

§ 2 方差和标准差

Chapter 4: Expected Values (期望值)

- ➤ The Expected Value of a Random Variable(随机变量的期望)
- ➤ Variance and Standard Deviation(方差和标准差)
- ➤ Covariance and Correlation Coefficient (协方差和相关系数)
- ➤ Conditional Expectation (条件期望)



Chapter 4: Expected Values (期望值)

- ➤ The Expected Value of a Random Variable(随机变量的期望)
- ➤ Variance and Standard Deviation(方差和标准差)
- ➤ Covariance and Correlation Coefficient (协分差和相关系数)
- ➤ Conditional Expectation (条件期望)



Background

Example: there are two watches with daily accuracies shown as follows:

Daily Accuracy (s)	-3	-2	-1	0	1	2	3
Probability (Watch A)	0.10	0.15	0.15	0.20	0.15	0.15	0.10
Probability (Watch B)	0.05	0.05	0.10	0.60	0.10	0.05	0.05

Which watch has better quality?

Answer: Assume that the accuracy of these two watches are X and Y, then

$$E(X) = \sum_{k=1}^{7} x_k P\{X = x_k\} = 0$$
$$E(Y) = \sum_{k=1}^{7} y_k P\{Y = y_k\} = 0$$

Therefore, the average accuracies from the two watches are the same.

Are they of the same quality?



Thinking based on the deviation (偏差) from the average

The deviation(偏差) of r.v. X is

If the deviation is small, the quality is stable

$$|X - E(X)|$$

Calculation based on absolute value is not convenient

Consider the squared deviation (平方偏差)

$$[X - E(X)]^2$$

Squared deviation is still a r.v.

Consider the average of squared deviation

$$E(X-E(X))^2$$

Definition: For r.v. X, if the following exists

$$Var(X) \triangleq D(X) \triangleq E(X - E(X))^{2}$$

Then D(X) is the Variance (方差) of r.v. X, $\sqrt{D(X)}$ is the Standard Deviation (标准差)



Example: Assume that the scores of shooters A and B are X and Y respectively. The frequency functions are:

x_k	8	9	10
p_k	0.15	0.40	0.45

y_k	8	9	10
p_k	0.35	0.10	0.55

Please evaluate these shooters.

Answer: First find out the expected values:

$$E(X) = 8 \times 0.15 + 9 \times 0.40 + 10 \times 0.45 = 9.3$$

$$E(Y) = 8 \times 0.35 + 9 \times 0.10 + 10 \times 0.55 = 9.2$$

And then find out the variance

$$Var(X) \triangleq D(X) \triangleq E(X - E(X))^{2}$$

$$D(X) = (8 - 9.3)^2 \times 0.15 + (9 - 9.3)^2 \times 0.40 + (10 - 9.3)^2 \times 0.45 = 0.51$$

$$D(Y) = (8 - 9.2)^2 \times 0.35 + (9 - 9.2)^2 \times 0.10 + (10 - 9.2)^2 \times 0.55 = 0.86$$

Based on the results, shooter A is not only better than shooter B in terms of shooting level, but also the score is more stable.



Example: there are two watches with daily accuracies shown as follows:

Daily Accuracy (s)	-3	-2	-1	0	1	2	3
Probability (Watch A)	0.10	0.15	0.15	0.20	0.15	0.15	0.10
Probability (Watch B)	0.05	0.05	0.10	0.60	0.10	0.05	0.05

Which watch has better quality?

Answer: Assume that the accuracies of the two watches A, B are X and Y, respectively, then

$$E(X) = \sum_{k=1}^{7} x_k P\{X = x_k\} = 0$$

$$E(Y) = \sum_{k=1}^{7} y_k P\{Y = y_k\} = 0$$

$$D(X) = E(X - E(X))^2 = E(X^2) = 3.3$$

$$D(Y) = E(Y - E(Y))^{2} = E(Y^{2}) = 1.5$$

As above, the average accuracies are the same, but their variances are different. So, Watch B is better than Watch A because it is more stable.





Explanation: Expected value — The average value of a r.v.

Variance — The average deviation between a r.v. and its average.



The calculation of variance

$$D(X) = E(X - E(X))^{2}$$

is the expected value of $g(X) = (X - E(X))^2$, then

 \bigcirc Assume that the <u>frequency</u> function of X is

$$P\{X = x_i\} = p_k, k = 1, 2, \dots$$

then
$$D(X) = \sum_{k=1}^{\infty} (x_k - E(X))^2 \cdot p_k$$

Assume that the probability density of X is f(x), then

$$D(X) = \int_{-\infty}^{\infty} (x - E(X))^{2} f(x) dx$$



$$D(X) = E(X - E(X))^{2} = E(X^{2}) - [E(X)]^{2}$$



Example: Assume that $X \sim P(\lambda)$, find D(X)

Answer: Based on the calculation from the last section, we have $E(X) = \lambda$

$$E(X^{2}) = E[X(X - 1)] + E(X)$$

$$= \sum_{k=0}^{\infty} k(k - 1) \frac{\lambda^{k}}{k!} e^{-\lambda} + \lambda$$

$$= \lambda^{2} e^{-\lambda} \sum_{k=2}^{\infty} \frac{\lambda^{k-2}}{(k-2)!} + \lambda$$

$$= \lambda^{2} + \lambda$$

$$\therefore D(X) = E(X^{2}) - [E(X)]^{2}$$

$$= \lambda^{2} + \lambda - \lambda^{2}$$

$$= \lambda$$



Example: Assume that $X \sim U(a, b)$, find D(X)

Answer: Based on the calculation from the last section, we have $E(X) = \frac{a+b}{2}$, the density of X is

$$f(x) = \begin{cases} \frac{1}{b-a}, & a < x < b \\ 0, & \text{otherwise} \end{cases}$$

$$\therefore D(X) = E(X^2) - [E(X)]^2$$

$$= \frac{a^2 + ab + b^2}{3} - \left(\frac{a+b}{2}\right)^2$$

$$= \frac{(b-a)^2}{12}$$



Example: Assume that the density function of X is

$$f(x) = \begin{cases} \frac{1}{\theta} e^{-x/\theta}, x > 0\\ 0, & x \le 0 \end{cases}$$

Find D(X).

Answer: Since it is the exponential distribution, then $E(X) = \theta$, so:

$$D(X) = E(X^{2}) - [E(X)]^{2}$$

$$= \int_{0}^{\infty} x^{2} \frac{1}{\theta} e^{-x/\theta} dx - \theta^{2}$$

$$= -x^{2} e^{-x/\theta} \Big|_{0}^{\infty} + 2 \int_{0}^{\infty} x e^{-x/\theta} dx - \theta^{2}$$

$$= 2\theta^{2} - \theta^{2} = \theta^{2}$$



Basic properties of variance (方差的基本性质)

- If $X =_{a.e.} c$ (constant), then D(X) = 0.
- If c is a constant, then $D(cX) = c^2 D(X)$.

Proof:
$$D(cX) = E(cX - E(cX))^{2}$$
$$= E(cX - cE(X))^{2}$$
$$= c^{2}E(X - E(X))^{2}$$
$$= c^{2}D(X)$$



Basic properties of variance(方差的基本性质)

- If $X =_{a.e.} c$ (constant), then D(X) = 0.
- If c is a constant, then $D(cX) = c^2 D(X)$.
- \bigcirc For r.v.s X and Y:

$$D(X + Y) = D(X) + D(Y) + 2E[(X - E(X))(Y - E(Y))]$$

Proof:
$$D(X + Y) = E[(X + Y) - E(X + Y)]^2$$

 $= E[(X - E(X)) + (Y - E(Y))]^2$
 $= E[X - E(X)]^2 + E[Y - E(Y)]^2$
 $+2E[(X - E(X))(Y - E(Y))]$
 $= D(X) + D(Y) + 2E[(X - E(X))(Y - E(Y))]$



Basic properties of variance (方差的基本性质)

- If $X =_{a.e.} c$ (constant), then D(X) = 0.
- If c is a constant, then $D(cX) = c^2 D(X)$.
- For r.v.s X and Y:

$$D(X + Y) = D(X) + D(Y) + 2E[(X - E(X))(Y - E(Y))]$$

If *X* and *Y* are independent, then D(X + Y) = D(X) + D(Y)

Proof: X and Y are independent

$$\therefore X - E(X)$$
 and $Y - E(Y)$ are independent



If X and Y are independent, is the following right?

$$D(X - Y) = D(X) + D(Y)$$





Example: Assume that $X \sim b(n, p)$, find D(X)

Answer: Since binomial distribution is from n-fold Bernoulli trials, then

$$X = X_1 + X_2 + \dots + X_n$$

where

$$X_i = \begin{cases} 1, & \text{A happened in the } i \text{th Bernoulli trial} \\ 0, & \overline{A} \text{ happened in the } i \text{th Bernoulli trial} \end{cases} (i = 1, 2, ..., n)$$

And $X_1, X_2, ..., X_n$ are independent and identically distributed (i.i.d.) with frequency functions:

$$P\{X_i = 1\} = p, P\{X_i = 0\} = 1 - p.$$

$$\therefore E(X) = E(X_1) + E(X_2) + \dots + E(X_n)$$

$$= nE(X_1) = np$$

$$D(X) = D(X_1) + D(X_2) + \dots + D(X_n)$$

$$= nD(X_1) = n[(1 - p)^2 \cdot p + (0 - p)^2 \cdot (1 - p)]$$

$$= np(1 - p)$$



Example: Assume that $X \sim N(\mu, \sigma^2)$, find D(X)

Answer: Based on the last section $E(X) = \mu$, then

$$D(X) = \int_{-\infty}^{\infty} (x - \mu)^2 f(x) dx$$

$$= \frac{1}{\sqrt{2\pi}\sigma} \int_{-\infty}^{\infty} (x - \mu)^2 e^{-\frac{(x - \mu)^2}{2\sigma^2}} dx$$

$$= \frac{\sigma^2}{\sqrt{2\pi}} \int_{-\infty}^{\infty} t^2 e^{-\frac{t^2}{2}} dt$$

$$= \frac{\sigma^2}{\sqrt{2\pi}} \int_{-\infty}^{\infty} (-t) de^{-\frac{t^2}{2}}$$

$$= \frac{\sigma^2}{\sqrt{2\pi}} \left[(-t) e^{-\frac{t^2}{2}} \Big|_{-\infty}^{\infty} + \int_{-\infty}^{\infty} e^{-\frac{t^2}{2}} dt \right]$$

$$= \frac{\sigma^2}{\sqrt{2\pi}} \sqrt{2\pi} = \sigma^2$$

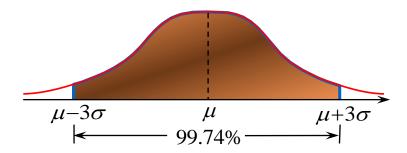


 3σ Principle: Most values of the normal r.v. are in the range $(\mu - 3\sigma, \mu + 3\sigma)$

If
$$X \sim N(\mu, \sigma^2)$$
, then

$$P\{|X - \mu| < 3\sigma\} = 0.9974$$

$$P\{|X - \mu| \ge 3\sigma\} = 0.0026$$





Generally, how to find the following probability of r.v. X:

$$P\{|X - \mu| \ge \varepsilon\}$$



where $\mu = E(X)$, $\varepsilon > 0$ are constants.

Theorem: Chebyshev's inequality (切比雪夫不等式)

If $\mu \triangleq E(X)$ and $\sigma^2 = D(X)$ both exist, then for $\forall \varepsilon > 0$, we have:

$$P\{|X - \mu| \ge \varepsilon\} \le \frac{\sigma^2}{\varepsilon^2}$$



Theorem: Chebyshev's inequality (切比雪夫不等式)

If $\mu \triangleq E(X)$ and $\sigma^2 = D(X)$ both exist, then for $\forall \varepsilon > 0$, we have:

$$P\{|X - \mu| \ge \varepsilon\} \le \frac{\sigma^2}{\varepsilon^2}$$

Proof: Only prove for continuous r.v.s, discrete r.v.s are similar. Let f(x) be the density function of X, then:

$$P\{|X - \mu| \ge \varepsilon\} = \int_{|x - \mu| \ge \varepsilon} f(x) dx$$

$$\le \int_{|x - \mu| \ge \varepsilon} \frac{(x - \mu)^2}{\varepsilon^2} f(x) dx$$

$$\le \frac{1}{\varepsilon^2} D(X) = \frac{\sigma^2}{\varepsilon^2}$$

Or we can let $Y = (X - \mu)^2$, then $E(Y) = \sigma^2$, applying the Markov Inequality:

$$P\{|X - \mu| \ge \varepsilon\} = P\{Y \ge \varepsilon^2\} \le \frac{E(Y)}{\varepsilon^2} = \frac{\sigma^2}{\varepsilon^2}$$



Theorem: Chebyshev's inequality

If $\mu \triangleq E(X)$ and $\sigma^2 = D(X)$ both exist, then for $\forall \varepsilon > 0$, we have:

$$P\{|X - \mu| \ge \varepsilon\} \le \frac{\sigma^2}{\varepsilon^2}$$

Equivalently, we have:

$$P\{|X - \mu| < \varepsilon\} \ge 1 - \frac{\sigma^2}{\varepsilon^2}$$

Let $\varepsilon = 3\sigma$ and 4σ , we have:

$$P\{|X - \mu| < 3\sigma\} \ge 1 - \frac{1}{9} = 88.90\%$$

$$P\{|X - \mu| < 4\sigma\} \ge 1 - \frac{1}{16} = 93.75\%$$

Inferencing result from Chebyshev's inequality:

If
$$\sigma^2 = 0$$
, then $P\{X = \mu\} = 1$.

Even for ordinary r.v.s, the reliability of the 3σ rule is close to 90%.



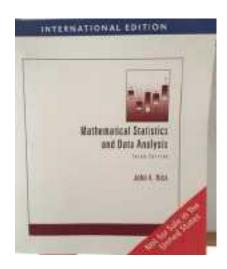
Summary Expected values and Variances

<i>X</i> ~	E(X)	D(X)
$X \sim p(\lambda)$	λ	λ
$X \sim b(n, p)$	np	<i>np(1 – p)</i>
$X \sim U(a, b)$	$\frac{a+b}{2}$	$x = \frac{(b-a)^2}{12}$
$X \sim exp(1/\theta)$	$oldsymbol{ heta}$	θ^2
$X \sim N(\mu, \sigma^2)$	μ	σ^2





Homework



P170: 49, 50, 55



Supplementary Questions:

- 1. Suppose that X and Y are independent random variables. E(X) = 3, E(Y) = 1, D(X) = 4, D(Y) = 9. If Z = 5X 2Y + 15, compute E(Z), D(Z).
- 2. Suppose that X_i (i = 1,2,3,4) are mutually independent to each other. (X_i) = 2i, $D(X_i) = 5 i$. If $Z = 2X_1 X_2 + 3X_3 0.5X_4$, compute E(Z) and D(Z).