

# Time series forecasting

## Multivariate time series

Julien JACQUES

Université Lumière Lyon 2

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## Time series regression models

# Time series regression models

Let assume that you want to explain your serie  $x_t$  according to  $k$  features  $z_{1t}, \dots, z_{kt}$ :

$$x_t = c + \beta_1 z_{1t} + \dots + \beta_k z_{kt} + \epsilon_t.$$

Usual linear regression model assume that the error  $\epsilon_t$  are independent and identically distributed according:  $\epsilon_t \sim \mathcal{N}(0, \sigma^2)$ .

Such model can be estimated with the usual `lm` function or with

```
?tslm
```

# Time series regression models

In addition to the effect of external features, times series often contain:

- ▶ a **trend**. A linear model including a linear trend can be written:

$$x_t = \underbrace{c + \beta_0 t}_{\text{trend}} + \underbrace{\beta_1 z_{1t} + \dots + \beta_k z_{kt}}_{\text{covariates}} + \epsilon_t.$$

- ▶ a **seasonal pattern** of period  $T$ . Corresponding regression model is:

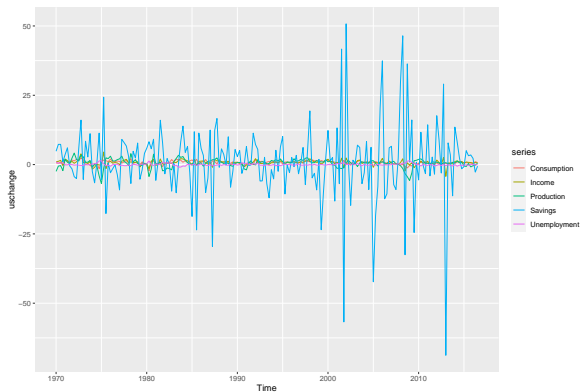
$$x_t = \underbrace{c + \beta_0 t}_{\text{trend}} + \underbrace{\delta_2 d_{2t} + \dots + \delta_T d_{Tt}}_{\text{seasonal effect}} + \underbrace{\beta_1 z_{1t} + \dots + \beta_k z_{kt}}_{\text{covariates}} + \epsilon_t.$$

where  $d_{2t}, \dots, d_{Tt}$  are the dummy notations for the  $T - 1$  days of the period:  $d_{jt} = 1$  if  $j = t$  and 0 otherwise. Note that the effect of the first day  $d_{1t}$  is included in the intercept, so  $d_{jt}$  is the additional effect of day  $j$  in comparison with day 1.

# Time series regression models

Let's go back to the uschange time series

```
library(fpp2)
autoplot(uschange)
```



# Time series regression models

We want to predict Consumption using other times series

```
fit=tslm(Consumption~Income+Production+Unemployment+Savings,data=uschange)
summary(fit)
```

```
##
## Call:
## tslm(formula = Consumption ~ Income + Production + Unemployment +
##       Savings, data = uschange)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -0.88296 -0.17638 -0.03679  0.15251  1.20553
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   0.26729    0.03721   7.184 1.68e-11 ***
## Income        0.71449    0.04219  16.934 < 2e-16 ***
## Production    0.04589    0.02588   1.773  0.0778 .
## Unemployment -0.20477    0.10550  -1.941  0.0538 .
## Savings      -0.04527    0.00278 -16.287 < 2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.3286 on 182 degrees of freedom
## Multiple R-squared:  0.754, Adjusted R-squared:  0.7486
## F-statistic: 139.5 on 4 and 182 DF,  p-value: < 2.2e-16
```

# Time series regression models

We can add a trend and a seasonal pattern

```
fit=tslm(Consumption~Income+Production+Unemployment+Savings+trend+season,data=uschange)
summary(fit)
```

```
##
## Call:
## tslm(formula = Consumption ~ Income + Production + Unemployment +
##       Savings + trend + season, data = uschange)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -0.88653 -0.15100 -0.00713  0.14232  1.10178
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  0.4535889  0.0717294   6.324   2e-09 ***
## Income       0.7093775  0.0419836  16.897  <2e-16 ***
## Production   0.0389018  0.0264104   1.473   0.1425
## Unemployment -0.2396921  0.1096766  -2.185   0.0302 *
## Savings      -0.0450622  0.0027690 -16.274  <2e-16 ***
## trend        -0.0010066  0.0004616  -2.181   0.0305 *
## season2      -0.1294052  0.0669461  -1.933   0.0548 .
## season3      -0.0602444  0.0671966  -0.897   0.3712
## season4      -0.1495544  0.0675787  -2.213   0.0282 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.322 on 178 degrees of freedom
## Multiple R-squared:  0.769, Adjusted R-squared:  0.7586
## F-statistic: 74.06 on 8 and 178 DF, p-value: < 2.2e-16
```



## Feature selection

As for any multivariate regression model, we have to select which are the best features to include in the model.

Comparison between models can be done with usual criteria (AIC, AICc, BIC, adjusted  $R^2$ , ...)

Those criterion can be obtained as follows:

```
CV(fit)
```

##	CV	AIC	AICc	BIC	AdjR2
##	0.1141794	-413.0495738	-411.7995738	-380.7384876	0.7585933

## Feature selection

In the previous model we have seen that Production is not significant in the model.

We can remove it and compare the model to the previous one

```
fit2=tslm(Consumption~Income+Unemployment+Savings+trend+season,data=usc  
CV(fit)
```

##	CV	AIC	AICc	BIC	AdjR2
##	0.1141794	-413.0495738	-411.7995738	-380.7384876	0.7585933

```
CV(fit2)
```

##	CV	AIC	AICc	BIC	AdjR2
##	0.1136653	-412.7840112	-411.7670620	-383.7040336	0.7570159

There is no evident difference between these models (better CV, AIC, AICc and BIC, but worse  $\text{Adj}R^2$ ).

## Feature selection

Stepwise selection procedure should be used to correctly select the best set of features.

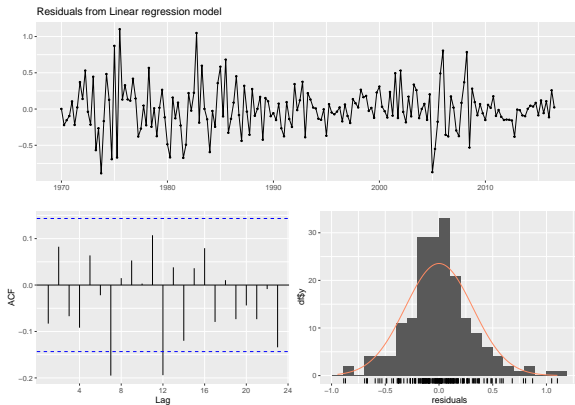
To the best of my knowledge, such procedures are not available for the `tslm` function.

But you can use the `lm` function (with time and season as covariates), and use usual stepwise function for `lm` as `step` or `stepAIC`.

# Checking the residuals

Usual checking of linear model can/should be done:

```
checkresiduals(fit,test=FALSE,plot=TRUE)
```



## Checking the residuals

including test of non correlation of the residuals

```
checkresiduals(fit,test='LB',plot=FALSE)
```

```
##  
##  Ljung-Box test  
##  
## data:  Residuals from Linear regression model  
## Q* = 13.5, df = 8, p-value = 0.09577  
##  
## Model df: 0.    Total lags used: 8
```

Here the residual are correlated, which means that this regression model (which assumes independent residuals) is not appropriated.

## Dynamic regression model

# Dynamic regression model

We saw in the previous model:

$$x_t = c + \beta_1 z_{1t} + \dots + \beta_k z_{kt} + \epsilon_t.$$

that the residuals  $\epsilon_t$  are not independent.

**Dynamic regression model** modelizes the **residuals with an  $ARIMA_{p,d,q}$  model**

# Dynamic regression model

We saw in the previous model:

$$x_t = c + \beta_1 z_{1t} + \dots + \beta_k z_{kt} + \epsilon_t.$$

that the residuals  $\epsilon_t$  are not independent.

**Dynamic regression model** modelizes the **residuals with an  $ARIMA_{p,d,q}$  model**

The choice of the orders  $p, d, q$  can be done by examining the residuals or automatically with the `auto.arima` function.



# Dynamic regression model

Let's try for instance an  $ARIMA_{1,0,2}$ :

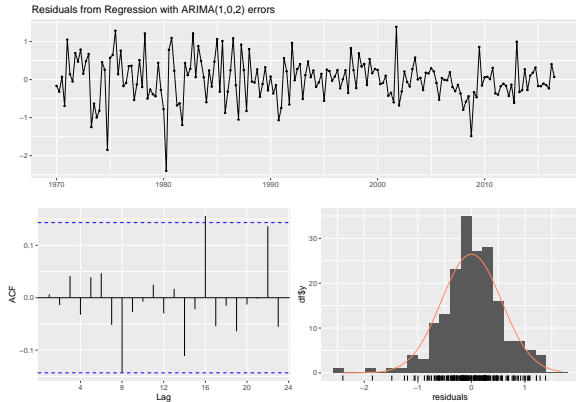
```
fit=Arima(uschange[, 'Consumption'], xreg=uschange[, 'Income'], order=c(1,0,2))  
summary(fit)
```

```
## Series: uschange[, "Consumption"]  
## Regression with ARIMA(1,0,2) errors  
##  
## Coefficients:  
##          ar1          ma1          ma2  intercept          xreg  
##          0.6922   -0.5758   0.1984          0.5990   0.2028  
## s.e.    0.1159    0.1301   0.0756          0.0884   0.0461  
##  
## sigma^2 = 0.3219:  log likelihood = -156.95  
## AIC=325.91   AICc=326.37   BIC=345.29  
##  
## Training set error measures:  
##              ME          RMSE          MAE          MPE          MAPE          MASE  
## Training set 0.001714366 0.5597088 0.4209056 27.4477 161.8417 0.6594731  
##              ACF1  
## Training set 0.006299231
```

# Dynamic regression model

We can now check the residuals:

```
checkresiduals(fit, test=FALSE)
```



# Dynamic regression model

and test their autocorrelation:

```
checkresiduals(fit,plot=FALSE)
```

```
##  
##  Ljung-Box test  
##  
## data:  Residuals from Regression with ARIMA(1,0,2) errors  
## Q* = 5.8916, df = 5, p-value = 0.3169  
##  
## Model df: 3.    Total lags used: 8
```

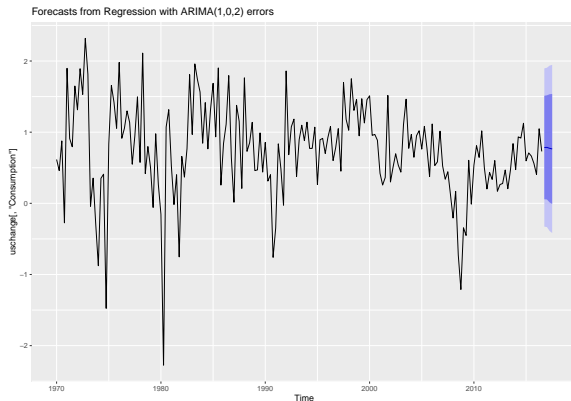
It seems that all the auto-correlations of the residuals have been modelled with this model.

# Dynamic regression model

The model being validated, we can forecast the future !

**Warning:** since we use covariate, we should have the value of the covariate for the future !

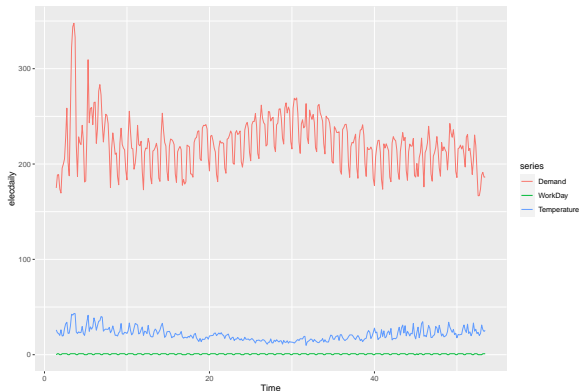
```
autoplot(forecast(fit,xreg=rep(mean(uschange[,2]),4)))
```



## Exercise: Electricity demand

Try to find the best model for forecasting electricity (using or not covariates) demand

```
autoplot(elecdaily)
```



Forecasting efficiency will be evaluated on the last 7 days, and will assume that we dispose of a forecasting of the Temperature for the next 7 days (WorkDay are of course also known).

Grouped time series models: VAR models

# VAR models

- ▶ Data : bivariate time series  $(X_{1,t}, X_{2,t})$ .
- ▶ We want to **forecast both time series**
- ▶ The idea is that each time series can help in forecasting the other one.
- ▶ **Vectoriel Auto-Regressive** model  $VAR_1$  :

$$X_{1,t} = c_1 + \epsilon_{1,t} + a_{1,1}X_{1,t-1} + a_{1,2}X_{2,t-1},$$

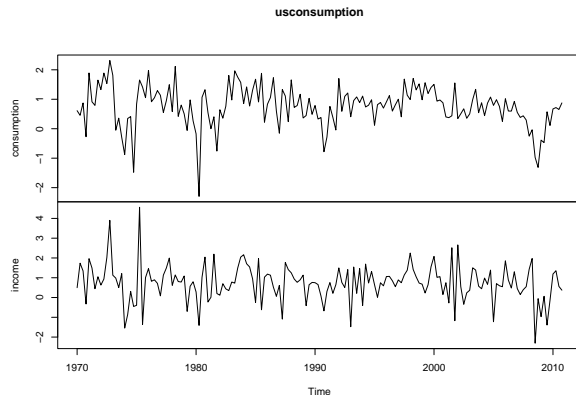
$$X_{2,t} = c_2 + \epsilon_{2,t} + a_{2,1}X_{1,t-1} + a_{2,2}X_{2,t-1},$$

- ▶ High order model can be also considered  $VAR_p$

# Data usconsumption

We will work with data usconsumption

```
library(fpp)
data(usconsumption)
plot(usconsumption)
```





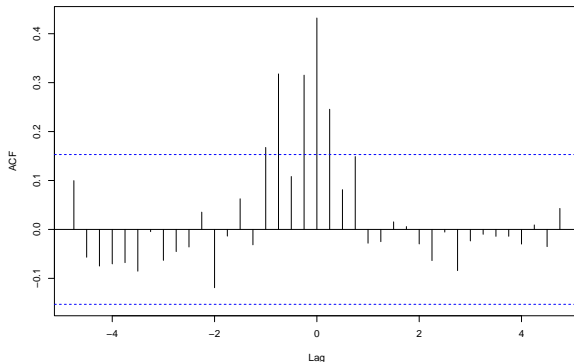
# Cross correlation

We can compute the cross-covariance (or cross-correlation) between these two time series:

$$\widehat{ccov}_n(h) = \frac{1}{n-h} \sum_{t=1}^{n-h} (x_t - \bar{x}_n)(y_{t+h} - \bar{y}_n),$$

```
ccf(usconsumption[, "consumption"], usconsumption[, "income"])
```

usconsumption[, "consumption"] & usconsumption[, "income"]



# Cross correlation test

It is possible to test the significance of the cross correlation with the Dalla et al. 2020) procedure:

```
library(testcorr)
cc.test(usconsumption[, "consumption"], usconsumption[, "income"], max.lag = 1,
        plot = FALSE,)
```

```
##
```

```
## Tests for zero cross-correlation of x and y
```

```
##
```

##	Lag	CC	Stand. CB(95%)	Robust CB(95%)	Lag	t	p-value	t-tilde
##	---:	-----:	-----:	-----:	---:	-----:	-----:	-----:
##	-1	0.315	(-0.153, 0.153)	(-0.175, 0.175)	-1	4.037	0.000	3.536
##	0	0.432	(-0.153, 0.153)	(-0.240, 0.240)	0	5.533	0.000	3.528
##	1	0.246	(-0.153, 0.153)	(-0.189, 0.189)	1	3.145	0.002	2.550

# Data usconsumption

We choose two last years as test data

```
us_app=ts(usconsumption[1:156,],start=c(1970,1),end=c(2008,4),frequency = 4)  
us_test=ts(usconsumption[157:164,],start=c(2009,1),end=c(2010,4),frequency = 4)
```

## $VAR_p$ model

Function VARselect allows to choose the best  $VAR_p$  model according to some criteria (among which AIC), for  $1 \leq p \leq lag.max$

```
library(vars)  
VARselect(us_app, lag.max=8, type="const", season=4)
```

The option type allows to introduce a trend ("const", "trend", "both", "none"), the option season for a seasonal pattern and exogen for external covariates.

## VAR<sub>p</sub> model

```
library(vars)
VARselect(us_app, lag.max=8, type="const")
```

```
## $selection
## AIC(n)   HQ(n)   SC(n) FPE(n)
##      5      1      1      5
##
## $criteria
##              1              2              3              4              5
## AIC(n) -1.2314131 -1.217775 -1.2505936 -1.2505388 -1.2690326
## HQ(n)  -1.1820445 -1.135494 -1.1354000 -1.1024328 -1.0880142
## SC(n)   -1.1099045 -1.015261 -0.9670735 -0.8860130 -0.8235011
## FPE(n)   0.2918831  0.295903  0.2863752  0.2864365  0.2812579
##
##              7              8
## AIC(n) -1.1932426 -1.1788984
## HQ(n)  -0.9463992 -0.8991427
## SC(n)   -0.5856995 -0.4903497
## FPE(n)   0.3036602  0.3082446
```

## VAR<sub>p</sub> model

Estimation of an VAR<sub>5</sub>

```
var <- VAR(us_app, p=5,type = "const",season = 4,exogen=NULL)
summary(var)
```

```
##
```

```
## VAR Estimation Results:
```

```
## =====
```

```
## Endogenous variables: consumption, income
```

```
## Deterministic variables: const
```

```
## Sample size: 151
```

```
## Log Likelihood: -306.307
```

```
## Roots of the characteristic polynomial:
```

```
## 0.7553 0.7553 0.7247 0.7247 0.7229 0.6198 0.6198 0.5715 0.571
```

```
## Call:
```

```
## VAR(y = us_app, p = 5, type = "const", season = 4L, exogen =
```

```
##
```

```
##
```

```
## Estimation results for equation consumption:
```

```
## =====
```

```
## consumption = consumption.l1 + income.l1 + consumption.l2 + i
```

## $VAR_p$ model

We check that the residual are a white noise

```
serial.test(var, lags.pt=10, type="PT.asymptotic")
```

```
##
```

```
##  Portmanteau Test (asymptotic)
```

```
##
```

```
## data:  Residuals of VAR object var
```

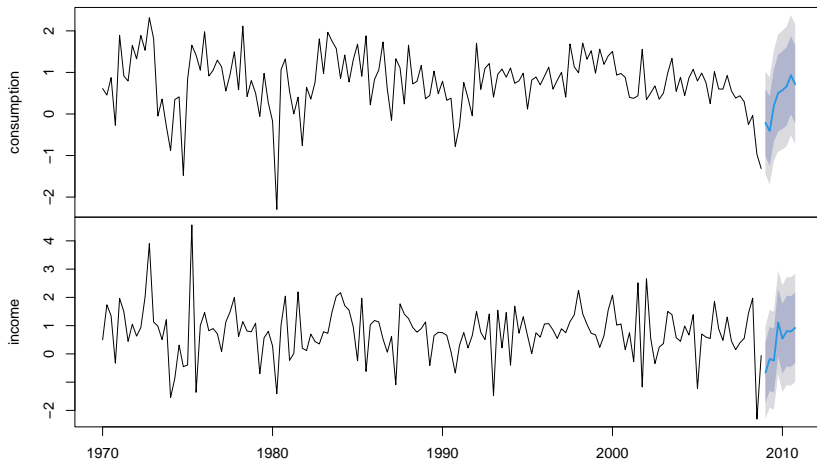
```
## Chi-squared = 14.26, df = 20, p-value = 0.8171
```

# $VAR_p$ model

## Forecasting

```
fcst <- forecast(var,h=8)  
plot(fcst, xlab="Year")
```

Forecasts from VAR(5)





# Data usconsumption

Forecasting efficiency with a  $VAR_5$

```
print(sqrt(mean(us_test[,1]-fcst$forecast$consumption$mean)^2))
```

```
## [1] 0.02851079
```

```
print(sqrt(mean(us_test[,2]-fcst$forecast$income$mean)^2))
```

```
## [1] 0.2568282
```

# Data usconsumption

Forecasting of consumption and income separately

```
mod1=auto.arima(us_app[,1])  
pred1=forecast(mod1,h =8)  
mod2=auto.arima(us_app[,2])  
pred2=forecast(mod2,h =8)  
print(sqrt(mean(us_test[,1]-pred1$mean)^2))
```

```
## [1] 0.04460754
```

```
print(sqrt(mean(us_test[,2]-pred2$mean)^2))
```

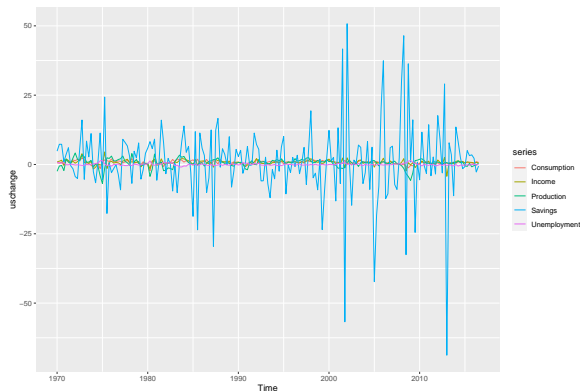
```
## [1] 0.6388838
```

Quality of prediction is lower when each times series is used separately.

## Exercise: Data uschange

Try to find the best forecasting model for the 5 uschange time series

```
autoplot(uschange)
```



Forecasting efficiency will be evaluated on 2016 data, and compare to forecasting each time series separately.

## To go further

Facebook develop a kind of automatic time series modelling which appear to be relatively efficient (but I never test it). If you are interested in fast forecast, you can test it :

<https://cran.r-project.org/web/packages/prophet/index.html>