# ADVANCED MACHINE LEARNING EXERCISE #1 SOLUTION

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- (a) For a set of 3 variables we must look at 3 kinds of conditional independencies, w.l.o.g (all other permutations are symmetrical):
  - 1.  $X_1 \perp X_2$
  - 2.  $X_1 \perp X_2 | X_3$
  - 3.  $X_1 \perp X_2, X_3$

## Type 1 CI - hold

Starting with the first kind of CI, for any values for  $x_1, x_2$  it holds that  $q(x_1, x_2) = \frac{1}{6} + \frac{1}{12} = \frac{1}{4}$  and for any  $k \in [3]$ , and any value for  $x_k q(x_k) = \frac{1}{6} + \frac{1}{12} + \frac{1}{6} + \frac{1}{12} = \frac{1}{2}$ . Therefore  $q(x_1, x_2) = q(x_1)q(x_2)$  which implies that for all  $i, j \colon X_i \perp X_j \in I(q)$ .

## Type 2 CI - don't hold

We shall show an example for which this CI doesn't hold:

We will show  $p(x_1|x_3)p(x_2|x_3) \neq p(x_1,x_2|x_3)$  for the case:  $x_1,x_2,x_3=0$ 

$$p(x_1, x_2 | x_3) = \frac{p(x_1 = 0, x_2 = 0, x_3 = 0)}{p(x_3 = 0)} = \frac{\frac{1}{12}}{\frac{1}{2}} = \frac{1}{6}$$
$$p(x_1 | x_3) = p(x_2 | x_3) = \frac{\frac{1}{12} + \frac{1}{6}}{\frac{1}{2}} = \frac{1}{2}$$
$$\frac{1}{4} = p(x_1 | x_3)p(x_2 | x_3) \neq p(x_1, x_2 | x_3) = \frac{1}{6}$$

#### Type 3 CI - don't hold

We shall show an example for which this CI doesn't hold:

We will show  $p(x_1, x_2, x_3) \neq p(x_1)p(x_2, x_3)$  for the case:  $x_1, x_2, x_3 = 0$ 

$$p(x_1 = 0, x_2 = 0, x_3 = 0) = \frac{1}{12}$$

$$p(x_1 = 0) = \frac{1}{2}$$

$$p(x_2 = 0, x_3 = 0) = \frac{1}{12} + \frac{1}{6} = \frac{1}{4}$$

$$\frac{1}{12} = p(x_1 = 0, x_2 = 0, x_3 = 0) \neq p(x_1 = 0)p(x_2 = 0, x_3 = 0) = \frac{1}{8}$$

Overall:  $I(q) = \{(X_1 \perp X_2), (X_2 \perp X_3), (X_3 \perp X_1)\}.$ 

- (b) No, proof by contradiction. Let us assume there exists a DAG G where  $I_{LM}(G) = I(q)$ , so  $X_i \perp X_j \in I_{LM}(G)$  which means  $Pa(i) = \emptyset$  and  $ND(i) = \{X_j\}$  but this is true for any  $i, j \in \{(1, 2), (2, 3), (3, 1)\}$  and hence the contradiction is obvious (there are no parents but every node has one decendant for example).  $\square$
- (c) No, proof by contradiction. Let us assume there exists an undirected graph G where  $I_{sep}(G) = I(q)$ , so because for all  $i, j \colon X_i \perp X_j \in I_{sep}(G)$  there are no edges in G. because there are no routes from  $X_1$  to  $X_2, X_3$  at all it holds that  $X_1 \perp X_2, X_3 \in I_{sep}(G)$ . But we know that  $X_1 \perp X_2, X_3 \notin I(q)$  a contradiction to  $I_{sep}(G) = I(q)$ .  $\square$

# Old Question 1

For the second type of CI: it is clear that for any permutation (i, j, k) of [3] it holds that

$$q(x_i|x_k)q(x_j|x_k) = \frac{1}{4}$$

$$\neq \begin{cases} \frac{1}{6} & x_1 \oplus x_2 \oplus x_3 = 0\\ \frac{1}{3} & x_1 \oplus x_2 \oplus x_3 = 1 \end{cases} = 2q(x_1, x_2, x_3) = q(x_i, x_j|x_k)$$

which implies  $X_i \perp X_j | X_k \notin I(q)$ . For the Third kind of CI:

$$q(x_i) = \frac{1}{2}$$

$$\neq \begin{cases} \frac{1}{3} & x_1 \oplus x_2 \oplus x_3 = 0\\ \frac{2}{3} & x_1 \oplus x_2 \oplus x_3 = 1 \end{cases} = 4q(x_1, x_2, x_3) = q(x_i | x_j, x_k)$$

which implies  $X_i \perp X_j, X_k \notin I(q)$ . so overall we get  $I(q) = \{X_1 \perp X_2, X_2 \perp X_3, X_3 \perp X_1\}$ .

- (b) No, proof by contradiction. Let us assume there exists a DAG G where  $I_{LM}(G) = I(q)$ , so  $X_i \perp X_j \in I_{LM}(G)$  which means  $Pa(i) = \emptyset$  and  $ND(i) = \{X_j\}$  but this is true for any  $i, j \in \{(1, 2), (2, 3), (3, 1)\}$  and hence the contradiction is obvious (there are no parents but every node has one decendant for example).  $\square$
- (c) No, proof by contradiction. Let us assume there exists an undirected graph G where  $I_{sep}(G) = I(q)$ , so  $X_i \perp X_j \in I_{sep}(G)$  which means there are no edges in G but in this case for example  $X_1 \perp X_2, X_3 \in I_{sep}(G)$  where  $X_1 \perp X_2, X_3 \notin I_{sep}(G)$  and hence  $I_{sep}(G) \neq I(q)$ , a contradiction.  $\square$

We are given a positive distribution p(w, x, y, z). It is known that:

- 1. We are given  $(X \perp Y|Z, W)$ , equivalently: p(x|y, z, w) = p(x|z, w)
- 2. We are given  $(X \perp W|Z,Y)$ , equivalently: p(x|w,z,y) = p(x|z,y)

From 1,2 we have: p(x|z, w) = p(x|z, y) (\*)

We will prove that  $(X \perp Y, W|Z)$ , that is: p(x|y, w, z) = p(x|z):

From 1, we have p(x|y,z,w)=p(x|z,w). Multiplying both sides by p(y|z,w) gives:

$$p(y|z, w)p(x|y, z, w) = p(y|z, w)p(x|z, w)$$

Equivalently:

$$p(x, y|w, z) = p(y|w, z)p(x|w, z)$$

Multiplying both sides by p(w|z) gives:

$$p(w|z)p(x,y|w,z) = p(w|z)p(y|w,z)p(x|w,z)$$

Equivalently:

$$p(x, y, w|z) = p(w|z)p(y|w, z)p(x|w, z)$$

Summing over y gives:

$$p(x, w|z) = p(w|z)p(x|w, z)$$

From (\*):

$$p(x, w|z) = p(w|z)p(x|y, z)$$

Summing over w gives:

$$p(x|z) = p(x|y,z)$$

From 2:

$$p(x|z) = p(x|w, z, y) \Rightarrow X \perp Y, W|Z$$

We define  $S = Pa(i) \cup Ch(i) \cup \bigcup_{j \in Ch(i)} Pa(j)$  and argue that S is the Markov

Blanket for some Bayesian network p on G. We first show that  $X_i \perp X_{\bar{S}\setminus i}|X_S \in I_{d-sep}(G)$ .

More specifically we will show that:  $X_i$  is d-separated from  $X_{\bar{S}\setminus i}$  given  $X_S$  by showing that there is no active trail from  $X_i$  to some node  $j \in X_{\bar{S}\setminus i}$  given  $X_S$  (definition 6).

There are 3 possible cases for an undirected path between i and j:

- An undirected path between i and j via  $k \in Ch(i)$  and some  $l \in Pa(k)$ . Then because  $l \in S$  and l is not the descendent of the child in a v-structure the path is not active.
- An undirected path between i and j via  $k \in Ch(i)$  but not via some  $l \in Pa(k)$ , then  $k \in S$  and k is not the child in a v-structure in the path or a descendent of one so the path is not active.
- An undirected path between i and j via  $k \in Pa(i)$ . In this case k which is in S cannot be the child in a v-structure or a descendent of one so the path is not active.

Overall, there is no active trail between  $X_i$  and  $X_{\bar{S}\setminus i}$  given  $X_S$  and we get  $X_i \perp X_{\bar{S}\setminus i}|X_S \in I_{d-sep}(G) \subseteq I(p)$  by definition 6 and Proposition 2.3.1. We now have left to show the minimalism of S. If we remove  $k \in Pa(i) \cup Ch(i)$ , it is clear that  $X_i$  and  $X_k$  are dependant given  $X_S$  and hence  $X_i \perp X_{\bar{S}\setminus i}|X_S \notin I(p)$ . If we remove  $k \in \bigcup_{i \in Ch(i)} Pa(i)$ , we have an active trail i and k given  $X_S$  (via

the mutual child) and hence  $X_i \perp X_{\bar{S}\backslash i}|X_S \notin I_{d-sep}(G)$ . We know that except for measure zero set, all Baseyian networks on G will hold  $I(p) = I_{d-sep}(G)$ , and therefore in any case of removal we get  $X_i \perp X_{\bar{S}\backslash i}|X_S \notin I(p)$  so S is indeed minimal.  $\square$ 

# Question 4

- (a) By Theorem 4: as p is positive it suffices to show that  $I_{pair}(G) \subseteq I(p)$ . Indeed, if  $X_i \perp X_j | X_{V \setminus \{i,j\}} \in I_{pair}(G)$  then by definition of G we get  $ij \notin E$  and hence  $X_i \perp X_j | X_{V \setminus \{i,j\}} \in I(p)$  since otherwise edge ij was supposed to be added to G. Therefore  $I_{sep}(G) \subseteq I(p)$ .  $\square$
- (b) Towards contradiction, let us assume that G is not a minimal I-map for p. Therefore, there exists edge ij such that it can be removed and still  $I_{sep}(G) \subseteq I(p)$ . On one hand, be the definition of G, we have  $X_i \perp X_j | X_{V \setminus \{i,j\}} \notin I(p)$ . On the other hand, since now  $ij \notin E$ , we get  $X_i \perp X_j | X_{V \setminus \{i,j\}} \in I_{pair}(G)$ , and hence  $X_i \perp X_j | X_{V \setminus \{i,j\}} \in I_{sep}(G)$  so therefore  $I_{sep}(G) \nsubseteq I(p)$ , a contradiction. We thus conclude G is indeed minimal I-map for p.  $\square$

Obviously,  $I_{sep}(G) = \{X_1 \perp X_3 | X_{2,4}, X_2 \perp X_4 | X_{1,3}\}$ . First, we show that  $X_1 \perp X_3 | X_{2,4} \in I(p)$ . Because  $p(x_3 = 0) = 1/2$ , the following tables show  $p(x_1 | x_{2,3,4}) = p(x_1 | x_{2,4})$  and hence  $X_1 \perp X_3 | X_{2,4} \in I(p)$ :

Each cell is the probability of  $p(x_1|x_{2,3,4})$ :

$x_1 x_{2,3,4}$	0,0,0	1,0,0	0,1,0	1,1,0	0,0,1	1,0,1	0,1,1	1,1,1
0	1/2	0	undefined	0	1	unedfined	1	1/2
1	1/2	1	undefined	1	0	undefined	0	1/2

Each cell is the probability of  $p(x_1|x_{2,4})$ :

$x_1 x_{2,4}$	0,0	1,0	0,1	1,1
0	1/2	0	1	1/2
1	1/2	1	0	1/2

 $X_2 \perp X_4 | X_{1,3} \in I(p)$  is showed in the same way. Therefore we have  $I_{sep}(G) \subseteq I(p)$ . To end our proof, we now show that p is not a Markov network with respect to G. Towards a contradiction, let us assume that p is indeed a Markov network with respect to G. We know that  $\mathcal{C}(G) = \{\{X_1, X_4\}, \{X_1, X_2\}, \{X_3, X_4\}, \{X_2, X_3\}\}$ . Hence, by the definition of Markov network,  $p(\mathbf{x})$  can be written as:

$$p(\mathbf{x}) = Z^{-1} \prod_{c \in \mathcal{C}(G)} \phi_c(x_c) = Z^{-1} \phi_{12}(x_{1,2}) \phi_{23}(x_{2,3}) \phi_{34}(x_{3,4}) \phi_{41}(x_{4,1})$$

Let us inspect the following equations:

$$p(0,0,0,0) = \phi_{12}(0,0)\phi_{23}(0,0)\phi_{34}(0,0)\phi_{41}(0,0) = Z/8 \tag{1}$$

$$p(0,0,1,1) = \phi_{12}(0,0)\phi_{23}(0,1)\phi_{34}(1,1)\phi_{41}(1,0) = Z/8 \tag{2}$$

$$p(1,1,1,0) = \phi_{12}(1,1)\phi_{23}(1,1)\phi_{34}(1,0)\phi_{41}(0,1) = Z/8$$
 (3)

$$p(0,0,1,0) = \phi_{12}(0,0)\phi_{23}(0,1)\phi_{34}(1,0)\phi_{41}(0,0) = 0 \tag{4}$$

From (4), we know that  $\phi_{12}(0,0)=0 \lor \phi_{23}(0,1)=0 \lor \phi_{34}(1,0)=0 \lor \phi_{41}(0,0)=0$ , but if any one of those is true, at least one out of (1), (2), (3) will be equal to  $0 \ne Z/8$ , a contradiction. We therefore conclude that p is not a Markov network with respect to G.  $\square$ 

G = (E, V) is a tree graph. p(x) is a markov network on G. We will show that for any assignment  $x_1, ..., x_n$  it holds that:

$$p(x_{[n]}) = \prod_{i=1}^{n} p(x_i) \prod_{ij \in E} \frac{p(x_i x_j)}{p(x_i) p(x_j)}$$

Note: we use the notation  $x_{[n]}$  to denote:  $x_1, ..., x_n$ 

## **Proof:**

We will prove the claim by induction over the size of V. Base case |V| = 1 trivially holds:

$$p(x_1) = \prod_{i=1}^{1} p(x_i) \prod_{ij \in \emptyset} \frac{p(x_i x_j)}{p(x_i)p(x_j)}$$

The case for |V|=2 is similar (and will enable us to assume  $n\geq 3$  in all cases henceforth):

$$p(x_1, x_2) = \prod_{i=1}^{2} p(x_i) \prod_{i \neq i \in \{1, 2\}} \frac{p(x_i x_j)}{p(x_i) p(x_j)}$$

Let us assume that the claim holds for any graph where |V| < n, and we now prove for |V| = n. Let G = (V, E) be some tree graph where |V| = n, and p(x) a markov network on this graph. It holds for any distribution that:

$$p(x_{[n]}) = p(x_n | x_{[n-1]}) p(x_{[n-1]})$$
(5)

In any tree graph there exists a leaf. Assume w.l.o.g that  $x_n$  is a leaf. That is  $x_n$  has only one neighbour. Assume w.l.o.g that this neibbour is  $x_{n-1}$ . By Theorem 4.1 (if p factorizes according to G then  $I_{sep}(G) \subseteq I(p)$ ) and the fact that  $I_{LM}(G) \subseteq I_{sep}(G)$ :

$$p(x_n|x_{\lceil n-1\rceil}) = p(x_n|x_{n-1})$$

Replacing this in (5) we arrive at:

$$p(x_{[n]}) = p(x_n | x_{n-1}) p(x_{[n-1]})$$
(6)

Note that the sub-graph resulting from removing  $x_n$  from G is also a tree as  $x_n$  was a leaf so the remaining graph is still connected and has no cycles.

Claim:  $p(x_{[n-1]})$  is a markov network on the tree graph G' resulting from G by removing the leaf  $x_n$ . Proof:

By definition of markov network and the fact that all cliques in a tree are of size 2 (or else we will have cycles), corresponding exactly to edges:

$$p(x_{[n-1]}) = \sum_{x_n} p(x_{[n-1]}, x_n) = \sum_{x_n} \frac{1}{Z} \prod_{ij \in E} \phi_{ij}(x_i, x_j) = \frac{1}{Z} \prod_{ij \in E \setminus (n, n-1)} \phi_{ij}(x_i x_j) \sum_{x_n} \phi_{n, n-1}(x_n, x_{n-1})$$
(7)

In the last equality we used the fact that all  $x_i$  where  $i \neq n$  are constants within the sum. We would like to arrange the expression as a product of  $\phi_{ij}(x_i,x_j)$  for  $ij \in E \setminus (n,n-1)$ . To acheive this we will "push" the value of  $\sum_{x_n} \phi_{n.n-1}(x_n,x_{n-1})$  into one of the  $\phi_{ij}$ . We need all functions to depend only on their 2 parameters so we have to choose one where  $x_{n-1}$  is one. Recognize that because  $n \geq 3$  and  $x_{n-1}$  is connected to the rest of the graph, there exists some variable, w.l.o.g  $x_{n-2}$ , s.t:  $(x_{n-1},x_{n-2}) \in E$ . Now we define:

$$\phi'_{n-1,n-2}(x_{n-1},x_{n-2}) = \phi_{n-1,n-2}(x_{n-1},x_{n-2}) \sum_{x_n} \phi_{n,n-1}(x_n,x_{n-1})$$

For ease of notation, for all other pairs denote  $\phi'_{ij} = \phi_{ij}$  We arrive at:

$$\prod_{ij \in E \setminus (n,n-1)} \phi_{ij}(x_i x_j) \sum_{x_n} \phi_{n.n-1}(x_n,x_{n-1}) = \prod_{ij \in E \setminus (n,n-1), (n-1,n-2)} \phi_{ij}(x_i x_j) \cdot (\phi_{n-1.n-2}(x_{n-1},x_{n-2}) \sum_{x_n} \phi_{n.n-1}(x_n,x_{n-1})) = \prod_{ij \in E \setminus (n,n-1)} \phi'_{ij}(x_i,x_j)$$

And for Z

$$Z = \sum_{x_{[n]}} \prod_{ij \in E} \phi_{ij}(x_i, x_j) = \sum_{x_{[n-1]}} \sum_{x_n} \prod_{ij \in E} \phi_{ij}(x_i, x_j) = \sum_{x_{[n-1]}} \prod_{ij \in E \setminus (x_{n-1}, x_n)} \phi_{ij}(x_i, x_j) \sum_{x_n} \phi_{n-1, n}(x_{n-1}, x_n) = \sum_{x_{[n-1]}} \prod_{ij \in E \setminus (x_{n-1}, x_n)} \phi'_{ij}(x_i, x_j)$$

Finally arriving at:

$$p(x_{[n-1]}) = \frac{1}{Z} \prod_{ij \in E \setminus (n,n-1)} \phi_{ij}(x_i x_j) \sum_{x_n} \phi_{n.n-1}(x_n, x_{n-1}) = \frac{1}{\sum_{x_{[n-1]}} \prod_{ij \in E \setminus (x_{n-1}, x_n)} \phi'_{ij}(x_i, x_j)} \prod_{ij \in E \setminus (n.n-1)} \phi'_{ij}(x_i, x_j)$$

So  $p(x_{[n-1]})$  is a markov network on the resulting tree!

#### We continue with the induction:

By the induction hypothesis, on the Tree resulting from removing  $x_n$  from G:

$$\begin{split} p(x_{[n]}) &= p(x_n|x_{n-1})p(x_{[n-1]}) = \\ p(x_n|x_{n-1}) \prod_{i \in [n-1]} p(x_i) \prod_{ij \in E \backslash (n,n-1)} \frac{p(x_i,x_j)}{p(x_i)p(x_j)} = \\ &\frac{p(x_n,x_{n-1})}{p(x_{n-1})} \prod_{i \in [n-1]} p(x_i) \prod_{ij \in E \backslash (n,n-1)} \frac{p(x_i,x_j)}{p(x_i)p(x_j)} = \\ p(x_n) \prod_{i \in [n-1]} p(x_i) \prod_{ij \in E \backslash (n,n-1)} \frac{p(x_i,x_j)}{p(x_i)p(x_j)} \cdot \frac{p(x_n,x_{n-1})}{p(x_n)p(x_{n-1})} = \\ &\prod_{i \in [n]} p(x_i) \prod_{ij \in E} \frac{p(x_i,x_j)}{p(x_i)p(x_j)} \end{split}$$

# Old Question 6

First, w.l.o.g we assume that  $\forall ij \in E, i < j$ . Proof by induction on the size of V. For |V| = 1 the claim holds trivially (as there are no edges in the graph, the markov network p(x) is in the required form). Let us assume the claim holds  $\forall k \in [n-1]$  and we now prove for n. Let there be G, some graph with n nodes, and p(x) a markov network on this graph. We know that for any distribution:

$$p(x_{[n]}) = p(x_n | x_{[n-1]}) p(x_{[n-1]})$$
(8)

From the inductive hypothesis:

$$p(x_{[n-1]}) = \prod_{i=1}^{n-1} p(x_i) \prod_{\substack{ij \in E \\ i,j \in [n-1]}} \frac{p(x_i, x_j)}{p(x_i)p(x_j)}$$
(9)

By Theorem 4.1 (if p factorizes according to G then  $I_{sep}(G) \subseteq I(p)$ ) and the fact that  $I_{LM}(G) \subseteq I_{sep}(G)$ :

$$p(x_n|x_{[n-1]}) = p(x_n|x_{Nbr(n)}) = \frac{p(x_n, x_{Nbr(n)})}{p(x_{Nbr(n)})}$$

G is a tree a graph, hence there are no circles and therefore, w.l.o.g Nbr(n) < n-1. So, again by the inductive hypothesis:

$$p(x_{n}, x_{Nbr(n)}) = \prod_{i \in Nbr(n) \cup \{n\}} p(x_{i}) \prod_{\substack{ij \in E \\ i,j \in Nbr(n) \cup \{n\}}} \frac{p(x_{i}, x_{j})}{p(x_{i})p(x_{j})}$$

$$p(x_{Nbr(n)}) = \prod_{i \in Nbr(n)} p(x_{i}) \prod_{\substack{ij \in E \\ i,j \in Nbr(n)}} \frac{p(x_{i}, x_{j})}{p(x_{i})p(x_{j})}$$

$$\Rightarrow p(x_{n}|x_{[n-1]}) = p(x_{n}) \prod_{in \in E} \frac{p(x_{i}, x_{j})}{p(x_{i})p(x_{j})}$$
(10)

By (5), (6), (7) we get:

$$p(x_n) = \prod_{i=1}^{n} p(x_i) \prod_{ij \in E} \frac{p(x_i, x_j)}{p(x_i)p(x_j)}$$