | 0.115 | 0.216 | 0.352 | 0.584 | 6.251 | 7.815 | 9.348 | 11.345 | 12.838 |
| 4 | 0.207 | 0.297 | 0.484 | 0.711 | 1.064 | 7.779 | 9.488 | 11.143 | 13.277 | 14.860 |
| 5 | 0.412 | 0.554 | 0.831 | 1.145 | 1.610 | 9.236 | 11.070 | 12.833 | 15.086 | 16.750 |
| 6 | 0.676 | 0.872 | 1.237 | 1.635 | 2.204 | 10.645 | 12.592 | 14.449 | 16.812 | 18.548 |
| 7 | 0.989 | 1.239 | 1.690 | 2.167 | 2.833 | 12.017 | 14.067 | 16.013 | 18.475 | 20.278 |
| 8 | 1.344 | 1.646 | 2.180 | 2.733 | 3.490 | 13.362 | 15.507 | 17.535 | 20.090 | 21.955 |
| 9 | 1.735 | 2.088 | 2.700 | 3.325 | 4.168 | 14.684 | 16.919 | 19.023 | 21.666 | 23.589 |
| 10 | 2.156 | 2.558 | 3.247 | 3.940 | 4.865 | 15.987 | 18.307 | 20.483 | 23.209 | 25.188 |
| 11 | 2.603 | 3.053 | 3.816 | 4.575 | 5.578 | 17.275 | 19.675 | 21.920 | 24.725 | 26.757 |
| 12 | 3.074 | 3.571 | 4.404 | 5.226 | 6.304 | 18.549 | 21.026 | 23.337 | 26.217 | 28.300 |
| 13 | 3.565 | 4.107 | 5.009 | 5.892 | 7.042 | 19.812 | 22.362 | 24.736 | 27.688 | 29.819 |
| 14 | 4.075 | 4.660 | 5.629 | 6.571 | 7.790 | 21.064 | 23.685 | 26.119 | 29.141 | 31.319 |
| 15 | 4.601 | 5.229 | 6.262 | 7.261 | 8.547 | 22.307 | 24.996 | 27.488 | 30.578 | 32.801 |
| 16 | 5.142 | 5.812 | 6.908 | 7.962 | 9.312 | 23.542 | 26.296 | 28.845 | 32.000 | 34.267 |
| 17 | 5.697 | 6.408 | 7.564 | 8.672 | 10.085 | 24.769 | 27.587 | 30.191 | 33.409 | 35.718 |
| 18 | 6.265 | 7.015 | 8.231 | 9.390 | 10.865 | 25.989 | 28.869 | 31.526 | 34.805 | 37.156 |
| 19 | 6.844 | 7.633 | 8.907 | 10.117 | 11.651 | 27.204 | 30.144 | 32.852 | 36.191 | 38.582 |
| 20 | 7.434 | 8.260 | 9.591 | 10.851 | 12.443 | 28.412 | 31.410 | 34.170 | 37.566 | 39.997 |
| 21 | 8.034 | 8.897 | 10.283 | 11.591 | 13.240 | 29.615 | 32.671 | 35.479 | 38.932 | 41.401 |
| 22 | 8.643 | 9.542 | 10.982 | 12.338 | 14.041 | 30.813 | 33.924 | 36.781 | 40.289 | 42.796 |
| 23 | 9.260 | 10.196 | 11.689 | 13.091 | 14.848 | 32.007 | 35.172 | 38.076 | 41.638 | 44.181 |
| 24 | 9.886 | 10.856 | 12.401 | 13.848 | 15.659 | 33.196 | 36.415 | 39.364 | 42.980 | 45.559 |
| 25 | 10.520 | 11.524 | 13.120 | 14.611 | 16.473 | 34.382 | 37.652 | 40.646 | 44.314 | 46.928 |
| 26 | 11.160 | 12.198 | 13.844 | 15.379 | 17.292 | 35.563 | 38.885 | 41.923 | 45.642 | 48.290 |
| 27 | 11.808 | 12.879 | 14.573 | 16.151 | 18.114 | 36.741 | 40.113 | 43.195 | 46.963 | 49.645 |
| 28 | 12.461 | 13.565 | 15.308 | 16.928 | 18.939 | 37.916 | 41.337 | 44.461 | 48.278 | 50.993 |
| 29 | 13.121 | 14.256 | 16.047 | 17.708 | 19.768 | 39.087 | 42.557 | 45.722 | 49.588 | 52.336 |
| 30 | 13.787 | 14.953 | 16.791 | 18.493 | 20.599 | 40.256 | 43.773 | 46.979 | 50.892 | 53.672 |
| 40 | 20.707 | 22.164 | 24.433 | 26.509 | 29.051 | 51.805 | 55.758 | 59.342 | 63.691 | 66.766 |
| 50 | 27.991 | 29.707 | 32.357 | 34.764 | 37.689 | 63.167 | 67.505 | 71.420 | 76.154 | 79.490 |
| 60 | 35.534 | 37.485 | 40.482 | 43.188 | 46.459 | 74.397 | 79.082 | 83.298 | 88.379 | 91.952 |
| 70 | 43.275 | 45.442 | 48.758 | 51.739 | 55.329 | 85.527 | 90.531 | 95.023 | 100.425 | 104.215 |
| 80 | 51.172 | 53.540 | 57.153 | 60.391 | 64.278 | 96.578 | 101.879 | 106.629 | 112.329 | 117.155 |
| 90 | 59.196 | 61.754 | 65.647 | 69.126 | 73.291 | 107.565 | 113.145 | 118.136 | 124.116 | 128.299 |
| 100 | 67.328 | 70.065 | 74.222 | 77.929 | 82.358 | 118.498 | 124.342 | 129.561 | 135.807 | 140.169 |

TURN OVER