CS 240: Reasoning Under Uncertainty (Fall 21)

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Lecture 12

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#### 1 Introduction

In this lecture, we discussed several specific continuous random variables. They are uniform continuous random variables, exponential random variables, normal (Gaussian) random variables, and the standard normal random variable.

### 2 Continuous Random Variables

#### 2.1 Uniform Continuous Random Variable

A uniform continuous random variable has a uniform probability density in [a, b]. Its expected value:

$$E(X) = \frac{b+a}{2}$$

Its variance:

$$Var(X) = \frac{1}{12}(b-a)^2$$

Its PDF is:

$$f_X(x) = \begin{cases} \frac{1}{b-a} & a \le x \le b\\ 0 & otherwise \end{cases}$$

Its CDF:

$$F_X(x) = \begin{cases} 0 & x < a \\ \frac{x-a}{b-a} & a \le x \le b \\ 1 & x > b \end{cases}$$

# 2.2 Exponential Random Variable

An exponential random variable characterizes the (time/space) interval between events in a Poisson point process. Suppose we have a Poisson random variable X. We know that  $E(X) = \lambda$  which is the expected number of times an event occurs internally with a unit length is  $\lambda$ . Now the question is what is the expected interval between two events? Let's use the random variable Y to represent the size of the interval between the two events.

Its expected value should be:

$$E(Y) = \frac{1}{\lambda}$$

Its variance:

$$Var(Y) = \frac{1}{\lambda^2}$$

Its PDF is:

$$f_X(x) = \begin{cases} \lambda e^{-\lambda x} & 0 < x \\ 0 & otherwise \end{cases}$$

Its CDF:

$$F_X(x) = \begin{cases} 1 - e^{-\lambda x} & x \ge 0\\ 0 & otherwise \end{cases}$$

#### 2.3 Normal Random Variable

Let X be a normal random variable with  $E(X) = \mu$  and  $Var(X) = \sigma^2$ . Its probability density function is:

$$f_X(x) = \frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

Its cumulative distribution function is:

$$F_X(x) = \frac{1}{2}(1 + erf(\frac{x - \mu}{\sigma\sqrt{2}}))$$

where  $erf(x) = \frac{1}{\sqrt{\pi}} \int_{-x}^{x} e^{-t^2} dt$ 

The **probability mass** of an interval [a, b] is the definite integral:

$$P(a < X < b) = \int_{a}^{b} \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{(x-\mu)^{2}}{2\sigma^{2}}} = F_{X}(b) - F_{X}(a)$$

#### 2.4 Standard Normal Random Variable

Let X be a normal random variable with E(X) = 0 and  $Var(X) = \sigma^2 = 1$ .

Its probability density function is:

$$f_X(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$$

Its cumulative distribution function is:

$$F_X(x) = \frac{1}{2\sqrt{\pi}} \int_{-\infty}^x e^{-\frac{t^2}{2}} dt$$

### 2.5 Standardizing a Normal Variable

For a given normal random variable X with mean  $\mu$  and variance  $\sigma^2$ , you can standardize it by defining a new random variable Y given by

$$Y = \frac{X - \mu}{\sigma}$$

Where Y is a standard normal variable with mean 0 and variance 1.

This is because Y is a linear function of X (Y = aX + b), we know Y perceives the normality.

#### 2.6 Practice Problems

- 1. Suppose that Y has a uniform distribution over the interval (0, 1).
  - a) Find F(y).
  - b) Show that  $P(a \le Y \le a + b)$  (for  $a \ge 0$ ,  $b \ge 0$ , and  $a + b \le 1$ ) depends only upon the value of b.
- 2. A random variable Y has a uniform distribution over the interval (a, b). Derive the variance of Y.
- 3. Let X be a continuous random variable with PDF  $f_X(x) = \frac{\lambda}{2} e^{-\lambda |x|}$ .
  - a) Verify that  $f_X$  is a valid PDF.
  - b) Evaluate E(X) and Var(X)
- 4. Let Z denote a normal random variable with mean 0 and standard deviation 1.
  - a) Find P(Z > 1)
  - b) Find P(-1 < Z < 1)
  - c) Find P(0.3 < Z < 1.5)
- 5. If Z is a standard normal random variable, find the value  $z_0$  such that
  - a)  $P(Z > z_0) = 0.5$
  - b)  $P(Z < z_0) = 0.8643$
  - c)  $P(-z_0 < Z < z_0) = 0.90$
  - d)  $P(-z_0 < Z < z_0) = 0.99$
- 6. A soft-drink machine can be regulated so that it discharges an average of  $\mu$  ounces per cup. If the ounces of fill are normally distributed with a standard deviation of 0, compute the setting for  $\mu$  so that 8-ounce cups will overflow only 1 percent of the time.

# 3 Answer

1. a)

$$f_X(x) = \begin{cases} 0 & x < 0 \\ 1 & 0 \le x \le 1 \\ 0 & x > 1 \end{cases}$$

$$F_X(x) = \begin{cases} 0 & x < 0 \\ x & 0 \le x \le 1 \\ 1 & x > 1 \end{cases}$$

b) 
$$P(a < X < a + b) = F_X(a + b) - F_X(a) = b$$

2.

$$f_X(x) = \begin{cases} 0 & x < a \\ \frac{1}{b-a} & a \le x \le b \\ 0 & x > b \end{cases}$$

$$\begin{split} E(X) &= \int_a^b x \frac{1}{b-a} dx = \frac{b^2 - a^2}{2(b-a)} = \frac{a+b}{2} \\ Var(X) &= E(X^2) - (E(X))^2 = \int_a^b x^2 \frac{1}{b-a} dx - (\frac{a+b}{2})^2 = \frac{b^3 - a^3}{3(b-a)} - \frac{(a+b)^2}{4} = \frac{(b-a)^2}{12} \end{split}$$

3. a) 
$$P(-\infty < X < \infty) = \int_{-\infty}^{\infty} \frac{\lambda}{2} e^{-\lambda |x|} dx = \int_{0}^{\infty} \lambda e^{-\lambda x} dx = 1$$
  
b)  $E(X) = \int_{-\infty}^{\infty} \frac{\lambda}{2} x e^{-\lambda |x|} dx = 0$  (symmetry)

Or you can use integration by parts:

$$\begin{split} E(X) &= \int_{-\infty}^{\infty} \frac{\lambda}{2} x e^{-\lambda |x|} dx \\ &= \int_{-\infty}^{0} \frac{\lambda}{2} x e^{\lambda x} dx + \int_{0}^{\infty} \frac{\lambda}{2} x e^{-\lambda x} dx \end{split}$$

Let 
$$A = \int_{-\infty}^{0} \frac{\lambda}{2} x e^{\lambda x} dx$$
,  $B = \int_{0}^{\infty} \frac{\lambda}{2} x e^{-\lambda x}$ 

We now show that A=-B using integration by parts (  $\int uv'=uv-\int u'v$  ) To compute A let  $u=x,\ u'=1,\ v'=\frac{\lambda}{2}e^{\lambda x},\ v=\frac{1}{2}e^{\lambda x}$ 

$$A = \int_{-\infty}^{0} \frac{\lambda}{2} x e^{\lambda x} dx$$

$$= \left[ \frac{1}{2} x e^{\lambda x} \right]_{-\infty}^{0} - \int_{-\infty}^{0} \frac{1}{2} e^{\lambda x} dx$$

$$= -\int_{-\infty}^{0} \frac{1}{2} e^{\lambda x} dx$$

$$= -\left[ \frac{1}{2\lambda} e^{\lambda x} \right]_{-\infty}^{0}$$

$$= -\frac{1}{2\lambda}$$

To compute B let  $u=x,\,u'=1,\,v'=\frac{\lambda}{2}e^{-\lambda x},\,v=-\frac{1}{2}e^{-\lambda x}$ 

$$\begin{split} B &= \int_0^\infty \frac{\lambda}{2} x e^{-\lambda x} dx \\ &= [-\frac{1}{2} x e^{-\lambda x}]_0^\infty - \int_0^\infty -\frac{1}{2} e^{-\lambda x} dx \\ &= \int_0^\infty \frac{1}{2} e^{-\lambda x} dx \\ &= [-\frac{1}{2\lambda} e^{-\lambda x}]_0^\infty \\ &= \frac{1}{2\lambda} \end{split}$$

Thus, 
$$E(X) = A + B = 0$$

$$Var(X)=E(X^2)-E(X)^2=\int_{-\infty}^{\infty}\frac{\lambda}{2}x^2e^{-\lambda|x|}dx=\int_{0}^{\infty}\lambda x^2e^{-\lambda|x|}dx=\frac{2}{\lambda^2}$$
 (use integrate by parts twice)

- 4. a) P(Z > 1) = 0.15866
  - b) P(-1 < Z < 1) = (0.5 P(Z > 1)) + (0.5 P(Z < -1)) = 1 2 \* 0.15866 = 0.68268
  - c) P(0.3 < Z < 1.5) = (0.5 P(Z > 1.5)) (0.5 P(Z > 0.3)) = 0.38209 0.06681 = 0.31528
- 5. a)  $z_0 = 0$ 
  - b)  $z_0 = 1.1$
  - c)  $P(Z < z_0) = 0.9 + (1 0.9)/2 = 0.95, z_0 = 1.64$
  - d)  $P(Z < z_0) = 0.995, z_0 = 2.58$
- 6.  $P(Z < \frac{8-\mu}{0.3}) = 0.99, \, \mu = 7.301$