## Lecture 14

## § 1 Non-homogeneous Poisson process

标准 Poisson process (A) 具有无记忆性,且 interarrival times 为 i.i.d. Exp(A).

由于这是一个强 assumption,我们需要一个不具有无记忆性的 more general/flexible model

- 1. Definition: Non-homogeneous Poisson process (非市农海松过程)
  - 一个 counting process  $\{Nt: t \ge 0\}$  被称为 a non-homogeneous Poisson process with a continuous intensity function  $\lambda(t): t \ge 0$ ,若其满足
    - 0 N10) = 0
    - ② {Nt.t≥D多有 independent increments
    - ③ 对于足够小的 positive h≥D,有
      - $P \mid N_{t+h} N_t \ge 2$  = O(h)
      - PINtih-Nt=1 =  $\lambda(t)h + o(h)$
      - · (Pf Nt+h-Nt=0) = 1- \(\lambda(t)\) h + O(h) ) (若前两条满足,则自然满足)

注: 类比 Poisson process, 有 N++n-Nt ~ Ber (λ+H·h), 唯一的区划在于用λ+t/) 替换λ

21 Lemma: Poisson approximation (non-homogeneous version)

 $P_1 + \cdots + P_n \rightarrow \lambda > 0$  (pn  $\rightarrow \lambda > 0$ ), when  $n \rightarrow \infty$ 

 $\mathbb{R} \setminus S_n = X_1 + \cdots + X_n$  converges to  $Poi(\lambda)$ , when  $n \to \infty$ 

3. Theorem: increments \$5 distribution

定义  $m_t > \int_0^t \lambda(u) du$ , for  $t \ge 0$ .

- 对于一个 non-homogeneous Poisson process with intensity function  $\lambda(t)$ , 对于  $0 \le s < t$ ,有  $N_t N_s \sim Poi(m(t) m(s)) = Poi(\int_s^t \lambda(u) du)$
- 注: ① The increment distribution 不仅取决于 length t-s,还取决于起始时间s和结束时间t,因此 increment 不为 stationary
  - D interarrival time 不服从指数分布
  - 图 若 S=0,则有 Nt~ Poi (m(t)),此时 m(t)被称为 the mean value function of the non-homogeneous Poisson process

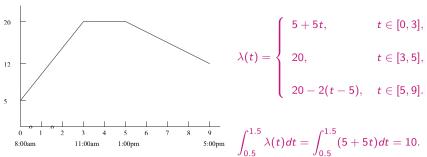
## 证明

$$N(t) - N(0) = \sum_{i=1}^{n} (N(t_i) - N(t_{i-1}))$$

Ber (
$$\lambda(t_{i-1})$$
 · ( $t_{i-t_{i-1}}$ )   
 $\lambda(t_{i-1})$  · ( $t_{i-t_{i-1}}$ ) →  $\lambda(u)$  du as  $n \to \infty$ 

因此, 
$$N(t) - N(0) \sim Poi(\int_0^t \lambda(u) du)$$
  
 $N(t) - N(s) \sim Poi(\int_s^t \lambda(u) du)$ 

- ► From 8 until 11 A.M. customers seem to arrive, on the average, at a steadily increasing rate that starts with an initial rate of 5 customers per hour at 8 A.M. and reaches a maximum of 20 customers per hour at 11 A.M.
- ► From 11 A.M. until 1 P.M. the (average) rate seems to remain constant at 20 customers per hour.
- However, the (average) arrival rate then drops steadily from 1 P.M. until closing time at 5 P.M. at which time it has the value of 12 customers per hour.



- ▶ If we assume that the numbers of customers arriving at Siegbert's stand during disjoint time periods are independent, then what is a good probability model for the preceding?
- ▶ What is the probability that no customers arrive between 8:30 A.M. and 9:30 A.M. on Monday morning?
- ► What is the expected number of arrivals in this period?

$$\begin{array}{llll}
\text{D} & \int_{0.5}^{15} \lambda(t) \, dt = \int_{0.5}^{15} (5+5t) \, dt = 5t + \frac{5}{2} t^2 \Big|_{0.5}^{1.5} = 10 \\
N_{1.5} - N_{0.5} \sim P_{01} (10) \\
P(N_{1.5} - N_{0.5} = 0) = \frac{10^{0} e^{-10}}{0!} = e^{-10}
\end{array}$$

13/2

6. The capital of a bank grows proportionally with time t: at time t, the bank holds ct units of capital, where c is a positive constant. The bank has to undergo stress tests, which occur according to a Poisson process with rate λ. The bank passes a stress test if it holds at least a certain amount of capital at the moment of the test. The required amounts of capitals are independently distributed as Pareto distribution, with probability density function given by:

$$f(y) = \frac{24}{(2+y)^4}, \quad y > 0.$$

- (a) Suppose that a stress test occurs at time s. Determine the probability that the bank fails this stress test.

  [4 marks]

  [8 (2+c5)<sup>-3</sup>
- (b) Let N(t) denote the number of stress tests the bank fails before time t. Find the distribution of N(t). [7 marks]
   κ(t) ~ for ζ λ Cl 4/(2 + c t)²
   (c) Determine the probability that the bank can pass all the stress tests which have ever
- (c) Determine the probability that the bank can pass all the stress tests which have ever occurred.

  [4 marks]

  [7 1/2]

  [7 1/2]

  [7 1/2]

  [7 1/2]

(a) 
$$P(s) = P(Y > cs) = \int_{cs}^{\infty} 24(2+y)^{-4} dy = \frac{8}{(2+cs)^3}$$

- (b) DN(0) > 0
  - ② independent increment 显然满足
  - B 全 N(t) 表示 stress test 的 出现数,则 N(t+h)-N(t) ≤ N(t+h)-N(t)
    - ·  $P(N(t+h)-N(t)\geq 2) \leq P(\bar{N}(t+h)-\bar{N}(t)\geq 2) = o(h)$

$$P(N(t+h)-N(t)=1) = P(\overline{N}(t+h)-\overline{N}(t)=1, fail)$$

$$= P(\overline{N}(t+h)-\overline{N}(t)=1) \cdot P(fail | \overline{N}(t+h)-\overline{N}(t)=1)$$

$$= (\lambda h + o(h)) \cdot P(t)$$

$$= \lambda P(t) h + o(h)$$

因此 1Nt3为一个 non-homogeneous Poisson process with intensity function 入(七)= 2p(七)

因此 Nt 服从铂松马布, parameter 为

$$m_t = \int_0^t \lambda(s) \, ds = \int_0^t \lambda \cdot \frac{8}{(z+cs)^3} \, ds = -\frac{4\lambda}{c(z+cs)^2} \Big|_0^t = \frac{\lambda}{c} \Big\{ 1 - \frac{4}{(z+ct)^2} \Big\}$$

₽P

$$N_{t} \sim P_{0i} \left( \frac{\lambda}{c} \left\{ 1 - \frac{4}{(z+ct)^{2}} \right\} \right)$$
(c)  $P(N_{t}=0) = e^{-m(t)}$ 

$$\lim_{t\to\infty} P(N_{t}=0) = \lim_{t\to\infty} \exp\{-\frac{\lambda}{c}\{1 - \frac{4}{(z+ct)^{2}}\}\} = e^{-\frac{\lambda}{c}}$$