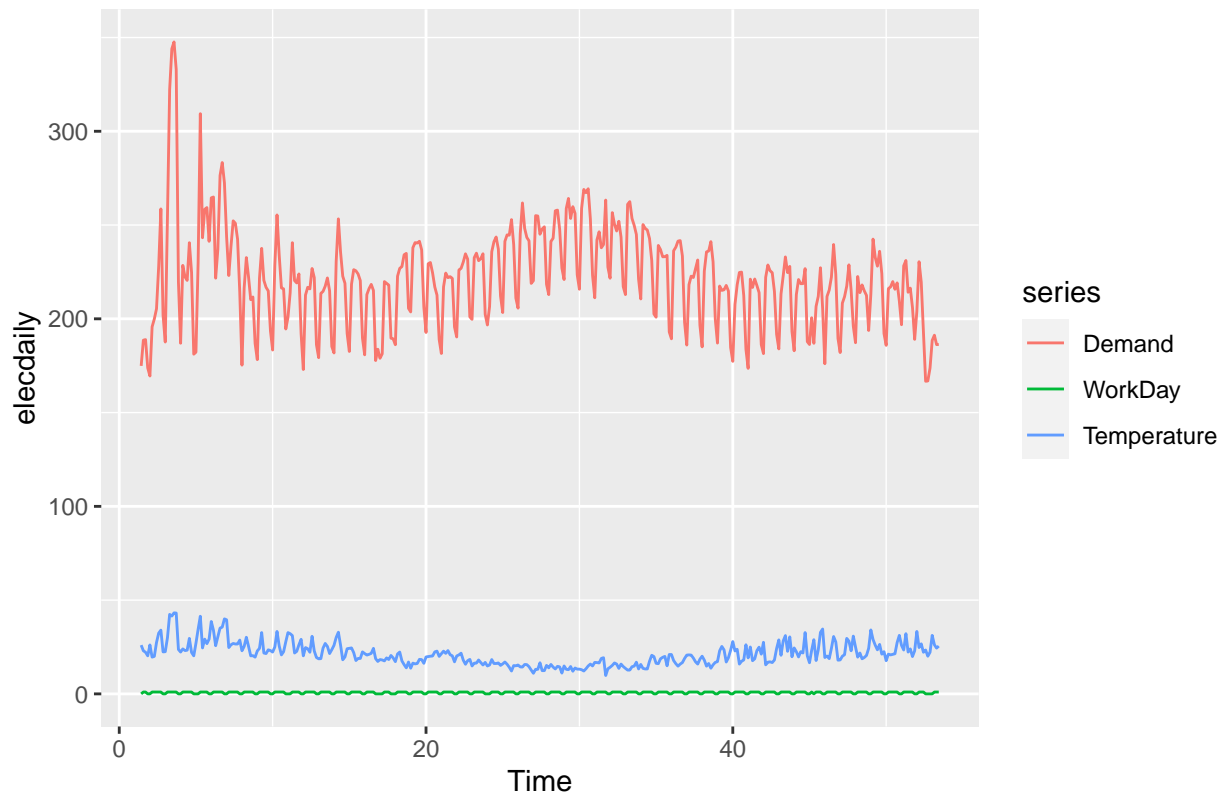


Electricity Demand

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```
library(fpp2)
autoplot(elecdaily)
```



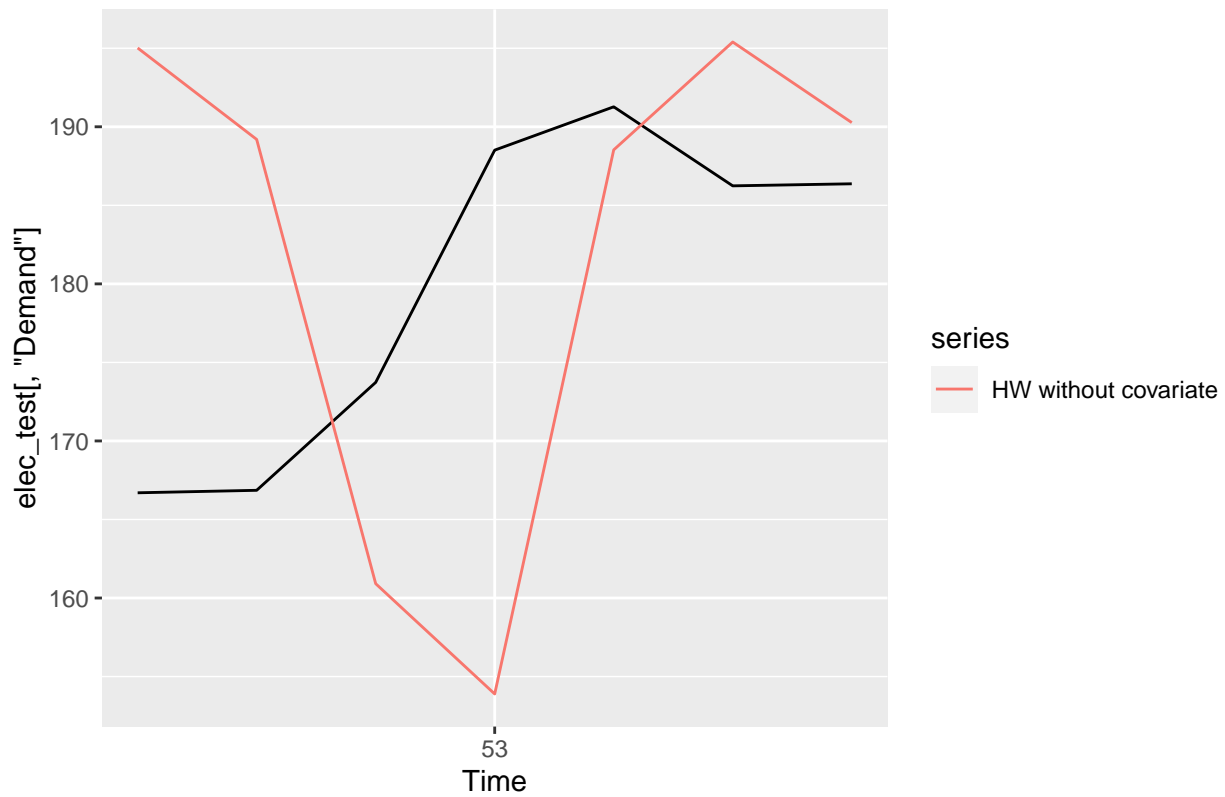
We split into train and test

```
elec_train=window(elecdaily,start=c(1,1),end=c(52,4))
elec_test=window(elecdaily,start=c(52,5),end=c(53,4))
```

Forecasting without covariates

First, we start with an HoltWinters exponential smoothing

```
fit=hw(elec_train[, "Demand"])
prev=forecast(fit,h=7)
autoplot(elec_test[, "Demand"])+autolayer(prev$mean,series="HW without covariate")
```

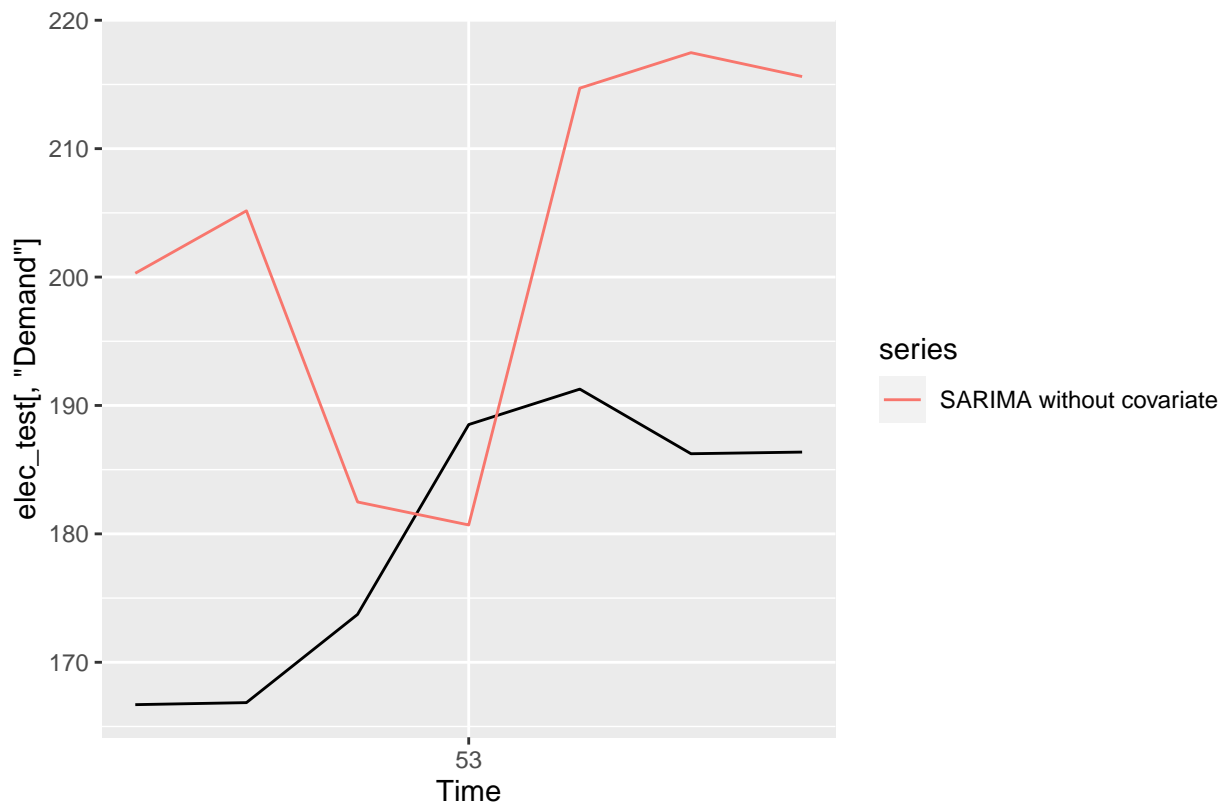


```
print(sqrt(mean((prev$mean-elec_test[,\"Demand\"])^2)))
```

```
## [1] 19.89454
```

and a SARIMA model:

```
fit=auto.arima(elec_train[,\"Demand\"])\nprev=forecast(fit,h=7)\nautoplot(elec_test[,\"Demand\"])+autolayer(prev$mean,series=\"SARIMA without covariate\")
```



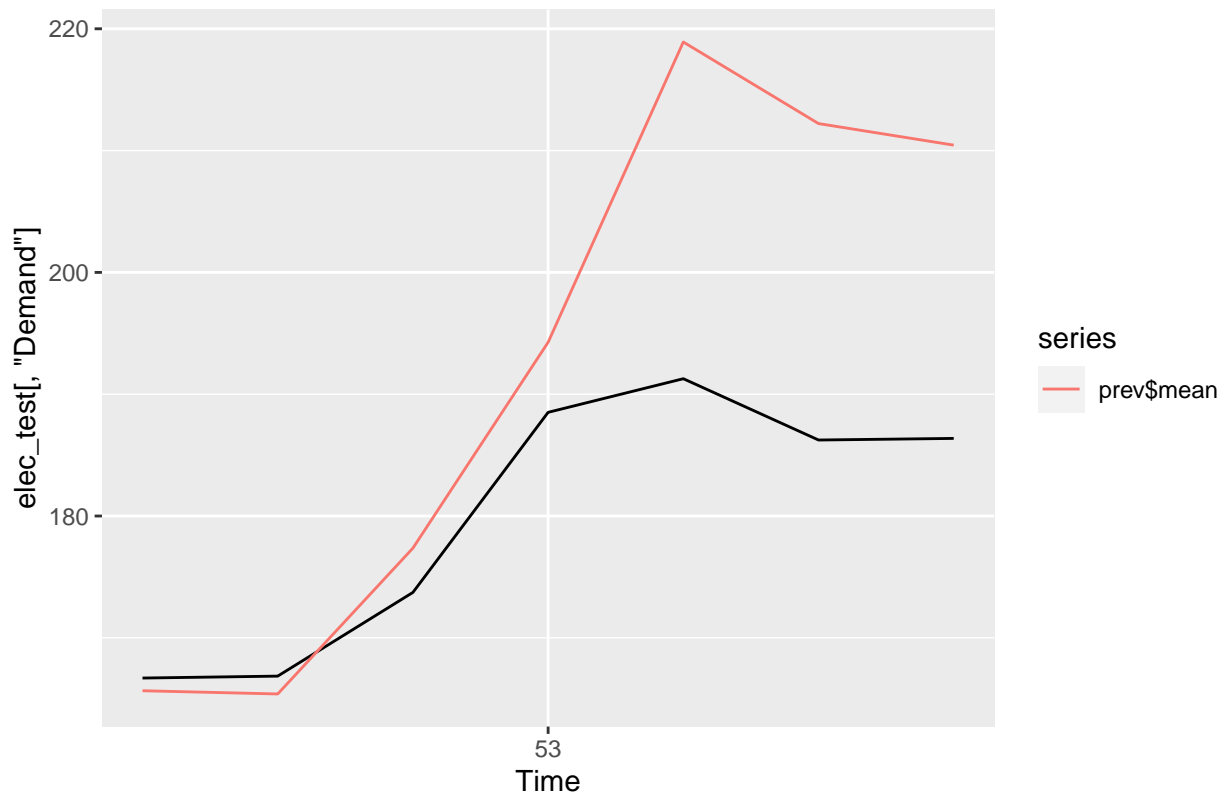
```
print(sqrt(mean((prev$mean-elec_test[, "Demand"])^2)))
```

```
## [1] 27.02949
```

Forecasting with covariates

We will use a dynamic regression model for forecasting electricity demand, using temperature and workday as external covariates. The order of the ARIMA model for the residual part is automatically selected

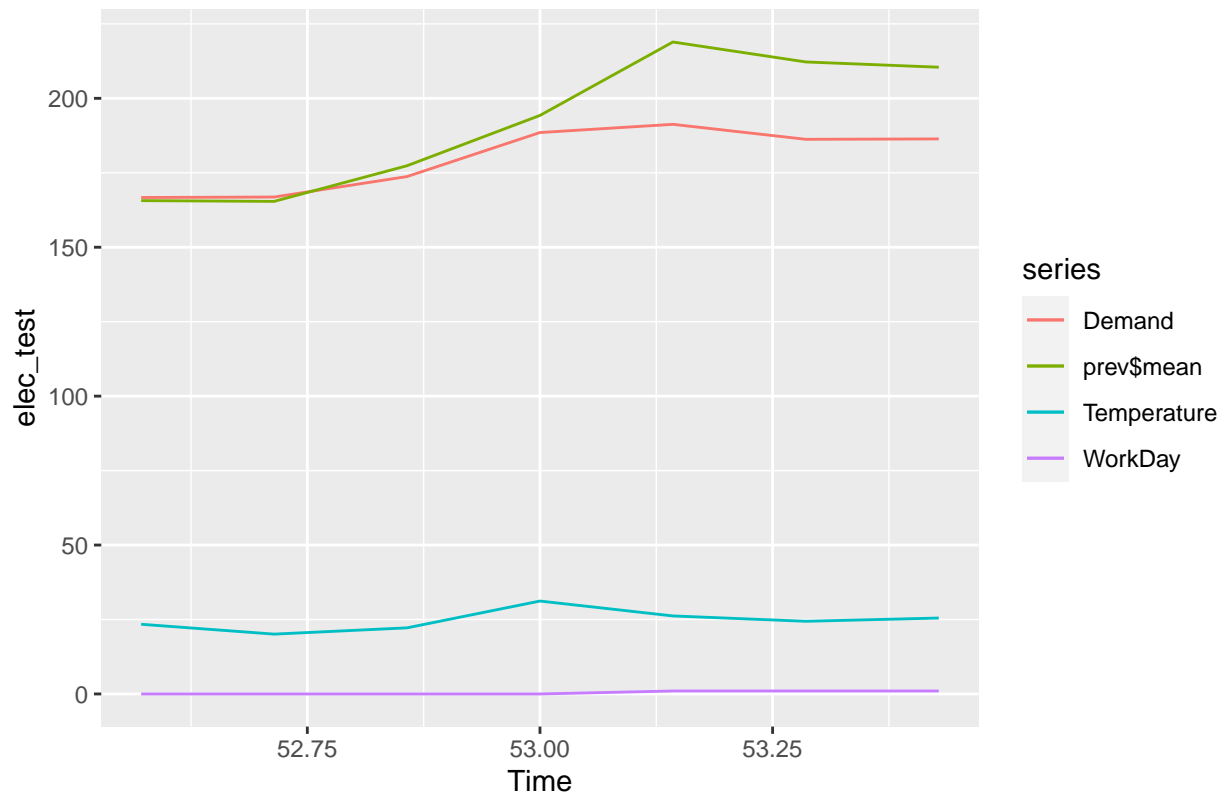
```
fit=auto.arima(elec_train[, "Demand"], xreg=elec_train[, 2:3])
prev=forecast(fit, h=7, xreg=elec_test[, 2:3])
autoplot(elec_test[, "Demand"])+autolayer(prev$mean)
```



```
print(sqrt(mean((prev$mean-elec_test[, "Demand"])^2)))
```

```
## [1] 17.18817
```

```
autoplot(elec_test)+autolayer(prev$mean)
```



ing covariates allows us to improve the forecasting.

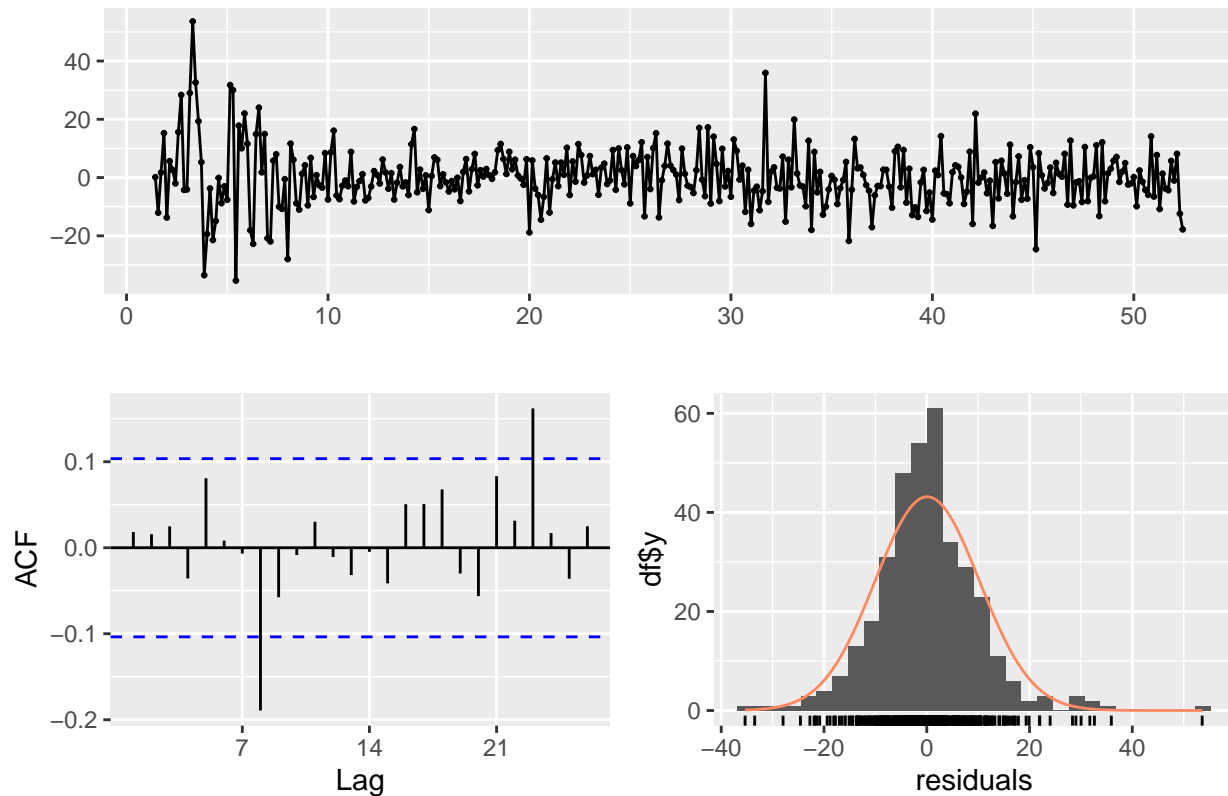
But if we check the residual, there is still some autocorrelations:

```
summary(fit)
```

```
## Series: elec_train[, "Demand"]
## Regression with ARIMA(5,1,1)(2,0,1)[7] errors
##
## Coefficients:
##          ar1      ar2      ar3      ar4      ar5      ma1      sar1      sar2
##          0.9003  -0.2229  0.0952  -0.1140  0.1567  -0.9625  -0.5076  0.1984
## s.e.      0.0687   0.0791  0.0752   0.0723  0.0597   0.0319   0.2783  0.0599
##          sma1  WorkDay  Temperature
##          0.6309  31.3165    1.4803
## s.e.      0.2813   1.3671    0.1422
##
## sigma^2 = 106.3: log likelihood = -1334.54
## AIC=2693.08  AICc=2693.99  BIC=2739.62
##
## Training set error measures:
##              ME      RMSE      MAE      MPE      MAPE      MASE
## Training set 0.09565789 10.13746 7.353345 -0.1447896 3.283934 0.5146259
##              ACF1
## Training set 0.0182663
```

```
checkresiduals(fit)
```

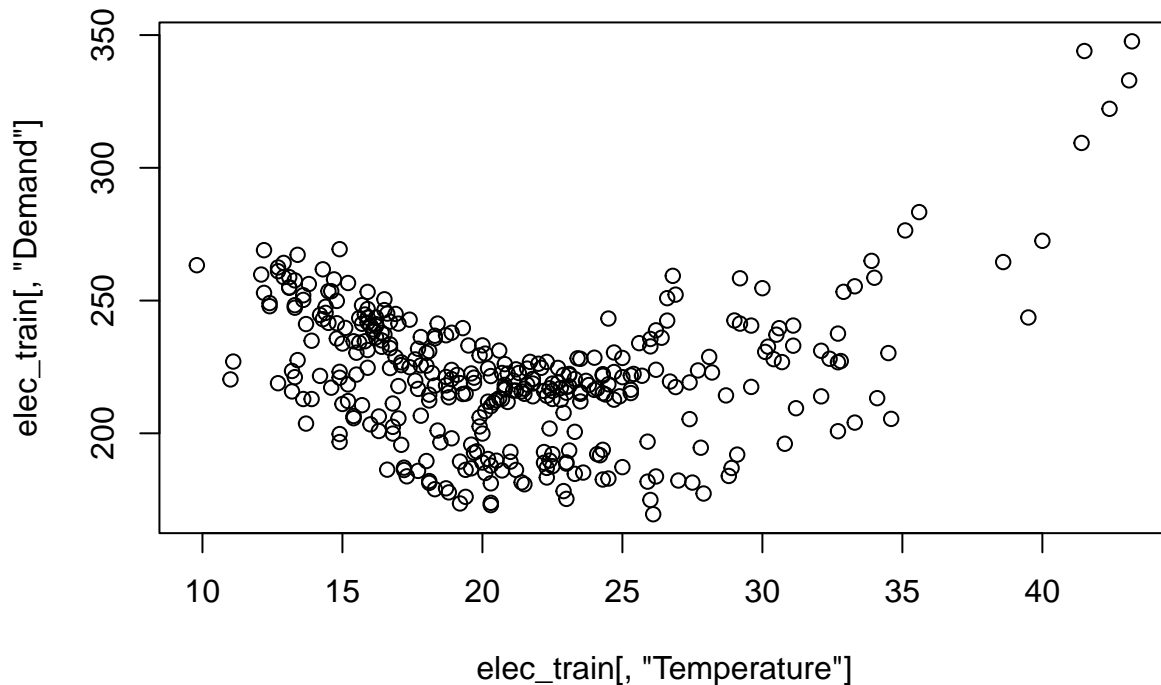
Residuals from Regression with ARIMA(5,1,1)(2,0,1)[7] errors



```
##
##  Ljung-Box test
##
## data:  Residuals from Regression with ARIMA(5,1,1)(2,0,1)[7] errors
## Q* = 18.529, df = 5, p-value = 0.002352
##
## Model df: 9.   Total lags used: 14
```

We can try to find a better model manually. Let's have a look to the relationship between Demand and Temperature

```
plot(elec_train[, "Temperature"], elec_train[, "Demand"])
```



The link seems to be quadratic. We introduced $Temperature^2$ in the model.

```
elec_train=cbind(Demand=elec_train[,1],WorkDay=elec_train[,2],Temp=elec_train[,3],SquareTemp=elec_train[,4])
```

Let's start by removing the effect of covariate.

```
fit2=tslm(Demand~WorkDay+Temp+SquareTemp+trend+season,data=elec_train)
summary(fit2)
```

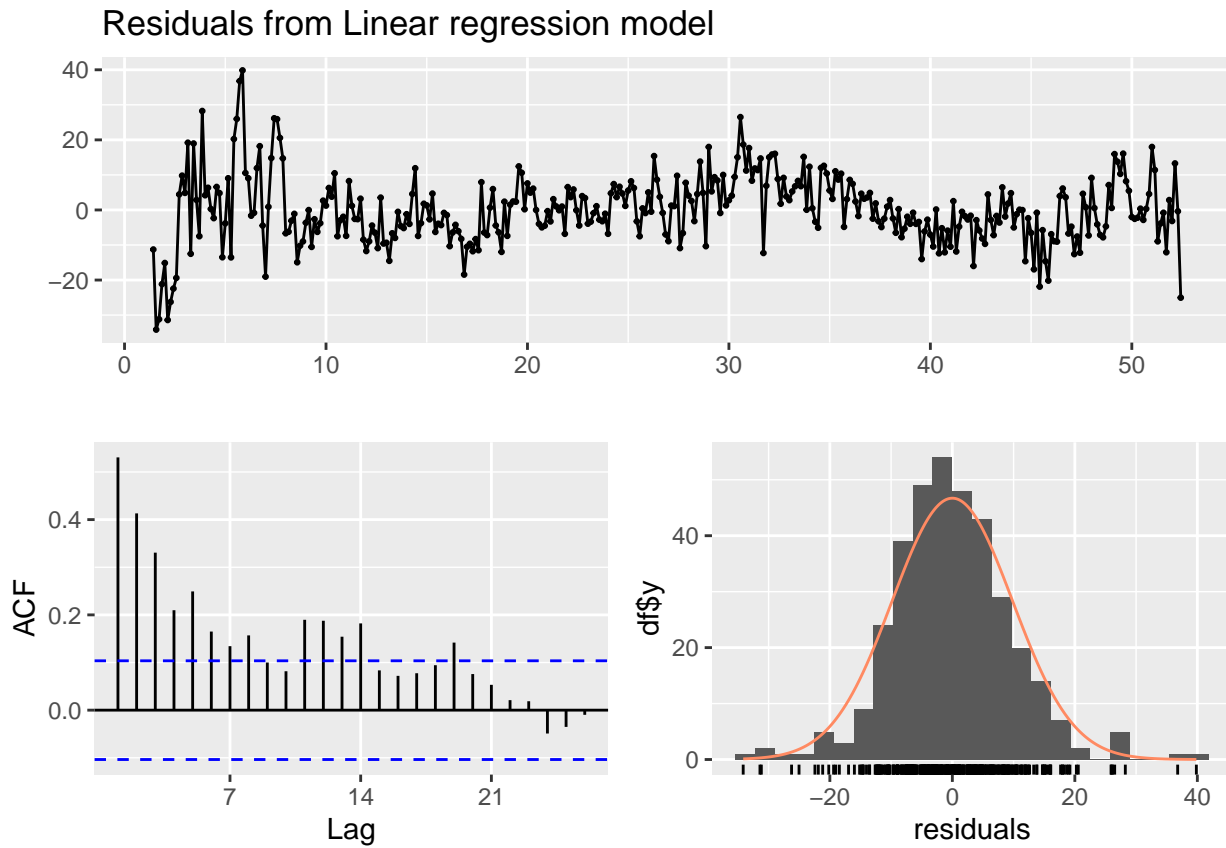
```
##
## Call:
## tslm(formula = Demand ~ WorkDay + Temp + SquareTemp + trend +
##       season, data = elec_train)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -34.149  -6.200  -0.524   5.293  39.840
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  357.217153   5.704688  62.618 < 2e-16 ***
## WorkDay       38.891730   3.650921  10.653 < 2e-16 ***
## Temp        -14.849365   0.472118 -31.453 < 2e-16 ***
## SquareTemp    0.315740   0.009600  32.889 < 2e-16 ***
## trend        -0.015952   0.005301  -3.009  0.00281 **
## season2       0.878590   3.902562   0.225  0.82201
## season3       0.850769   4.091306   0.208  0.83539
## season4       1.621769   4.085401   0.397  0.69164
## season5       1.171490   4.151346   0.282  0.77796
## season6      -1.909570   4.025997  -0.474  0.63558
## season7       9.160792   1.976629   4.635 5.07e-06 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
##
## Residual standard error: 9.977 on 347 degrees of freedom
## Multiple R-squared:  0.8593, Adjusted R-squared:  0.8552
## F-statistic: 211.9 on 10 and 347 DF,  p-value: < 2.2e-16
```

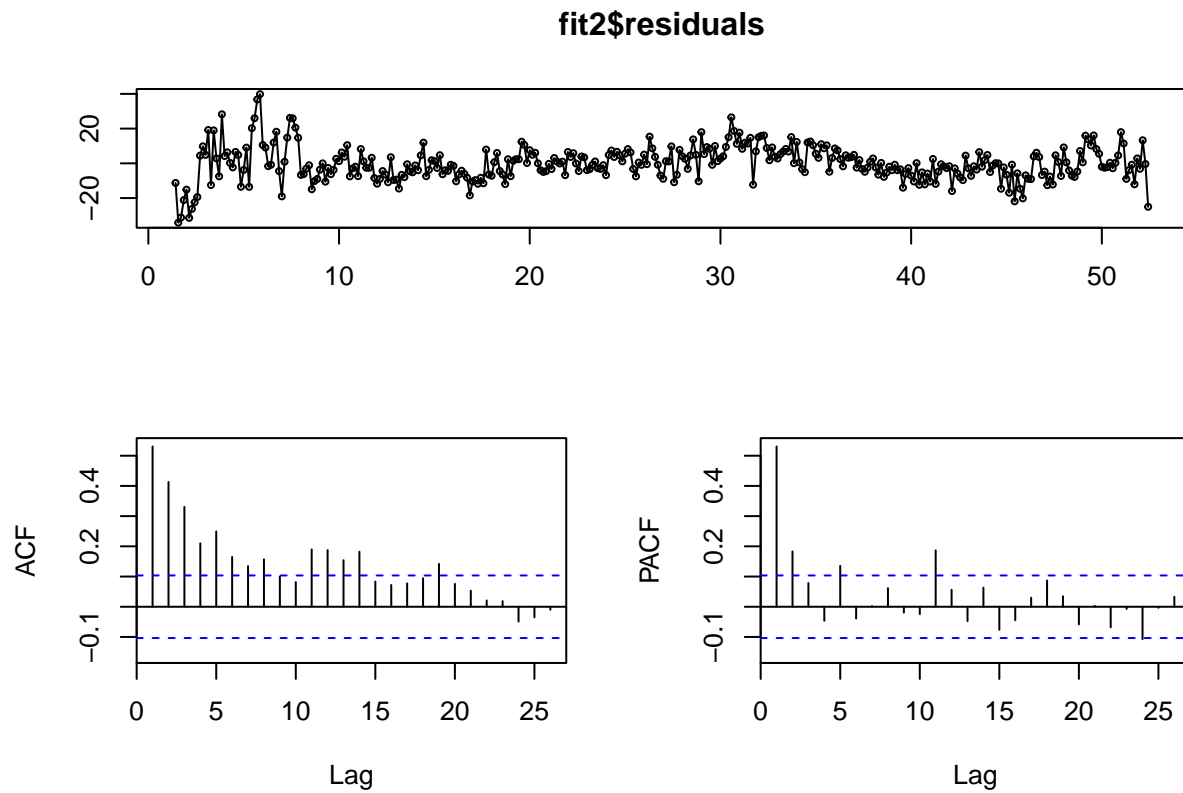
All the feature seems significant.

Let's now have a look to the residual

```
checkresiduals(fit2)
```



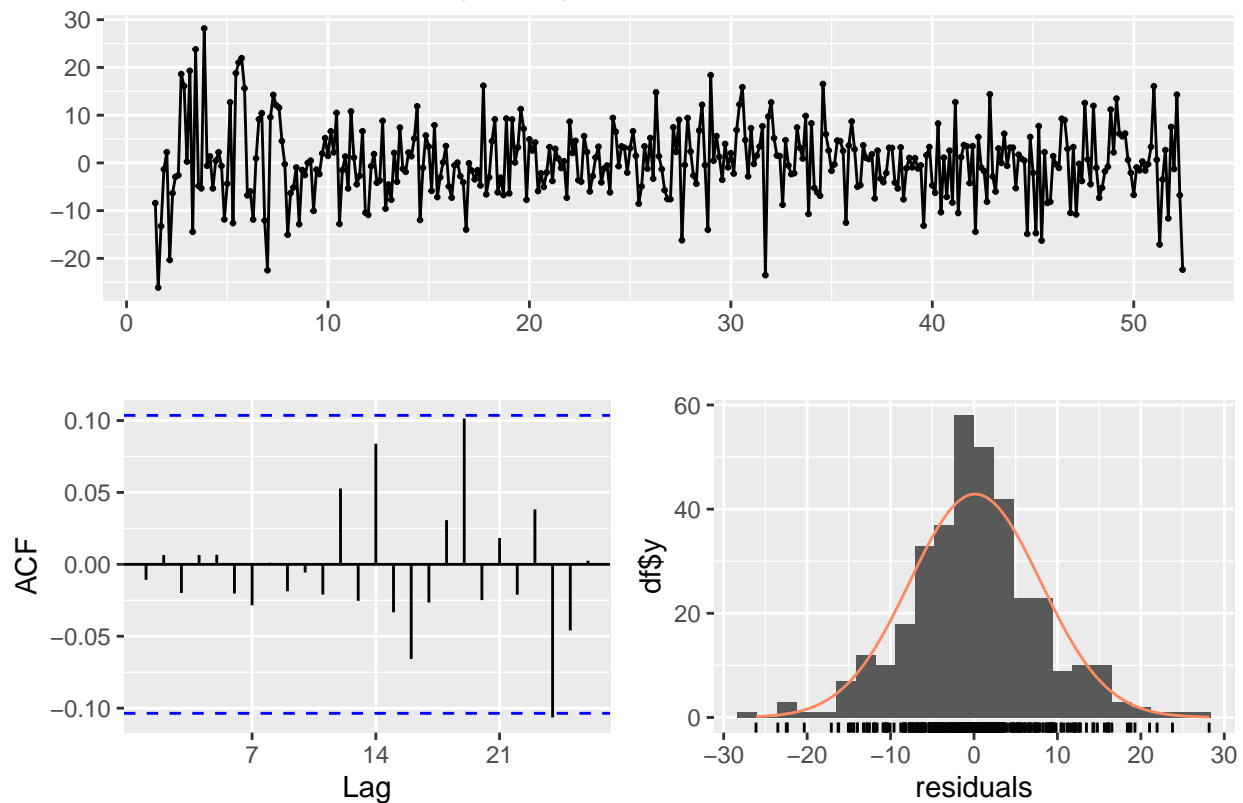
```
##
## Breusch-Godfrey test for serial correlation of order up to 14
##
## data: Residuals from Linear regression model
## LM test = 138.73, df = 14, p-value < 2.2e-16
tsdisplay(fit2$residuals)
```

The ACP and PACF look like those of an AR_{11} model: exponential decrease of the ACF and significant PCA at lag 11. We can test it:

```
tmp=fit2$residuals
fit3=Arima(tmp,order=c(11,0,0))
checkresiduals(fit3)
```

Residuals from ARIMA(11,0,0) with non-zero mean

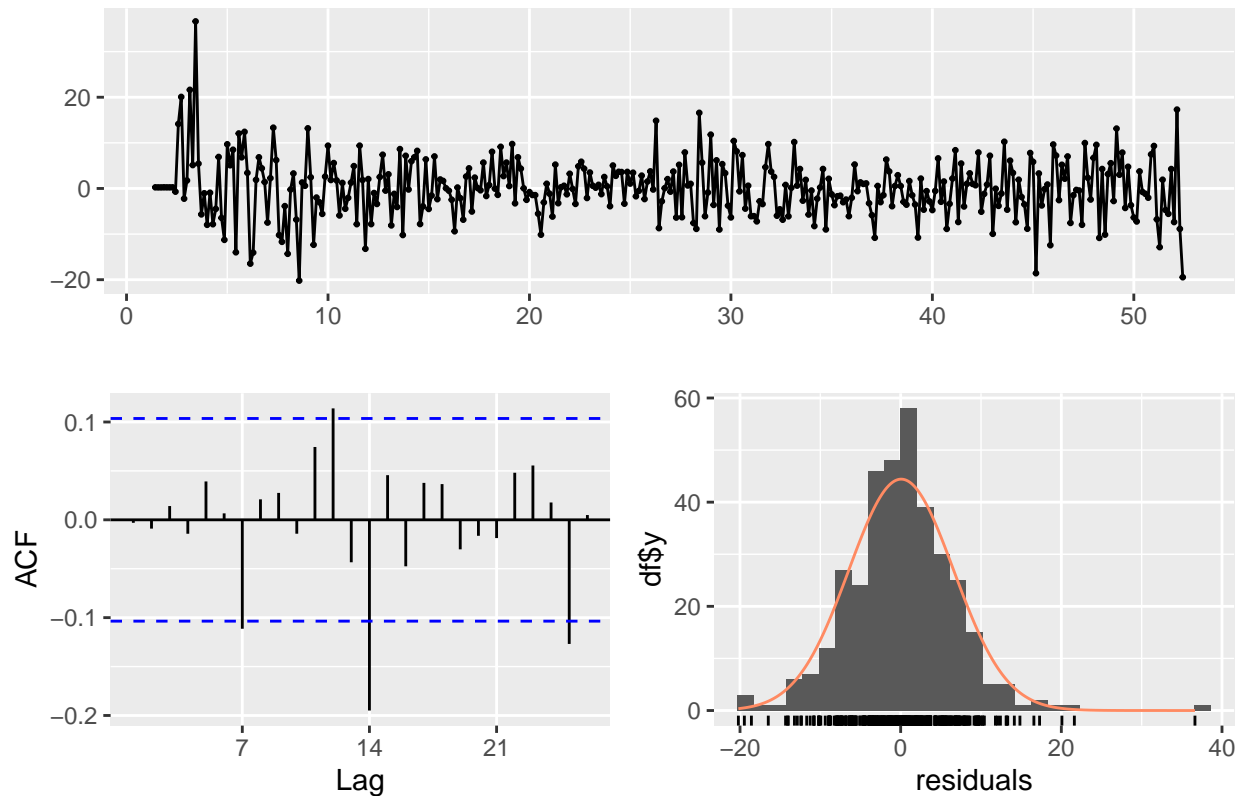


```
##
##  Ljung-Box test
##
## data:  Residuals from ARIMA(11,0,0) with non-zero mean
## Q* = 4.8998, df = 3, p-value = 0.1793
##
## Model df: 11.    Total lags used: 14
```

The correspondant residuals are uncorrelated. We can now go back to the initial series, and try to propose the following model (where `seasonal=c(0,1,0)` is for taking into account season+trend):

```
fit=Arima(elec_train[, "Demand"], xreg=elec_train[, 2:4], order=c(11,0,0), seasonal = c(0,1,0))
checkresiduals(fit)
```

Residuals from Regression with ARIMA(11,0,0)(0,1,0)[7] errors

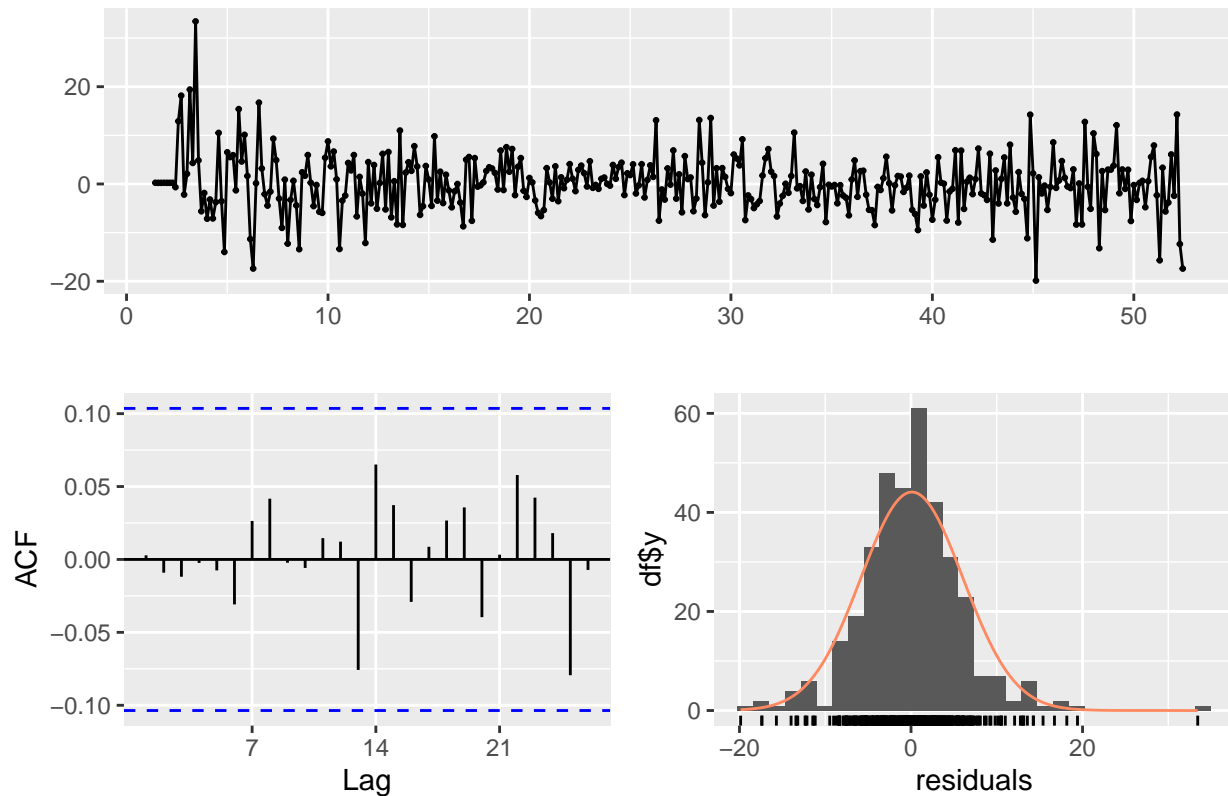


```
##
##  Ljung-Box test
##
## data:  Residuals from Regression with ARIMA(11,0,0)(0,1,0)[7] errors
## Q* = 27.651, df = 3, p-value = 4.299e-06
##
## Model df: 11.    Total lags used: 14
```

Residual have significant ACF at periodic lag (14). We will add a second order MA in the seasonal pattern:

```
fit=Arima(elec_train[, "Demand"], xreg=elec_train[, 2:4], order=c(11,0,0), seasonal = c(0,1,2))
checkresiduals(fit)
```

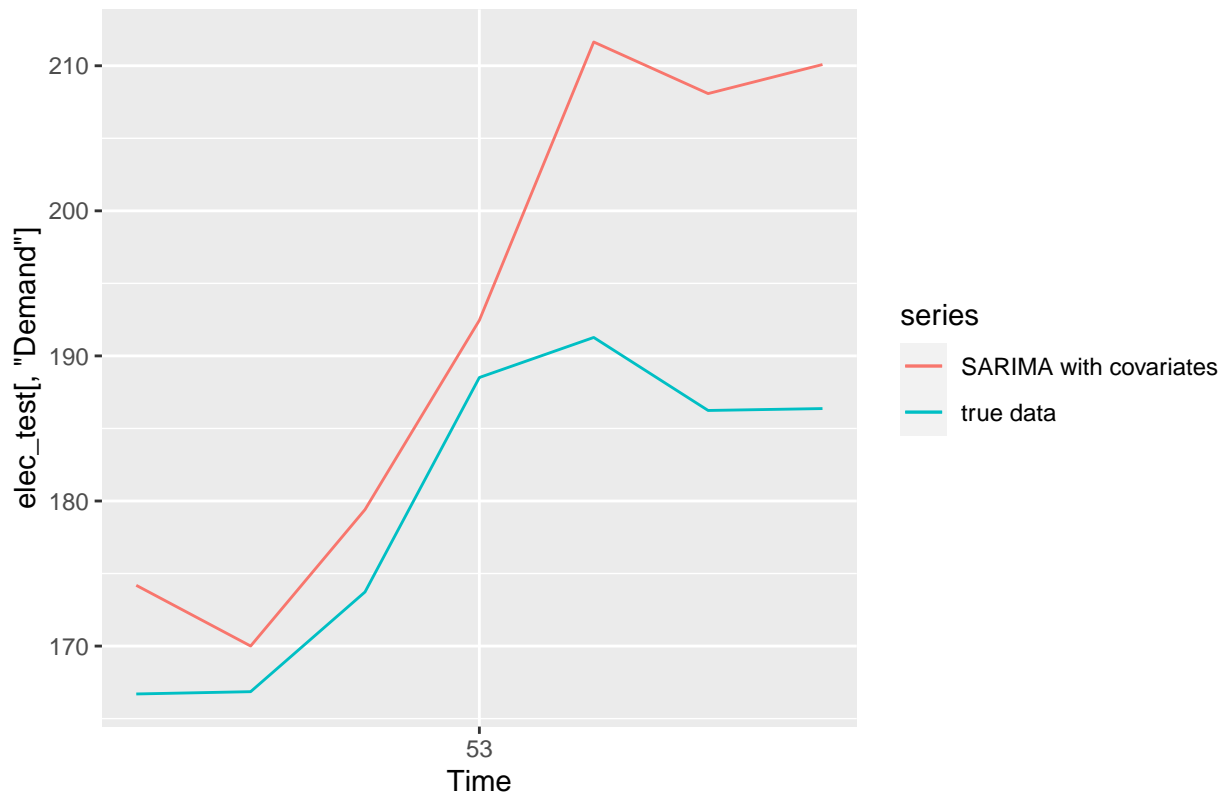
Residuals from Regression with ARIMA(11,0,0)(0,1,2)[7] errors



```
##
##  Ljung-Box test
##
## data:  Residuals from Regression with ARIMA(11,0,0)(0,1,2)[7] errors
## Q* = 6.0722, df = 3, p-value = 0.1082
##
## Model df: 13.    Total lags used: 16
```

Now residuals are uncorrelated. We can then perform forecasting:

```
elec_test=cbind(Demand=elec_test[,1],WorkDay=elec_test[,2],Temp=elec_test[,3],SquareTemp=elec_test[,3]^2)
prev=forecast(fit,h=7,xreg=elec_test[,2:4])
autoplot(elec_test[, "Demand"], series="true data")+autolayer(prev$mean,series="SARIMA with covariates")
```



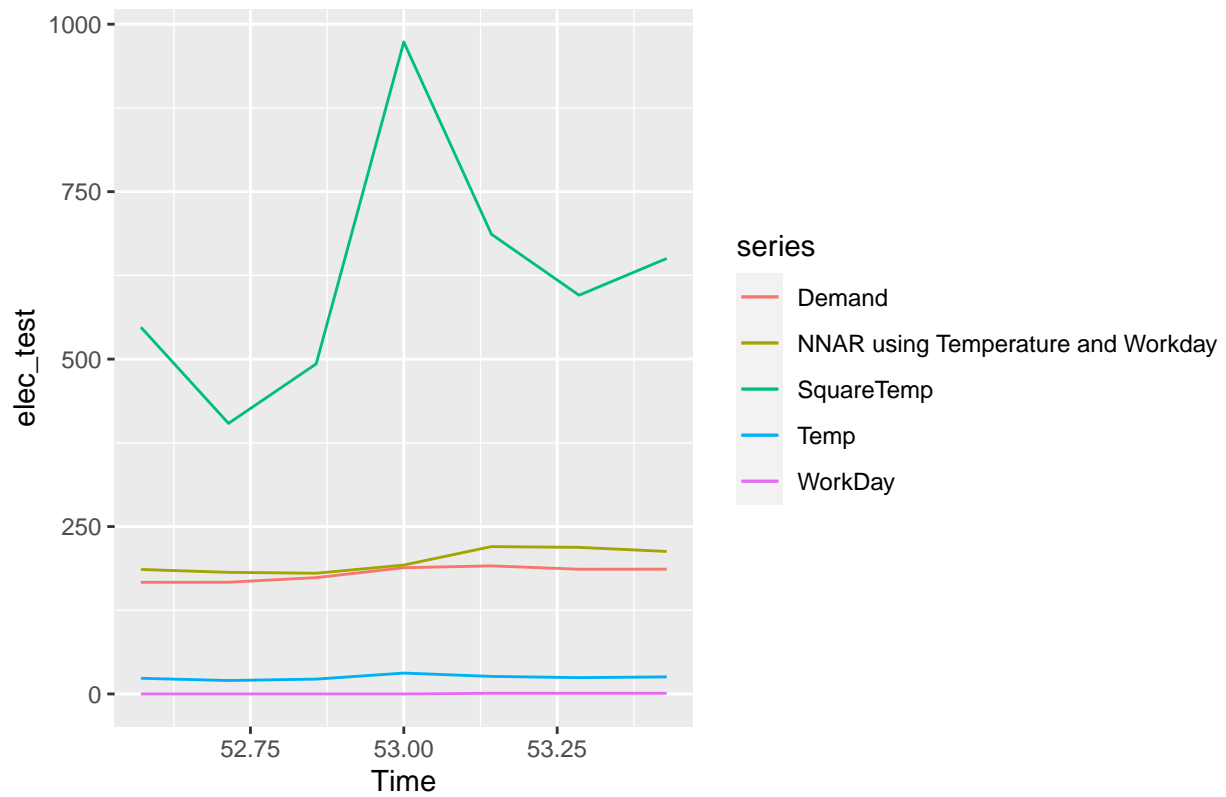
```
print(sqrt(mean((prev$mean-elec_test[, "Demand"])^2)))
```

```
## [1] 14.96637
```

The result are better than those obtained with the `auto.arima` function.

Finally, we can compare with a NNAR model with covariates, but it does not improve the forecast.

```
fit=nnetar(elec_train[, "Demand"], xreg=elec_train[, 2:4])
prev=forecast(fit, h=7, xreg=elec_test[, 2:4])
autoplot(elec_test)+autolayer(prev$mean, series="NNAR using Temperature and Workday")
```



```
print(sqrt(mean((prev$mean-elec_test[, "Demand"])^2)))
```

```
## [1] 21.50412
```