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Example 17: A batch of 50 fuses is known to have 10 defective fuses. If 10 fuses are selected at random, find the probability that all will be good.

Solution: 10 fuses can be selected from 50 fuses in ${}^{50}C_{10}$ ways, so the total number of outcomes is

$$n = {}^{50}C_{10} = \frac{50!}{10!(50-10)!}$$

Let A denotes the event 'all 10 fuses will be good'.

$$\therefore m(A) = {}^{50-10}C_{10} = {}^{40}C_{10} = \frac{40!}{10!(40-10)!}.$$

$$\therefore P(A) = \frac{m(A)}{n} = \frac{40!}{10!30!} \cdot \frac{10!40!}{50!} = \frac{(40!)^2}{30!50!}, \text{ this is the required probability.}$$

1.5 AXIOMATIC DEFINITION OF PROBABILITY (W.B.U.T. 2005)

Let E be a given random experiment and S be the corresponding event space. Let A be an event associated with the random experiment E , then probability of A , denoted by $P(A)$, is a real number which satisfies the following axioms:

$$(A_1) P(A) \geq 0 \text{ for every } A \in S$$

$$(A_2) P(S) = 1$$

(A_3) if $A_1, A_2, \dots, A_n, \dots$ be a finite or infinite number of pairwise mutually exclusive events, i.e., $A_i A_j = \emptyset$ for $i \neq j$ and $A_i, A_j \in S$, then $P(A_1 + A_2 + \dots + A_n + \dots) = P(A_1) + P(A_2) + \dots + P(A_n) + \dots$

The entire mathematical structure of the theory of probability can be logically built up on the basis of these three axioms.

Deductions from Axiomatic Definition

Theorem 1: Probability of an impossible event is zero, i.e., $P(\emptyset) = 0$.

Proof: Now, $S + \emptyset = S$.

Both S and \emptyset are disjoint, i.e.,

$$S \emptyset = \emptyset$$

$$\therefore P(S + \emptyset) = P(S)$$

$$\Rightarrow P(S) + P(\emptyset) = P(S) \quad [\text{By } (A_3)]$$

$$\Rightarrow P(\emptyset) = 0.$$

Theorem 2: Probability of the complementary event \bar{A} of A is $P(\bar{A}) = 1 - P(A)$. Also, $0 \leq P(A) \leq 1$.

Proof: A and \bar{A} are disjoint events.

Moreover, $A + \bar{A} = S$

$$\therefore P(A + \bar{A}) = P(S)$$

$$\Rightarrow P(A) + P(\bar{A}) = P(S) \quad [\text{By } A_3]$$

$$\Rightarrow P(A) + P(\bar{A}) = 1 \quad [\text{By } A_2]$$

$$\Rightarrow P(\bar{A}) = 1 - P(A)$$

Again since $P(A) \geq 0$

[By A_1]

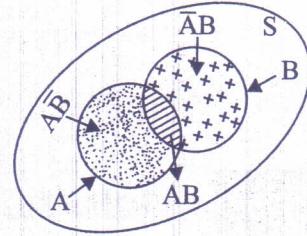
$$\begin{aligned}\therefore \quad & 0 \leq P(\bar{A}) \leq 1 \\ \Rightarrow \quad & 0 \leq P(A) \leq 1.\end{aligned}$$

Theorem 3: For any two events A and B ,

$$P(\bar{A}B) = P(B) - P(AB).$$

Proof: Now, AB and $\bar{A}B$ are disjoint.

$$\begin{aligned}\text{Also, } \quad & AB + \bar{A}B = B \\ \Rightarrow \quad & P(AB + \bar{A}B) = P(B) \\ \Rightarrow \quad & P(AB) + P(\bar{A}B) = P(B) \quad [\text{By } A_3] \\ \Rightarrow \quad & P(\bar{A}B) = P(B) - P(AB).\end{aligned}$$



Note: $\bar{A} \bar{B} + (A + B) = S$.

Theorem 4 (Addition Theorem): For any two events A and B (may not be mutually exclusive),

$$P(A + B) = P(A) + P(B) - P(AB). \quad (\text{W.B.U.T. 2004, 2007})$$

Proof: From the figure of Theorem 3, we have $A + B = A + \bar{A}B$, also A and $\bar{A}B$ are disjoint.

$$\begin{aligned}\therefore \quad P(A + B) &= P(A + \bar{A}B) \\ &= P(A) + P(\bar{A}B) \quad [\text{By } A_3] \\ &= P(A) + P(B) - P(AB) \quad [\text{By Theorem 3}]\end{aligned}$$

Theorem 5 (Extension of Theorem 4): For any three events A, B, C

$$P(A + B + C) = P(A) + P(B) + P(C) - P(AB) - P(BC) - P(CA) + P(ABC).$$

$$\begin{aligned}\text{Proof: } \quad \text{LHS} &= P(A + B + C) = P\{(A + B) + C\} \\ &= P(A + B) + P(C) - P\{(A + B)C\} \quad [\text{By Theorem 4}] \\ &= P(A) + P(B) - P(AB) + P(C) - P(AC + BC) \quad [\text{By Theorem 4}] \\ &= P(A) + P(B) + P(C) - P(AB) - \{P(AC) + P(BC) - P(ABC)\} \\ &\quad [\text{By Theorem 4 and since } (AC)(BC) = ABC] \\ &= \text{RHS.}\end{aligned}$$

Deduction of Classical Definition

Let the event space (of a random experiment E) contains n distinct event points (i.e., simple events) u_1, u_2, \dots, u_n .

$$\begin{aligned}\therefore \quad & u_1 + u_2 + \dots + u_n = S \\ \Rightarrow \quad & P(u_1 + u_2 + \dots + u_n) = P(S) \\ \Rightarrow \quad & P(u_1) + P(u_2) + \dots + P(u_n) = 1.\end{aligned}$$

Let the simple events have equal probability,

$$\text{i.e., } P(u_1) = P(u_2) = \dots = P(u_n) = \frac{1}{n}.$$

Now suppose A be an event connected to the given random experiment E . If A contains m ($\leq n$) event points of S , then without loss of generality we can write

$$\begin{aligned} A &= u_1 + u_2 + \dots + u_m \\ \therefore P(A) &= P(u_1 + u_2 + \dots + u_m) \\ &= P(u_1) + P(u_2) + \dots + P(u_m) \\ &= \frac{1}{n} + \frac{1}{n} + \dots + \frac{1}{n} \\ &\rightarrow m \text{ terms} \leftarrow \\ &= \frac{m}{n} = \frac{m(A)}{n} \end{aligned}$$

which is the classical definition of probability deduced from axiomatic definition of probability.

Theorem 6: For any two events A and B , $P(AB) \leq P(A) \leq P(A + B) \leq P(A) + P(B)$.

(W.B.U.T. 2004, 2008)

Proof: Now, $AB, A\bar{B}$ are disjoint and $AB + A\bar{B} = A$

(See the Fig. of Th. 3)

$$\begin{aligned} \therefore P(AB + A\bar{B}) &= P(A) \\ \Rightarrow P(AB) + P(A\bar{B}) &= P(A) \end{aligned}$$

Since $0 \leq P(A\bar{B}) \leq 1$, so $P(AB)$ can never exceed $P(A)$

i.e., $P(AB) \leq P(A)$... (1)

Again, $P(A + B) = P(A) + P(B) - P(AB)$ (By Theorem 4)
 $= P(A) + P(\bar{A}B)$ (By Theorem 3)

Since $0 \leq P(\bar{A}B) \leq 1$, so $P(A) \leq P(A + B)$... (2)

Now, $P(A + B) = P(A) + P(B) - P(AB) \leq P(A) + P(B)$ [$\because 0 \leq P(AB) \leq 1$] ... (3)

Combining (1), (2) and (3), we get

$$P(AB) \leq P(A) \leq P(A + B) \leq P(A) + P(B).$$

Note: Similarly, $P(AB) \leq P(B) \leq P(A + B) \leq P(A) + P(B)$

Theorem 7 (Boole's Inequality): For any n events A_1, A_2, \dots, A_n connected to a random experiment E ,

$$P(A_1 + A_2 + \dots + A_n) \leq P(A_1) + P(A_2) + \dots + P(A_n)$$

Proof: We have for any two events A_1, A_2 ,

$$P(A_1 + A_2) = P(A_1) + P(A_2) - P(A_1 A_2) \leq P(A_1) + P(A_2) \quad [\because P(A_1 A_2) \geq 0] \quad \dots (1)$$

Therefore the given inequality is true for $n = 1, 2$.

Let us assume that the given inequality is true for any positive integer $m \geq 2$,

i.e., $P(A_1 + A_2 + \dots + A_m) \leq P(A_1) + P(A_2) + \dots + P(A_m)$... (2)

$$\begin{aligned} \therefore P(A_1 + A_2 + \dots + A_m + A_{m+1}) &= P[(A_1 + A_2 + \dots + A_m) + A_{m+1}] \\ &\leq P(A_1 + A_2 + \dots + A_m) + P(A_{m+1}) \quad [\text{By (1)}] \\ &\leq P(A_1) + P(A_2) + \dots + P(A_m) + P(A_{m+1}) \quad [\text{By (2)}] \end{aligned}$$

Thus the given inequality is true for $m + 1$ whenever it is true for m . But we have already seen that this inequality is true for 2, hence it is true for $2 + 1 = 3, 3 + 1 = 4, 4 + 1 = 5$, etc. Hence the given inequality is true for any positive integer n .

Theorem 8 (Bonferroni's Inequalities): For any n events A_1, A_2, \dots, A_n connected to a random experiment E ,

$$(i) P(A_1 A_2 \dots A_n) \geq 1 - \sum_{i=1}^n P(\bar{A}_i)$$

$$(ii) P(A_1 A_2 \dots A_n) \geq \sum_{i=1}^n P(A_i) - (n-1). \quad (\text{W.B.U.T. 2003})$$

Proof: (i) We have by Boole's inequality,

$$\begin{aligned} & P(\bar{A}_1 + \bar{A}_2 + \dots + \bar{A}_n) \leq P(\bar{A}_1) + P(\bar{A}_2) + \dots + P(\bar{A}_n) \\ \Rightarrow & P(\overline{A_1 A_2 \dots A_n}) \leq P(\bar{A}_1) + P(\bar{A}_2) + \dots + P(\bar{A}_n) \quad [\text{By De Morgan's law}] \\ \Rightarrow & 1 - P(A_1 A_2 \dots A_n) \leq \sum_{i=1}^n P(\bar{A}_i) \\ \Rightarrow & 1 - \sum_{i=1}^n P(\bar{A}_i) \leq P(A_1 A_2 \dots A_n) \\ \Rightarrow & P(A_1 A_2 \dots A_n) \geq 1 - \sum_{i=1}^n P(\bar{A}_i) \end{aligned}$$

(ii) We know that $P(\bar{A}_i) = 1 - P(A_i)$

$$\therefore 1 - \sum_{i=1}^n P(\bar{A}_i) = 1 - \sum_{i=1}^n [1 - P(A_i)] = \sum_{i=1}^n P(A_i) - (n-1)$$

$$\text{Using (i): } P(A_1 A_2 \dots A_n) \geq \sum_{i=1}^n P(A_i) - (n-1).$$

Example: In an examination 60%, 70% and 75% students have passed in the subjects Physics, Chemistry and Mathematics respectively. Find the minimum percentage of students who have passed in all the subjects.

Solution: Let A, B and C be the events defined as follows:

$A \equiv$ student chosen passes the subject Physics,

$B \equiv$ student chosen passes the subject Chemistry,

$C \equiv$ student chosen passes the subject Mathematics.

$$\text{Given: } P(A) = \frac{60}{100} = \frac{3}{5}, P(B) = \frac{70}{100} = \frac{7}{10}, P(C) = \frac{75}{100} = \frac{15}{20}.$$

$$\therefore P(A) + P(B) + P(C) = \frac{3}{5} + \frac{7}{10} + \frac{15}{20} = \frac{12 + 14 + 15}{20} = \frac{41}{20}.$$

Now, applying Bonferroni's inequality, we have $P(\text{student chosen passes all subjects})$

$$= P(ABC) \geq P(A) + P(B) + P(C) - (3-1)$$

$$= \frac{41}{20} - 2 = \frac{1}{20} = 0.05$$

Hence, the minimum value of $P(ABC)$ is 0.05. Therefore, at least 5% students may pass in all the subjects.

1.6 CONDITIONAL PROBABILITY

Let A and B are any two events connected to a given random experiment E . The **conditional probability** of the event A on the hypothesis that the event B has occurred is denoted and defined by

$$P(A/B) = \frac{P(AB)}{P(B)}, \text{ provided } P(B) \neq 0.$$

Similarly, we can define

$$P(B/A) = \frac{P(AB)}{P(A)}, \text{ provided } P(A) \neq 0.$$

Note: The existence of conditional probabilities are guaranteed by Theorem 6, Art. 1.5.

Example 1: Let E be the random experiment of throwing of a die with event space S . Let A denotes the event ‘even face’ and B denotes the event ‘multiple of 3’.

∴

$$S = \{1, 2, 3, 4, 5, 6\}, A = \{2, 4, 6\}, B = \{3, 6\}, AB = \{6\}$$

∴

$$P(A) = \frac{3}{6} = \frac{1}{2}, P(B) = \frac{2}{6} = \frac{1}{3}, P(AB) = \frac{1}{6}.$$

∴

$$P(A/B) = \frac{P(AB)}{P(B)} = \frac{1/6}{1/3} = \frac{1}{2}.$$

$$P(B/A) = \frac{P(AB)}{P(A)} = \frac{1/6}{1/2} = \frac{1}{3}.$$

Example 2: Let A and B are two events such that $P(A \cup B) = 7/8$, $P(A \cap B) = 1/4$ and $P(A^c) = 5/8$. Find the conditional probability of A on the hypothesis that B does not occur. (W.B.U.T. 2003)

Solution: Given: $P(A \cup B) = 7/8$, $P(A \cap B) = 1/4$, $P(A^c) = 5/8$.

We have to find

$$P(A/B^c) = \frac{P(A \cap B^c)}{P(B^c)} = \frac{P(A) - P(A \cap B)}{1 - P(B)} \quad \dots(1)$$

Now,

$$P(A) = 1 - P(A^c) = 1 - \frac{5}{8} = \frac{3}{8}.$$

We know that,

$$P(A \cup B) = P(A) + P(B) - P(A \cap B).$$

⇒

$$\frac{7}{8} = \frac{3}{8} + P(B) - \frac{1}{4}$$

⇒

$$P(B) = \frac{7}{8} - \frac{3}{8} + \frac{1}{4} = \frac{7-3+2}{8} = \frac{3}{4}.$$

∴ Required probability

$$= P(A/B^c) = \frac{P(A) - P(A \cap B)}{1 - P(B)} \quad [\text{By (1)}]$$

$$= \frac{\frac{3}{8} - \frac{1}{4}}{1 - \frac{3}{4}} = \frac{\frac{1}{8}}{\frac{1}{4}} = \frac{1}{2}.$$

Theorem 1: If A_1, A_2, \dots, A_n are pairwise mutually exclusive events one of which certainly occurs, then

$$P(B) = \sum_{i=1}^n P(A_i) P(B/A_i)$$

where B is any event connected to the same random experiment provided the conditional probabilities are defined.

Proof: Here $A_i A_j = \emptyset$, for all $i \neq j$; $i, j = 1, 2, \dots, n$.

$A_1 + A_2 + \dots + A_n$ = certain event $= S$, since one of A_1, A_2, \dots, A_n certainly occurs.

$$\therefore B = SB = (A_1 + A_2 + \dots + A_n)B = A_1 B + A_2 B + \dots + A_n B.$$

$$\text{Also, } (A_i B)(A_j B) = (A_i A_j)B = \emptyset B = \emptyset \quad (i \neq j; i, j = 1, 2, \dots, n)$$

$$\begin{aligned} \therefore P(B) &= P(A_1 B + A_2 B + \dots + A_n B) \\ &= P(A_1 B) + P(A_2 B) + \dots + P(A_n B) \end{aligned}$$

$$= \sum_{i=1}^n P(A_i B)$$

$$= \sum_{i=1}^n P(A_i) P(B/A_i) \quad \left[\because P(B/A_i) = \frac{P(A_i B)}{P(A_i)}, i = 1, 2, \dots, n \right]$$

$$\therefore P(B) = \sum_{i=1}^n P(A_i) P(B/A_i) \quad \dots(1)$$

Notes: (i) This theorem is known as the **theorem on total probability**.

$$(ii) \text{ By definition: } P(A_i/B) = \frac{P(A_i B)}{P(B)} = \frac{P(A_i) P(B/A_i)}{P(B)}$$

$$\text{Using (1): } P(A_i/B) = \frac{P(A_i) P(B/A_i)}{\sum_{r=1}^n P(A_r) P(B/A_r)}, \text{ provided } P(B) \neq 0.$$

This is known as **Baye's theorem**.

(W.B.U.T. 2012)

Example 3: A bag contains 4 white and 2 black balls and a second bag contains 3 of each colours. A bag is selected at random and a ball is then taken out at random from the bag chosen. What is the probability that the ball selected is a white?

Solution: Let A_i ($i = 1, 2$) be the event of selecting i^{th} bag and B be the event of selecting a white ball.

$$\therefore P(A_1) = P(A_2) = \frac{1}{2}; P(B/A_1) = \frac{4}{6} = \frac{2}{3}; P(B/A_2) = \frac{3}{6} = \frac{1}{2}.$$

$$\therefore P(B) = P(A_1) P(B/A_1) + P(A_2) P(B/A_2)$$

$$= \frac{1}{2} \cdot \frac{2}{3} + \frac{1}{2} \cdot \frac{1}{2} = \frac{1}{3} + \frac{1}{4} = \frac{7}{12}.$$

Example 4: Suppose that there is a chance for newly constructed house to collapse whether the design is faulty or not. The chance that the design is faulty is 10%. The chance that the house collapse if the design is faulty is 95% and otherwise it is 45%. It is seen that the house collapsed. What is the probability that it is due to faulty design?

Solution: Let A , \bar{A} and B denote the following events:

$A \equiv$ Design is faulty, $\bar{A} \equiv$ design is not faulty and $B \equiv$ the house collapses.

Given, $P(A) = 0.1, P(B/A) = 0.95, P(B/\bar{A}) = 0.45.$

$$\therefore P(\bar{A}) = 1 - P(A) = 0.9.$$

\therefore Required probability $= P(A/B)$

$$\begin{aligned} &= \frac{P(A) P(B/A)}{P(A) P(B/A) + P(\bar{A}) P(B/\bar{A})} \\ &= \frac{0.1 \times 0.95}{0.1 \times 0.95 + 0.9 \times 0.45} = 0.19. \end{aligned}$$

[By Baye's theorem]

Example 5: The chance that a doctor will diagnose a certain disease correctly is 60%. The chance that a patient will die by his treatment after correct diagnosis is 40% and the chance of death by wrong diagnosis is 70%. A patient of the doctor, who had the disease, dies. What is the probability that the disease was diagnosed correctly?

Solution: Let A_1 denotes the event 'the disease was diagnosed correctly by the doctor' and A_2 denotes the event 'a patient who has the disease dies'.

$$\text{Given, } P(A_1) = \frac{60}{100} = \frac{3}{5}, P(\bar{A}_1) = 1 - P(A_1) = 1 - \frac{3}{5} = \frac{2}{5},$$

$$P(A_2/A_1) = \frac{40}{100} = \frac{2}{5} \text{ and } P(A_2/\bar{A}_1) = \frac{70}{100} = \frac{7}{10}.$$

Required probability

$= P(\text{Disease was diagnosed correctly given that the patient dies})$

$$= P(A_1/A_2) = \frac{P(A_1) P(A_2/A_1)}{P(A_1) P(A_2/A_1) + P(\bar{A}_1) P(A_2/\bar{A}_1)}$$

[By Baye's theorem]

$$= \frac{\frac{3}{5} \times \frac{2}{5}}{\frac{3}{5} \times \frac{2}{5} + \frac{2}{5} \times \frac{7}{10}} = \frac{12}{12+14} = \frac{6}{13}.$$

Example 6: There were three candidates A, B and C for the position of a manager whose chances of getting the appointment are in the proportion 4 : 2 : 3 respectively. The probability that A, if selected, would launch a new product in the market is 0.3. The probabilities of B and C doing the same are 0.5 and 0.8 respectively. What is the probability that the new product was launched in the market by C?

(W.B.U.T. 2006)

Solution: Let X, Y, Z respectively denote the events that A, B, C are selected as a manager and N denotes the event that a new product is launched in the market.

$$\text{Given, } P(X) = \frac{4}{9}, P(Y) = \frac{2}{9}, P(Z) = \frac{3}{9},$$

$$P(N/X) = \text{Pr. that a new product is launched in market by } A = 0.3$$

$$P(N/Y) = \text{Pr. that a new product is launched in market by } B = 0.5$$

$$P(N/Z) = \text{Pr. that a new product is launched in market by } C = 0.8$$

Using Baye's theorem, the required probability is

$$\begin{aligned} P(Z/N) &= \frac{P(Z) P(N/Z)}{P(X) P(N/X) + P(Y) P(N/Y) + P(Z) P(N/Z)} \\ &= \frac{\frac{3}{9} \times 0.8}{\frac{4}{9} \times 0.3 + \frac{2}{9} \times 0.5 + \frac{3}{9} \times 0.8} = \frac{24}{12+10+24} = \frac{24}{46} = \frac{12}{23}. \end{aligned}$$

Example 7: The probabilities of X , Y and Z becoming the Principal of a college are respectively 0.3, 0.5 and 0.2. The probabilities that 'Student Aid-Fund' will be introduced in the college if X , Y and Z become Principal, are 0.4, 0.6 and 0.1 respectively. Given that 'Student Aid-Fund' has been introduced, find the probability that Y has been appointed as the Principal. (W.B.U.T. 2004)

Solution: Let A , B and C respectively denote the events that X , Y and Z are becoming the Principal and F denotes the event that 'Student Aid-Fund' is introduced.

$$\text{Given: } P(A) = 0.3, P(B) = 0.5, P(C) = 0.2$$

$$P(F/A) = P(\text{Student Aid-Fund will be introduced if } X \text{ becomes Principal}) = 0.4.$$

$$\text{Similarly, } P(F/B) = 0.6 \text{ and } P(F/C) = 0.1.$$

Required probability

$$\begin{aligned} &= P(B/F) \\ &= P(Y \text{ becomes Principal given Student Aid-Fund has been introduced}). \end{aligned}$$

Using Baye's theorem,

$$\begin{aligned} P(B/F) &= \frac{P(B) P(F/B)}{P(A) P(F/A) + P(B) P(F/B) + P(C) P(F/C)} \\ &= \frac{0.5 \times 0.6}{0.3 \times 0.4 + 0.5 \times 0.6 + 0.2 \times 0.1} = \frac{30}{12+30+2} = \frac{30}{44} = \frac{15}{22}. \end{aligned}$$

Example 8: Three identical boxes, I, II and III contain respectively 4 white and 3 red balls, 3 white and 7 red balls, 2 white and 3 red balls. A box is chosen at random and a ball is drawn out of it. If the ball is found to be white, what is the probability that box II is selected? (W.B.U.T. 2007)

Solution: Let A_1 , A_2 and A_3 denote the events that the ball is drawn from boxes I, II and III respectively. Clearly the events A_1 , A_2 and A_3 are mutually exclusive and exhaustive.

$$\therefore P(A_1) = P(A_2) = P(A_3) = \frac{1}{3}.$$

Also let A denote the event that the drawn ball is white.

$$\therefore P(A/A_1) = \frac{4}{7}, P(A/A_2) = \frac{3}{10} \text{ and } P(A/A_3) = \frac{2}{5}.$$

Using Baye's theorem, the required probability is

$$P(A_2/A) = \frac{P(A_2) P(A/A_2)}{P(A_1) P(A/A_1) + P(A_2) P(A/A_2) + P(A_3) P(A/A_3)}$$