

Gaussian Processes: Master Normality

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Content

- Multivariate Gaussian Distribution
- Gaussian process
- Gaussian process regression
- Hyperparameter optimization
- Benefits & Difficulties



Univariate Gaussian Density

A random variable x with density $x \sim \mathcal{N}(\mu, \sigma^2)$

$$f_x(x) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right)$$

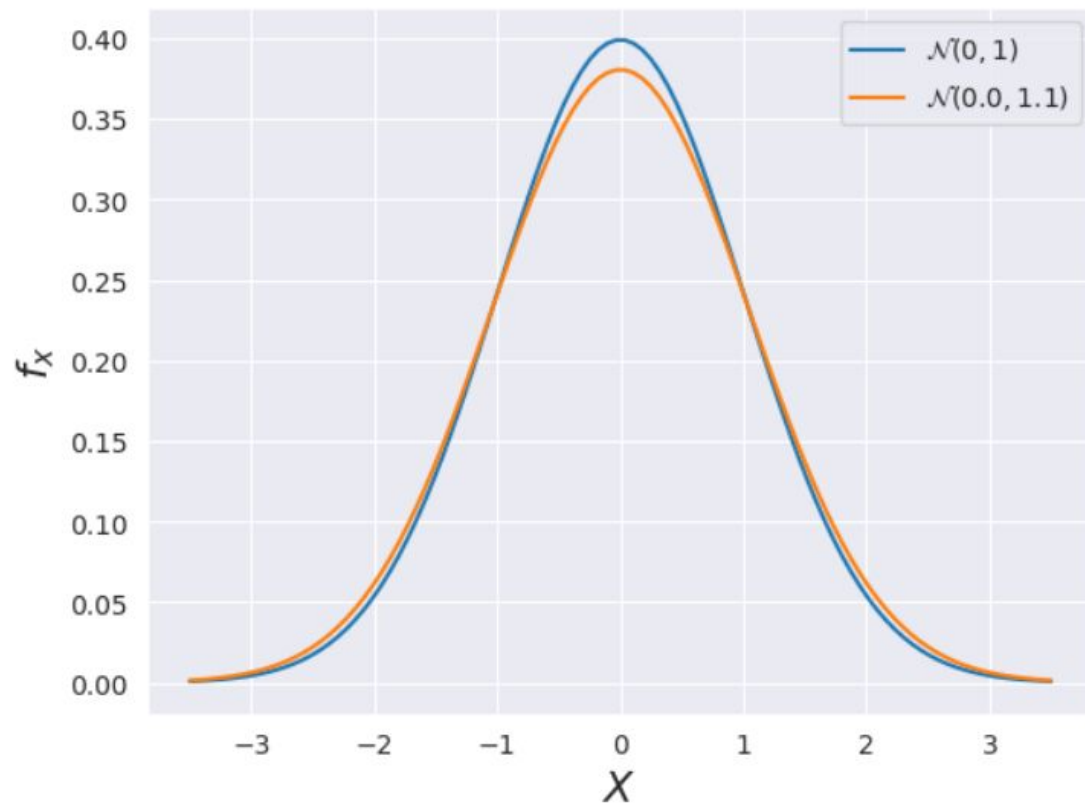
Where:

- x is the random variable.
- $\mu = \mathbb{E}[X]$ is the mean, representing the central tendency of the distribution.
- $\sigma^2 = \text{Var}(X)$ is the variance, determining the spread or dispersion of the distribution.



μ : 0.00

σ^2 : 1.10



Multivariate Gaussian Distribution

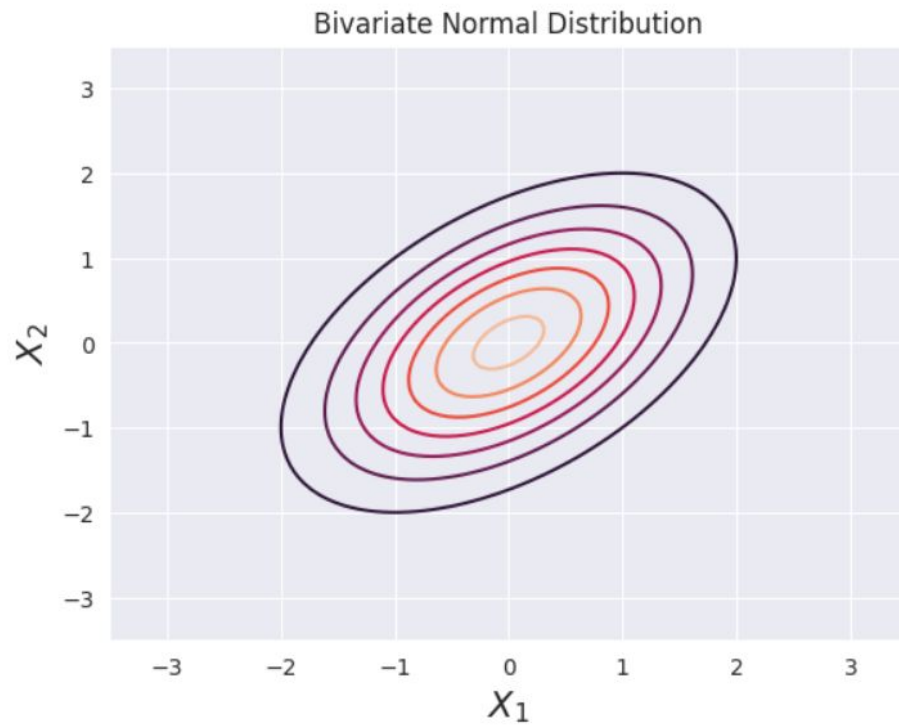
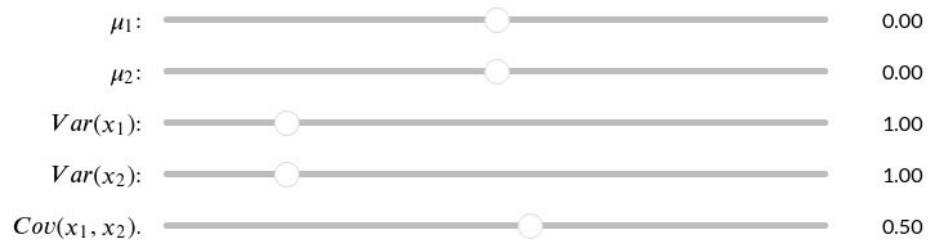
A random variable $\mathbf{x} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ presents a density function

$$f_{\mathbf{x}}(\mathbf{x}) = \frac{1}{(2\pi)^{d/2} |\boldsymbol{\Sigma}|^{1/2}} \exp\left(-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})\right)$$

Where:

- \mathbf{x} : d -dimensional vector representing the random variables.
- $\boldsymbol{\mu}$: mean vector, representing the expected value of each random variable.
- $\boldsymbol{\Sigma}$: covariance matrix, represents the relationships between random variables.





Covariance Matrix

For n random variables X_1, X_2, \dots, X_n , the multivariate covariance matrix Σ is:

$$\Sigma = \begin{bmatrix} \text{Var}(x_1) & \text{Cov}(x_1, x_2) & \dots & \text{Cov}(x_1, x_n) \\ \text{Cov}(x_2, x_1) & \text{Var}(x_2) & \dots & \text{Cov}(x_2, x_n) \\ \vdots & \vdots & \ddots & \vdots \\ \text{Cov}(x_n, x_1) & \text{Cov}(x_n, x_2) & \dots & \text{Var}(x_n) \end{bmatrix}$$

- **Symmetry:** $\text{cov}(x, x') = \text{cov}(x', x)$ for all x and x' .
- **Positive Semi-definite:** $x^T \Sigma x \geq 0$ for any vector $x \neq 0$.



Conditioning

Given the two random vectors \mathbf{x}_A and \mathbf{x}_B , the conditional probability of \mathbf{x}_A is defined as,

$$p(\mathbf{x}_A|\mathbf{x}_B) = \frac{p(\mathbf{x}_A, \mathbf{x}_B)}{p(\mathbf{x}_B)}$$

defined for $p(\mathbf{x}_B) > 0$



Exercise: Gaussian Conditioning

Assume an n -dimensional random vector has a normal distribution,

$$N\left(\begin{bmatrix} \mathbf{x} \\ \mathbf{y} \end{bmatrix}, \begin{bmatrix} \mu_X \\ \mu_Y \end{bmatrix}, \begin{bmatrix} A & C \\ C^T & B \end{bmatrix}\right)$$

where \mathbf{x}_1 and \mathbf{x}_2 are two subvectors of respective dimensions p and q with $p + q = n$.
Then, conditional distribution of \mathbf{y} given \mathbf{x} is also normal with mean vector

$$\mu_{\mathbf{y}|\mathbf{x}} = \mu_Y + C^T A^{-1}(\mathbf{x} - \mu_X)$$

and covariance matrix

$$\Sigma_{\mathbf{y}|\mathbf{x}} = B - C^T A^{-1} C$$

Proof:

The joint density of \mathbf{x} is:

$$p(\mathbf{x}, \mathbf{y}) = \frac{1}{(2\pi)^{n/2} |\Sigma|^{1/2}} \exp\left[-\frac{1}{2} Q(\tilde{\mathbf{x}})\right]$$

where Q is defined as

$$Q(\tilde{\mathbf{x}}) = (\mathbf{x} - \mu_X)^T A^{-1} (\mathbf{x} - \mu_X) + (\mathbf{y} - \mu_Y - C^T A^{-1} (\mathbf{x} - \mu_X))^T B^{-1} (\mathbf{y} - \mu_Y - C^T A^{-1} (\mathbf{x} - \mu_X))$$

$$Q(\tilde{\mathbf{x}}) = (\tilde{\mathbf{x}} - \tilde{\boldsymbol{\mu}})^T \Sigma^{-1} (\tilde{\mathbf{x}} - \tilde{\boldsymbol{\mu}}) = [(\mathbf{x} - \mu_X)^T, (\mathbf{y} - \mu_Y)^T] \begin{bmatrix} A & C \\ C^T & B \end{bmatrix} \begin{bmatrix} \mathbf{x} - \mu_X \\ \mathbf{y} - \mu_Y \end{bmatrix}$$

Here we have assumed

$$\Sigma^{-1} = \begin{bmatrix} \tilde{A} & \tilde{C} \\ \tilde{C}^T & \tilde{B} \end{bmatrix}$$

where

$$\tilde{A} = (A - CB^{-1}C^T)^{-1}C^T A^{-1}$$

$$\tilde{B} = (B - C^T A^{-1}C)^{-1}CB^{-1}$$

$$\tilde{C} = -A^{-1}C(B - C^T A^{-1}C)^{-1} = \tilde{C}^T$$

Substituting into $Q(\tilde{\mathbf{x}})$ to get:

$$Q(\tilde{\mathbf{x}}) = (\mathbf{x} - \mu_X)^T A^{-1}(\mathbf{x} - \mu_X) + [(\mathbf{y} - \mu_Y) - C^T A^{-1}(\mathbf{x} - \mu_X)]^T (B - C^T A^{-1}C)^{-1}[(\mathbf{y} - \mu_Y) - C^T A^{-1}(\mathbf{x} - \mu_X)]$$

Now the joint distribution can be written as:

$$p(\mathbf{x}, \mathbf{y}) = \frac{1}{(2\pi)^{n/2} |A|^{1/2}} \exp \left[-\frac{1}{2} Q(\tilde{\mathbf{x}}) \right] = N(\mathbf{x} | \mu_X, A) \cdot N(\mathbf{y} | b, M)$$

$$p(\mathbf{x}, \mathbf{y}) = \frac{1}{(2\pi)^{n/2} |A|^{1/2}} \exp \left[-\frac{1}{2} Q(\tilde{\mathbf{x}}) \right] = N(\mathbf{x} | \mu_X, A) \cdot N(\mathbf{y} | b, M)$$

The conditional distribution of \mathbf{y} given \mathbf{x} is

$$\begin{aligned} p(\mathbf{y} | \mathbf{x}) &= \frac{p(\mathbf{x}, \mathbf{y})}{p(\mathbf{x})} \\ &= \frac{1}{(2\pi)^{q/2} |M|^{1/2}} \exp \left[-\frac{1}{2} (\mathbf{y} - b)^T M^{-1} (\mathbf{y} - b) \right] \end{aligned}$$

with

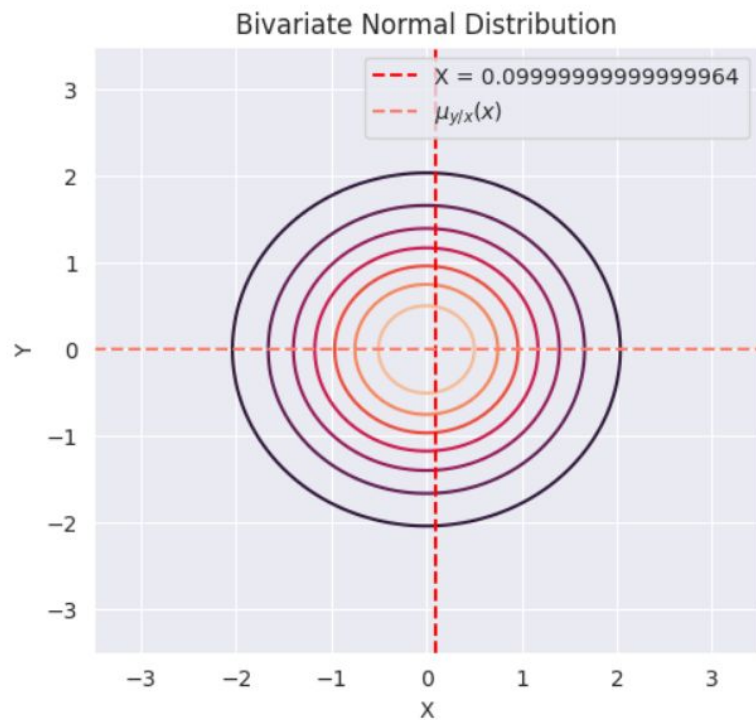
$$b = \mu_Y + C^T A^{-1} (\mathbf{x} - \mu_X)$$

$$M = B - C^T A^{-1} C$$

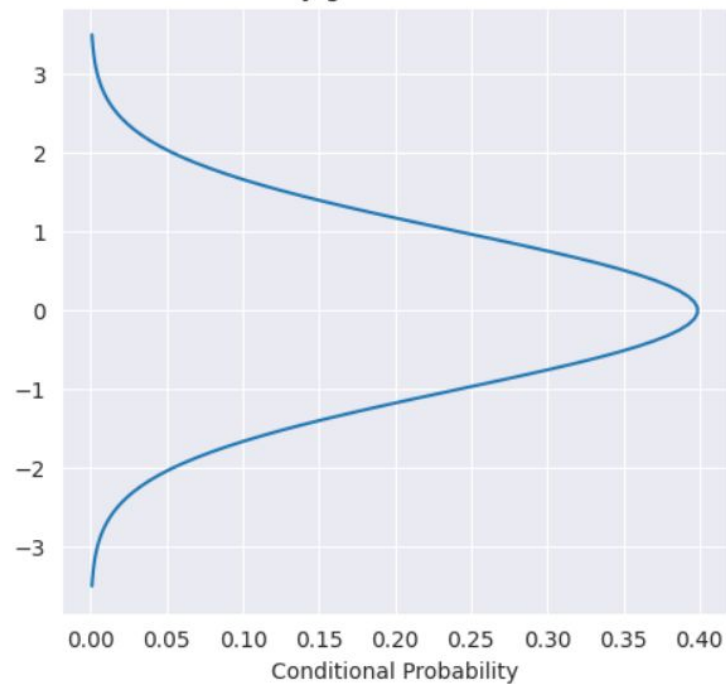
Consider $n = 2$, then,

$$b = \mu_Y + \frac{\text{Cov}(x, y)}{\text{Var}(x)} (x - \mu_X)$$

$$M = \text{Var}(y) - \frac{\text{Cov}(x, y)^2}{\text{Var}(x)}$$



Conditional Probability given $X = 0.09999999999999964$



Marginalisation

Given the two random vectors \mathbf{x}_A and \mathbf{x}_B , the marginal probability of \mathbf{x}_A is given by,

$$p(\mathbf{x}_A) = \int p(\mathbf{x}_A, \mathbf{x}_B) d\mathbf{x}_B$$



Exercise: Gaussian Marginalisation

Let \mathbf{x} and \mathbf{y} be jointly Gaussian random vector with dimension m and n , respectively.

$$\begin{bmatrix} \mathbf{x} \\ \mathbf{y} \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix} \mu_X \\ \mu_Y \end{bmatrix}, \begin{bmatrix} A & C \\ C^T & B \end{bmatrix} \right)$$

show that $x \sim \mathcal{N}(\mu_X, A)$

Solution :

$$\mathbf{x} = A \begin{bmatrix} \mathbf{x} \\ \mathbf{y} \end{bmatrix} = A \left(\begin{bmatrix} \mu_X \\ \mu_Y \end{bmatrix} + MZ \right) = A \begin{bmatrix} \mu_X \\ \mu_Y \end{bmatrix} + AMZ$$

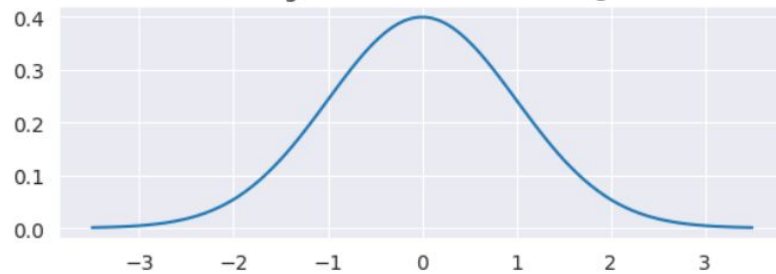
where,

$$A = [I_{m,m}, \mathbf{0}_{m,n}]$$

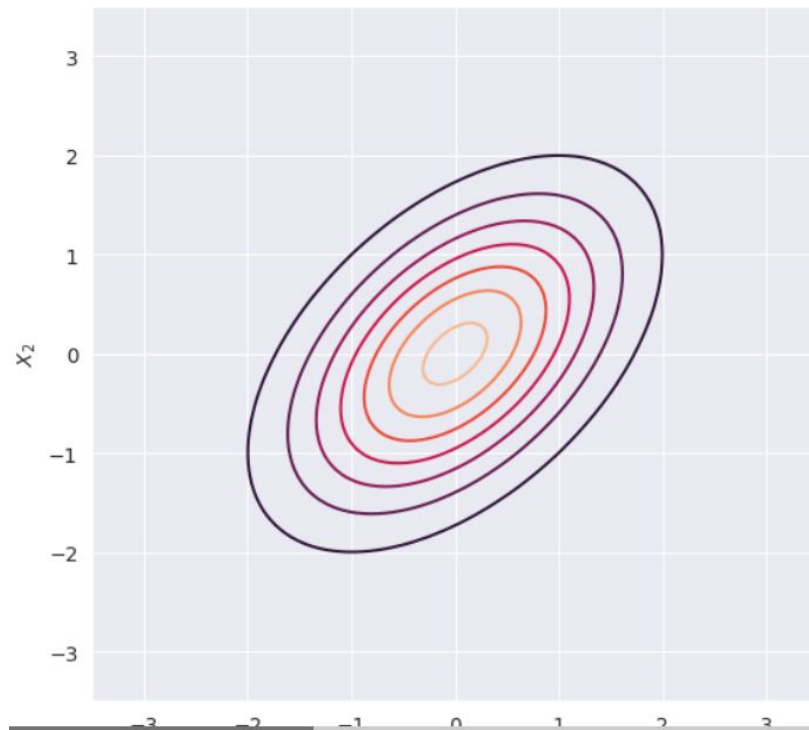
Therefore, \mathbf{x} is normally distributed with $\mathbb{E}[\mathbf{x}] = A \begin{bmatrix} \mu_X \\ \mu_Y \end{bmatrix} = \mu_X$ and

$$Cov(\mathbf{x}) = A \begin{bmatrix} A & C \\ C^T & B \end{bmatrix} A^T = A.$$

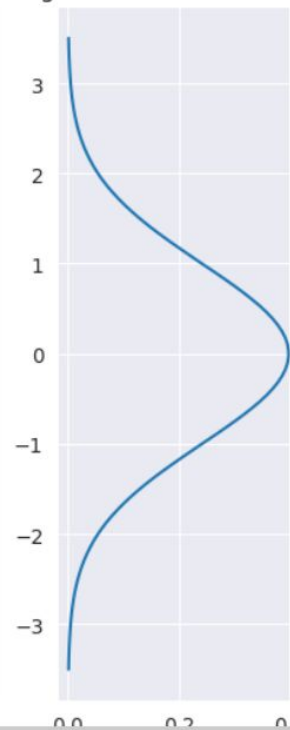
Marginalized Distribution of X_1



Bivariate Normal Distribution



Marginalized Distribution of X_2



Gaussian Normal Samples

Given a Cholesky decomposition of the covariance matrix to obtain the lower triangular matrix L ,

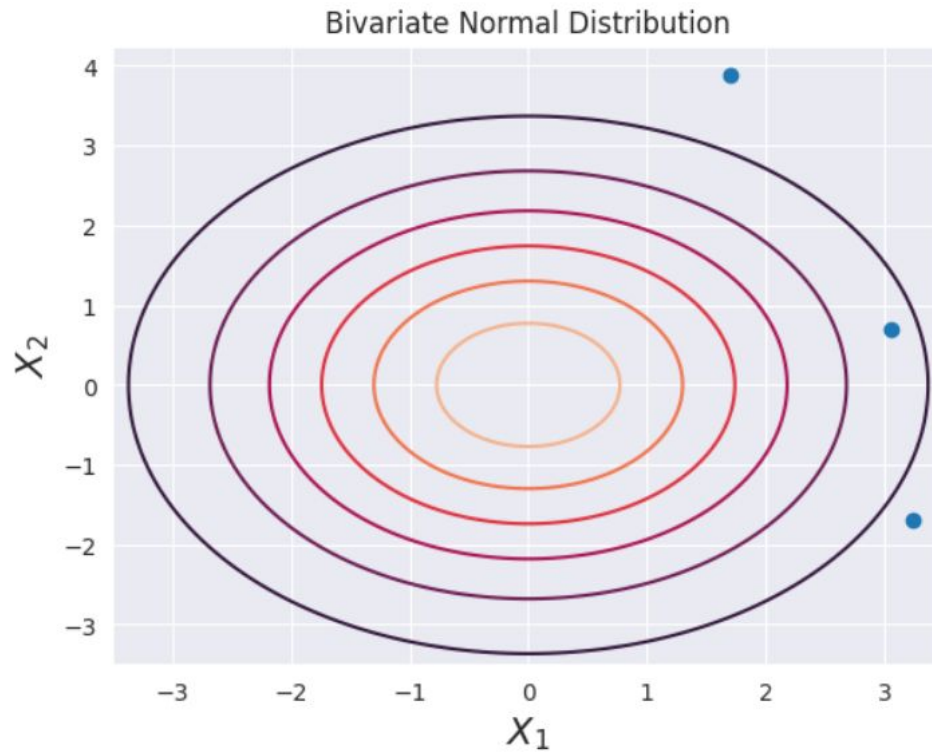
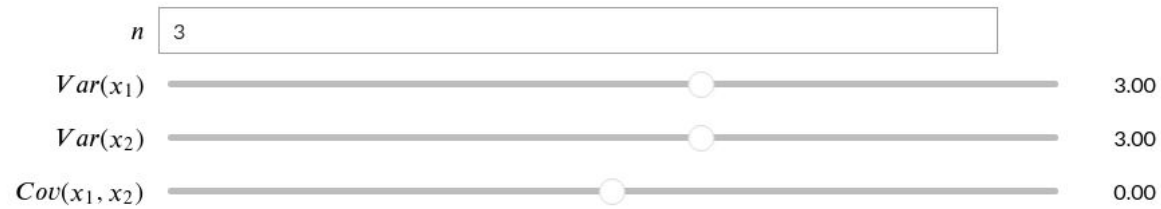
$$\Sigma = LL^T$$

Then, you can generate samples from a standard normal distribution as:

$$\mathbf{x} = \mu + L\mathbf{z}$$

where $\mathbf{z} \sim N(\mathbf{0}, I)$. For a single dimension we have, $x = \mu_x + \sigma_x z$





Summary

$$\begin{bmatrix} \mathbf{x} \\ \mathbf{y} \end{bmatrix} \sim \mathcal{N} \left(\mu = \begin{bmatrix} \mu_x \\ \mu_y \end{bmatrix}, \Sigma = \begin{bmatrix} A & C \\ C^T & B \end{bmatrix} \right)$$

Marginalisation & Conditioning:

$$\mathbf{x} \sim \mathcal{N}(\mu_x, A)$$

$$\mathbf{x}|\mathbf{y} \sim \mathcal{N}(\mu_x + CB^{-1}(\mathbf{y} - \mu_y), A - CB^{-1}C^T)$$

Sampling:

$$\begin{bmatrix} \mathbf{x} \\ \mathbf{y} \end{bmatrix} = \mu + L\mathbf{z}, \text{ where } \Sigma = LL^T$$



Gaussian Process

It is a collection of random variables, where any finite number of variables have a joint Gaussian distribution. A Gaussian process (GP) is defined by its mean function $m(x)$ and covariance function $k(x, x')$ as,

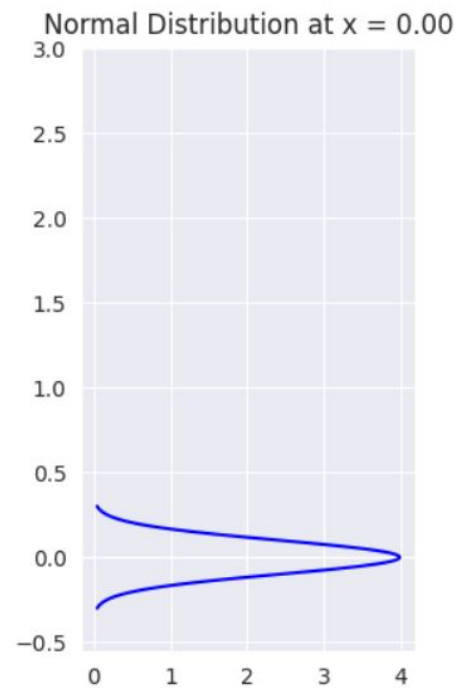
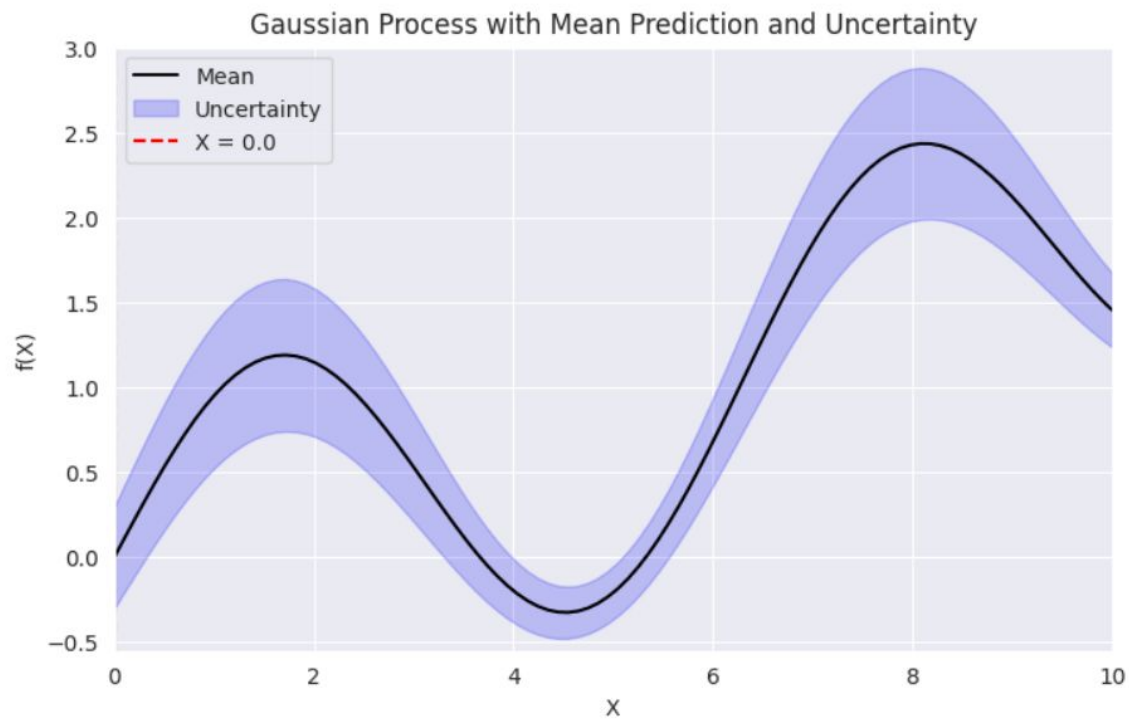
$$f(x) \sim GP(m(x), k(x, x'))$$

-Mean Function: $m(x) = \mathbb{E}[f(x)]$

-Covariance Function: $k(x, x') = \mathbb{E}[(f(x) - m(x))(f(x') - m(x'))]$



X: 0.00



Mean Function

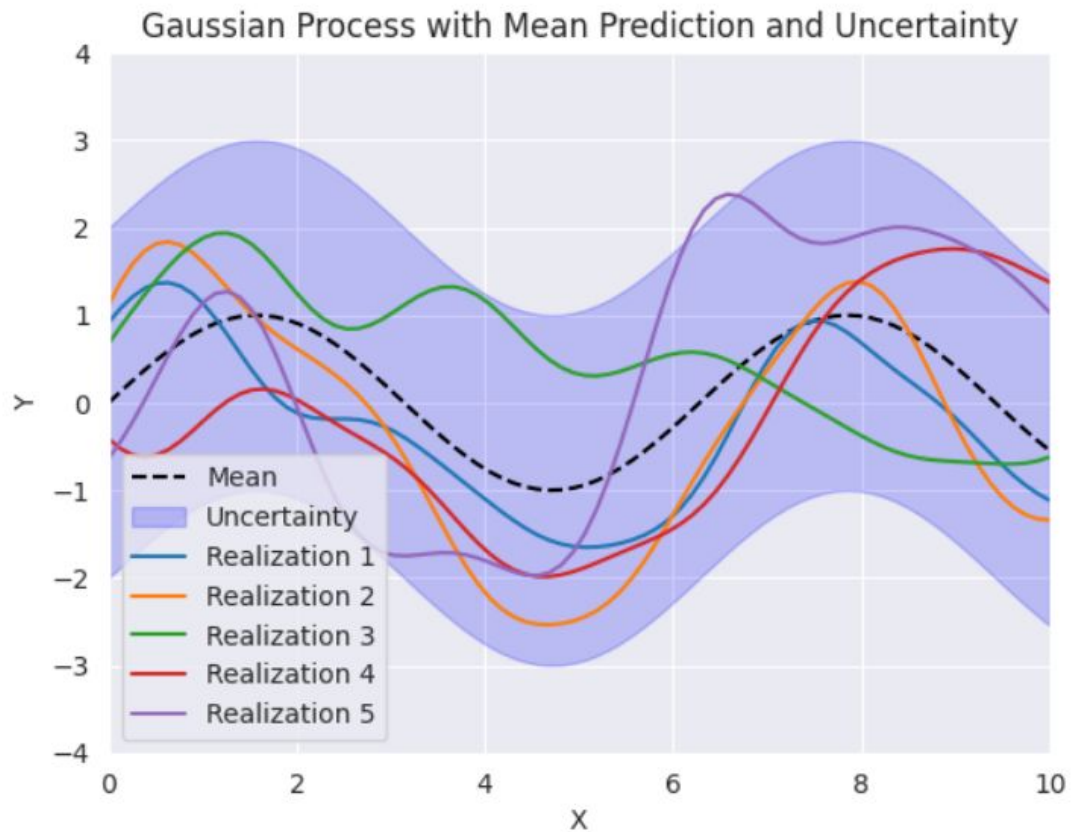
The mean function represents the expected value of the process at any given point.

- **Zero Mean Function:** The simplest assumption is to assume that it is zero everywhere, i.e., $m(x) = 0$ for all x .
- **Non-Zero Mean Function:** Prior knowledge, basis functions, etc.



Sin Mean Function

Linear Mean Function



Covariance Function

Define:

- Similarity/correlation between data points
- Smoothness & Periodicity

Properties:

- Symmetric & Positive Semi-definite

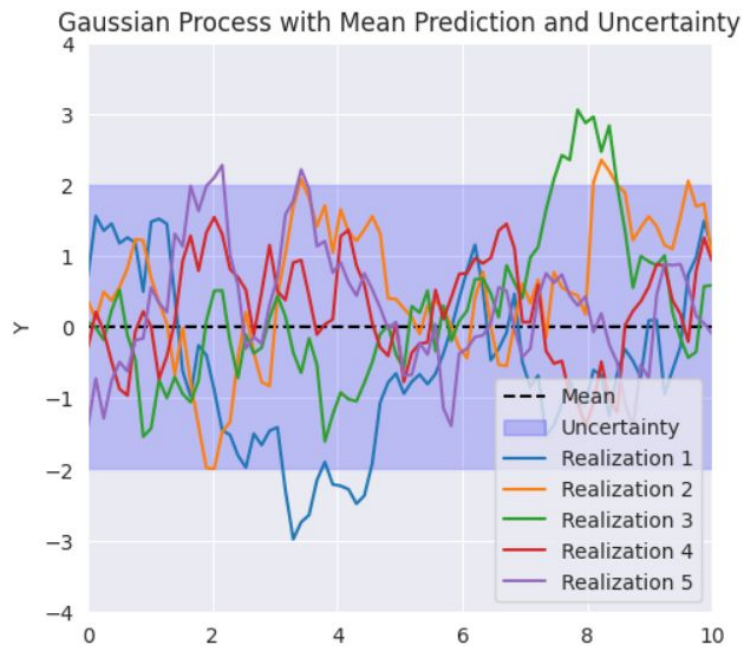
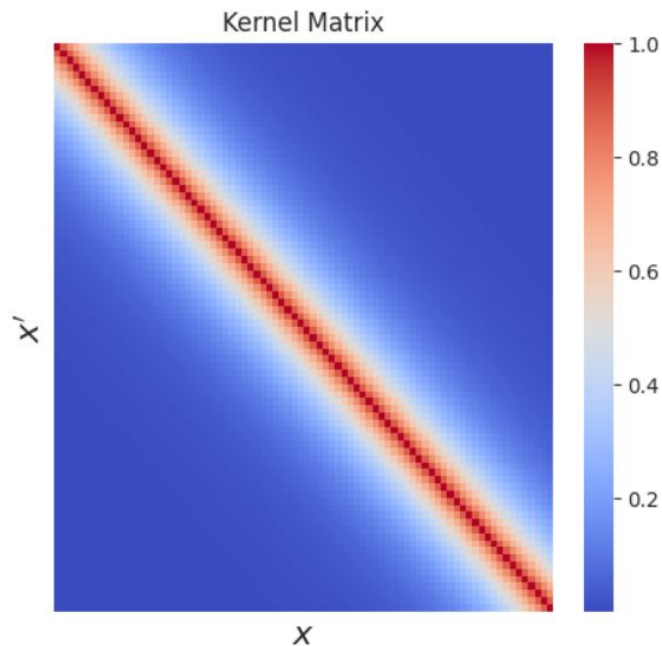


Gaussian Kernel	l	1
Linear Kernel	σ_v^2	0.
White Noise Kernel	σ^2	1

σ^2	1
σ_b^2	1

Matérn Kernel:

$$k_{\text{Matern}}(x, x') = \sigma^2 \frac{2^{1-\nu}}{\Gamma(\nu)} \left(\frac{\sqrt{2\nu}}{l} |x - x'| \right)^\nu K_\nu \left(\frac{\sqrt{2\nu}}{l} |x - x'| \right)$$



Combining Kernels

- **Summing Kernels:** The resulting covariance allows to capture various patterns simultaneously.

$$k_{\text{sum}}(x, x') = k_1(x, x') + k_2(x, x') + \dots + k_n(x, x')$$

- **Multiplying Kernels:** This approach is useful for modeling interactions between different patterns present in the data.

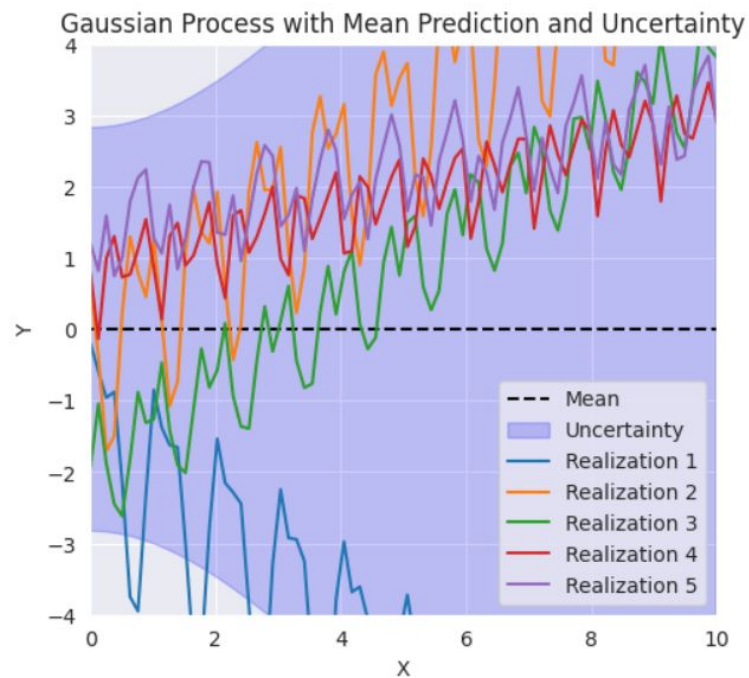
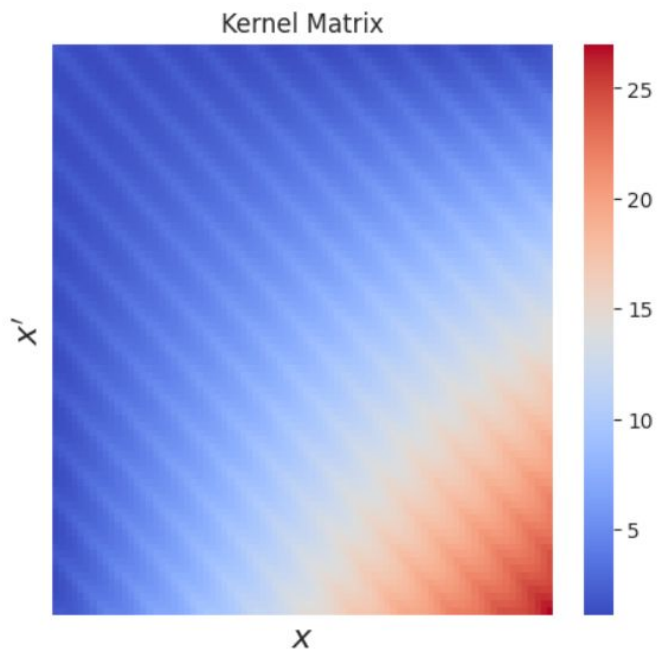
$$k_{\text{mult}}(x, x') = k_1(x, x') \times k_2(x, x') \times \dots \times k_n(x, x')$$



Operation addition



$Operation(k_{Linear}, k_{Periodic})$



Predictive Distribution

Given the noise-free observations \mathbf{f} , the joint distribution of observed locations and test points X and X^* is,

$$\begin{bmatrix} \mathbf{f} \\ \mathbf{f}^* \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix} \mathbf{m} \\ \mathbf{m}^* \end{bmatrix}, \begin{bmatrix} K(X, X) & K(X, X^*) \\ K(X^*, X) & K(X^*, X^*) \end{bmatrix} \right)$$

The predictive distribution is obtained by conditioning on the observed data:

$$\mathbf{f}^* | \mathbf{f} \sim \mathcal{N}(\mu^*, \Sigma^*)$$

$$\mu^* = \mathbf{m}^* + K(X^*, X)^T K(X, X)^{-1} (\mathbf{f} - \mathbf{m})$$

$$\Sigma^* = K(X^*, X^*) - K(X^*, X)^T K(X, X)^{-1} K(X, X^*)$$



Predictive Distribution using Noisy Observation

Consider, $y(x) = f(x) + \epsilon$, where $\epsilon \sim \mathcal{N}(0, \sigma_v^2)$. Therefore,

$$\begin{bmatrix} \mathbf{y} \\ \mathbf{f}^* \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix} \mathbf{m} \\ \mathbf{m}^* \end{bmatrix}, \begin{bmatrix} K(X, X) + \sigma_v^2 I & K(X, X^*) \\ K(X^*, X) & K(X^*, X^*) \end{bmatrix} \right)$$

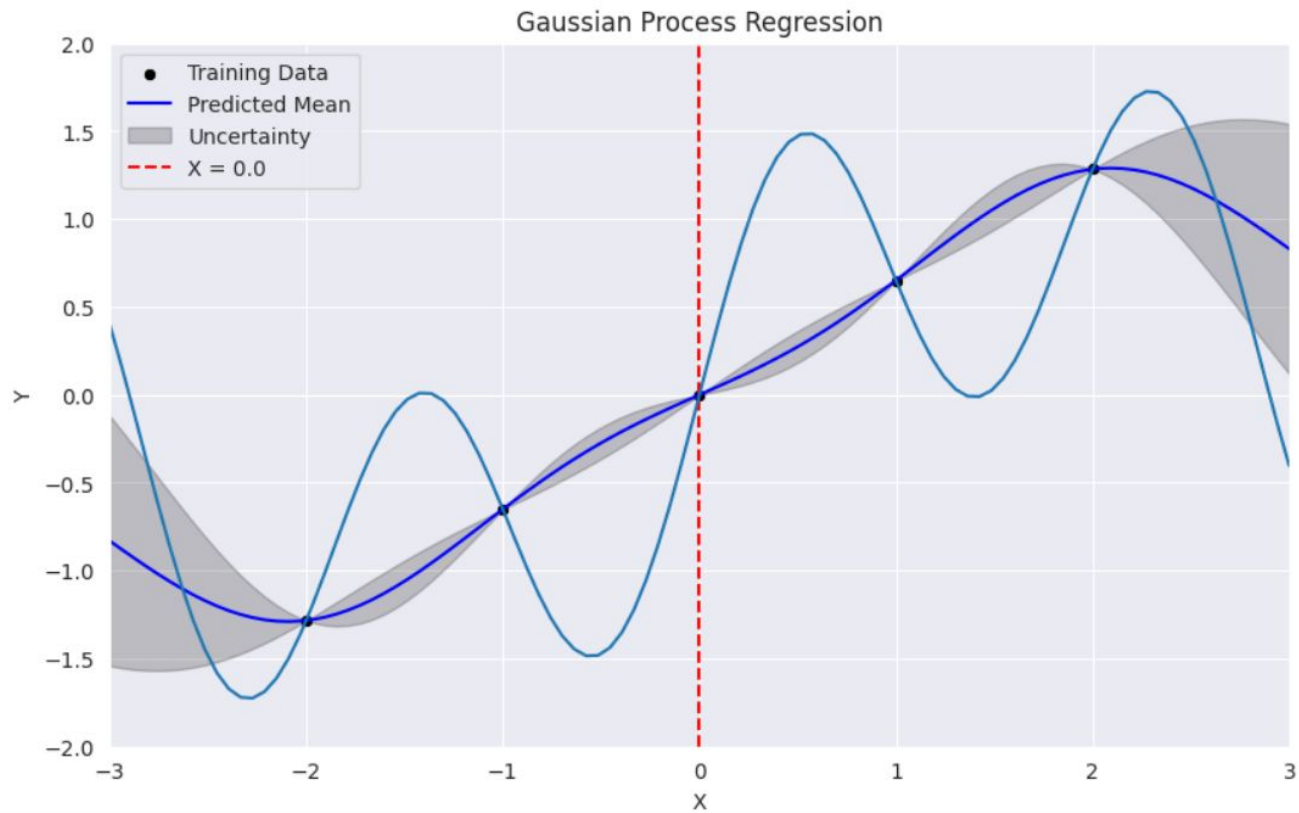
and,

$$\mathbf{f}^* | \mathbf{y} \sim \mathcal{N}(\mu^*, \Sigma^*)$$

$$\mu^* = \mathbf{m}^* + K(X^*, X)^T [K(X, X) + \sigma_v^2 I]^{-1} (\mathbf{y} - \mathbf{m})$$

$$\Sigma^* = K(X^*, X^*) - K(X^*, X)^T [K(X, X) + \sigma_v^2 I]^{-1} K(X, X^*)$$





Learning Hyperparameters

Given the marginal likelihood of the observed data.

$$p(\mathbf{f}^*|X, \mathbf{y}, X^*, \theta) = \frac{p(\mathbf{y}, \mathbf{f}^*|X^*, X, \theta)}{\underbrace{p(\mathbf{y}|X, \theta)}_{\text{marginal likelihood}}} = \frac{p(\mathbf{y}, \mathbf{f}^*|X^*, X, \theta)}{\int p(\mathbf{y}|\mathbf{f})p(\mathbf{f}|X, \theta)d\mathbf{f}}$$

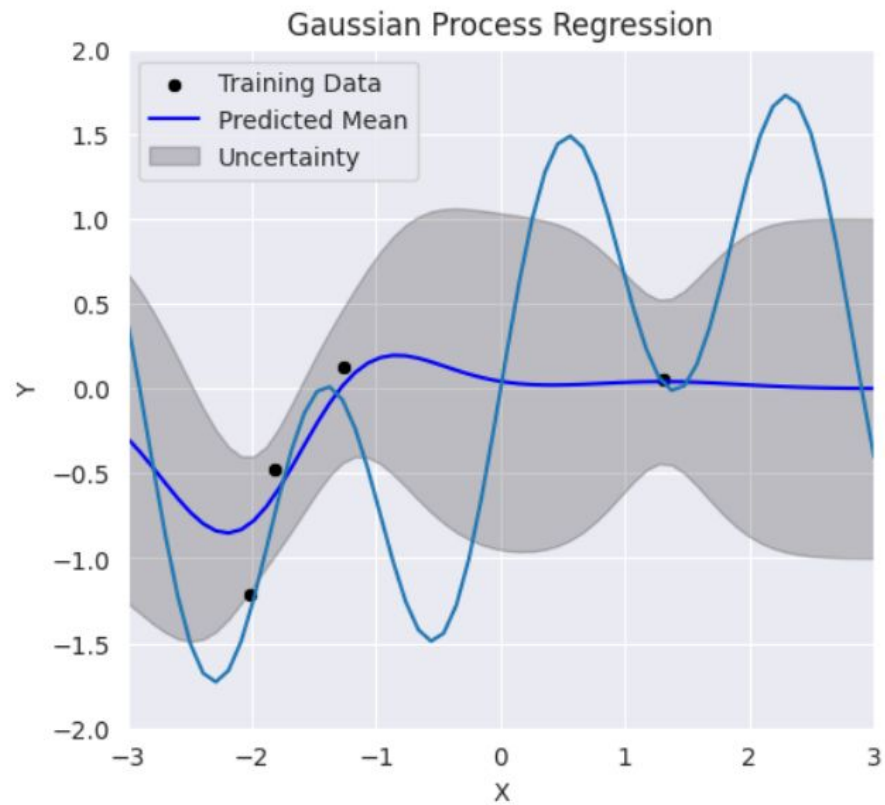
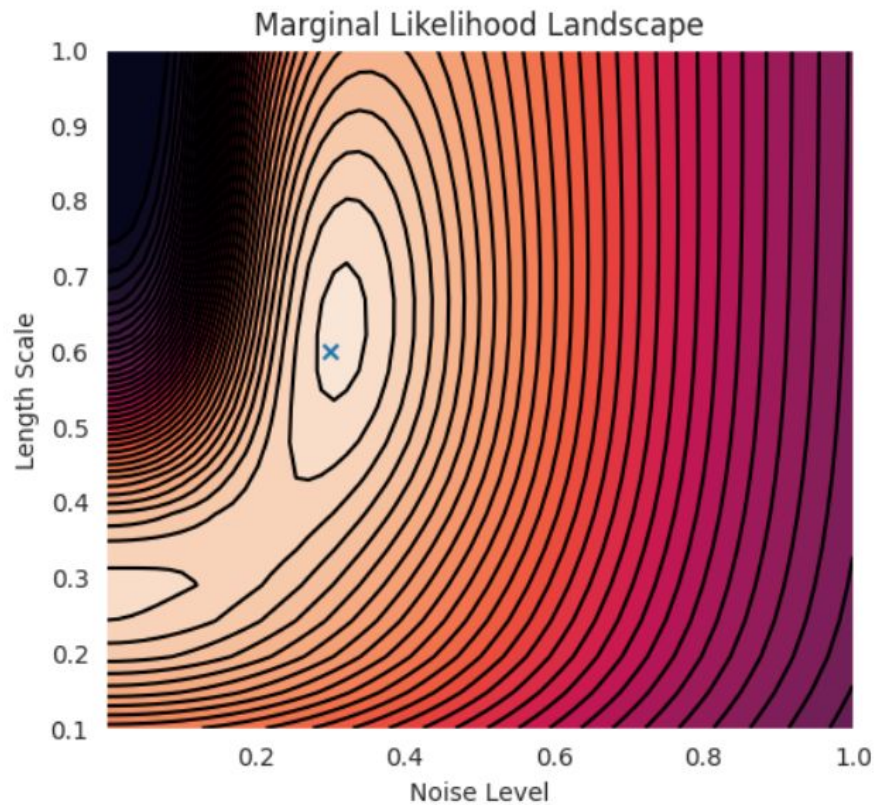
where,

$$\log p(\mathbf{y}|X, \theta) = -\frac{1}{2}(\mathbf{y}^\top (K_\theta(X, X) + \sigma_n^2 I)^{-1} \mathbf{y} + \log |K_\theta(X, X) + \sigma_n^2 I| + n \log(2\pi))$$

Aim:

$$\theta^* = \operatorname{argmax}_\theta (-\log p(\mathbf{y}|X, \theta))$$





Benefits

- **Flexibility:** GPs can model complex relationships between inputs and outputs without imposing a specific functional form.
- **Uncertainty Estimation:** GPs provide not only point predictions but also estimate uncertainty in predictions.



Difficulties

- Presents a $\mathcal{O}(n^3)$ computational cost and $\mathcal{O}(n^2)$ in memory.
 - Sparse approximation.
- Challenging in high number of dimensions.
 - Structural assumptions
- The marginal likelihood is often multi-modal, i.e, local optima.
 - Random start points, using prior distributions, marginalise over hyperparameters.

