EC 607, Set 10

Edward Rubin

Prologue

Schedule

Last time

- Introduction to selection-on-unobservables designs
- Instrumental variables (IV) and two-stage least squares (2SLS)

Today

- Regression discontinuity[†]
- Project reminder!

[†] These notes largely follow notes from Michael Anderson, Imbens and Lemieux (2008), and notes from Teppei Yamamoto.

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In addition, while RD is all the rage in modern applied econometrics, Thistlewaite and Campbell wrote about it back in 1960.

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The variable X_i need not be randomly assigned—we will assume it is not (i.e., X_i correlates with Y_{0i} and Y_{1i}).

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The variable X_i need not be randomly assigned—we will assume it is not (i.e., X_i correlates with Y_{0i} and Y_{1i}).

We will assume that Y_{0i} and Y_{1i} vary smoothly in X_i .

Examples

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- Election runoffs are triggered if "winner" is below 50%.
- Antidiscrimination laws only apply to firms with >15 employees.
- Prisoners are eligible for early parole if some score exceeds a threshold.
- An individual is eligible for Medicare if her age is at least 65.
- You get a ticket if your speed exceeds the speed limit.
- Fifteen-percent discount at Sizzler if your age exceeds 60.
- Counties with $PM_{2.5} > 35 \mu g/m^3$ are out of attainment.

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In some cases, "treatment" is definite once we exceed the threshold.

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E.g., a politician wins an election when the difference between her vote share and her competitor's vote share exceeds zero.

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In fuzzy RDs, the probability of treatment $\Pr(D_i = 1)$ discontinuously jumps at the threshold c, but it does not move from 0 to 1.

E.g., crossing some GRE threshold discontinuously increases your chances of getting into some grad schools (but doesn't guarantee admittance).

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To estimate the causal effect of D_i on Y_i , we compare the mean of Y_i just above the threshold to the mean of Y_i just below the threshold.

More formally

We can write the comparison of means at the threshold as

$$\lim_{x\downarrow c} E[\mathrm{Y}_i \mid \mathrm{X}_i = x] - \lim_{x\uparrow c} E[\mathrm{Y}_i \mid \mathrm{X}_i = x]$$

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Assumption $E[\mathbf{Y}_{1i} \mid \mathbf{X}_i = x]$ and $E[\mathbf{Y}_{0i} \mid \mathbf{X}_i = x]$ are continuous in x.

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I.e., Because we don't observe \mathbf{Y}_{0i} for treated individuals, we extrapolate $E[\mathbf{Y}_{0i} \mid \mathbf{X}_i = c - \varepsilon]$ to $E[\mathbf{Y}_{0i} \mid \mathbf{X}_i = c + \varepsilon]$ for small ε .

Estimation

Thus, we estimate

$$au_{ ext{SRD}} = \lim_{x\downarrow c} E[ext{Y}_i \mid ext{X}_i = x] - \lim_{x\uparrow c} E[ext{Y}_i \mid ext{X}_i = x]$$

as the diffrence between two regression functions estimated "near" c.

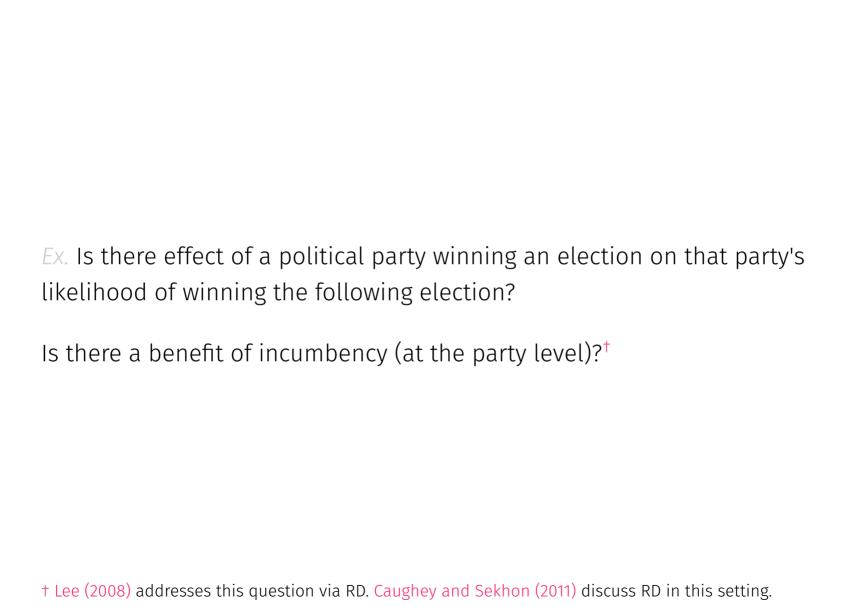
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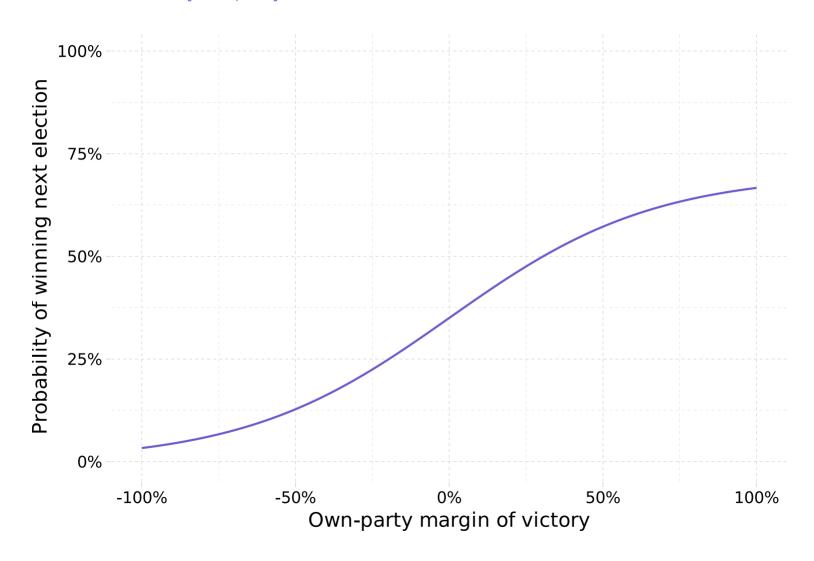
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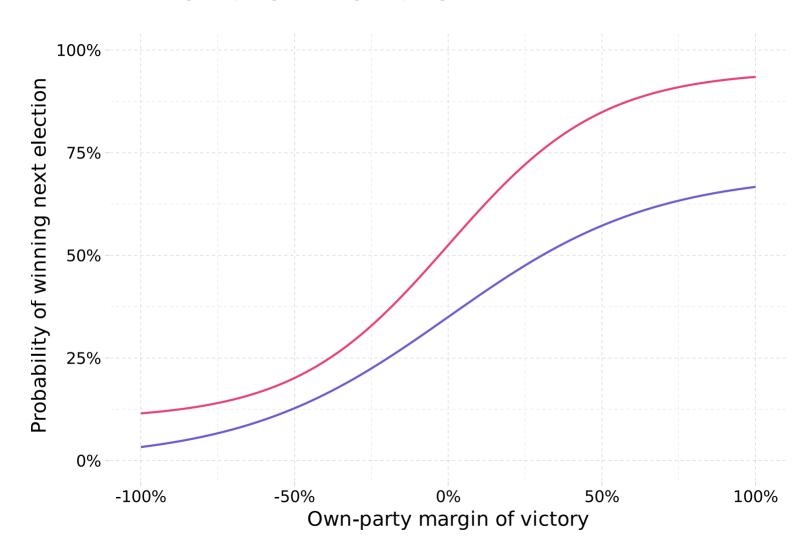
We must stay "near" to c to minimize the bias from extrapolating $E[Y_{0i} \mid X_i = c - \varepsilon]$ to $E[Y_{0i} \mid X_i = c + \varepsilon]$ (and assuming continuity).



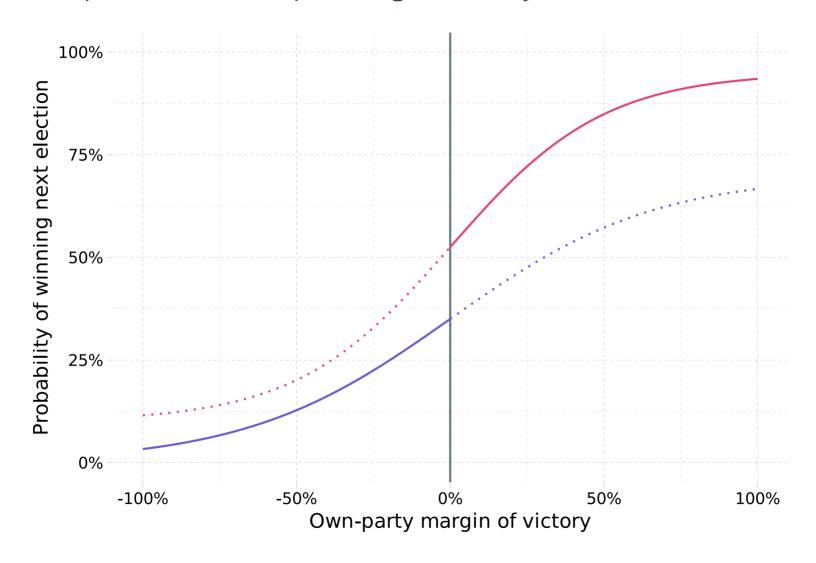
Let's start with $E[\mathbf{Y}_{0i} \mid \mathbf{X}_i]$



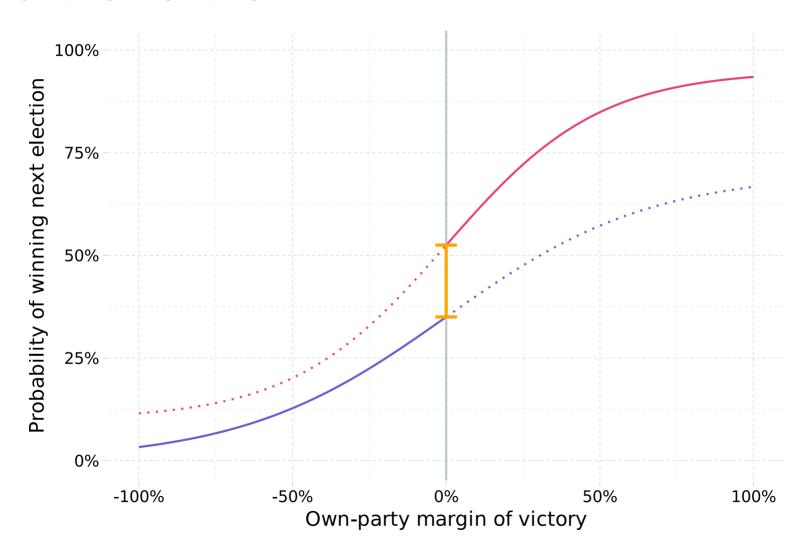
Let's start with $E[Y_{0i} \mid X_i]$ and $E[Y_{1i} \mid X_i]$.



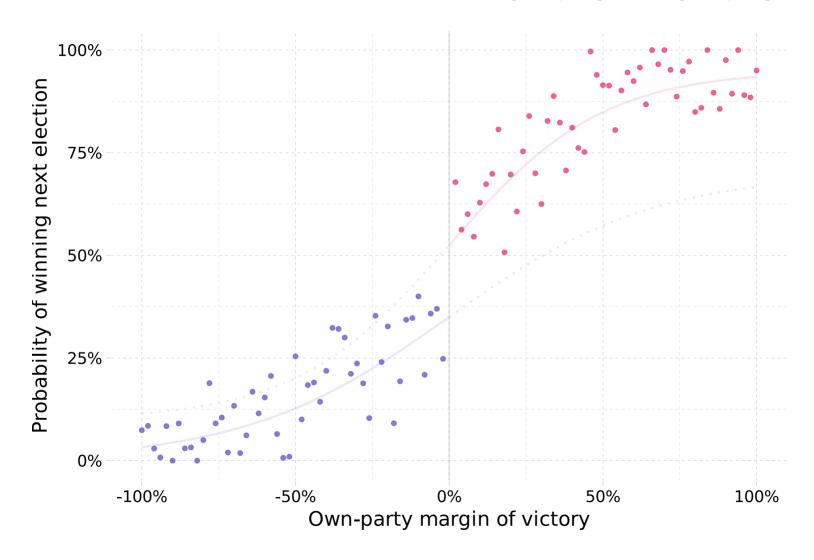
You only win an election if your margin of victory exceeds zero.



 $E[Y_{1i} \mid X_i] - E[Y_{0i} \mid X_i]$ at the discontinuity gives τ_{SRD} .



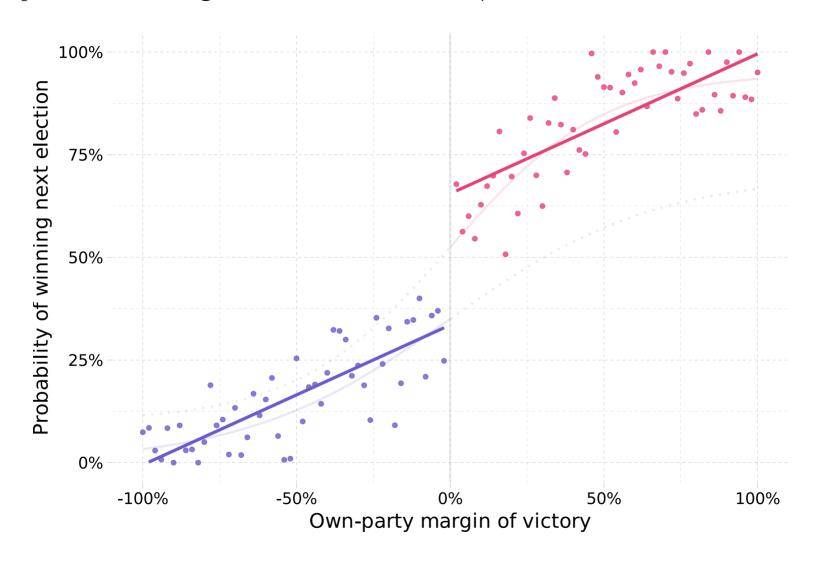
Real data are a bit trickier. We must estimate $E[Y_{1i} \mid X_i]$ and $E[Y_{0i} \mid X_i]$.



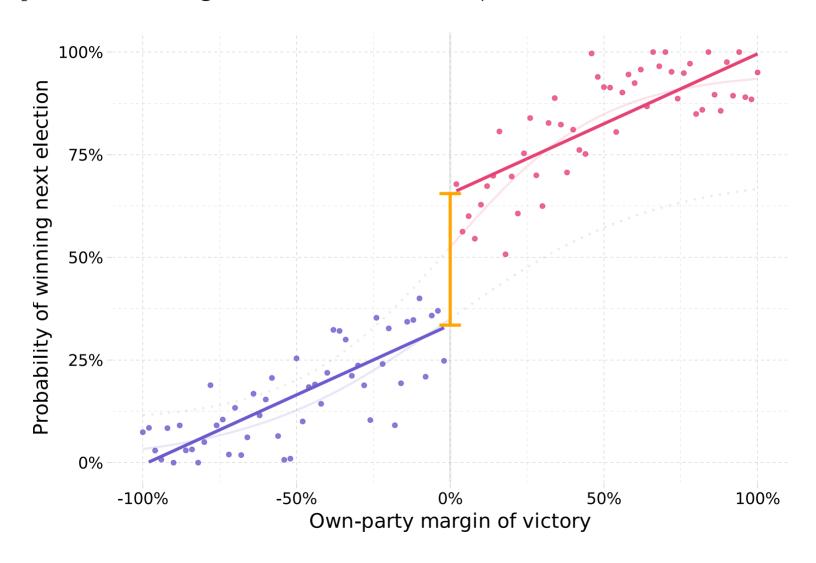
Questions

- 1. How should we estimate $E[Y_{1i} \mid X_i]$ and $E[Y_{0i} \mid X_i]$?
- 2. How much data should we use—i.e., what is the right bandwidth size?

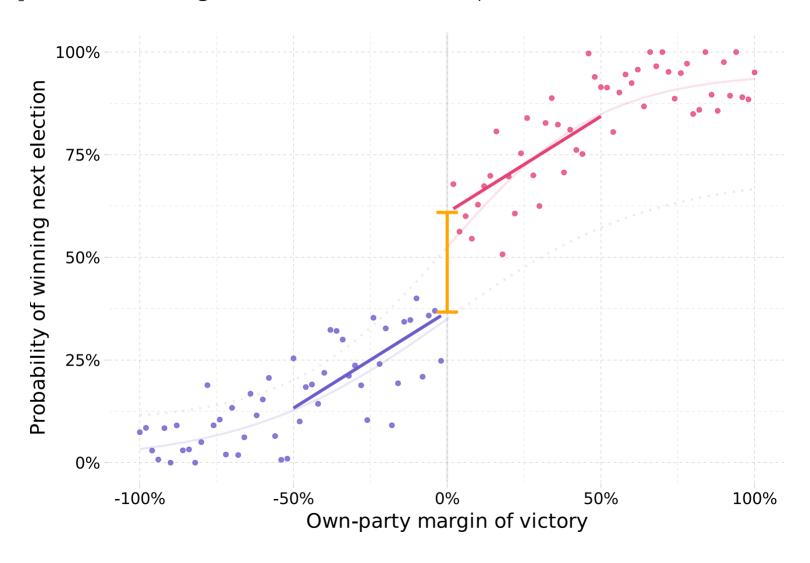
Option 1a Linear regression with constant slopes (and all data)



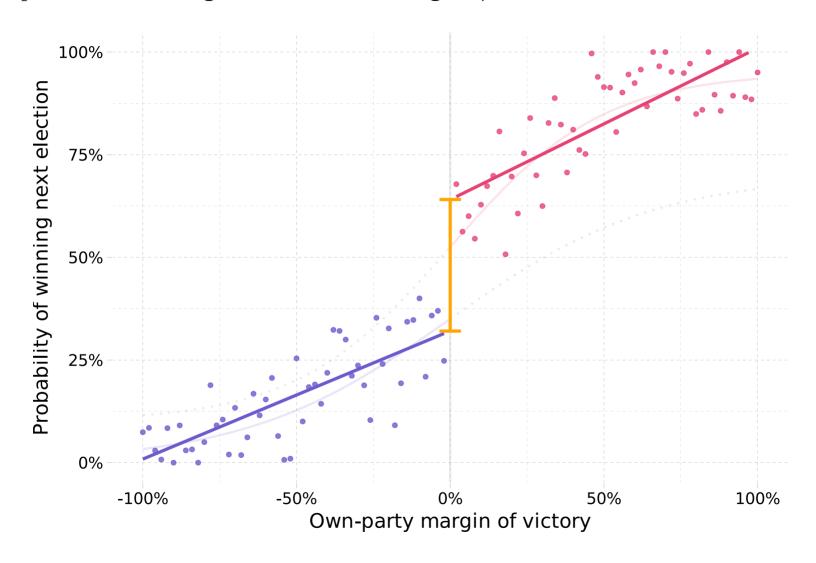
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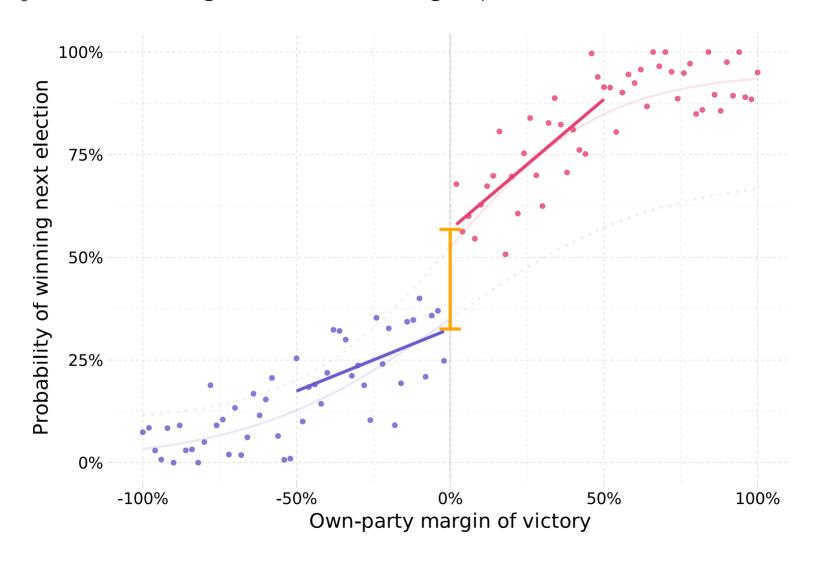
Option 1b Linear regression with constant slopes; limited to +/- 50%.



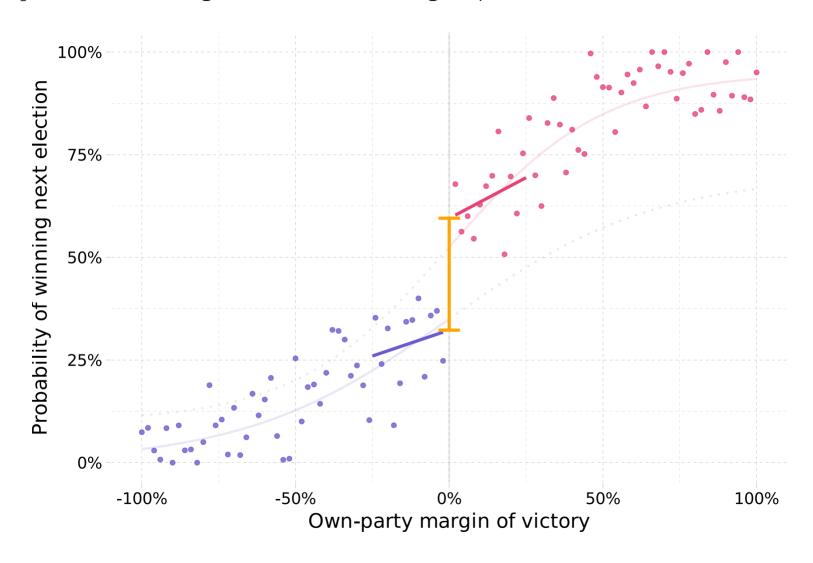
Option 2a Linear regression with differing slopes (and all data)



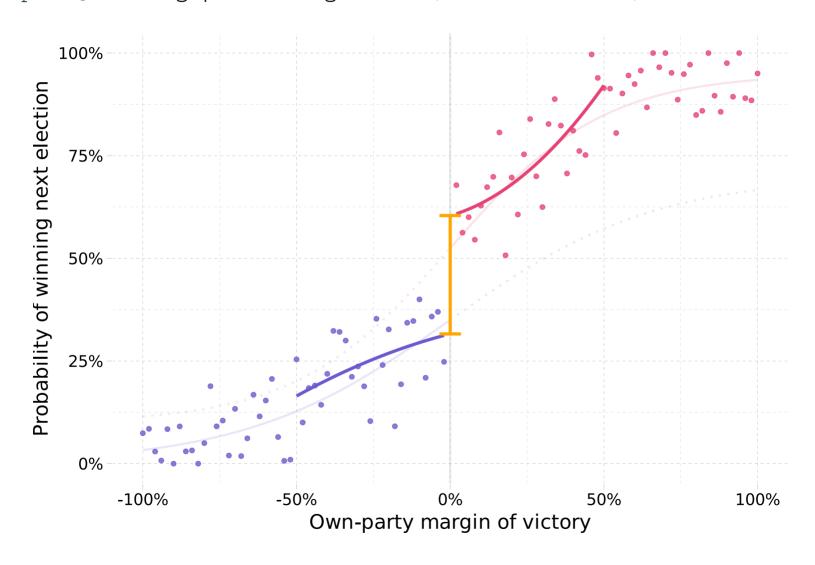
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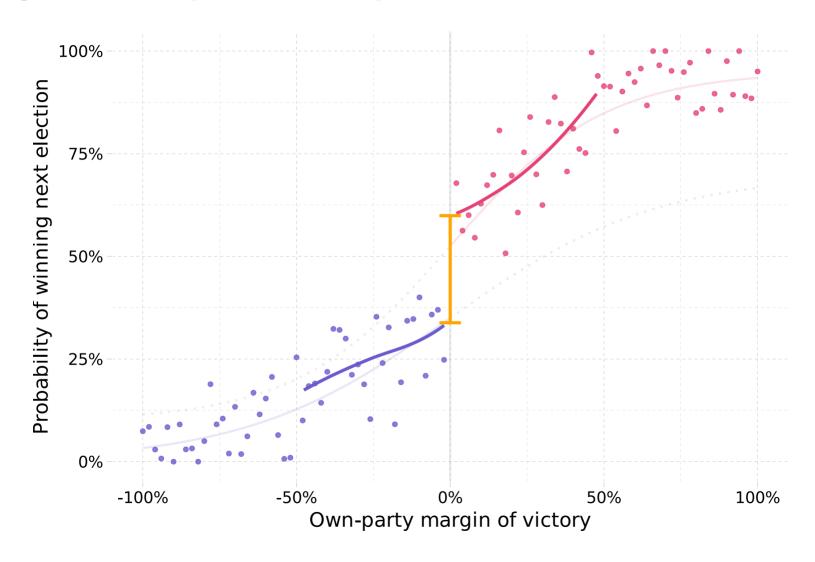
Option 2c Linear regression with differing slopes; limited to +/- 25%.



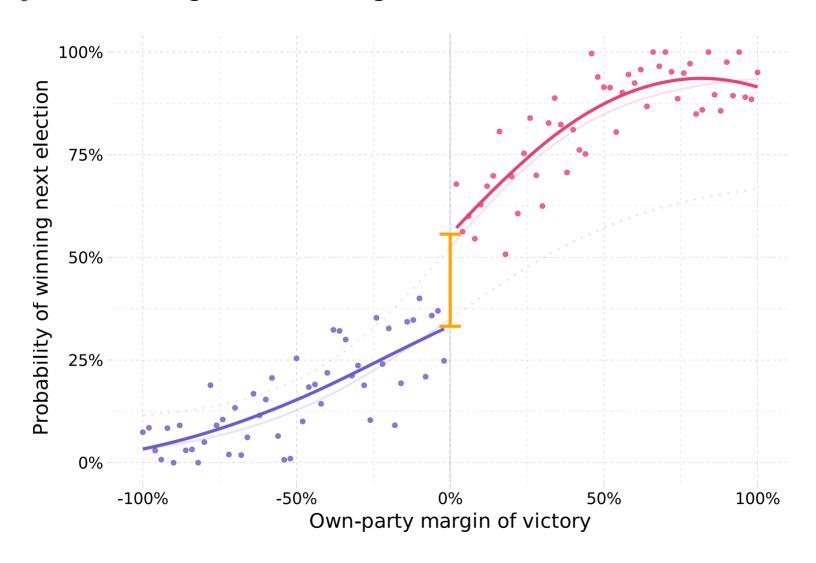
Option 3 Differing quadratic regressions (limited to +/- 50%).

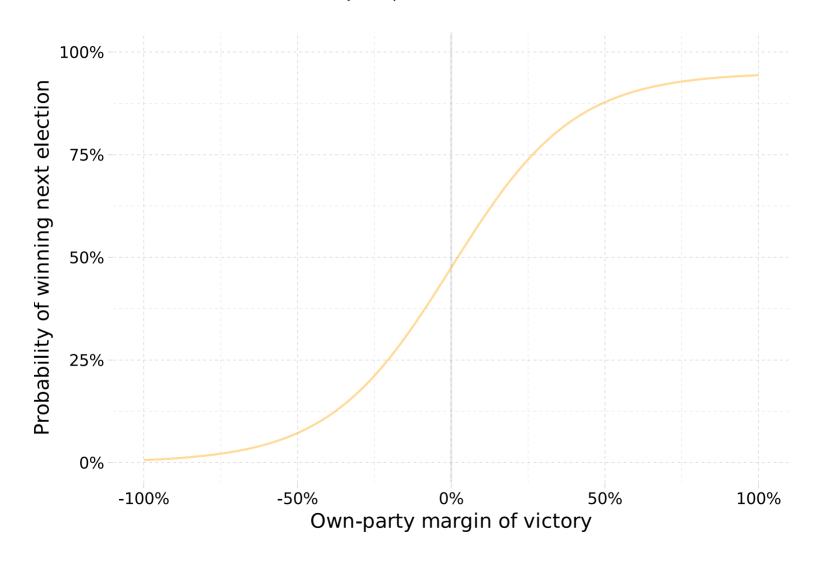


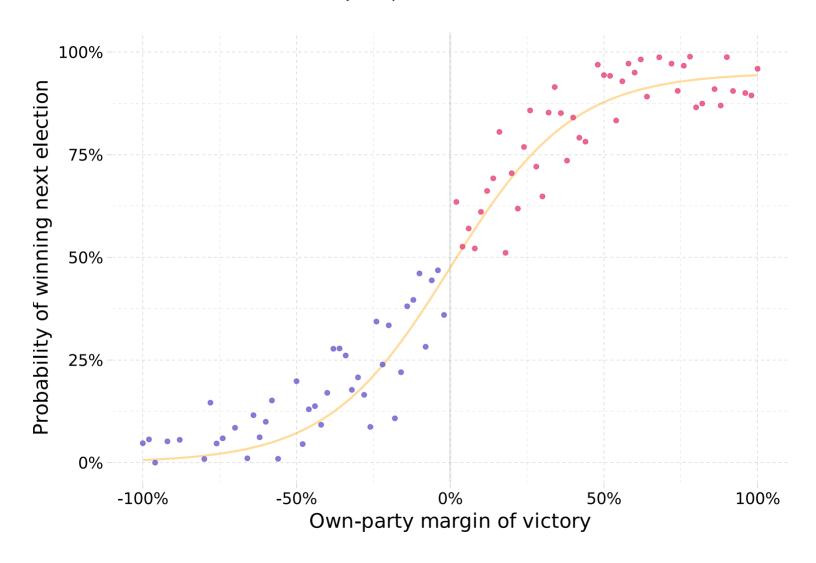
Option 4a Differing local (LOESS) regressions (limited to +/- 50%).

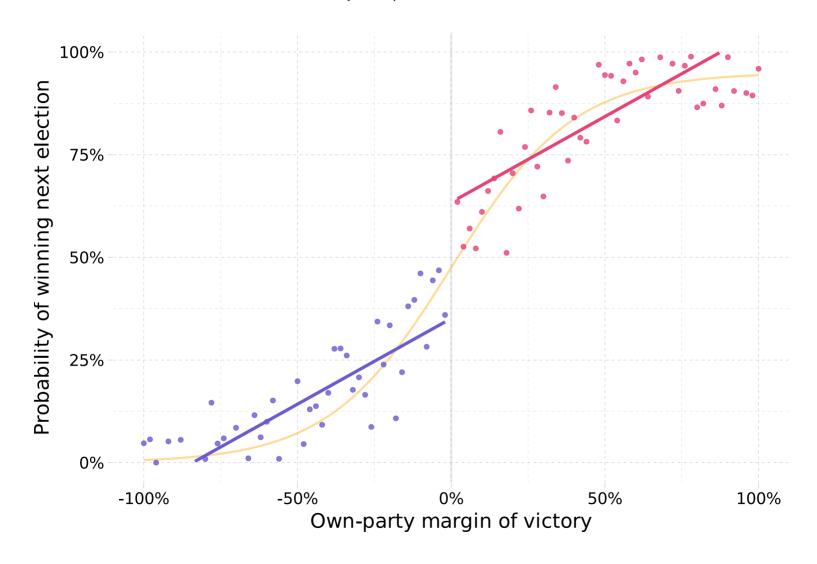


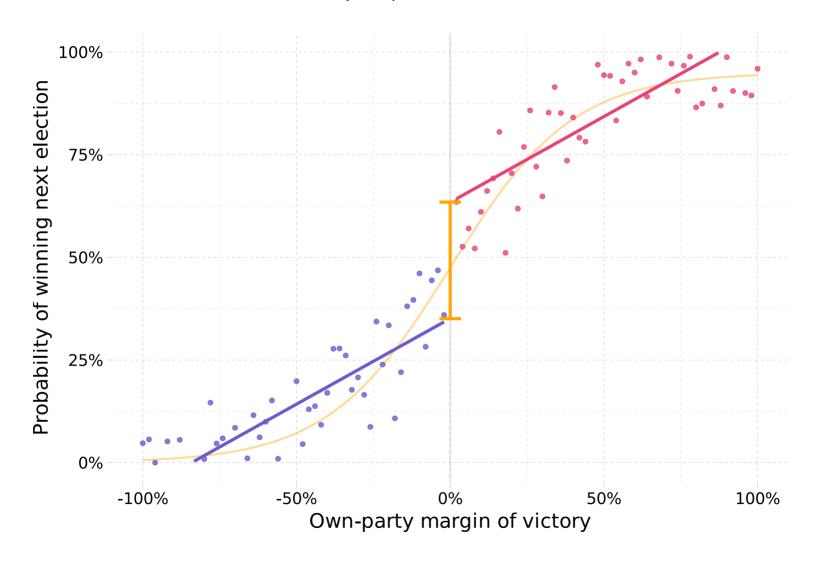
Option 4b Differing local (LOESS) regressions (all data).

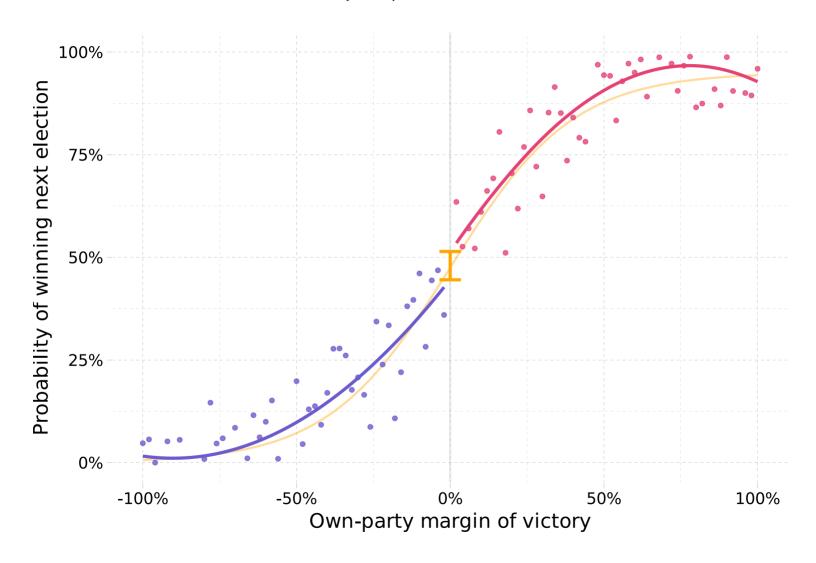




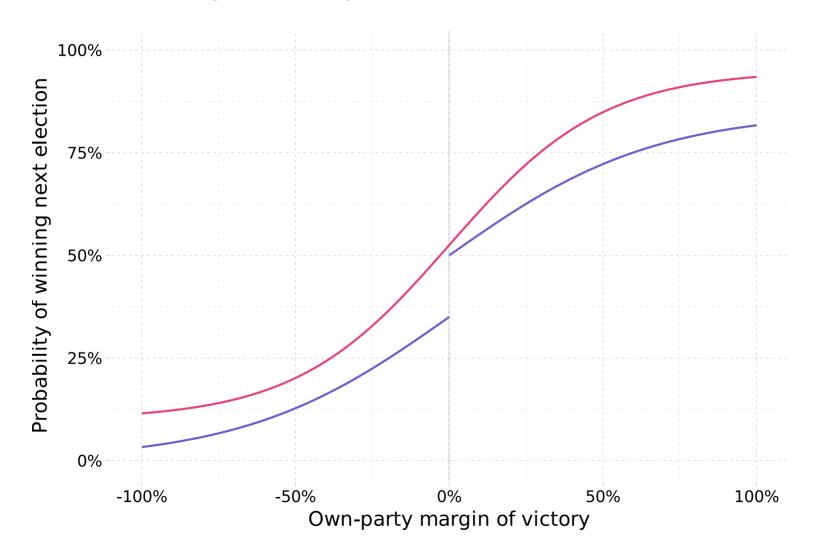








The continuity of $E[Y_{0i} \mid X_i = x]$ (in x) is also very important. No sorting.



In practice

Gelman and Imbens (2018) on functional form:

We argue that controlling for global high-order polynomials in regression discontinuity analysis is a flawed approach with three major problems: it leads to noisy estimates, sensitivity to the degree of the polynomial, and poor coverage of confidence intervals. We recommend researchers instead use estimators based on local linear or quadratic polynomials or other smooth functions.

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See Imbens and Kalyanaraman (2012) for optimal bandwidth selection.

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- 2. **Recode** X_i (the "forcing variable") as deviation from c, i.e., $\widetilde{X}_i = X_i c$
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- 3. Determine a model to **estimate** $Eig[Y_i \mid \widetilde{\mathbf{X}}_iig]$ for $\widetilde{\mathbf{X}}_i$ above and below 0
 - \circ Linear with common slopes for $Eig[\mathrm{Y}_i \mid \widetilde{\mathrm{X}}_i < 0 ig]$ and $Eig[\mathrm{Y}_i \mid \widetilde{\mathrm{X}}_i > 0 ig]$
 - Linear/quadratic/polynomial with differing slopes
 - LOESS, kernel regression, etc.

Estimation: Linear, common slope

Assumptions

- 1. $E[Y_{0i}|X_i=x]$ is linear in x, i.e., $E[Y_{0i}|X_i]=\alpha+\beta X_i$
- 2. Treatment effect does not depend upon \mathbf{X}_i , i.e., $E[\mathbf{Y}_{1i} \mathbf{Y}_{0i} \mid \mathbf{X}_i] = au$

where (1) comes from linearity and (2) comes from common slopes.

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$$E[\mathrm{Y}_i \mid \mathrm{X}_i, \, \mathrm{D}_i] = \mathrm{D}_i \, {\color{red} E[\mathrm{Y}_{1i} \mid \mathrm{X}_i]} + (1 - \mathrm{D}_i) \, {\color{red} E[\mathrm{Y}_{0i} \mid \mathrm{X}_i]}$$

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$$\implies E[Y_{1i} \mid X_i] = \tau + E[Y_{0i} \mid X_i] = \tau + \alpha + \beta X_i$$

Recall our definition of $Y_i = D_i Y_{1i} + (1 - D_i) Y_{0i}$.

$$egin{aligned} E[\mathbf{Y}_i \mid \mathbf{X}_i, \, \mathbf{D}_i] &= \mathbf{D}_i \, oldsymbol{E}[\mathbf{Y}_{1i} \mid \mathbf{X}_i] + (1 - \mathbf{D}_i) \, oldsymbol{E}[\mathbf{Y}_{0i} \mid \mathbf{X}_i] \ &= lpha + au \mathbf{D}_i + eta \mathbf{X}_i = lpha + au \mathbf{D}_i + eta \left(\widetilde{\mathbf{X}}_i + c\right) = \widetilde{lpha} + au \mathbf{D}_i + eta \widetilde{\mathbf{X}}_i \end{aligned}$$

which we can estimate by regressing Y_i on D_i and \widetilde{X}_i .

Estimation: Linear, differing slopes

```
Assumption E[Y_{0i}|X_i=x] and E[Y_{1i}|X_i=x] are linear in x, i.e., E[Y_{0i}\mid X_i]=\alpha_0+\beta_0X_i and E[Y_{1i}\mid X_i]=\alpha_1+\beta_1X_i
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$$E[Y_i \mid X_i, D_i]$$

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$$E[\mathrm{Y}_i \mid \mathrm{X}_i, \, \mathrm{D}_i] = \mathrm{D}_i \, {\color{red} E[\mathrm{Y}_{1i} \mid \mathrm{X}_i]} + (1 - \mathrm{D}_i) \, {\color{red} E[\mathrm{Y}_{0i} \mid \mathrm{X}_i]}$$

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au is the LATE at $\widetilde{\mathbf{X}}_i = 0 \; (\mathbf{X}_i = c)$.

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au is the LATE at $\widetilde{\mathbf{X}}_i=0$ $(\mathbf{X}_i=c)$. Estimate: Regress \mathbf{Y}_i in $\widetilde{\mathbf{X}}_i$, \mathbf{D}_i , and $\mathbf{D}_i\widetilde{\mathbf{X}}_i$.

[†] See Appendix for omitted steps.

Setup

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Formally,

$$0 < \lim_{x \downarrow c} \Pr(\mathrm{D}_i = 1 \mid \mathrm{X}_i = x) - \lim_{x \uparrow c} \Pr(\mathrm{D}_i = 1 \mid \mathrm{X}_i = x) < 1$$

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Ex., Exceeding a minimum GRE requirement for graduate school.

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The treatment effect defined by a fuzzy RD is the ratio of (1) to (2)

$$au_{ ext{FRD}} = rac{\lim_{x\downarrow c} E[ext{Y}_i \mid ext{X}_i = x] - \lim_{x\uparrow c} E[ext{Y}_i \mid ext{X}_i = x]}{\lim_{x\downarrow c} E[ext{D}_i \mid ext{X}_i = x] - \lim_{x\uparrow c} E[ext{D}_i \mid ext{X}_i = x]}$$

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Accordingly, fuzzy RDs are going to have the **same requirements and interpretation as IV**.

More formally

Let $D_i(x^*)$ denote the **potential treatment status** of *i* with threshold x^* .

Why write potential treatment status D_i a function of the threshold?

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This is our monotonicity assumption for fuzzy RDs. If we raise x^* from c to $c+\epsilon$, no one joins treatment—no defiers.

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Compliance

Our compliers in this setting are individuals such that

$$\lim_{x^{*}\downarrow \mathrm{X}_{i}}\mathrm{D}_{i}\left(x^{*}
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Thus, $au_{
m FRD}$ can be a *very local* LATE.

General

RD analyses hinge on their graphical analyses.

If the discontinuity is not graphically apparent, most people are not going to care about the results of a few tortured regressions.

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If the discontinuity is not graphically apparent, most people are not going to care about the results of a few tortured regressions.

You're arguing you know that treatment assignment changes across the threshold. If your reader/viewer cannot see it, they're likely not going to believe your regression tables.[†]

[†] This skepticism may be well founded. We know RDs are sensitive to functional form—and researchers have been known to *p-hack*.

Three figures

Most RD analyses will have some subset of three types of figures.

1. **Outcomes** by the running/forcing variable (X_i)

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- 3. **Density** of running/forcing variable (X_i) Is there evidence of sorting into treatment (across the threshold)?

Outcomes by running variable

These figures tend to show the average value of the outcome Y_i at evenly spaced bins of the running variable X_i .

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- 1. Binwidth (h)
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that yield
$$K=K_0+K_1$$
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$$b_k = c - (K_0 - k + 1) imes h$$

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We then calculate summaries for each bin.

Outcomes by running variable

The bin's **number of observations**, N_k

$$N_k = \sum_{i=1}^N \mathbb{I}\left\{b_k < \mathrm{X}_i \leq b_{k+1}
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The **average outcome** in the bin, \overline{Y}_k

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Outcomes by running variable

We then plot $\overline{\mathbf{D}}_k$ against the midpoint of each bin.

Outcomes by running variable

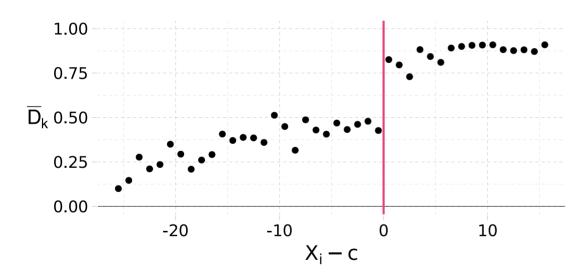
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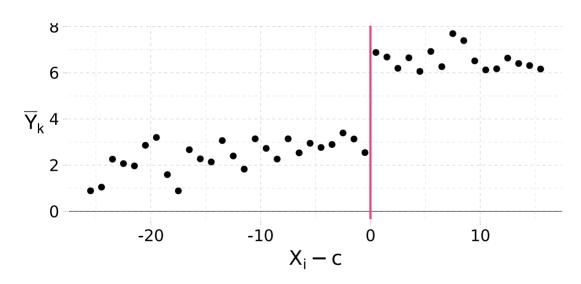
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Outcomes by running variable

And then plot $\overline{\mathbf{Y}}_k$ against the midpoint of each bin.

Q Does crossing c clearly affect our outcome Y_i ? (Fuzzy RD reduced form)



Covariates by running variable

Now we apply the same approach to covariates (\mathbf{Z}_i) .

Covariates by running variable

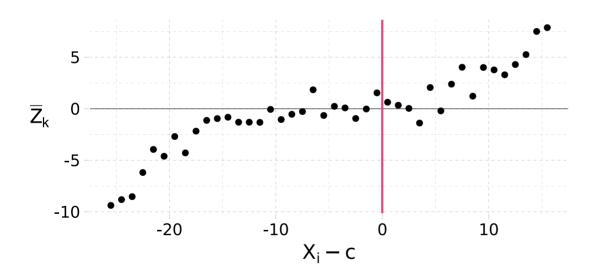
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Now we apply the same approach to covariates (\mathbf{Z}_i) .

 \mathbf{Q} Are covariates **smooth** across \mathbf{c} ? If not, your RD may be invalid.



Density of running variable

Finally we looking for other violations of smoothness—particularly in form gaming the threshold.

In other words: Are individuals **bunching** just above or just below the threshold?

If so, folks just below the threshold don't give us the clean counterfactual that we want for the folks just above the threshold.

McCrary (2008) suggests testing the density of X_i at c.

Density of running variable

Effectively, we can plot N_k at the midpoint of each bin.

Density of running variable

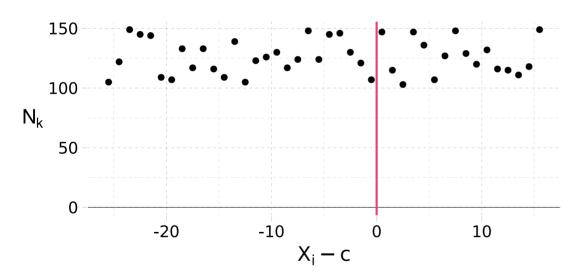
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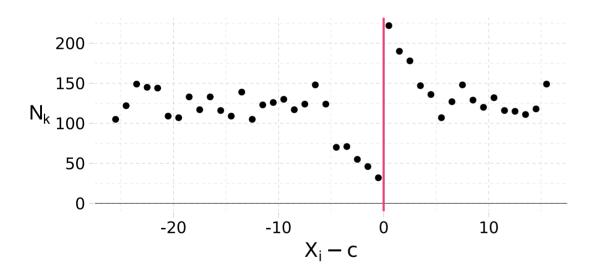
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Density of running variable

Likely bunching (problem)

Q Is the distribution of X_i smooth across c?



Additional points

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Again, if these graphs are not clear and convincing, it's going to be hard to make the case that you have a true/credible discontinuity.

Appendix

Estimation: Linear, differing slopes

Definitions of terms that magically appear

•
$$\widetilde{\alpha} = \alpha_0 + \beta_0 c$$

•
$$\tau = (\alpha_1 - \alpha_0) + (\beta_1 - \beta_0) c$$

•
$$\tilde{\beta} = (\beta_1 - \beta_0)$$

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- 2. Framework
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- 4. Density of X_i

Sharp RDs

- 1. Setup
- 2. (Semi) Formally
- 3. Estimation
- 4. Examples
- 5. In practice
- 6. More estimation

Fuzzy RDs

- 1. Setup
- 2. As IV
- 3. Somewhat formally