Chapter 0: Review

Chapter 2: Simple Linear Regression

$$\begin{split} \mathbf{E}[\mathbf{y}|\mathbf{x}] &= \mu_{y|x} = & \mathbf{E}[\beta_0 + \beta_1 x + \epsilon] = \beta_0 + \beta_1 x & \mathbf{V}[\mathbf{y}|\mathbf{x}] = \sigma_{y|x}^2 = & \mathbf{V}[\beta_0 + \beta_1 x + \epsilon] = \sigma^2 & \hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x} & \hat{\beta}_1 = \frac{S_{xy}}{S_{xx}} = \frac{\sum_{i=1}^n (x_i - \bar{x})y_i}{\sum_{i=1}^n (x_i - \bar{x})^2} \\ \mathbf{E}[\hat{\beta}_1] &= \sum_{i=1}^n c_i \mathbf{E}[y_i] = \beta_0 \sum_{i=1}^n c_i + \beta_0 \sum_{i=1}^n c_i x_i = \beta_1 & \mathbf{V}[\hat{\beta}_1] = \sum_{i=1}^n c_i^2 (\sigma^2) = \sigma^2 \frac{\sum_{i=1}^n (x_i - \bar{x})^2}{S_{xx}^2} = \frac{\sigma^2}{S_{xx}} \\ \mathbf{E}[\hat{\beta}_0] &= \beta_0 & \mathbf{V}[\hat{\beta}_0] = \sigma^2 (\frac{1}{n} + \frac{\bar{x}^2}{S_{xx}}) = V[\bar{y} - \beta_1 \bar{x})] = V[\bar{y}] + x^2 V[\hat{\beta}_1] - cov(\bar{y}, \hat{\beta}_1) & \mathbf{c}_i = \frac{x - \bar{x}}{S_{xx}} \\ \mathbf{SS}_{res} &= \sum_{i=1}^n (y_i - \hat{y}_i)^2 = \sum_{i=1}^n \epsilon_i^2 & \mathbf{SS}_{T} = \sum_{i=1}^n y_i^2 - n\bar{y}^2, n - 1 \text{ df} & \mathbf{SS}_{Reg} = \hat{\beta}_1 S_{xy}, \text{ if df} = 1, \text{ then } = MS_{Res} \\ \mathbf{MS}_{res} &= \sigma^2 = \frac{SS_{res}}{n-2} \end{split}$$

Hypothesis Testing (Regression)

$$\begin{aligned} & \textbf{Reject } \textbf{H}_0 \textbf{ if } |\textbf{t}_0| \geq \textbf{t}_{\frac{\alpha}{2},n-2} \textbf{ where } t_0 = \frac{\hat{\beta}_1 - \beta_{10}}{\sqrt{\frac{MS_{res}}{S_{xx}}}} \end{aligned} & \textbf{Failing to reject } \textbf{H}_0 \textbf{: } \beta_i = 0 \textbf{ implies no rlshp between } \textbf{x} \textbf{ and } \textbf{y}. & \textbf{E}[y_i] = \beta_1 \textbf{x} + \beta_0 \end{aligned} \\ & F_0 = \frac{MS_{Reg}}{MS_{res}} = t_0^2 & \textbf{Reject if } F_0 > F_{\alpha}, \textbf{1}, n-1 & \textbf{CI: } \hat{\beta}_1 - \textbf{t}_{\frac{\alpha}{2},n-2} se(\hat{\beta}_{10}) < \hat{\beta}_{10} < \hat{\beta}_1 + \textbf{t}_{\frac{\alpha}{2},n-2} se(\hat{\beta}_{10}) & \textbf{se}(\hat{\beta}_1) = \sqrt{\frac{MS_{res}}{S_{xx}}}, \textbf{se}(\hat{\beta}_0) = \sqrt{V[\hat{\beta}_0]} \end{aligned} \\ & \textbf{R}^2 = \textbf{1} - \frac{SS_{res}}{SS_T} = \frac{SS_{Reg}}{SS_T} & \textbf{R}^2_{adj} = \textbf{1} - \frac{SS_{res}(n-1)}{SS_T(n-k-1)} \end{aligned}$$

Chapter 3: Multiple Linear Regression

 $y = \underset{n \times 1}{x} \times \underset{p \times 1}{\beta} + \underset{n \times 1}{\epsilon} \text{ where p } = k+1, \text{ p is the total number of betas (or parameters), k is the number of regressor variables.}$ $\epsilon \sim N(0, \sigma^2 I) \text{ where I is the identity matrix whatever size } E[y] = x\beta \qquad V[y] = V[\epsilon] = \sigma^2 I \qquad y \sim N(x\beta, \sigma^2 I)$

Least Square Estimate for β and σ^2

$$S(\beta) = \sum_{i=1}^{n} \epsilon^2 = \epsilon' \epsilon = (y - x\beta)'(y - x\beta) = y'y - 2\beta'x'y + \beta'x'x\beta$$

$$\hat{\beta} = (x'x)^{-1}x'y \qquad E[\hat{\beta}] = E[(x'x)^{-1}x'y] = E[(x'x)^{-1}x'y] = E[(x'x)^{-1}x'(x\beta + \epsilon)] = \beta \qquad V[\hat{\beta}] = (x'x)^{-1}\sigma^2 = c\sigma^2 \qquad V[\hat{\beta}_j] = c_{jj}\sigma^2 \qquad E[\beta_j] = \beta_j$$

$$\hat{\beta}_j \sim N(\beta_j, c_{jj}\sigma^2) \ \hat{y} = x\hat{\beta} = x(x'x)^{-1}x'y = Hy \qquad E[\hat{y}] = E[x\hat{\beta}] = x\beta \qquad V[\hat{y}] = V[x\hat{\beta}] = xV[\hat{\beta}]x' = x(x'x)^{-1}x'\sigma^2 = H\sigma^2$$

$$\hat{y} \sim N(x\beta, H\sigma^2) \qquad \hat{y}_j \sim N(x_j\beta, h_{jj}\sigma^2), \text{ where } h_{jj} = x'_j(x'x)^{-1}x_j \qquad x_j = [x_{j0}, x_{j1}, \dots x_{jk}] \text{ and} \qquad \hat{\epsilon} = y - \hat{y} = y - Hy = (I - H)y$$

$$\hat{\sigma}^2(estimator) = \frac{SS_{res}}{n-p} = MS_{res} \text{ where } p = k + 1 = \text{the number of parameters (i.e. } \beta \text{ 's: } \beta_0, \beta_1, \dots \beta_k)$$

$$SS_{res}(\mathbf{n} - \mathbf{p}) = (y - x\hat{\beta})'(y - x\hat{\beta}) = y'y - 2\hat{\beta}'X'y + \hat{\beta}'x'x\hat{\beta} = y'y - \hat{\beta}'x'y \qquad SS_{Reg}(\mathbf{k}) = \hat{\beta}'x'y - \frac{(\sum_{i=1}^n y_i)^2}{n} \qquad SS_T(\mathbf{n} - \mathbf{1}) = y'y - \frac{(\sum_{i=1}^n y_i)^2}{n}$$

$$MS_{res} = \frac{SS_{res}}{n-k-1} \qquad MS_{Reg} = \frac{SS_{Reg}}{k} \qquad MS_T = \frac{SS_{Reg}}{SS_{Res}} = \frac{MS_{Reg}}{mS_{res}}$$

$$\text{error } = (\mathbf{I} - \mathbf{H})y = (\mathbf{I} - \mathbf{H})\epsilon \qquad E[\mathbf{M}S_{Res}] = \sigma^2 \qquad E[\mathbf{M}S_{Reg}] = \sigma^2 + \frac{\beta'}{k\sigma^2} \frac{x'_c x_c \beta^*}{k\sigma^2} \text{ where } \beta^* = (\beta_1, \beta_2, \dots \beta_k) \text{ and } x_c \text{ is the center}$$

$$\text{We reject } H_0 \text{ if } F_0 > F_{\alpha,k,n-k-1}$$

$$\text{Testing Individual Coefficients: If } H_0: \beta_j = 0 \text{ is not rejected then delete it: } t_0 = \frac{\hat{\beta}_j}{\sqrt{n^2 c_{ij}}} = \frac{\hat{\beta}_j}{se(\hat{\beta}_i)} \qquad \text{reject if } |t_0| > t_{\frac{\alpha}{2},n-k-1}$$

Confidence Intervals

 $\sigma^2\mathbf{known}:\ \hat{\beta}_j\sim N(\beta_j,c_{jj}\sigma^2)\longrightarrow \frac{\hat{\beta}_j-\beta_j}{\sqrt{c_{jj}\sigma^2}}\sim N(0,1)\ \text{ or, if variance is unknown},\ \hat{\beta}_j\sim N(\beta_j,c_{jj}MS_{res})\longrightarrow \frac{\hat{\beta}_j-\beta_j}{\sqrt{c_{jj}MS_{res}}} \text{ or } \frac{\hat{\beta}_j-\beta_j}{se(\hat{\beta}_j)}\sim t_{n-p}$ Then the variance estimator is $\hat{\sigma}^2=MS_{res}=\frac{SS_{res}}{n-p}\sim\chi^2_{n-p}$ So, the (1- α) **confidence interval** for β_j is $\hat{\beta}_j\pm t_{\frac{\alpha}{2},n-p}se(\hat{\beta}_j)$ $\hat{y}_j\sim N(x_j\beta_j,h_{jj}\sigma^2)$, so $\frac{\hat{y}_j-x_j\beta}{\sqrt{h_{jj}MS_{res}}}\sim t_{n-p} \text{ where } \sigma^2\text{ estimates } MS_{res}$ A 1- α confidence interval for $E[y_0|x_0]$ is $\hat{y}_0\pm t_{\frac{\alpha}{2},n-p}\sqrt{x'_0(x'x)^{-1}x_0\hat{\sigma}^2}$ or $\hat{y}_0\pm t_{\frac{\alpha}{2},n-p}\sqrt{x'_0(x'x)^{-1}x_0MS_{res}}$

Chapter 4: Model Testing

Properties of residuals: mean 0, $MS_{res} = \sum_{i=1}^{n} \frac{(\epsilon_i - \bar{\epsilon})^2}{n-p} = \sum_{i=1}^{n} \frac{\epsilon_i^2}{n-p} = \frac{SS_{res}}{n-p}$ Scaling Residuals: Standardized Residuals: $d_i = \frac{\epsilon_i}{\sqrt{MS_{res}}}$ Studentized: $r_i = \frac{\epsilon_i}{\sqrt{MS_{res}}(1-h_{ii})}$, $\mathbf{V}[\epsilon_i] = \sigma^2(1-h_{ii})$, $\mathbf{cov}(\epsilon_i, \epsilon_j) = -\sigma^2 h_{ij}$ Other model testing: plot \mathbf{x}_i and \mathbf{x}_j : linear rln means high corr.

Formal test for lack of fit: Assuming everything is tested and ideal, to test for linearity, we use: $SS_{res} = SS_{PE} + SS_{LOF}$ Formal $\mathbf{E}[MS_{LOF}] = \mathbf{E}[MS_{PE}] = \sigma^2$, where m is num regressors, n is num samples $\mathbf{V}[\bar{y}] = \frac{p\sigma^2}{n}$ (indpure e)