# FLS 6441 - Methods III: Explanation and Causation

Week 1 - Review of Regression

Jonathan Phillips

March 2019

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Course Website

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- 10. Matching (23rd May)
- 11. Comparative Cases and Process Tracing (30th May)

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- 12. Generalizability, Reproducibility and Mechanisms (6th June)

## Course Schedule

Wednesday 18h - Submit Replication Task

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- ► Wednesday 18h Submit Replication Task
- ► Thursday 14h-16h Class (105)

What Does Regression NOT Do?

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- ► Thursday 14h-16h Class (105)
- ► Thursday 16.15-17.30 Class (105) OR Lab (122)

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- ► Thursday 14h-16h Class (105)
- ► Thursday 16.15-17.30 Class (105) OR Lab (122)
- ► Friday 10h-12h Office Hours (DCP 2061)
  - Or just send me an email

► Replication Tasks - 40%

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  - 8 best grades out of 10 tasks

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- ► Participation 20%

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- ► Tip: Pick a simple causal question and dataset

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Guide to 'Smart' Regression

# Today's Objectives

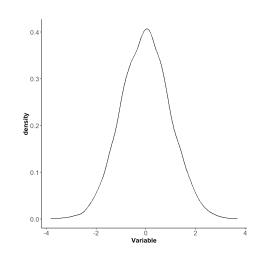
- 1. What Does Regression Actually Do?
- 2. Guide to 'Smart' Regression
- 3. What Does Regression NOT Do?

# Section 1

What Does Regression Actually Do?

## Data

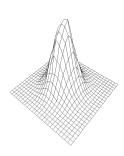
► We work with variables, which VARY!



## Data

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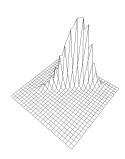
Variable_1	Variable_2
-0.44	0.63
0.06	0.68
-0.21	-0.02
0.44	-0.25
1.29	0.46
-0.38	-0.81
-1.04	0.24
-0.16	1.84
1.29	0.06
-0.10	-0.18



## Data

► We work with variables, which VARY!

Variable_1	Variable_2
-0.86	-0.69
-0.35	-0.44
1.27	0.42
-0.35	-0.22
-0.43	-0.56
0.05	-0.04
0.69	0.70
1.27	1.07
0.22	-0.00
-0.28	-0.13



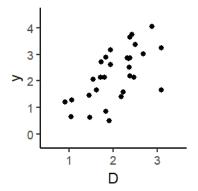
# What Does Regression Actually Do?

- 1. Regression as Least Squares
- 2. Regression as Conditional Expectation
- 3. Regression as (Partial) Correlation

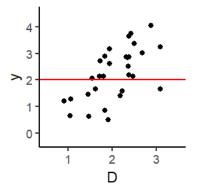
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- ▶  $y_i = \alpha + \beta D_i + \epsilon_i$

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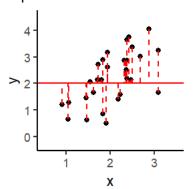


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Slope = 0

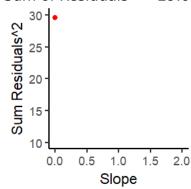


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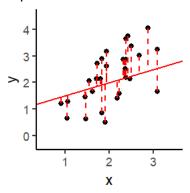
X

Sum of Residuals $^2 = 29.6$ 

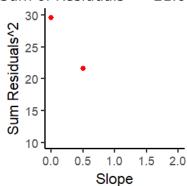


- Regression identifies the line through the data that minimizes the sum of squared vertical distances
- $\rightarrow y_i = \alpha + \beta D_i + \epsilon_i$

Slope = 0.5



Sum of Residuals $^2 = 21.6$ 

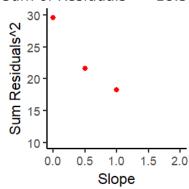


- Regression identifies the line through the data that minimizes the sum of squared vertical distances
- $\rightarrow y_i = \alpha + \beta D_i + \epsilon_i$

Slope = 13

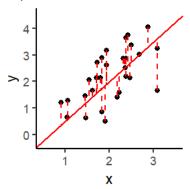
X

# Sum of Residuals $^2 = 18.3$

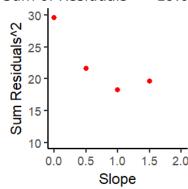


- Regression identifies the line through the data that minimizes the sum of squared vertical distances
- $\rightarrow y_i = \alpha + \beta D_i + \epsilon_i$

Slope = 1.5



Sum of Residuals $^2 = 19.6$ 

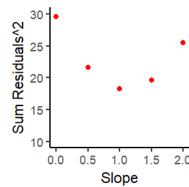


- Regression identifies the line through the data that minimizes the sum of squared vertical distances
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Slope = 2

X

Sum of Residuals $^2 = 25.5$ 

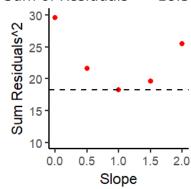


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Slope = 13

X

Sum of Residuals $^2 = 18.3$ 

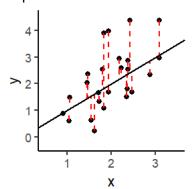


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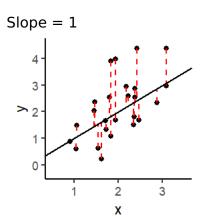
$$y_i = \alpha + \beta D_i + \epsilon_i$$

Slope = 1

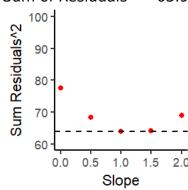


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$$\rightarrow y_i = \alpha + \beta D_i + \epsilon_i$$



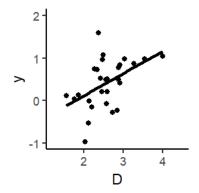
## Sum of Residuals $^2 = 63.9$



- Dummy control variables remove variation associated with specific levels or categories
  - ► The same for Fixed Effects

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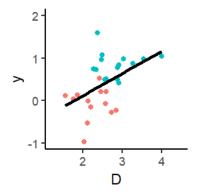
$$y_{ij} = \alpha + \beta_1 D_{ij} + \epsilon_i$$



Ignoring the dummy control variable, the slope coefficient is 1

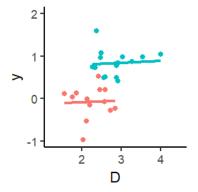
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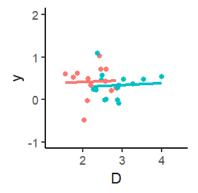
But the data points really represent two very different groups, blues and reds

- ► Dummy control variables *remove variation* associated with specific levels or categories
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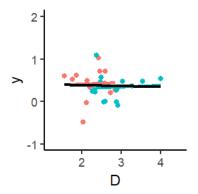
What if we ran the regression for each group *separately*?

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Dummy control variables remove the average Y differences between blues and reds

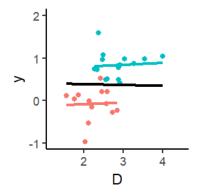
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The new regression line for the full data now has a slope of zero

- ▶ Dummy control variables remove variation associated with specific levels or categories
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$$y_{ij} = \alpha + \beta_1 D_{ij} + \beta_2 X_j + \epsilon_i$$

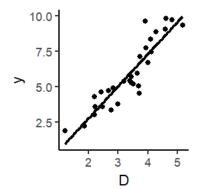


Equivalently, dummy control variables restrict comparisons to **within the same group**:

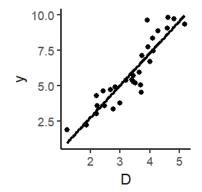
- 1. How much does *D* affect *Y* within the blue group? 0
- How much does D affect Y within the red group? 0
- What's the average of (1) and
   (2) (weighted by the number of units in each group)?

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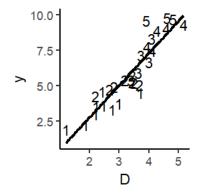


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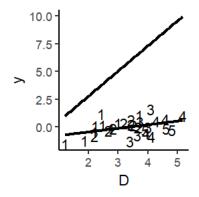
The coefficient  $\beta_1$  is 2.267 Real effect = 1

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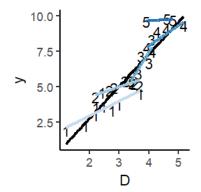
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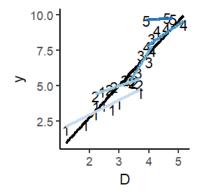
The coefficient  $\beta_1$  is 1.024 Real effect = 1

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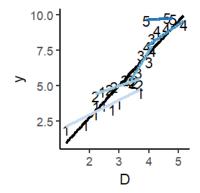
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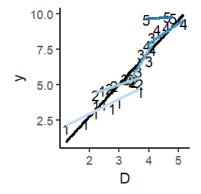
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- ► Then average these slopes,  $\beta_1 \sim 1$

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- ► Then average these slopes,  $\beta_1 \sim 1$
- ► Impossible with truly continuous variables
- ► So regression uses linearity to fill in the gaps

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- ► Conditional on D, What is our expectation (mean value) of *y*?

$$y_i = \alpha + \beta_1 D_i + \epsilon_i$$
  
$$E(y) = \alpha + \beta_1 D$$

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$$y_i = \alpha + \beta_1 D_i + \epsilon_i$$

$$Attitude_i = \alpha + \beta_1 Income_i + N(0, \sigma^2)$$

Conditional on a specific value of D, what is our expectation (mean value) of y?

$$y_i = \alpha + \beta_1 D_i + \epsilon_i$$

$$Attitude_i = \alpha + \beta_1 Income_i + N(0, \sigma^2)$$

$$Attitude_i = 2.235 - 0.000818 * Income_i + N(0, 2.38)$$

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$$(Attitude_i | Income_i = 3000) = 2.235 - 0.000818 * 3000 + N(0, 2.38)$$

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$$E(Attitude | Income = 3000) = -0.22$$

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- ► E(Attitude|Income)
  - ▶ When income is 3000, the average attitude is -0.22

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- ► E(Attitude|Income)
  - ▶ When income is 3000, the average attitude is -0.22
  - When income is 6000, the average attitude is -2.67

$$y_i = \alpha + \beta_1 D_i + \epsilon_i$$

$$Attitude_i = \alpha + \beta_1 Income_i + N(0, \sigma^2)$$

$$Attitude_i = 2.235 - 0.000818 * Income_i + N(0, 2.38)$$

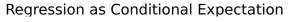
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  - ▶ When income is -1000, the average attitude is 3.05

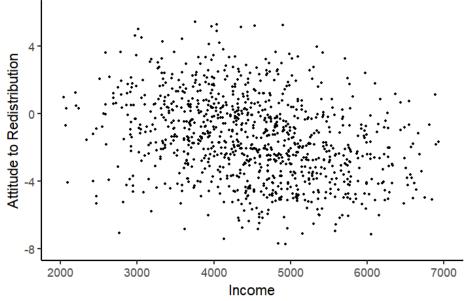
$$y_i = \alpha + \beta_1 D_i + \epsilon_i$$

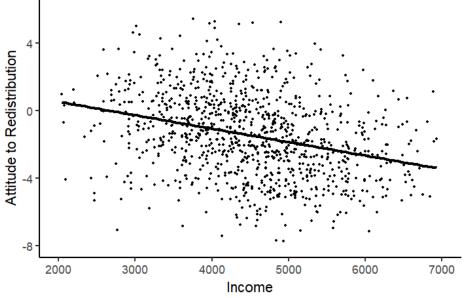
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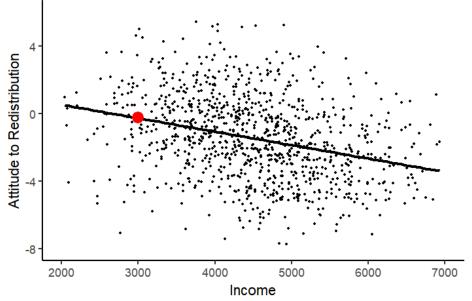
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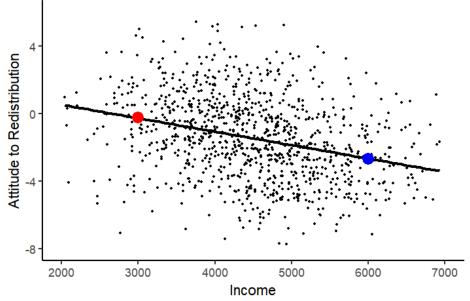
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  - When income is 6000, the average attitude is -2.67
  - ▶ When income is -1000, the average attitude is 3.05
- ► *E*(*Attitude*|Income, Age, Gender, Municipality)

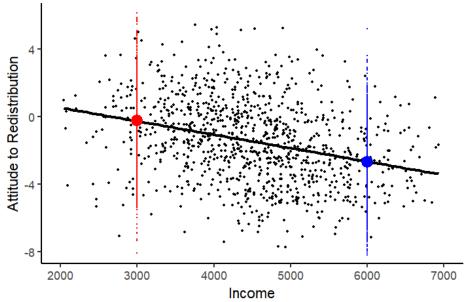












 $\blacktriangleright$  How do we work out the conditional expectation? We estimate the  $\beta$ s

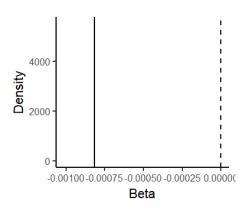
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- $\blacktriangleright$  Regression **estimates a distribution** for each  $oldsymbol{eta}$

- ► How do we work out the conditional expectation? We estimate the  $\beta$ s
- ▶ But we **NEVER** know the exact value of  $\beta$
- Regression estimates a distribution for each β
  - ightharpoonup That's why every eta comes with a standard error

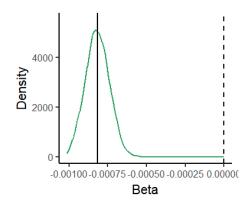
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	Dependent variable:
	redist
income	-0.000818*** (0.000078)
Constant	2.234719*** (0.361477)
Observations	1,000
Note:	*p<0.1; **p<0.05; ***p<0.01



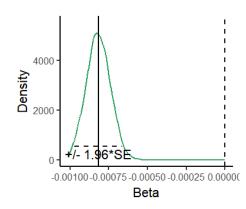
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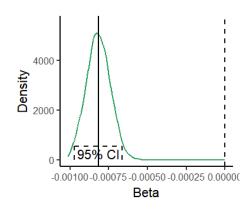
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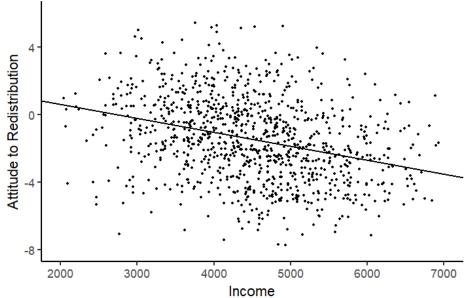


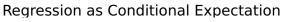
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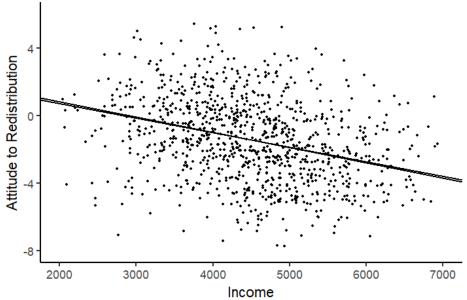
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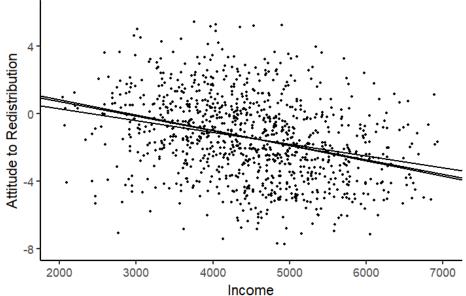


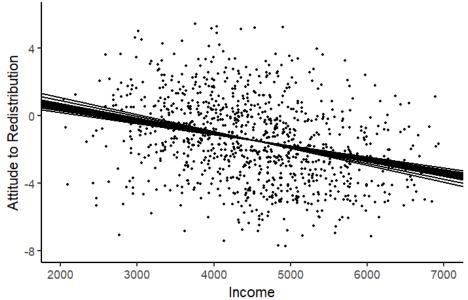


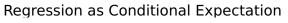


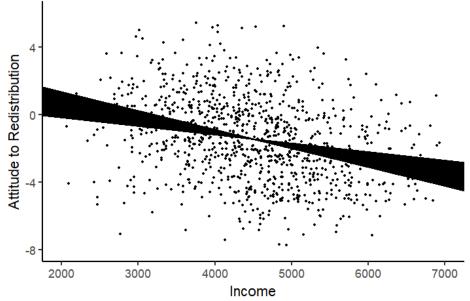


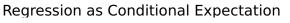


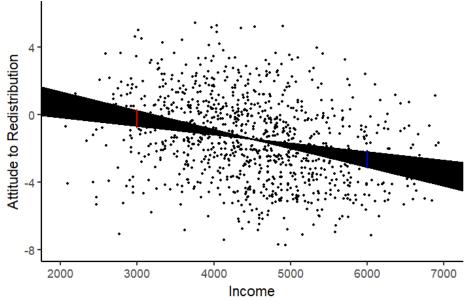


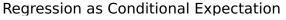


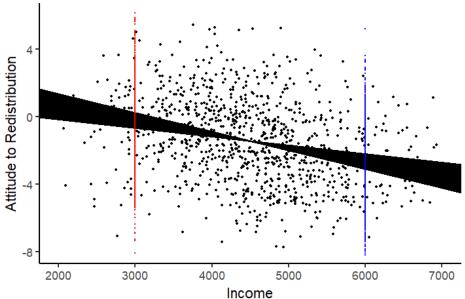








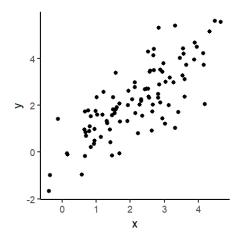




Regression with two variables is very similar to calculating correlation:

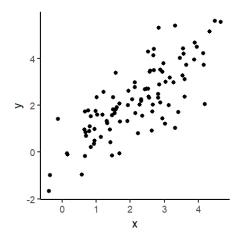
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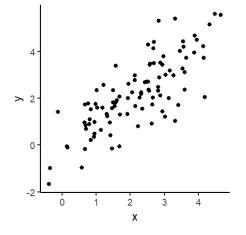
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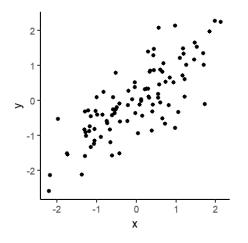


- ► Correlation is 0.781
- ► Regression Results:

-		term	estimate
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	2	X	1.008

Regression with two variables is very similar to calculating correlation:

$$\blacktriangleright \hat{\beta} = cor(x, y) * \frac{\sigma_Y}{\sigma_X}$$



- ► Correlation is 0.781
- ► It's **identical** if we standardize both variables first  $(\frac{(x_i \bar{x})}{\sigma_x})$
- Standardized Regression Results:

term		estimate
1	(Intercept)	0.000
2	Χ	0.781

► Regression with **multiple** variables is very similar to calculating **partial** correlation

- Regression with multiple variables is very similar to calculating partial correlation

#### Regression as (Partial) Correlation

- Regression with multiple variables is very similar to calculating partial correlation
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- Just a small difference in the denominator (how we standardize the measure)

#### Regression as (Partial) Correlation

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- Just a small difference in the denominator (how we standardize the measure)

$$\beta_{x_1} = \frac{r_{yx_1} - r_{yx_2} r_{x_1 x_2}}{1 - r_{x_1 x_2}^2}$$

$$r_{yx_1|x_2} = \frac{r_{yx_1} - r_{yx_2} r_{x_1 x_2}}{\sqrt{(1 - r_{yx_2}^2)(1 - r_{x_1 x_2}^2)}}$$

► There is no magic in regression, it's just 'extra' correlation

# Section 2

Guide to 'Smart' Regression

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- 3. There are fundamental best practices that apply to all the methodologies

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- 8. **Predict Meaningful Comparisons:** To communicate your findings

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- ► What measure of income should we use?
  - Pre-tax, post-tax, after government benefits?
- ▶ It depends on the theory we are testing

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- ► We may as well throw the Qatar data away

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  - Do you want a lot more, more, the same, less, or a lot less redistribution?
- ▶ Count -> Poisson
  - In the past year, how many times have you complained about redistribution?

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- ► Removing *ALL* the variation between countries
  - If rich countries have stronger attitudes to redistribution, we control for this
  - So we can ask whether richer people have stronger attitudes
- ► Our question becomes: How do variations within income in the same country affect attitudes to redistribution?

#### 6. Errors Structure

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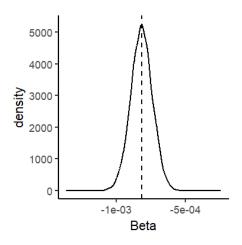
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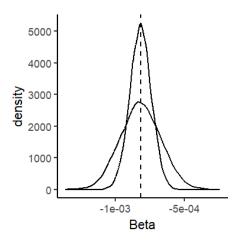
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- So we don't really have 2 observations, we have closer to 1 'independent' observation
- ► So the standard errors for our  $\beta$ 's are *over-confident* (too small)
- We need to adjust for these dependencies with clustered standard errors
  - Created by the underlying structure of the data
  - Or by our data sampling process



► The distribution of our estimated betas suggests we're pretty confident β is close to −0.0008175



With clustered SEs, the wider distribution of our betas suggests we're less confident β is close to -0.0008175

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  - 1. The scale of the explanatory variable

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- ► Basic OLS:  $y_i = \alpha + \beta D_i + \epsilon$ 
  - A 1 [unit of D] change in the explanatory variable is associated with a  $\beta$  [unit of y] change in the outcome, holding other variables constant

- ▶ Difficult! It depends on:
  - 1. The scale of the explanatory variable
  - 2. The scale of the outcome
  - 3. The regression model we used
  - 4. The presence of any interaction
- ▶ Basic OLS with log outcome:  $log(y_i) = \alpha + \beta D_i + \epsilon$ 
  - ► A 1 [unit of D] change in the explanatory variable is associated with a  $100 * (e^{\beta} 1)$ % change in the outcome, holding other variables constant

- ► Difficult! It depends on:
  - 1. The scale of the explanatory variable
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  - 3. The regression model we used
  - 4. The presence of any interaction
- ▶ Basic OLS with log treatment:  $y_i = \alpha + \beta log(D_i) + \epsilon$ 
  - A 1% change in the explanatory variable is associated with a  $\beta * log(\frac{101}{100})$  change in the outcome, holding other variables constant

- ► Difficult! It depends on:
  - 1. The scale of the explanatory variable
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  - 3. The regression model we used
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- ► **Logit:**  $Pr(y_i = 1) = logit^{-1}(\alpha + \beta D_i + \epsilon)$ 
  - ► A 1 [unit of D] change in the explanatory variable is associated with a  $\beta$  change in the log-odds of  $y_i = 1$ , holding other variables constant

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  - A 1 [unit of D] change in the explanatory variable is associated with a  $100*(e^{\beta}-1)$ % change in the odds (relative probability,  $\frac{\rho}{1-\rho}$ ) of  $y_i=1$ , holding other variables constant

- ▶ Difficult! It depends on:
  - 1. The scale of the explanatory variable
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- ▶ Multinomial:  $Pr(y_i = C) = \alpha + \beta D_i + \epsilon$ 
  - ► A 1 [unit of D] change in the explanatory variable is associated with a  $100 * (e^{\beta c} 1)\%$  change in the odds (relative probability,  $\frac{p}{1-p}$ ) of moving from the baseline category to the outcome category C, holding other variables constant

- ► Difficult! It depends on:
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- ▶ Ordered Multinomial:  $Pr(y_i = C) = \alpha + \beta D_i + \epsilon$ 
  - ► A 1 [unit of D] change in the explanatory variable is associated with a  $100 * (e^{\beta c} 1)\%$  change in the odds (relative probability,  $\frac{p}{1-p}$ ) of moving up one unit on the outcome scale, holding other variables constant

- ▶ Difficult! It depends on:
  - 1. The scale of the explanatory variable
  - 2. The scale of the outcome
  - 3. The regression model we used
  - 4. The presence of any interaction
- ▶ **OLS with Interaction:**  $y_i = \alpha + \beta_1 D_i + \beta_2 X_i + \beta_3 D_i * X_i + \epsilon$

$$\frac{\partial y}{\partial D} = \beta_1 + \beta_3 X$$

- ▶  $\beta_1$  is the effect of D when X = 0: May not make sense!
- ▶ Insert values for X and see how the marginal effect changes

### **OLS with Interaction:**

Redist<sub>i</sub> =  $\alpha + \beta_1$ Gender<sub>i</sub> +  $\beta_2$ Income<sub>i</sub> +  $\beta_3$ Gender<sub>i</sub> \* Income<sub>i</sub> +  $\epsilon_i$ 

### **OLS with Interaction:**

+ 
$$\beta_3$$
Gender<sub>i</sub> \* Income<sub>i</sub> +  $\epsilon_i$ 

$$\frac{\partial Redist}{\partial Gender} = \beta_1 + \beta_3 * Income$$

 $Redist_i = \alpha + \beta_1 Gender_i + \beta_2 Income_i$ 

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	Dependent variable:
	redist
gender1	-2.942614*** (0.700510)
income	0.079980 * * * (0.000110)
gender1:income	0.000986*** (0.000152)
Constant	0.112903 (0.454926)
Observations	1,000
Note:	*p<0.1; **p<0.05; ***p<0.01

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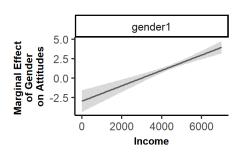
			ge	nder1		
#	5.0				4	
Marginal Effect of Gender on Attitudes	2.5 -					
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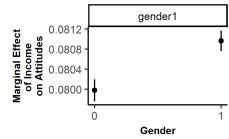
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  - And p-values are arbitrary

- ► The coefficient on the regression of income on attitude to redistribution is -0.000818
  - ► So??? What do we learn from this?
  - Coefficients are hard to interpret, and depend on how we measure each variable
  - And p-values are arbitrary
- Better to make specific predictions of how changes in D produce changes in Y

$$Attitude_i = \alpha + \beta_1 \text{ Income}_i + \epsilon_i$$

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### If Income is 3000:

$$Attitude_i = 2.235 - 0.000818 * 3000 + N(0, 2.378)$$

$$Attitude_i = -0.219 + N(0, 2.378)$$

$$Attitude_i = \alpha + \beta_1 \text{ Income}_i + \epsilon_i$$

 $Attitude_i = 2.235 - 0.000818 \text{ Income}_i + N(0, 2.378)$ 

### If Income is 6000:

$$Attitude_i = 2.235 - 0.000818 * 6000 + N(0, 2.378)$$

$$Attitude_i = -2.673 + N(0, 2.378)$$

$$Attitude_i = \alpha + \beta_1 \text{ Income}_i + \epsilon_i$$

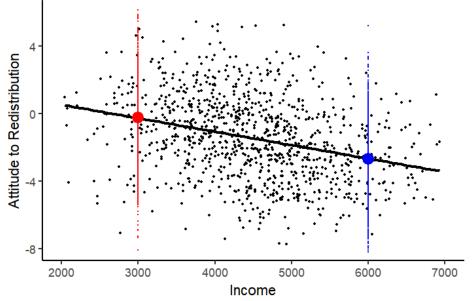
$$Attitude_i = 2.235 - 0.000818 \text{ Income}_i + N(0, 2.378)$$

# Increasing Income from 3000 to 6000:

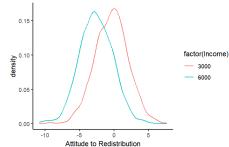
$$\Delta Attitude_i = (2.235-0.000818*6000)-(2.235-0.000818*3006)$$

$$\Delta Attitude_i = -2.673-(-0.219)$$

$$\Delta Attitude_i = -2.454$$



### **Predicted Values:**



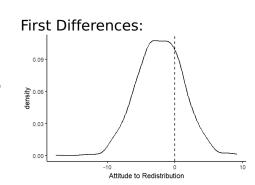
Guide to 'Smart' Regression

# 8. Predictions from Regressions

# 

Attitude to Redistribution

-10



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- ► Consider a binary outcome:  $Gender_i = \alpha + \beta Income_i + \epsilon_i$

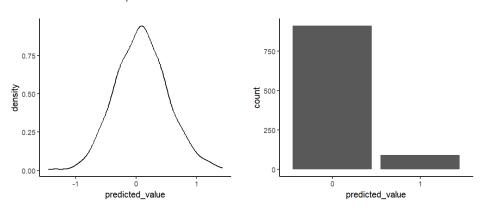
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- ► Compare the OLS and Logit regression tables:

	Dependent variable:
	gender
income	0.0003*** (0.00001)
Constant	-0.696*** (0.066)
Observations	1,000
Note:	*p<0.1: **p<0.05: ***p<0.01

	Dependent variable:
	gender
income	0.001*** (0.0001)
Constant	-6.360*** (0.457)
Observations	1,000
Note:	*p<0.1; **p<0.05; ***p<0.01

- ► The regression model matters because the wrong model makes non-sensical predictions
- ► Consider a binary outcome:  $Gender_i = \alpha + \beta Income_i + \epsilon_i$
- ► Compare the OLS and Logit **predictions** of gender for an income of R\$3000:



# Section 3

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  - 1. Explain anything
  - Make bad data better
  - 3. Tell you which theory is 'correct'
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- Even after following all this guidance, Regression does NOT:
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- These all require research design, theory and assumptions

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- ► More data will not help
- ► The problem is the content of data; it does not allow us to answer the causal question

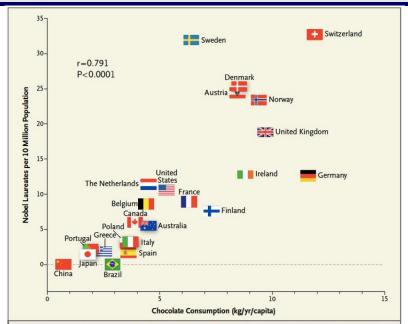
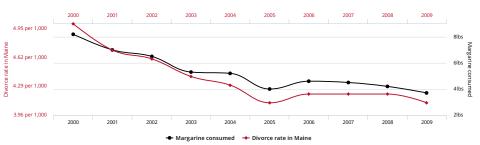


Figure 1. Correlation between Countries' Annual Per Capita Chocolate Consumption and the Number of Nobel Laureates per 10 Million Population.

#### Divorce rate in Maine

correlates with

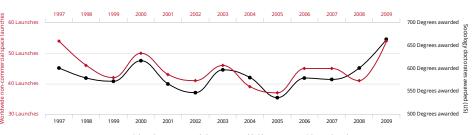
#### Per capita consumption of margarine



#### Worldwide non-commercial space launches

correlates with

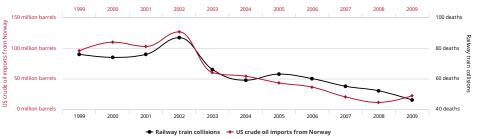
#### Sociology doctorates awarded (US)



#### US crude oil imports from Norway

correlates with

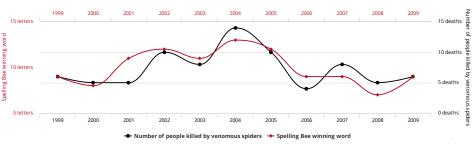
#### Drivers killed in collision with railway train



#### Letters in Winning Word of Scripps National Spelling Bee

correlates with

#### Number of people killed by venomous spiders



► Why is correlation (regression) not causation?

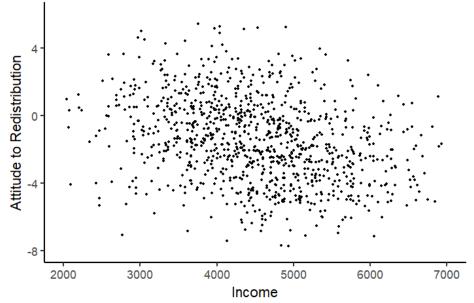
- ► Why is correlation (regression) not causation?
  - 1. Omitted Variable Bias

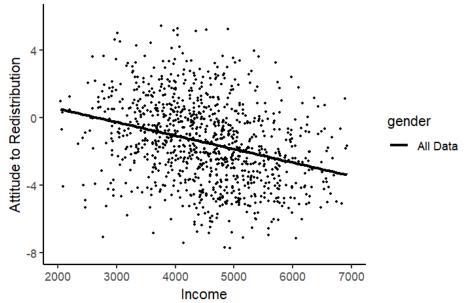
- Why is correlation (regression) not causation?
  - 1. Omitted Variable Bias
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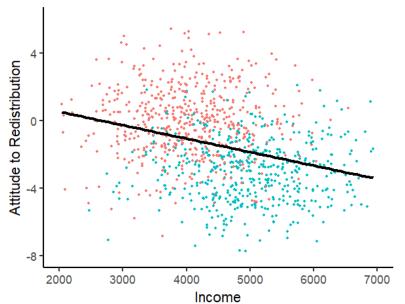
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  - 5. Lack of Overlap, Model Dependence



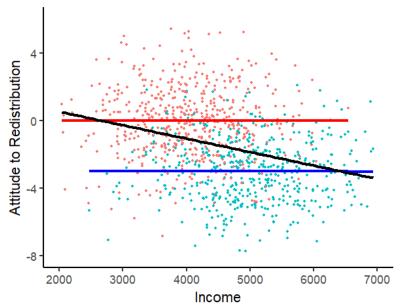




gender

0

1



0

1

#### 2. Reverse Causation

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	redist	
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gender1	-1.201*** (0.058)	
Constant	0.589*** (0.038)	
Observations	1,000	
Note:	*p<0.1; **p<0.05; ***p<0.01	

	Dependent variable:	
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- ► Remember, regression measures the *vertical* (not diagonal) distances to the regression line
  - It minimizes the prediction errors for Y
- But that doesn't mean it identifies the direction of causation!

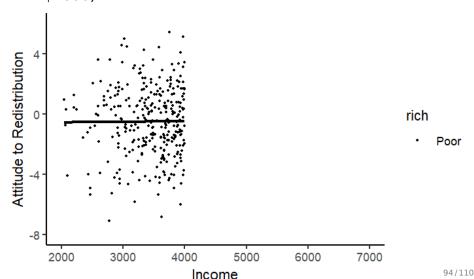
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- But negative attitudes to redistribution might also make you more likely to work in the private sector and cause you to receive a higher salary

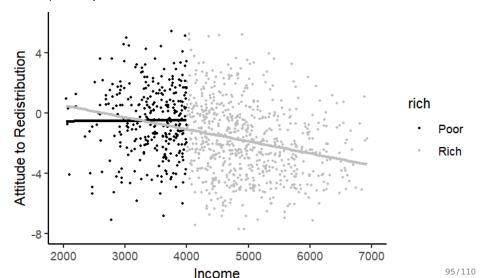
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- But negative attitudes to redistribution might also make you more likely to work in the private sector and cause you to receive a higher salary
- ▶ Both would look the same in a regression
- ► We cannot *explain* the relationship with a regression

 Imagine we do not see 'rich' units with high income (above R\$4000)



► Imagine we do not see 'rich' units with high income (above R\$4000)



- ► There are four selection risks:
  - 1. Selection into existence
  - 2. Selection into survival
  - 3. Selection into the dataset
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  - 2. Selection into survival
  - 3. Selection into the dataset
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- In each case, we don't see the full relationship between D and Y
- ► So our regression estimates are biased

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  - 1. Selection into existence:
    - Where do units (eg. political parties) come from?
    - Probably only parties that have a chance of success are formed
    - Does forming a party cause electoral success? Not for most people!

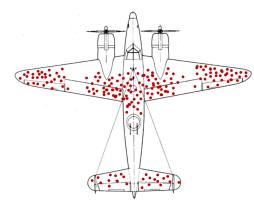
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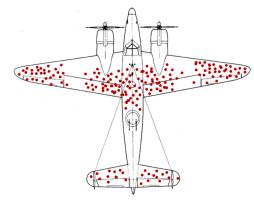


► Where would additional armour protect bombers?

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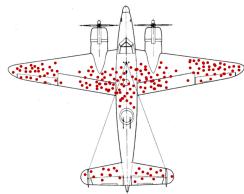


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#### 2. Selection into survival:

 Certain types of units disappear, so the units we see don't tell the full story



- Where would additional armour protect bombers?
- ► Returned bombers got hit
- ▶ But we do not know where bombers that did not return got hit

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    - ► Only units with particular values of *D* and *Y* enter the dataset
    - Eg. If survey respondents who refuse are different from those who respond - the anti-redistribution poor may dislike answering surveys

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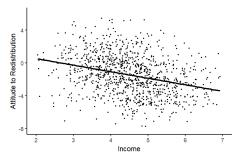
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    - Applying treatment to the others would probably have a very different effect

► What happens if we measure our variables wrongly?

## Effects of Measurement Error

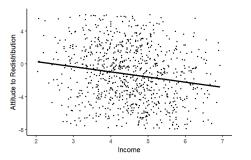
	Measured with <b>Bias</b>	Measured with <b>Random</b>
		Noise
Outcome Variable	Coefficient biased	No bias but wider stan- dard errors
Treatment Variable	Coefficient biased	Effect biased towards zero

- ► What happens if we measure our variables wrongly?
- ► No extra noise:



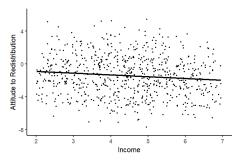
	Dependent variable:
	redist
income	-0.818*** (0.078)
Constant	2.235 * * * (0.361)
Observations	1,000
Note:	*p<0.1; **p<0.05; ***p<0.01

- ▶ What happens if we measure our variables wrongly?
- ► Noise in the **outcome variable**:



	Dependent variable:	
	redist	
income	-0.831*** (0.144)	
Constant	2.272*** (0.665)	
Observations	1,000	
Note:	*p<0.1; **p<0.05; ***p<0.01	

- ▶ What happens if we measure our variables wrongly?
- ► Noise in the **explanatory** variable:



	Dependent variable:
	redist
income	-0.187*** (0.037)
Constant	-0.620*** (0.183)
Observations	1,000
Note:	*p<0.1; **p<0.05; ***p<0.01

# 5. Lack of Overlap

► Regression normally helps us pick appropriate comparisons

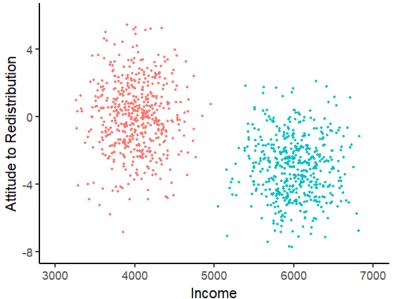
# 5. Lack of Overlap

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  - ► Eg. Controlling for gender, what is the effect of income on attitudes to redistribution?

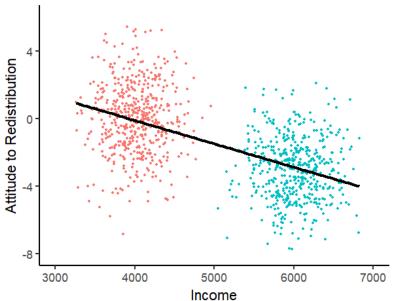
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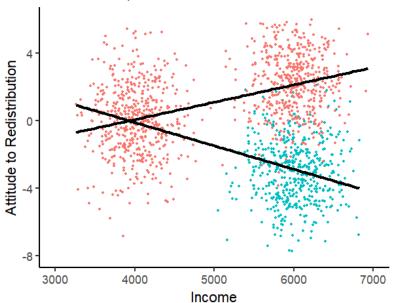
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- ► Regression *creates* comparisons for us
  - How? Using the functional form of the regression
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- Lack of overlap probably means we cannot explain outcomes with this data



gender





gender

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- A problem of multi-dimensionality
- And of model dependence our results depend on the functional form in our regression model

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  - We need to understand better how the data were produced
  - Explanation depends on research design, data selection, assumptions and qualitative evidence