FLS 6441 - Methods III: Explanation and Causation

Week 12 - Review & Frontiers

Jonathan Phillips

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Section 1

Review

Classification of Research Designs

- Correlation is not causation
 - And regresssion is just fancy correlation
- ► So how do we provide evidence of causation?

Classification of Research Designs

		Independence of Treatment Assignment	Researcher Controls Treatment Assignment?
Controlled Experiments	Field Experiments	√	✓
	Survey and Lab Experiments	√	√
Natural Experiments	Natural Experiments	√	
	Instrumental Variables	√	
	Discontinuities	√	
Observational Studies	Difference-in-Differences		
	Controlling for Confounding		
	Matching		
	Comparative Cases and Process Tracing		

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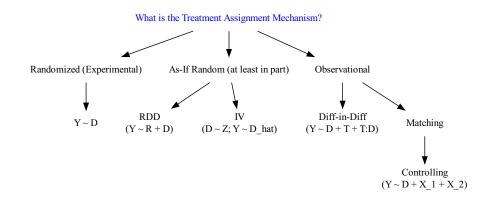
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- 12. Overlap in sample characteristics



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 - Parallel trends, no sorting, balance...

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- ► To help us interpret what we have learned

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- ► For Process Tracing: Causal Process Observations

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- Causal effects are probably highly heteregeneous who cares about the average?
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- ► Even if variable X has a causal effect, *how much* of the real world does it explain?
- Sometimes it's just not possible to show causation. That's OK!
 - We just need to recognize the evidence we have is not representative of everything that happens in the real world

Section 2

Frontiers

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- We also need to show that our estimate is reliable and not a 'chance' finding
- More importantly, that it is evidence in support of specific theory
- You don't want to publish a paper that someone contradicts next week!

- ► In general, we will trust our estimate more if it doesn't change even when we change our model
 - ► Not just direction and significance, but in the substantive effect size
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- Various formal tests, but best to plot overlap of confidence intervals from many models

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 - ightharpoonup Take a larger set of covariates, run your regression and store ho_F
 - ► Calculate $\frac{\beta_F}{\beta_R beta_F}$
- ► Eg. Nunn and Wantchekon (2011) argue that for unmeasured confounders to explain their estimated effect of the slave trade on trust, they would have to be 3 11 times larger than measured confounders

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- CRUCIAL: Our covariate is not randomly assigned, so the interpretation of causal effects is not causal, just descriptive

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- Note this does not mean that being a first-term mayor causes audits to be less effective

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- ► More details on this egap page

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- ► Common for regression discontinuities (alternative thresholds) and difference-in-differences (alternative times of treatment)

Table 2: The LPT effect on the PT electoral support in presidential elections (2002-2018)

	PT (2002)	PT (2006)	PT (2010)	PT (2014)	PT (2018)
		distribute		and the state of t	
LATE	-2.62	6.90***	4.87**	5.97***	5.59**
	(2.12)	(2.68)	(2.32)	(2.46)	(2.62)
BW est (h)	5.28	4.50	5.00	4.31	4.39
BW bias (b)	8.27	7.88	8.24	7.32	7.11
N Left	1711	1711	1711	1711	1711
N Right	3851	3851	3851	3851	3851
Eff N Left	351	303	334	289	295
Eff N Right	491	412	462	389	399
N clusters Left	523	506	521	478	466
N clusters Right	879	826	871	737	697

Note: * p < 0.1, ** p < 0.05, *** p < 0.01. RD local linear estimates using Calonico et al. (2014b) optimal bandwidth triangular kernel selection. Robust standard errors, clustered at the municipal level, in parenthesis. Controls: the expectation of schooling years, and share of households with the mid-school degree. N Left and N Right represent the total number of observation in the left and right sides of the cutoff. Eff N Left and Eff N Right are the number of cases within the bandwidth. BW est (h) is the Bandwidth used to compute the LATE (Local Average Treatment Effect). BW bias (b) is the Bandwidth used to compute the standard errors.

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$$\frac{Pr(D_i = 1 \& Z_i = 1 | X_i = 1)}{Pr(D_i = 1 \& Z_i = 1)}$$

Table 4.4.3
Complier characteristics ratios for twins and sex composition instruments

		Twins at Second Birth		First Two Children Are Same Sex		
Variable	$P[x_{1i} = 1] $ (1)	$P[x_{1i} = 1 D_{1i} > D_{0i}]$ (2)	$P[x_{1i} = 1 D_{1i} > D_{0i}] / P[x_{1i} = 1] $ (3)	$P[x_{1i} = 1 D_{1i} > D_{0i}]$ (4)	$P[x_{1i} = 1 D_{1i} > D_{0i}] / P[x_{1i} = 1] $ (5)	
Age 30 or older at first birth	.0029	.004	1.39	.0023	.995	
Black or hispanic	.125	.103	.822	.102	.814	
High school graduate	.822	.861	1.048	.815	.998	
College graduate	.132	.151	1.14	.0904	.704	

Notes: The table reports an analysis of complier characteristics for twins and sex composition instruments. The ratios in columns 3 and 5 give the relative likelihood that compliers have the characteristic indicated at left. Data are from the 1980 census 5 percent sample, including married mothers aged 21–35 with at least two children, as in Angrist and Evans (1998). The sample size is 254,654 for all columns.

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- But require additional assumptions: Sequential ignorability
 - ► That the mediator (mechanism) is independent of potential outcomes conditional on treatment
 - ► Hard!

► One practical approach is to run two regressions that recreates our DAG:

$$M_i = \alpha_1 + \beta_1 D_i + \epsilon_1$$

$$Y_i = \alpha_3 + \beta_3 D_i + \beta_4 M_i + \epsilon_3$$

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- ▶ Indirect effect of treatment = $\beta_4 * \beta_1$

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