## Sparse Prior

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## 1 Sparse Prior

We introduce sparse prior to improve the prediction accuracy of our Naive Bayes classifier. We assign each  $\theta_{y,i}^{(k)}(j)$  with a binary random variable  $Z_{y,i}^{(k)}(j)$ .  $\theta_{y,i}^{(k)}(j) = 0$  when  $Z_{y,i}^{(k)}(j) = 0$ . Fix y, k, j,  $\{\theta_{y,i}^{(k)}(j) : Z_{y,i}^{(k)}(j) = 1\}$  are sampled from Dirichlet distribution as before.

## 1.1 the Model

Fix y, k, denote i as position index for feature vector, and  $j = \{1, 2, ..., 8\}$  as class index, thus

$$\theta_{i,j} := \theta_{y,i}^{(k)}(j)$$
 $Z_{i,j} := Z_{y,i}^{(k)}(j)$ 

$$Z_{i,j} \sim Bernoulli(p_i)$$

$$\theta_{i,j} = \frac{\mu_{i,j} Z_{i,j}}{\sum_{j'=1}^{8} \mu_{ij'} Z_{ij'}}$$

$$\mu_{i,j} \sim Gamma(\alpha_{i,j}, 1)$$

In order to simulate posterior distribution of  $\theta_{ij}$ , we need to simulate joint distribution  $\mathbb{P}(p_i, Z_{ij}, \mu_{ij}|Data)$ . We use Gibbs sampler to simulate this distribution.

## 1.2 Gibbs Sampler

Fix y, k. Let X be a data matrix, and each row represents a feature vector x and Y(x) = y, the conditional distributions for this model are

1. 
$$\mu_{ij}, \forall i, j | Z_{i'j'}, \forall i'j', p_{i'}, \forall i', X$$
  
Fix  $i$ ,

• For j with  $Z_{ij} = 0$ , simulate  $\mu_{ij}$  from prior distribution.

- For j with  $Z_{ij} = 1$ , let  $\mathcal{J} = \{j : Z_{ij} = 1\}$ . First we can simulate  $\{\frac{\mu_{ij}}{\sum_{j' \in \mathcal{J}} \mu_{ij'}} : j \in \mathcal{J}\}$  given data, and this is a Dirichlet distribution with parameters  $(\alpha_{ij} + \# \text{ of occurrence of class } j \text{ for feature } i : j \in \mathcal{J}\}$ . Given  $\{\frac{\mu_{ij}}{\sum_{j' \in \mathcal{J}} \mu_{ij'}} : j \in \mathcal{J}\}$ , we can simulate  $\sum_{j \in \mathcal{J}} \mu_{ij}$  from  $Gamma(\sum_{j \in \mathcal{J}} \alpha_{ij}, 1)$ , where  $\alpha_{ij}$  are posterior Dirichlet parameters. Eventually we get  $\{\mu_{ij} : j \in \mathcal{J}\}$  given  $\{\frac{\mu_{ij}}{\sum_{j' \in \mathcal{J}} \mu_{ij'}} : j \in \mathcal{J}\}$  and  $\sum_{j \in \mathcal{J}} \mu_{ij}$ .
- 2.  $Z_{ij}, \forall ij | \mu_{i'j'}, p_{i'}, \forall i'j', X$

Let  $Z_i$  be a binary vector equals to  $(Z_{ij}: j=1,\ldots,8)$ . Since X is independent across columns (i.e features), We can sample  $(Z_i|\mu_{i'j'},p_{i'},\forall i'j',X)$  independently across different i's.

$$\mathbb{P}(Z_{i}|\mu_{i'j'}, p_{i'}, \forall i'j', X, Z_{i'}, i' \neq i)$$

$$\propto \mathbb{P}(X|\mu_{i'j'}, p_{i'}, \forall i'j', Z_{i}, Z_{i'}, i' \neq i) \times \mathbb{P}(Z_{i}|\mu_{i'j'}, p_{i'}, \forall i'j', Z_{i'}, i' \neq i)$$

$$\propto \mathbb{P}(X|\mu_{i'j'}, p_{i'}, \forall i'j', Z_{i}, Z_{i'}, i' \neq i)$$

$$\propto \mathbb{P}(X_{i}|\mu_{ij'}, p_{i}, \forall j', Z_{i})$$

where  $X_i$  indicates ith column of X. Since  $Z_i$  can only have  $2^8$  possible values, it is a discreate distribution and is easy to find out.

3.  $p_i, \forall i | Z_{i'j'}, \mu_{ij}, X \sim Beta(\alpha_i + \#(Z_{ij} = 1), \beta_i + \#(Z_{ij} = 0))$  where  $\alpha_i, \beta_i$  are parameters for prior distribution of  $p_i$ .

By sampling  $\mu_{ij}$ ,  $Z_{ij}$ ,  $p_i$  iteratively, we can eventually simulate joint distribution  $\mathbb{P}(p_i, Z_{ij}, \mu_{ij}|X)$ .