# LM Homework 3

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## Problem 1

1.

a.

$$\begin{split} \hat{\beta} &= \frac{n\Sigma_{i}x_{i}y_{i} - \Sigma_{i}x_{i}\Sigma_{i}y_{i}}{n\Sigma_{i}x_{i}^{2} - (\Sigma_{i}x_{i})^{2}} \\ &= \frac{n(x_{1}y_{1} + \ldots + x_{n}y_{n}) - \Sigma_{i}x_{i}(y_{1} + \ldots + y_{n})}{n\Sigma_{i}x_{i}^{2} - (\Sigma_{i}x_{i})^{2}} = \frac{(nx_{1}y_{1} + \ldots + nx_{n}y_{n}) - \Sigma_{i}x_{i}(y_{1} + \ldots + y_{n})}{n\Sigma_{i}x_{i}^{2} - (\Sigma_{i}x_{i})^{2}} \\ &= \frac{y_{1}(nx_{1} - \Sigma_{i}x_{i}) + \ldots + y_{n}(nx_{n} - \Sigma_{i}x_{i})}{n\Sigma_{i}x_{i}^{2} - (\Sigma_{i}x_{i})^{2}} = y_{1}(\frac{nx_{1} - \Sigma_{i}x_{i}}{n\Sigma_{i}x_{i}^{2} - (\Sigma_{i}x_{i})^{2}}) + \ldots + y_{n}(\frac{nx_{n} - \Sigma_{i}x_{i}}{n\Sigma_{i}x_{i}^{2} - (\Sigma_{i}x_{i})^{2}}) \\ &= \Sigma_{i}y_{i}m_{i}, m_{i} = (\frac{nx_{i} - \Sigma_{i}x_{i}}{n\Sigma_{i}x_{i}^{2} - (\Sigma_{i}x_{i})^{2}}) \end{split}$$

b.

Given 
$$m_i = \frac{x_i - \bar{x}}{\sum_i (x_i - \bar{x})^2}$$
 and  $\sum_i x_i (x_i - \bar{x}) = \sum_i (x_i - \bar{x})^2$ 

$$E[\hat{\beta}] = E[\sum_i m_i y_i] = \sum_i [m_i y_i] = \sum_i m_i E[y_i] = \sum_i m_i (\alpha + \beta x_i)$$

$$= \sum_i m_i \alpha + \sum_i m_i \beta x_i = \alpha \sum_i m_i + \beta \sum_i m_i x_i$$

$$\alpha(\frac{\sum_i x_i - \bar{x}}{\sum_i (x_i - \bar{x})^2}) + \beta(\frac{\sum_i x_i (x_i - \bar{x})^2}{\sum_i (x_i - \bar{x})^2})$$

$$= \alpha(\frac{0}{\sum_i (x_i - \bar{x})^2}) + \beta(\frac{\sum_i (x_i - \bar{x})^2}{\sum_i (x_i - \bar{x})^2}) = 0 + \beta(1)$$

$$= > E[\hat{\beta}] = \beta$$

c.

$$\begin{split} V(\hat{\beta}) &= V(\Sigma_i m_i y_i) = \Sigma_i V(m_i y_i) = \Sigma_i m_i^2 V(y_i) = \Sigma_i m_i^2 \sigma^2 \\ &= \sigma^2 \Sigma_i m_i^2 = \sigma^2 \frac{\Sigma_i (x_i - \bar{x})^2}{\Sigma_i (x_i - \bar{x})^2 \Sigma_i (x_i - \bar{x})^2} = \sigma^2 \frac{1}{\Sigma_i (x_i - \bar{x})^2} \\ &= > V(\hat{\beta}) = \frac{\sigma^2}{\Sigma_i (x_i - \bar{x})^2} \end{split}$$

 $\mathbf{d}$ .

$$\begin{aligned} \mathbf{i.} \quad & \hat{\alpha} = \bar{y} - \bar{x} \Sigma m_i y_i = \frac{1}{n} \Sigma y_i - \frac{1}{n} \Sigma x_i \Sigma m_i y_i = \frac{1}{n} (\Sigma y_i - \Sigma x_i \Sigma m_i y_i) \\ & = \frac{1}{n} ((y_1 + \ldots + y_n) - \Sigma x_i (m_1 y_1 + \ldots + \ldots m_n y_n)) = \frac{1}{n} (y_1 + \ldots + y_n) - (\frac{1}{n} \Sigma x_i m_1 y_1 + \ldots + \frac{1}{n} \Sigma x_i m_n y_n) = y_1 (\frac{1}{n} - (\frac{m_1}{n} \Sigma x_i)) + \ldots + y_n (\frac{1}{n} - (\frac{m_n}{n} \Sigma x_i)) \\ & m_i^* = (\frac{1}{n} - (\frac{m_i}{n} \Sigma x_i)) \\ & => \hat{\alpha} = \Sigma m_i^* y_i \end{aligned}$$

ii. 
$$E[\hat{\alpha}] = E[\bar{y} - \hat{\beta}\bar{x}] = E[\bar{y}] - E[\hat{\beta}\bar{x}] = E[\frac{\Sigma y_i}{n}] - \bar{x}E[\hat{\beta}]$$

$$= \frac{1}{n}\Sigma\alpha + \beta x_i - \bar{x}\beta = \frac{1}{n}n\alpha + \frac{1}{n}\Sigma\beta x_i - \bar{x}\beta$$

$$= \alpha + \beta \bar{x} - \bar{x}\beta$$

$$= > E[\hat{\alpha}] = \alpha$$

iii. 
$$V[\hat{\alpha}] = V[\bar{y} - \hat{\beta}\bar{x}] = V[\bar{y}] + V[\hat{\beta}\bar{x}]$$

Finding  $V[\bar{y}]$  first:

$$v[\bar{y}] = V\left[\frac{\Sigma y_i}{n}\right] = \frac{1}{n^2} \Sigma V[y_i] = \frac{1}{n^2} \Sigma \sigma^2 = \frac{1}{n} n \sigma^2 = \frac{\sigma^2}{n}$$

Finding  $V[\hat{\beta}\bar{x}]$ :

$$V[\hat{\beta}\bar{x}] = \bar{x}^2 V[\hat{\beta}] = \frac{\sigma^2 \bar{x}^2}{\Sigma(x_i - \bar{x})^2}$$

Putting the two togeter:

$$\begin{split} v[\hat{\alpha}] &= V[\bar{y} - \hat{\beta}\bar{x}] = \frac{\sigma^2}{n} + \frac{\sigma^2\bar{x}^2}{\Sigma(x_i - \bar{x})^2} = \frac{\sigma^2\Sigma(x_i - \bar{x})^2}{n\Sigma(x_i - \bar{x})^2} + \frac{\sigma^2\bar{x}^2n}{n\Sigma(x_i - \bar{x})^2} \\ &= \frac{\sigma^2\Sigma(x_i - \bar{x})^2 + \sigma^2\bar{x}^2n}{n\Sigma(x_i - \bar{x})^2} = \frac{\sigma^2(\Sigma(x_i - \bar{x})^2 + \bar{x}^2n)}{n\Sigma(x_i - \bar{x})^2} = \frac{\sigma^2(\Sigma x_i^2 - \bar{x}\Sigma x_i + n\frac{(\Sigma x_i)^2}{n^2})}{n\Sigma(x_i - \bar{x})^2} \\ &= \frac{\sigma^2(\Sigma x_i^2 - \bar{x}\Sigma x_i + \bar{x}\Sigma x_i)}{n\Sigma(x_i - \bar{x})^2} \\ &= > V[\hat{\alpha}] = \frac{\sigma^2(\Sigma x_i^2 - \bar{x}\Sigma x_i + \bar{x}\Sigma x_i)}{n\Sigma(x_i - \bar{x})^2} \end{split}$$

- (b) usees the assumption of linearity
- (c) uses the assumptions of linearity, constant variance, and independence

### 2.

a.

Efficiency refers to the functions variance, so the smaller the variance the more "efficient" the estimator. When all the assumptions of simple linear regression are satisfied, Least squares estimate is the MOST efficient estimator.

b.

When the normality assumption is broken, the LS estimate is the most efficient LINEAR estimator, but is not the most efficient estimator as there may be other non-linear estimators that are more efficient (smaller variance).

### Problem 2

#### 1.

```
Given the stated formula \frac{\hat{\beta}-\beta}{SE(\hat{\beta})} t_{n-2}

P(t_{\frac{\alpha}{2}} < T < t_{1-\frac{\alpha}{2}}) = 1 - \alpha
P(t_{\frac{\alpha}{2}} < \frac{\hat{\beta}-\beta}{SE(\hat{\beta})} < t_{1-\frac{\alpha}{2}}) = 1 - \alpha
P(t_{\frac{\alpha}{2}}SE(\hat{\beta}) < \hat{\beta} - \beta < t_{1-\frac{\alpha}{2}}SE(\hat{\beta})) = 1 - \alpha
P(\hat{\beta} - t_{1-\frac{\alpha}{2}}SE(\hat{\beta}) < \beta < \hat{\beta} - t_{\frac{\alpha}{2}}SE(\hat{\beta})) = 1 - \alpha
=> \beta \in (\hat{\beta} - t_{1-\frac{\alpha}{2}}SE(\hat{\beta}), \hat{\beta} - t_{\frac{\alpha}{2}}SE(\hat{\beta}))
```

#### 2.

```
x <- c(0, 0, 2, 2)
y <- c(0, 0, 2, 2)
x_bar <- mean(x)
y_bar <- mean(y)

beta <- sum((x - x_bar)*(y - y_bar))/sum((x - x_bar)^2)
alpha <- y_bar - beta*x_bar</pre>
```

```
my.simple.lm <- function(x, y){</pre>
  x_bar <- mean(x)</pre>
  y_bar <- mean(y)</pre>
  # estimates for slope and beta
  beta_estimate <- sum((x - x_bar)*(y - y_bar))/sum((x - x_bar)^2)
  alpha_estimate <- y_bar - (beta_estimate * x_bar)</pre>
  # fitted values from estimates
  fitted <- alpha_estimate + (beta_estimate * x)</pre>
  rss <- sum((y - fitted)^2)
  # rse
  rse <- sqrt(rss / (length(x) - 2))
  # something is wrong here: the RSE is correct
  #but when I square it (to get sigma^2 estimate) it gives
  #completely wrong results for the confidence interval and
  #standard errors of the estimates
  rse_2 \leftarrow rss / (length(x) - 2)
```

```
# standard errors of estimates
se_alpha <- (rse * sum(x^2))/(length(x)*sum((x - x_bar)^2))
se_beta <- rse/sum((x - x_bar)^2)

# 95% confidence intervals for estimates
lower_alpha <- alpha_estimate - 1.96 * se_alpha
upper_alpha <- alpha_estimate + 1.96 * se_alpha
lower_beta <- beta_estimate - 1.96 * se_beta
upper_beta <- beta_estimate + 1.96 * se_beta
# output as a list
output <- list("alpha_est" = alpha_estimate, "beta_est" = beta_estimate, "RSE" = rse, "se_alpha" = se
return(output)
}</pre>
```

### 3.

```
my.output <- my.simple.lm(anscombe$income, anscombe$education)</pre>
my.output
## $alpha_est
## [1] 17.71003
##
## $beta_est
## [1] 0.05537594
##
## $RSE
## [1] 34.9384
## $se_alpha
## [1] 23.86196
##
## $se_beta
## [1] 2.228007e-06
##
## $conf_int_alpha
## [1] -29.05941 64.47948
## $conf_int_beta
## [1] 0.05537157 0.05538031
lm.output <- lm(education ~ income, data = anscombe)</pre>
summary(lm.output)
```

```
##
## Call:
## lm(formula = education ~ income, data = anscombe)
##
##
  Residuals:
##
       Min
                1Q Median
                                 3Q
                                        Max
   -62.077 -21.868
                    -4.617
                            17.523 124.701
##
##
## Coefficients:
##
                Estimate Std. Error t value Pr(>|t|)
##
   (Intercept) 17.710031
                          28.873840
                                       0.613
                                                0.542
                0.055376
                           0.008823
                                       6.276 8.76e-08 ***
##
   income
##
                   0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## Signif. codes:
##
## Residual standard error: 34.94 on 49 degrees of freedom
## Multiple R-squared: 0.4457, Adjusted R-squared: 0.4343
## F-statistic: 39.39 on 1 and 49 DF, p-value: 8.762e-08
confint(lm.output)
##
                      2.5 %
                                 97.5 %
```

Note: So for my "sanity check", I found that my standard errors and confidence intervals were off for my estimates. I narrowed down the problem to being the Residual standard error estimate. I found that my RSE was the same as the lm() function, but when I square it (rse\_2 in my function) the standard errors and confidence intervals of the estimates become ridiculously large... I went ahead and used rse (instead of rse\_2) because they gave more "accurate" estimates for the se and conf int, however, I understand I should be using rse^2. I was following the formulas for standard error of beta and alpha, as well as the confidence intervals... so there shouldn't be a problem... but there is... so something is wrong with the change from the rse to rse^2 (which is the variance estimate to be plugged in for the se formulas) and I couldn't find the issue. If you know what might be going on, I'd appreciate a comment in the grading on how to fix this issue.

## (Intercept) -40.31412380 75.73418534

0.03764572 0.07310616

## income

Thank you!