

Missing Data

Week 3

Simulating Missing Data

Generate data from simple regression model.

Delete some observations by MCAR.

Study the effect of missing data when using:

- Complete cases
- Mean imputation
- Random imputation

R commands for simulating data and for regression

`runif` returns random variable from the Uniform(0, 1)

If different interval is desired use `runif(min = a, max = b)`

`rnorm` returns random variable from the Normal(0, 1)

Different mean and standard deviation can be requested with

`rnorm(mean = m, sd = s)`

`plot(x, y)`

is used to generate a scatterplot of the variables

R commands for simulating data and for regression

```
lm(formula, data)
```

fits "linear model" aka as regression to the data

Syntax is dependent variable $Y \sim \text{ind. Var1} + \text{ind. Var2} + \dots$

A formula has an implied intercept term. To remove this use

```
y ~ x - 1
```

`summary(lm)` provides the estimates of coefficients and various tests and statistics.

`predict()` provides the predicted values of the response

R commands for creating missing data

`mice` package:

```
ampute(data, prop = 0.5, mech = "MAR")
```

generates multivariate missing data

`prop` specifying the proportion of missingness (number of missing cases)

`data` complete data frame or matrix

Result is of type `mads` (multivariate amputed data set)

To extract the “destroyed” data use `$amp`

Outline of Course

- 1) **Missing Data Mechanisms. How are missing data generated and why should we care? Complete Case Analysis. Getting comfortable with R.**
- 2) **Simple missing data fixes: listwise deletion, available case, LVCF, mean imputation, dummy variable methods**
- 3) **More complicated missing data fixes: weighting, hotdecking, regression imputation**
- 4) **Building blocks and overview of multiple imputation (including regression imputation with noise)**
- 5) **Multiple imputation in practice (software in R, simple analyses, and diagnostics)**
- 6) **Multiple imputation in practice (more complicated models and considerations, more advanced diagnostics)**
- 7) **More advanced imputation and other missing data methods**

Notation

Let R be the **matrix** of variables R_1, \dots, R_p , corresponding to variables in our dataset, X_1, \dots, X_p , that indicate whether a given value of the corresponding X variable is observed ($= 1$) or missing ($= 0$)

| X_1 | X_2 | X_3 | X_4 |
|-------|-------|-------|-------|
| 0 | 2 | 5 | 2 |
| 0 | 1 | 3 | ? |
| 0 | 2 | 4 | 3 |
| 0 | 3 | 5 | ? |
| 1 | 5 | ? | ? |
| 1 | 4 | ? | 12 |
| 1 | 6 | 11 | ? |
| 1 | 6 | 12 | 16 |

| R_1 | R_2 | R_3 | R_4 |
|-------|-------|-------|-------|
| 1 | 1 | 1 | 1 |
| 1 | 1 | 1 | 0 |
| 1 | 1 | 1 | 1 |
| 1 | 1 | 1 | 0 |
| 1 | 1 | 0 | 0 |
| 1 | 1 | 0 | 1 |
| 1 | 1 | 1 | 0 |
| 1 | 1 | 1 | 1 |

Notation extension

Sometimes we may explicitly refer to fully observed variables by W_1, \dots, W_k and reserve the notation, X_1, \dots, X_p , only for the variables in our dataset with missing data.

| W_1 | W_2 | X_1 | X_2 |
|-------|-------|-------|-------|
| 0 | 2 | 5 | 2 |
| 0 | 1 | 3 | ? |
| 0 | 2 | 4 | 3 |
| 0 | 3 | 5 | ? |
| 1 | 5 | ? | ? |
| 1 | 4 | ? | 12 |
| 1 | 6 | 11 | ? |
| 1 | 6 | 12 | 16 |

| R_1 | R_2 |
|-------|-------|
| 1 | 1 |
| 1 | 0 |
| 1 | 1 |
| 1 | 0 |
| 0 | 0 |
| 0 | 1 |
| 1 | 0 |
| 1 | 1 |

More advanced missing data methods

Methods that throw away data

- Non-response weighting

Methods that don't throw away data

- Hotdecking
- Regression imputation
- Regression imputation with noise (next time)

Non-response weighting

Survey weighting background

A scientifically sound method for creating a roster of survey participants is to conduct a random sample in which n units are selected at random from a list (sample frame) of N population units, each with a known and non-zero selection probability, π_i . When $\pi_i = \pi = n/N$ for all i , the method is termed simple random sampling, but equal probabilities of selection is not mandatory — alternative, unequal sample designs can be more precise than simple random sampling depending on the population structure and estimator(s) of interest (Cochran, 1977).

Non-response weighting

- Suppose only one variable has missing data
- Calculate the probability $\Pr(R_i | X_i)$ of a value being missing using observed values from the other variables;
- Use these predicted probabilities to create survey weights of the form $1/\Pr(R_i | X_i)$ to make the complete case sample representative of the full sample once again;
- Typically we normalize by multiplying the weights by the overall (marginal) probability of missingness, $\Pr(R_i)$. This way the weights will sum to the number of people left in the complete case sample;
- This becomes more complicated when there is more than one variable with missing data
- Probabilities close to 1 or 0 can lead to crazy standard errors (due to extreme weights) though there are ways of “stabilizing” the weights

Non-response weighting

complete case dataset
plus weights

| X_1 | X_2 | X_3 | X_4 |
|-------|-------|-------|-------|
| 0 | 1 | 7 | 3 |
| 0 | 1 | 7 | 3 |
| 0 | 1 | 7 | 3 |
| 0 | 1 | 7 | 3 |
| 1 | 8 | 2 | 1 |
| 1 | 8 | 2 | ? |
| 1 | 8 | 2 | ? |
| 1 | 8 | 2 | ? |
| 1 | 8 | 2 | ? |
| 0 | 1 | 7 | ? |



| X_1 | X_2 | X_3 | X_4 | Weight |
|-------|-------|-------|-------|--------|
| 0 | 1 | 7 | 3 | .625 |
| 0 | 1 | 7 | 3 | .625 |
| 0 | 1 | 7 | 3 | .625 |
| 0 | 1 | 7 | 3 | .625 |
| 1 | 8 | 2 | 1 | 2.5 |

$$\Pr(R_4 = 1 \mid X_1 = 0, X_2 = 1, X_3 = 7) = .2$$

$$\Pr(R_4 = 1 \mid X_1 = 1, X_2 = 8, X_3 = 2) = .8$$

$$\Pr(R_4 = 1) = .5$$

$$\text{weight}(0, 1, 7) = .5/.2 = 2.5$$

$$\text{weight}(1, 8, 2) = .5/.8 = .625$$

note that Σ weights = 5 which is our
complete case sample size

Extensions when several variables have missing data

- When several (or many) variables have missing data a similar approach can be used.
- In this case the indicator to be predicted, R^{cc} , is for whether or not an observation belongs to the complete case sample (or not)
- We can model $\Pr(R^{cc} | W)$, where W denotes fully observed variables.
- Another option is to create weights to address drop-out over time in a panel dataset. Here the missing data indicator would indicate attrition from the sample (dropping out of the study).
- This is basically what is going on with survey weights for attrition in large public-use surveys.

Hotdecking

Hotdecking – basic idea

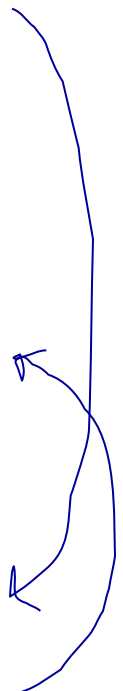
- Replaces missing values using other values found in the dataset
- For each person with a missing value on variable Y , find another person who has all the same values (or close to the same values) on observed variables X_1, X_2, X_3, \dots , and use that person's Y value.
- The R package VIM has a hotdeck command (we will get back to this when we begin using imputation packages)
- Stata has a hotdeck command (have to download) that randomly samples within strata defined by fully observed categorical variables

Hotdecking

| X_1 | X_2 | X_3 | X_4 |
|-------|-------|-------|-------|
| 0 | 3 | 8 | 2 |
| 0 | 4 | 9 | 3 |
| 1 | 2 | 7 | 3 |
| 0 | 4 | 6 | ? |
| 1 | 5 | 3 | 1 |
| 0 | 3 | ? | 2 |
| 0 | 4 | 6 | 2 |



| X_1 | X_2 | X_3 | X_4 |
|-------|-------|-------|-------|
| 0 | 3 | 8 | 2 |
| 0 | 4 | 9 | 3 |
| 1 | 2 | 7 | 3 |
| 0 | 4 | 6 | 2 |
| 1 | 5 | 3 | 1 |
| 0 | 3 | 8 | 2 |
| 0 | 4 | 6 | 2 |



Hotdecking – properties

- Typically thought of a nonparametric in that it avoids making modeling assumptions
- Census has used this technique for years
- Is potentially ok in terms of bias
- Gets complicated if many variables have missing data
- Can be problematic if there is too much missing data

Hotdecking

- Hotdecking actually comes in many varieties. The previous is a form of “exact matching” algorithm.
- However there are other ways to determine the “distance” between two observations.

Regression Imputation

- Suppose we deal with univariate missing data and we call the variable with missingness Y and the rest are denoted \mathbf{X}
- Let \mathbf{X}_{obs} be the subset of n_1 rows of \mathbf{X} for which Y is observed.
- Let \mathbf{X}_{mis} be the complementing subset of n_0 rows of \mathbf{X} for which Y is missing.
- The vector containing the n_1 observed data in Y is denoted by y_{obs}
- Finally, let the vector of n_0 imputed values in Y be indicated by \hat{Y}

Regression Imputation

- Within the complete cases X_{obs} , build a model that predicts the values Y
- Use this model within the n_0 cases with missing data X_{mis} to predict (impute) Y
- That is,

$$\hat{Y} = X_{\text{mis}}\hat{\beta}$$

where $\hat{\beta}$ is the LSE of β calculated from the n_1 cases of y_{obs} and X_{obs}

- This method becomes much more complicated when there are many variables with missing data
- At best, this method will underestimate standard errors

Regression imputation: Example

| | X_1 | X_2 | X_3 | Y |
|------------------|-------|-------|-------|-----|
| | 0 | 3 | 8 | 2 |
| | 0 | 4 | 9 | 3 |
| | 0 | 3 | 7 | 3 |
| X_{mis} | 0 | 4 | 6 | NA |
| | 1 | 5 | 10 | NA |
| | 1 | 3 | 7 | NA |
| | 1 | 2 | 5 | NA |
| | 1 | 2 | 7 | 6 |
| | 1 | 3 | 6 | 7 |
| | 1 | 4 | 8 | 9 |
| | 1 | 5 | 11 | 8 |

In this case we have $n_1 = 7$ and $n_0 = 4$

\dot{Y} to be imputed here

Regression imputation: Example

After running the `lm` function in R:

Coefficients:

| | Estimate | Std. Error | t value | Pr(> t) |
|-------------|----------|------------|---------|----------|
| (Intercept) | 1.6000 | 1.4652 | 1.092 | 0.3547 |
| x1 | 4.6000 | 0.5292 | 8.693 | 0.0032 |
| x2 | 1.4000 | 0.5292 | 2.646 | 0.0773 |
| x3 | -0.4500 | 0.3149 | -1.429 | 0.2483 |

Then for example, for observation 5:

$$1.6 + 1(4.6) + 5(1.4) + 10(-0.45) = 8.7$$

Remark: In mice this method is available as method `norm.predict`

| X_1 | X_2 | X_3 | X_4 |
|-------|-------|-------|-------|
| 0 | 3 | 8 | 2 |
| 0 | 4 | 9 | 3 |
| 0 | 3 | 7 | 3 |
| 0 | 4 | 6 | 4.5 |
| 1 | 5 | 10 | 8.7 |
| 1 | 3 | 7 | 7.25 |
| 1 | 2 | 5 | 6.75 |
| 1 | 2 | 7 | 6 |
| 1 | 3 | 6 | 7 |
| 1 | 4 | 8 | 9 |
| 1 | 5 | 11 | 8 |

Beyond Regression Imputation

- Regression imputation with noise (next time):

$$\dot{Y} = X_{\text{mis}}\hat{\beta} + \dot{\varepsilon}$$

where $\dot{\varepsilon}$ is randomly drawn from the normal distribution as $\dot{\varepsilon} \sim N(0, \hat{\sigma}^2)$. In mice this method is available as `method norm.obs`

- Bayesian multiple imputation:

$$\dot{Y} = X_{\text{mis}}\dot{\beta} + \dot{\varepsilon}$$

where $\dot{\varepsilon} \sim N(0, \dot{\sigma}^2)$ and $\dot{\beta}$ and $\dot{\sigma}^2$ are random draws from their posterior distribution, given the data. In mice this method is available as `method norm`