## Time series forecasting Multivariate time series

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Dynamic regression model

Machine learning model with covariates

Grouped time series models: VAR models

Let assume that you want to explain your serie  $x_t$  according to k features  $z_{1t}, \ldots, z_{kt}$ :

$$x_t = c + \beta_1 z_{1t} + \ldots + \beta_k z_{kt} + \epsilon_t.$$

Usual linear regression model assume that the error  $\epsilon_t$  are independent and identically distributed according:  $\epsilon_t \sim \mathcal{N}(0, \sigma^2)$ .

Such model can be estimated with the usual  ${\tt lm}$  function or with

?tslm

In addition to the effect of external features, times series often contain:

▶ a **trend**. A linear model including a linear trend can be written:

$$x_t = \underbrace{c + \beta_0 t}_{\text{trend}} + \underbrace{\beta_1 z_{1t} + \ldots + \beta_k z_{kt}}_{\text{covariates}} + \epsilon_t.$$

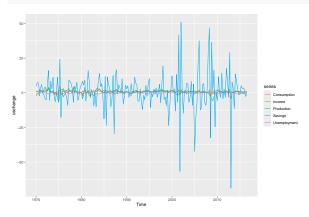
▶ a **seasonal pattern** of period *T*. Corresponding regression model is:

$$x_t = \underbrace{c + \beta_0 t}_{\text{trend}} + \underbrace{\delta_2 d_{2t} + \ldots + \delta_T d_{Tt}}_{\text{seasonal effect}} + \underbrace{\beta_1 z_{1t} + \ldots + \beta_k z_{kt}}_{\text{covariates}} + \epsilon_t.$$

where  $d_{2t},\ldots,d_{Tt}$  are the dummy notations for the T-1 days of the period:  $d_{jt}=1$  if j=t and 0 otherwise. Note that the effect of the first day  $d_{1t}$  is included in the intercept, so  $d_{jt}$  is the additional effet of day j in comparison with day 1.

Let's go back to the uschange time series

library(fpp2)
autoplot(uschange)



We want to predict Consumption using other times series

```
##
## Call:
## tslm(formula = Consumption ~ Income + Production + Unemployment +
      Savings, data = uschange)
##
## Residuals:
##
       Min
                1Q Median
                                         Max
## -0.88296 -0.17638 -0.03679 0.15251 1.20553
##
## Coefficients:
               Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) 0.26729 0.03721 7.184 1.68e-11 ***
            0.71449 0.04219 16.934 < 2e-16 ***
## Income
## Production 0.04589 0.02588 1.773 0.0778 .
## Unemployment -0.20477 0.10550 -1.941 0.0538 .
## Savings
            -0.04527 0.00278 -16.287 < 2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.3286 on 182 degrees of freedom
## Multiple R-squared: 0.754, Adjusted R-squared: 0.7486
## F-statistic: 139.5 on 4 and 182 DF, p-value: < 2.2e-16
```

We can add a trend and a seasonnal pattern

```
\label{lem:consumption-Income+Production+Unemployment+Savings+trend+season, data=uschange) summary (fit)
```

```
##
## Call:
## tslm(formula = Consumption ~ Income + Production + Unemployment +
      Savings + trend + season, data = uschange)
##
##
## Residuals:
                10 Median
##
       Min
## -0.88653 -0.15100 -0.00713 0.14232 1.10178
##
## Coefficients:
               Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) 0.4535889 0.0717294 6.324
                                             26-09 ***
## Income
           0.7093775 0.0419836 16.897 <2e-16 ***
## Production 0.0389018 0.0264104 1.473 0.1425
## Unemployment -0.2396921 0.1096766 -2.185 0.0302 *
## Savings
              -0.0450622 0.0027690 -16.274 <2e-16 ***
## trend
              -0.0010066 0.0004616 -2.181 0.0305 *
             -0.1294052 0.0669461 -1.933 0.0548 .
## season2
            -0.0602444 0.0671966 -0.897 0.3712
## season3
## season4
              -0.1495544 0.0675787 -2.213 0.0282 *
## ---
## Signif, codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.322 on 178 degrees of freedom
## Multiple R-squared: 0.769. Adjusted R-squared: 0.7586
## F-statistic: 74.06 on 8 and 178 DF, p-value: < 2.2e-16
```

#### Feature selection

As for any multivariate regression model, we have to select which are the best features to include in the model.

Comparison between models can be done with usual criteria (AIC, AICc, BIC, adjusted  $R^2, \ldots$ )

Those criterion can be obtained as follows:

```
CV(fit)

## CV AIC AICC BIC AdjR2
```

```
## CV AIC AICc BIC AdjR2
## 0.1141794 -413.0495738 -411.7995738 -380.7384876 0.7585933
```

#### Feature selection

In the previous model we have seen that Production is not significant in the model.

We can remove it and compare the model to the previous one CV(fit)

```
## CV AIC AICC BIC AdjR2
## 0.1141794 -413.0495738 -411.7995738 -380.7384876 0.7585933
fit2=tslm(Consumption~Income+Unemployment+Savings+trend+season,data=uscCV(fit2)
```

```
## CV AIC AICc BIC AdjR2
## 0.1136653 -412.7840112 -411.7670620 -383.7040336 0.7570159
```

There is no evident difference between these models (better CV, AIC, AICc and BIC, but worse  $AdjR^2$ ).

#### Feature selection

Stepwise selection procedure should be used to correctly select the best set of features.

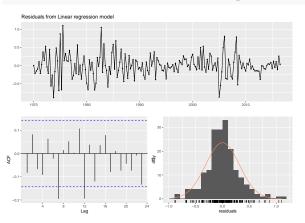
To the best of my knowledge, such procedures are not available for the tslm function.

But you can use the 1m function (with time and season as covariates), and use usual stepwise function for 1m as step or stepAIC.

#### Checking the residuals

Usual checking of linear model can/should be done:

checkresiduals(fit,test=FALSE,plot=TRUE)



#### Checking the residuals

including test of non correlation of the residuals

```
checkresiduals(fit,test='LB',plot=FALSE)
```

```
##
## Ljung-Box test
##
## data: Residuals from Linear regression model
## Q* = 13.5, df = 8, p-value = 0.09577
##
## Model df: 0. Total lags used: 8
```

Here the residual are correlated, which means that this regression model (which assumes independent residuals) is not appropriated.

We saw in the previous model:

$$x_t = c + \beta_1 z_{1t} + \ldots + \beta_k z_{kt} + \epsilon_t.$$

that the residuals  $\epsilon_t$  are not independent.

Dynamic regression model modelizes the residuals with an  $ARIMA_{p,d,q}$  model

We saw in the previous model:

$$x_t = c + \beta_1 z_{1t} + \ldots + \beta_k z_{kt} + \epsilon_t.$$

that the residuals  $\epsilon_t$  are not independent.

### Dynamic regression model modelizes the residuals with an $ARIMA_{p,d,q}$ model

The choice of the orders p, d, q can be done by examining the residuals or automatically with the auto.arima function.

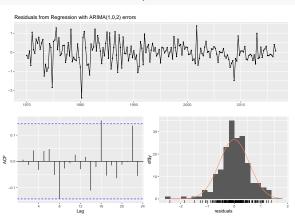
#### Let's try for instance an $ARIMA_{1,0,2}$ :

```
summarv(fit)
## Series: uschange[, "Consumption"]
## Regression with ARIMA(1,0,2) errors
##
## Coefficients:
           ar1
##
                    ma1
                           ma2 intercept
                                             xreg
       0.6922 -0.5758 0.1984
                                   0.5990 0.2028
##
## s.e. 0.1159 0.1301 0.0756 0.0884 0.0461
##
## sigma^2 = 0.3219: log likelihood = -156.95
## AIC=325.91 AICc=326.37 BIC=345.29
##
## Training set error measures:
##
                        ME
                               RMSE
                                          MAE
                                                  MPE
                                                         MAPE
                                                                   MASE
## Training set 0.001714366 0.5597088 0.4209056 27.4477 161.8417 0.6594731
##
                      ACF1
## Training set 0.006299231
```

fit=Arima(uschange[, 'Consumption'], xreg=uschange[, 'Income'], order=c(1,0,2))

We can now check the residuals:

checkresiduals(fit,test=FALSE)



and test their autocorrelation:

```
checkresiduals(fit,plot=FALSE)
```

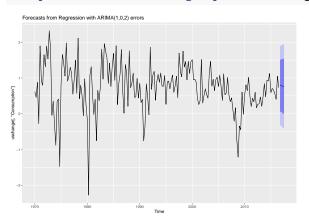
```
##
## Ljung-Box test
##
## data: Residuals from Regression with ARIMA(1,0,2) errors
## Q* = 5.8916, df = 5, p-value = 0.3169
##
## Model df: 3. Total lags used: 8
```

It seems that all the auto-correlations of the residuals have been modelled with this model.

The model being validated, we can forecast the future !

**Warning**: since we use covariate, we should have the value of the covariate for the future !

autoplot(forecast(fit,xreg=rep(mean(uschange[,2]),4)))



# Machine learning model with covariates

#### Machine learning model with covariates

Any machine learning model we saw before can include covariates.

▶ for Neural Network it is include in the nnetar function:

```
fit=nnetar(uschange[,'Consumption'],xreg=uschange[,'Income'])
```

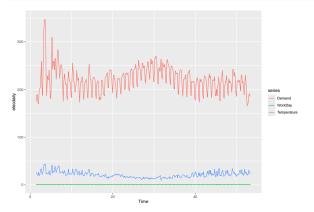
▶ for Random Forest, XGBoost, Support Vector Regression, you just have to add the covariates (at the same time, or at previous times...) as you enter the previous values of the series to forecast when you build the desing (inputs/output) matrix

	inputs			covariates z <sub>T</sub>		output
<i>x</i> <sub>1</sub>	• • •	$x_T$	• • •	ΖŢ	$z_{T+1}$	$X_{T+1}$
$X_{n-T}$		$x_{n-1}$		$z_{n-1}$	Zn	X <sub>n</sub>

#### Exercice: Electricty demand

Try to find the best model for forecasting electricity (using or not covariates) demand

#### autoplot(elecdaily)



Forecasting efficiency will be evaluated on the last 14 days, and will assume that we dispose of a forecasting of the Temperature for the next 14 days (WorkDay are of course also known).



#### VAR models

- ▶ Data : bivariate time series  $(X_{1,t}, X_{2,t})$ .
- We want to forecast both time series
- ► The idea is that each time series can help in forecasting the other one.
- Vectoriel Auto-Regressive model VAR<sub>1</sub>:

$$X_{1,t} = c_1 + \epsilon_{1,t} + a_{1,1}X_{1,t-1} + a_{1,2}X_{2,t-1},$$
  

$$X_{2,t} = c_2 + \epsilon_{2,t} + a_{2,1}X_{1,t-1} + a_{2,2}X_{2,t-1},$$

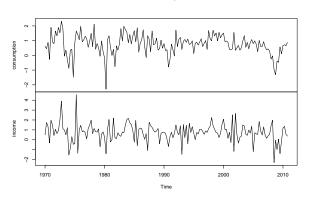
► High order model can be also considered VAR<sub>p</sub>

#### Data usconsumption

#### We will work with data usconsumption

```
library(fpp)
data(usconsumption)
plot(usconsumption)
```

#### usconsumption



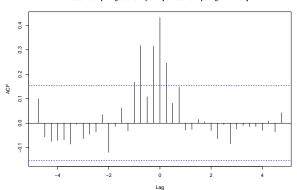
#### Cross correlation

We can compute the cross-covariance (or cross-correlation) between these two time series:

$$\widehat{ccov}_n(h) = \frac{1}{n-h} \sum_{t=1}^{n-h} (x_t - \bar{x}_n)(y_{t+h} - \bar{y}_n),$$

ccf(usconsumption[,"consumption"],usconsumption[,"income"])

#### usconsumption[, "consumption"] & usconsumption[, "income"]



#### Cross correlation test

It is possible to test the significancy of the cross correlation with the Dalla et al. 2020) procedure:

```
library(testcorr)
cc.test(usconsumption[,"consumption"],usconsumption[,"income"],max.lag = 1,
      plot = FALSE.)
##
## Tests for zero cross-correlation of x and y
##
  | Lag | CC | Stand. CB(95%) | Robust CB(95%) | Lag | t | p-value | t-tilde
##
## | -1| 0.315| (-0.153, 0.153)| (-0.175, 0.175)| -1| 4.037| 0.000| 3.536
     0 | 0.432 | (-0.153, 0.153) | (-0.240, 0.240) | 0 | 5.533 | 0.000 |
                                                               3.528
## |
     1 | 0.246 | (-0.153, 0.153) | (-0.189, 0.189) | 1 | 3.145 | 0.002 |
## |
                                                                2,550
```

#### Data usconsumption

We choose two last years as test data

```
 us\_app=ts(usconsumption[1:156,],start=c(1970,1),end=c(2008,4),frequency=4) \\ us\_test=ts(usconsumption[157:164,],start=c(2009,1),end=c(2010,4),frequency=4)
```

#### $VAR_p$ model

Function VARselect allows to choose the best  $VAR_p$  model according to some criteria (among which AIC), for  $1 \le p \le lag.max$ 

```
library(vars)
VARselect(us_app, lag.max=8, type="const",season=4)
```

The option type allows to introduce a trend ("const", "trend", "both", "none"), the option season for a seasonal pattern and exogen for external covariates.

#### $VAR_p$ model

```
library(vars)
VARselect(us_app, lag.max=8, type="const")
## $selection
## AIC(n) HQ(n) SC(n) FPE(n)
##
##
## $criteria
##
## AIC(n) -1.2314131 -1.217775 -1.2505936 -1.2505388 -1.2690326
## HQ(n) -1.1820445 -1.135494 -1.1354000 -1.1024328 -1.0880142
## SC(n) -1.1099045 -1.015261 -0.9670735 -0.8860130 -0.8235011
## FPE(n) 0.2918831 0.295903 0.2863752 0.2864365 0.2812579
##
                             8
## AIC(n) -1.1932426 -1.1788984
  HQ(n) -0.9463992 -0.8991427
## SC(n) -0.5856995 -0.4903497
## FPE(n) 0.3036602 0.3082446
```

#### VAR<sub>p</sub> model

##

##

Estimation of an VAR<sub>5</sub>

```
var <- VAR(us_app, p=5,type = "const",season = 4,exogen=NULL)
summary(var)</pre>
```

```
## VAR Estimation Results:
## ==========
```

## Endogenous variables: consumption, income
## Deterministic variables: const

```
## Sample size: 151
## Log Likelihood: -306.307
```

## Roots of the characteristic polynomial: ## 0.7553 0.7553 0.7247 0.7247 0.7229 0.6198 0.6198 0.5715 0.571

```
## 0.7555 0.7555 0.7247 0.7247 0.7229 0.6198 0.6198 0.5715 0.57
## Call:
## VAR(y = us_app, p = 5, type = "const", season = 4L, exogen = ##
```

## Estimation results for equation consumption:

```
## consumption = consumption.l1 + income.l1 + consumption.l2 + i
```

#### $VAR_p$ model

```
We check that the residual are a white noise
serial.test(var, lags.pt=10, type="PT.asymptotic")

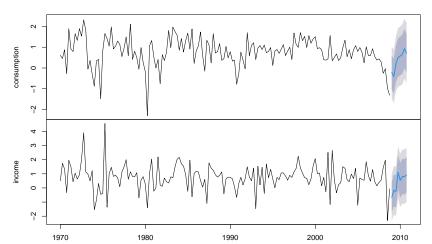
##
## Portmanteau Test (asymptotic)
##
## data: Residuals of VAR object var
## Chi-squared = 14.26, df = 20, p-value = 0.8171
```

#### $VAR_p$ model

#### Forecasting

```
fcst <- forecast(var,h=8)
plot(fcst, xlab="Year")</pre>
```

#### Forecasts from VAR(5)



#### Data usconsumption

## [1] 0.2568282

```
Forecasting efficiency with a VAR<sub>5</sub>

print(sqrt(mean(us_test[,1]-fcst$forecast$consumption$mean)^2))

## [1] 0.02851079

print(sqrt(mean(us_test[,2]-fcst$forecast$income$mean)^2))
```

#### Data usconsumption

Forecasting of consumption and income separately

```
mod1=auto.arima(us_app[,1])
pred1=forecast(mod1,h =8)
mod2=auto.arima(us_app[,2])
pred2=forecast(mod2,h =8)
print(sqrt(mean(us_test[,1]-pred1$mean)^2))
```

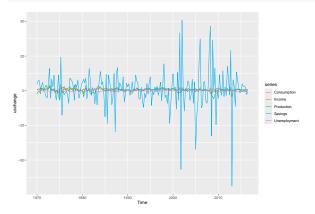
```
## [1] 0.04460754
print(sqrt(mean(us_test[,2]-pred2$mean)^2))
```

```
## [1] 0.6388838
```

Quality of prediction is lower when each times series is used separaterly.

#### Exercice: Data uschange

Try to find the best forecasting model for the 5 uschange time series autoplot(uschange)



Forecasting efficiency will be evaluated on 2016 data, and compare to forecasting each time series separately.

#### To go further

Facebook develop a kind of automatic time series modelling which appear to be relatively efficient (but I never test it). If you are interested in fast forecast, you can test it: https://cran.r-project.org/web/packages/prophet/index.html