

Chapter 2

Simulations

2.1 Introduction

In this chapter, we carried-out simulations with synthetic patient arrival datasets, and compares the posterior distribution results and prediction rate between Independent Bayes Model (IBM) and Hierarchical Bayes Model (HBM) in scenarios of with different anomalies and different hierarchy structures.

Sections 2.1 details functions used to synthesis simulation datasets. Sections 2.2 evaluates the use of different priors for Bayesian models to be used during the simulations. Section 2.3 explores the impact of size of anomaly towards anomaly detection at different levels of hierarchy. Section 2.4 and section 2.5 considers how complexity of the hierarchical structure affects anomaly detection at different levels of hierarchy.

2.2 Simulation Methods

Synthetic datasets were used for the simulation studies, synthetic datasets refer to data information that's artificially manufactured rather than generated from real-world events. Synthetic data was preferred over the alternative option, perturbed data, dataset that is generated by adding alteration and noise to real-word data. (Drechsler et al., 2008; Drechsler, 2011) Advantages of synthetic data includes: (1) Disclosure protection, as medical dataset almost always contained sensitive information about individuals.(2) Reduce the effects from possible background noises presented in real-word data. (3) Allows control over the hierarchical structure of datasets, note that real-world hierarchical data are often extremely complex and can have lots of levels, and complexity that comes with it, so we would prefer use of simpler structures for simulation studies.

	index	EventID	time.char	time.day	time.int	cat1	cat2	leaf	A	B	AA	AB	BA	BB
1	1	72627125	2006-01-01 01:50:09	2006-01-01	13149	A	A	AA	1	0	1	0	0	0
2	2	86476860	2006-01-01 04:50:19	2006-01-01	13149	A	B	AB	1	0	0	1	0	0
3	3	29910729	2006-01-01 06:25:41	2006-01-01	13149	A	A	AA	1	0	1	0	0	0
4	4	93841355	2006-01-01 06:57:32	2006-01-01	13149	B	A	BA	0	1	0	0	1	0
5	5	42626823	2006-01-01 07:01:39	2006-01-01	13149	A	B	AB	1	0	0	1	0	0
6	6	35263557	2006-01-01 08:22:15	2006-01-01	13149	A	B	AB	1	0	0	1	0	0
7	7	42034666	2006-01-01 08:44:04	2006-01-01	13149	A	A	AA	1	0	1	0	0	0
8	8	30904123	2006-01-01 09:52:56	2006-01-01	13149	B	A	BA	0	1	0	0	1	0
9	9	07098006	2006-01-01 09:58:03	2006-01-01	13149	B	A	BA	0	1	0	0	1	0
10	10	29415173	2006-01-01 10:50:32	2006-01-01	13149	A	A	AA	1	0	1	0	0	0
11	11	80281559	2006-01-01 10:55:57	2006-01-01	13149	B	B	BB	0	1	0	0	0	1
12	12	70606392	2006-01-01 11:54:54	2006-01-01	13149	B	A	BA	0	1	0	0	1	0
13	13	97423382	2006-01-01 12:09:38	2006-01-01	13149	B	B	BB	0	1	0	0	0	1
14	14	07248823	2006-01-01 13:51:19	2006-01-01	13149	B	A	BA	0	1	0	0	1	0
15	15	99860761	2006-01-01 15:11:29	2006-01-01	13149	A	B	AB	1	0	0	1	0	0
16	16	54655430	2006-01-01 17:16:19	2006-01-01	13149	A	A	AA	1	0	1	0	0	0

Figure 2.1: Entries of the raw dataset

Several custom R-functions were created to synthesis datasets used in this study, Descriptions of the functions are provided in Appendix A.1. There are several reasons for creating custom simulation synthesis functions. Firstly, a large number of simulated datasets with variable settings were expected to be required to perform our analyses, this required a large number of custom function arguments that existing packages may not contain. Secondly, synthetic datasets was simulated to mimic a health dataset in a pure theoretical setting. Thirdly, an emphasis was placed on the hierarchical nature of certain variable of dataset (for example, ICD codes and geographic location), brunch and leaf variable was designed to be generated with a custom structures during the simulation process. Lastly, as this is an exploration research it was unclear what functionalities were needed, existing statistical data simulation packages may become insufficient at later stage.

`simdata` and `adddaily.anomaly` are the two major function used in the data synthesis process. `simdata` function were used to synthesise a raw dataset (figure 2.1) that simulate a hospital record that you would typically found in hospital digital archives, where each row correspond to a single entry of hospital event with a unique hospital event identifier. Time, dates and various information about the hierarchical structure are automatically generated by the function. The default number of simulation is 1,000,000, and default period is a 12 year long period between 1 January 2006 and 31 December 2018, and a matrix that contains information of the hierarchical structure are manipulated and used to produce hierarchical structure with different characteristics.

The particular matrix used in the `simdata` function contains the theoretical value of each leaf of a two level hierarchical structure in a proportion out of 1000. As shown with the example in figure 2.2, if the theoretical value at of the leafs (at level 2) are

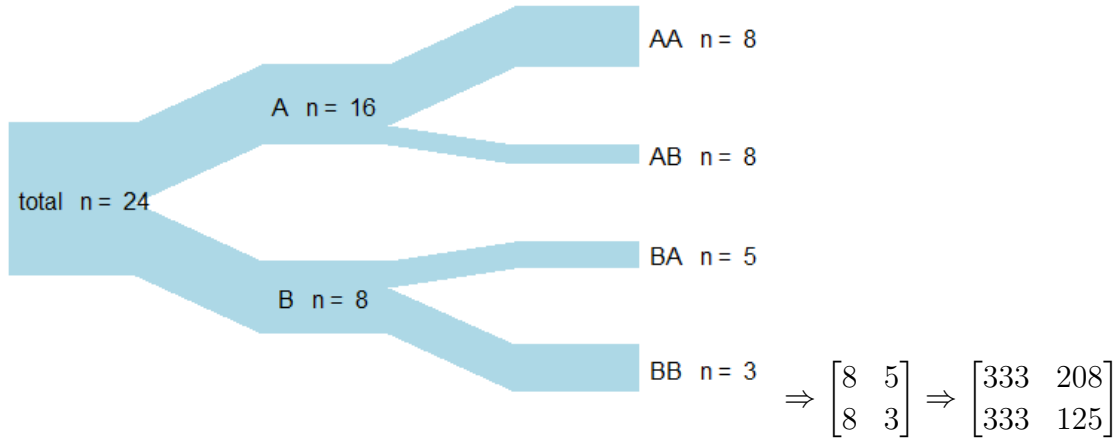


Figure 2.2: Conversion of a simple hierarchical structure to a matrix

8, 8, 5, and 3, you would represent this information with a matrix with numbers 333,333,208 and 125. each column correspond to a single group, therefore column sums of the matrix correspond to the count of a level 1 categories, and the sum of all numbers in the matrix equal to the total count. The parent levels are simply the sum of children levels, the complete hierarchical structure can be generated automatically just from numerical information of leafs. Note that the number of column of the matrix correspond to the maximum number of leafs within a level 1 category, for level 1 categories with less than maximum amount of leafs, 0 is used to represent no leaf. The default value of the matrix is 250, 250, 250 and 250, which represent a two level hierarchical of equal proportions, with 2 level 1 categories, each with 2 level 1 categories of equal proportions. This way of storing hierarchical information is simple and efficient but can not be used for hierarchical models that contain 3 or more levels, and may be a little confusing for other people at first, future improvements of the function could be allowing usage of a user-generated hierarchical time series (**hts**) object from the **hts** package instead of the matrix.

`adddaily.anomaly` were used to `addanomaly` function to add anomaly to daily count data, a hierarchical time-series dataset generated by tabulating the raw dataset using `tabulatedata` function. The default setting is a point anomaly of 1 day, on 2015-11-15. The amount of anomaly added in terms of percentage of total count, and amount of anomaly assigned to each leaf is varied.

For point anomalies The amount of anomaly assigned to each leaf is transcribed with a binding of 3 vector, a vector of percentage of anomaly to be added to total, a percentage vector of anomaly to be added to level 1 categories and a vector of

	day	total	A	B	AA	AB	BA	BB
1	2006-01-01	22	11	11	7	4	7	4
2	2006-01-02	23	13	10	9	4	4	6
3	2006-01-03	29	13	16	9	4	10	6
4	2006-01-04	28	16	12	7	9	4	8
5	2006-01-05	20	10	10	6	4	5	5
6	2006-01-06	28	10	18	4	6	7	11
7	2006-01-07	24	16	8	8	8	5	3
8	2006-01-08	23	13	10	7	6	6	4
9	2006-01-09	20	11	9	5	6	3	6

Figure 2.3: Entries of the daily count dataset

percentage of anomaly to be added to level 1 categories. For example, for a hierarchical time series with the structure as the example given in figure 2.2, if we want to add all anomaly to AA, the percentage vector of anomaly for total is $\{1\}$ and this will be the case for any hierarchical structure, the percentage vector of anomaly for level 1 categories is $\{1, 0\}$ and the percentage vector for level 2 categories is $\{1, 0, 0, 0\}$, so we just need to add a percentage vector $\{1, 1, 0, 1, 0, 0, 0\}$.

$$\begin{array}{ccccccc}
 P_{total} & P_{lv1} & & P_{lv2} & & & P \\
 [1] + [1 \ 0] + [1 \ 0 \ 0 \ 0] \Rightarrow [1 \ 1 \ 0 \ 1 \ 0 \ 0 \ 0]
 \end{array}$$

Figure 2.4: percentage vector

Details about setups used for the data synthesis process can be found in the start of each section, for each of our simulation tests, and the R codes used can be found in `91-createdata-test.R`, `92-createdata-anoamly.R`, `93-createdata-proportion.R` and `94-createdata-count.R` found in GitHub Repository <https://github.com/jungxue/research-masters-Jung>.

2.3 Simulation 1: Prior selection

2.3.1 Priors

For Bayesian inference to make sense a prior is essentially required, however there are different ways to generate the prior. For simulations, no external informations are available, therefore the informative prior approach is out of the question. However it is possible for us to assign relatively simple weak-informative and non-informative priors by generating probability distributions with vaguely defined parameters. And therefore, we proposed 6 different priors that express different hypothesised information about the prior.

A simple Poisson process can be used to model for a random, and mutually independent arrivals rate, with the likelihood model:

$$\begin{aligned} y_{i,t} &\sim \text{Poisson}(\mu_{i,t}) \\ \log(\mu_{i,t}) &= \log(\rho_{i,t}) \end{aligned} \tag{2.1}$$

For the Poisson model, our observations $y_{i,t}$ (y of i th category given time t) is hypothesised to follow the distribution of a Poisson process, with parameter $\mu_{i,t}$ as the expected number of arrivals per time-interval. For our simulations this time interval is setup to be a 12 year long period between 1 January 2006 and 31 December 2018. Priors can then assigned to $\mu_{i,t}$, this provides information to the parameter and will have significant impact in our calculations. Samples of the JAG codes used to build our models are presented in appendix A.1.8, and complete JAGS codes and R codes used to obtain our graphical results are available on GitHub repository: <https://github.com/jungxue/research-masters-Jung>.

Model 1 Non-informative Prior

Model 1 is a Independent Bayesian Model with no priors (IBM) that models each categories independently on all hierarchical levels. This model provides no information for the estimation of distribution parameter $\mu_{i,t}$, and therefore the prior can be defined as a non-informative prior, here in this thesis we call it the null prior. The likelihood model is:

$$\begin{aligned} y_{i,t} &\sim \text{Poisson}(\mu_{i,t}) \\ \log(\mu_{i,t}) &= \log(\rho_{i,t}) \end{aligned} \tag{2.2}$$

Model 2 Normal(1,0.3) Prior

Model 2 is a Independent Bayesian Model with $\rho \sim Normal(1, 0.3)$ priors that models each categories independently on all hierarchical levels. This model does contain information for the estimation of distribution parameter μ, t , and therefore the prior can be defined as a weak-informative prior. $T_{[\lambda > 0]}$ indicates a truncation of distribution at 0 in which λ values less than 0 is truncated, therefore λ should not contain any negative values. The likelihood model is:

$$\begin{aligned} y_{i,t} &\sim Poisson(\mu_{i,t}) \\ \log(\mu_{i,t}) &= \log(\rho_{i,t}\lambda_{i,t}) \end{aligned} \tag{2.3}$$

Priors for model parameters:

$$\lambda_{i,t} \sim Normal(\mu_{i,t}, \sigma_{i,t})T_{[\lambda > 0]} \tag{2.4}$$

Hyper-priors for model parameters:

$$\begin{aligned} \mu_{i,t} &\sim Normal(1, 0.1) \\ \sigma_{i,t}^2 &\sim Normal(0.3, 0.1) \end{aligned} \tag{2.5}$$

Model 3 Normal(1,0.1) Prior

Model 3 is a Independent Bayesian Model with $\rho \sim Normal(1, 0.1)$ priors that models each categories independently on all hierarchical levels. This model does contain information for the estimation of distribution parameter μ, t , and therefore the prior can be defined as a weak-informative prior. $T_{[\lambda > 0]}$ indicates a truncation of distribution at 0 in which λ values less than 0 is truncated, therefore λ should not contain any negative values. The likelihood model is:

$$\begin{aligned} y_{i,t} &\sim Poisson(\mu_{i,t}) \\ \log(\mu_{i,t}) &= \log(\rho_{i,t}\lambda_{i,t}) \end{aligned} \tag{2.6}$$

Priors for model parameters:

$$\lambda_{i,t} \sim Normal(\mu_{i,t}, \sigma_{i,t})T_{[\lambda > 0]} \tag{2.7}$$

Hyper-priors for model parameters:

$$\begin{aligned} \mu_{i,t} &\sim Normal(1, 0.1) \\ \sigma_{i,t}^2 &\sim Normal(0.1, 0.1) \end{aligned} \tag{2.8}$$

Model 4 Gamma(4,3) Prior

Model 4 is a Independent Bayesian Model with $\rho \sim \text{Gamma}(4, 3)$ priors that models each categories independently on all hierarchical levels. This model does contain information for the estimation of distribution parameter μ, t , and therefore the prior can be defined as a weak-informative prior. $T_{[\lambda > 0]}$ indicates a truncation of distribution at 0 in which λ values less than 0 is truncated, therefore λ should not contain any negative values. The likelihood model is:

$$\begin{aligned} y_{i,t} &\sim \text{Poisson}(\mu_{i,t}) \\ \log(\mu_{i,t}) &= \log(\rho_{i,t} \lambda_{i,t}) \end{aligned} \tag{2.9}$$

Priors for model parameters:

$$\lambda_{i,t} \sim \text{Gamma}(\alpha_{i,t}, \beta_{i,t}) T_{[\lambda > 0]} \tag{2.10}$$

Hyper-priors for model parameters:

$$\begin{aligned} \alpha_{i,t} &\sim \text{Normal}(4, 0.1) \\ \beta_{i,t} &\sim \text{Normal}(3, 0.1) \end{aligned} \tag{2.11}$$

Model 5 Laplace(1,1) Prior

Model 5 is a Independent Bayesian Model with $\rho \sim \text{Laplace}(1, 1)$ priors that models each categories independently on all hierarchical levels. This model does contain information for the estimation of distribution parameter μ, t , and therefore the prior can be defined as a weak-informative prior. $T_{[\lambda > 0]}$ indicates a truncation of distribution at 0 in which λ values less than 0 is truncated, therefore λ should not contain any negative values. The likelihood model is:

$$\begin{aligned} y_{i,t} &\sim \text{Poisson}(\mu_{i,t}) \\ \log(\mu_{i,t}) &= \log(\rho_{i,t} \lambda_{i,t}) \end{aligned} \tag{2.12}$$

Priors for model parameters:

$$\lambda_{i,t} \sim \text{Laplace}(\alpha_{i,t}, \beta_{i,t}) T_{[\lambda > 0]} \tag{2.13}$$

Hyper-priors for model parameters:

$$\begin{aligned} \alpha_{i,t} &\sim \text{Normal}(4, 0.1) \\ \beta_{i,t} &\sim \text{Normal}(3, 0.1) \end{aligned} \tag{2.14}$$

Model 6 Mixture Prior

Lastly, Model 6 is a Independent Bayesian Model with a mixture priors that models each categories independently on all hierarchical levels. It is a mixture of no information and weak information, in other words, we are proposing that for most of the time there are no variation, and for some of the time there are informations. The idea comes from the prior Berry and Berry (2004) where they proposed a scenario where majority of difference is zero and there are just some variations. However note that their prior is one sided and center at 0, our prior is two sided and center at 1. This model does contain information for the estimation of distribution parameter μ, t , and therefore the prior can be defined as a weak-informative prior. $T_{[\lambda > 0]}$ indicates a truncation of distribution at 0 in which λ values less than 0 is truncated, therefore λ should not contain any negative values. The likelihood model is:

$$\begin{aligned} y_{i,t} &\sim \text{Poisson}(\mu_{i,t}) \\ \log(\mu_{i,t}) &= \log(\rho_{i,t}\lambda_{i,t}) \end{aligned} \tag{2.15}$$

Priors for model parameters:

$$\lambda_{i,t} \sim \text{Normal}(\text{spike}_{i,t} + (1 - \text{spike}_{i,t})\text{slab}_{i,t}, 0.1)T_{[\lambda > 0]} \tag{2.16}$$

Hyper-priors for model parameters:

$$\begin{aligned} \text{spike}_{i,t} &\sim \text{Binomial}(0.9, 1) \\ \text{slab}_{i,t} &\sim \text{Normal}(\mu_{i,t}, \sigma_{i,t}) \end{aligned} \tag{2.17}$$

Hyper-hyper-priors for model parameters:

$$\begin{aligned} \mu_{i,t} &\sim \text{Normal}(1, 0.1) \\ \sigma_{i,t}^2 &\sim \text{Normal}(0.1, 0.1) \end{aligned} \tag{2.18}$$

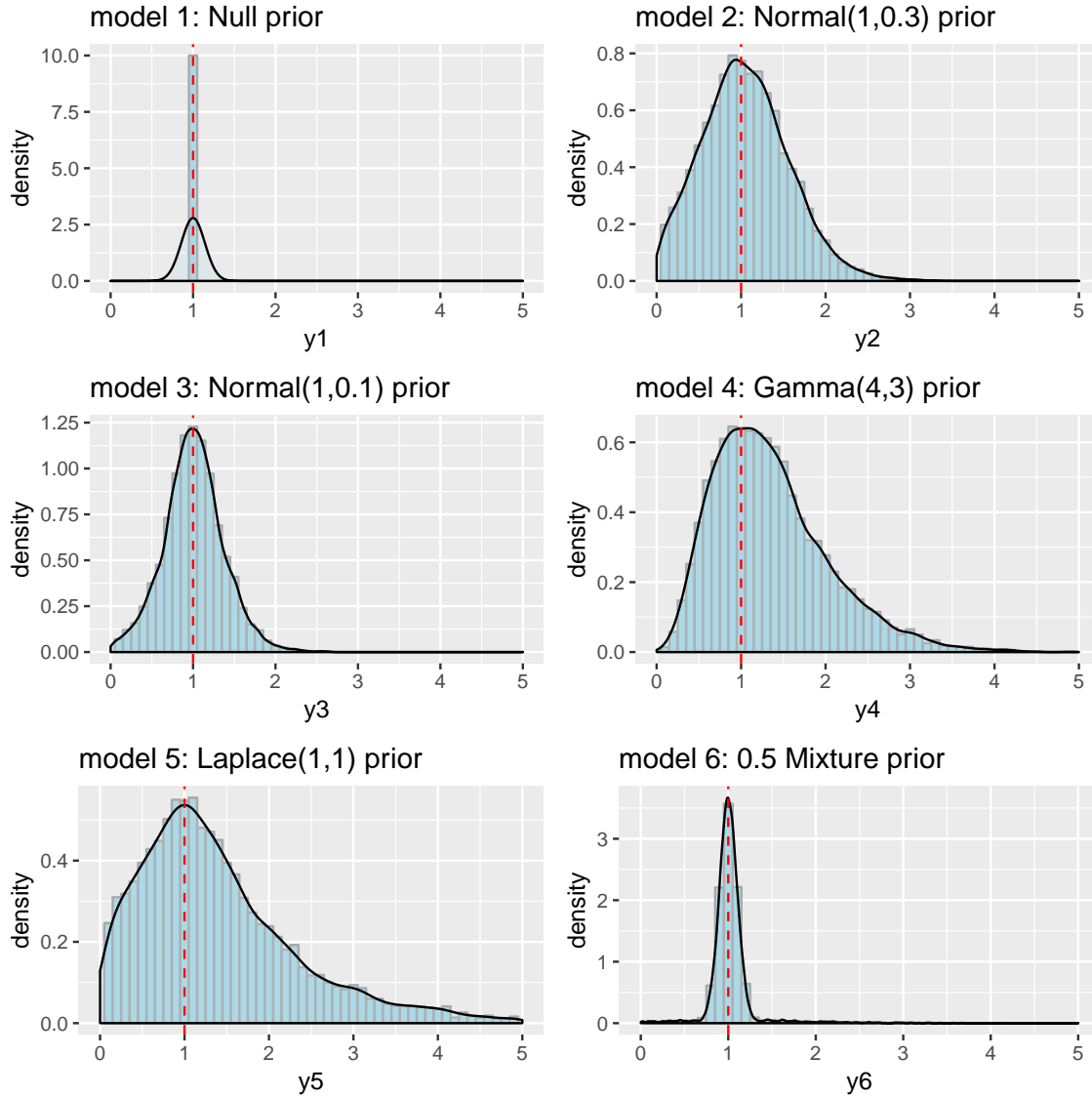


Figure 2.5: Histo-density plot of the prior probability distribution of proposed prior density models. The red dashed line indicates the center point of these distributions, which is 1.

Figure 2.5 show the distribution of λ , because we assign prior distribution to our μ by directly multiplying by λ , the distribution of λ are the equivalent to the distribution of our prior. As it is presented, our priors have different distributions that suggest a different believe. Model 1 is a non-informative prior, that contain no information, and hence no distribution is observed. Model 2 is a $Normal(1, 0.3)$ prior that seem reasonably, however becaus eit's left tail was truncated at 0, the symmetrical property has been lost. Model 3 is a $Normal(1, 0.1)$ that has a smaller variance, hence the distribution is a little bit taller, indicating a slight more observations tend to be close to the mean. Model 4 is a $Gamma(4, 3)$ prior, It skew towards the right hand

side, this translates into we are observing more positive anomalies. Model 5 is a $\text{Laplace}(4,3)$ prior, it also skew to the right but with a heavier tail and a much higher distribution at extreme positive values, This translates into information that we are consistently observing a large anomaly. Lastly we have the mixture model, Information from

2.3.2 Simulation setups

In section 2.3.2 we presents the results of the posterior distribution calculations for the synthesised data, with anomalies added on category AA on level 2 (leaf) of the hierarchy. Default setting for the data synthesis process is used and a 25% anomaly is added. Note the 25% anomaly meant 25% of the count of total category on top of the hierarchy, not an addition of 25% on Category AA.

2.3.3 Results

In section 2.3.2 we presents the results of the posterior distribution calculations for the synthesised data, with anomalies added on category AA on level 2 (leaf) of the hierarchy. Default setting for the data synthesis process is used and a 255 anomaly is added. Note the 25% anomaly meant 25% of the count of total category on top of the hierarchy, not an addition of 25% on Category AA.

	Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
M1:M2	-2.126	-1.574	-0.599	-0.868	-0.158	0.113
M1:M3	-4.456	-3.295	-2.069	-2.032	-0.650	0.189
M1:M4	-2.992	-1.729	-1.174	-1.303	-0.800	0.106
M1:M5	-2.645	-1.745	-0.764	-1.172	-0.590	-0.124
M1:M6	-3.383	-2.200	-1.645	-1.683	-1.047	-0.258
M2:M3	-4.555	-2.288	-1.884	-1.164	0.713	1.440
M2:M4	-2.393	-1.009	-0.552	-0.434	-0.155	2.232
M2:M5	-2.382	-0.494	-0.071	-0.304	0.267	0.782
M2:M6	-3.331	-1.571	-0.626	-0.815	-0.122	1.639
M3:M4	-2.450	-0.285	0.895	0.730	1.787	3.658
M3:M5	-1.020	-0.578	-0.168	0.860	2.345	3.677
M3:M6	-1.995	-1.322	-0.412	0.349	1.649	4.198
M4:M5	-1.884	-0.905	0.501	0.131	0.913	2.282
M4:M6	-2.858	-1.537	-0.829	-0.380	1.029	2.039
M5:M6	-3.259	-0.929	-0.672	-0.511	0.378	1.454

Table 2.1: DIC (with 1,000,000 iteration) summary distribution cross comparison results for model 1 to model 6

DIC is used to compare between models

Table 2.1 gives the comparison of DIC distributions. A important note is that initial dic results with 1000 iterations are very unstable, showing signs of slow convergence, this meant that the dic values for each model differ by a large margin each time dic is ran for different jags models, therefore it is very hard to decide which model is the best.

so the number of iterations was increased and increasing iterations to 1,000,000 does not still yield unstable DIC. The problem has been raised

and other research students from UoA stats lab also encountered

Quantifying uncertainty in marginal likelihood estimation via the stepping-stone sampling algorithm John Siry (Renate Meyer, Patricio Maturana Russell)

However, what has been observed is that at least for the numerous trials, model 1 always yields inferior DIC compared to all other models

	mean	sd	2.5%	50%	97.5%	Rhat	n.eff
model1	21.0615	0.0000	21.0615	21.0615	21.0615		0.0000
model2	24.4976	4.7805	15.7987	24.2398	34.1927	1.0000	3119.0000
model3	24.3373	4.9322	15.7804	23.9781	34.8648	1.0000	3009.0000
model4	24.1116	4.8248	15.5574	23.7915	34.4035	1.0000	3587.0000
model5	24.4041	4.9549	15.9001	24.1139	34.7204	1.0000	3152.0000
model6	24.8840	4.9712	16.4075	24.5082	35.4618	1.0000	2895.0000

Table 2.2: Posterior distributions of different models for Total, with added anomalies, and calculated with independent Bayes model

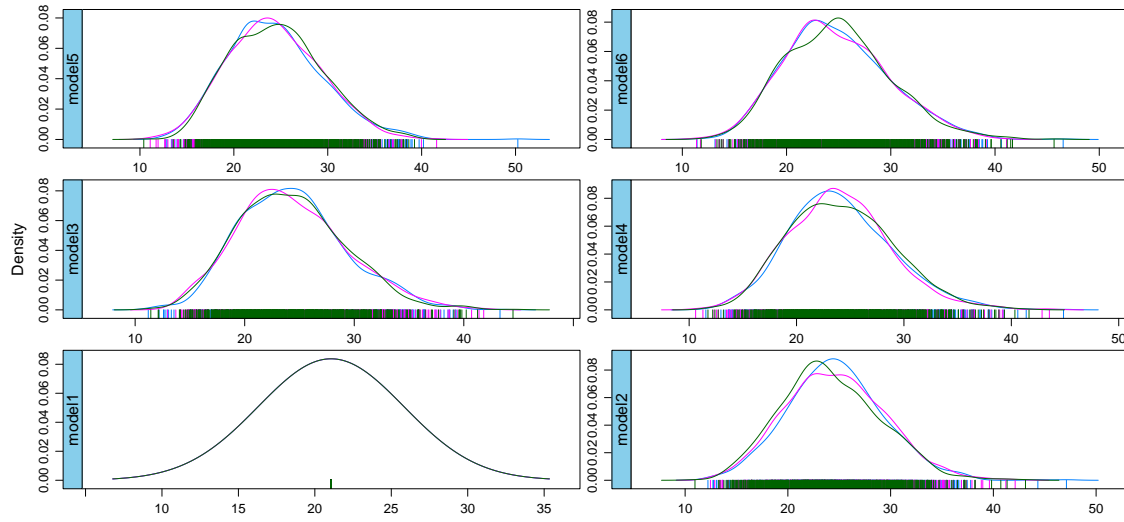


Figure 2.6

	mean	sd	2.5%	50%	97.5%	Rhat	n.eff
model1	10.5463	0.0000	10.5463	10.5463	10.5463		0.0000
model2	12.3481	3.5241	6.4608	12.0826	19.7636	1.0000	2833.0000
model3	12.2433	3.5058	6.3331	11.9460	20.1100	1.0000	2955.0000
model4	12.0331	3.2668	6.4515	11.7501	19.1090	1.0000	3322.0000
model5	12.3475	3.5362	6.4799	12.0049	19.9832	1.0000	2503.0000
model6	12.9588	3.6536	6.8353	12.6056	21.2272	1.0000	3000.0000

Table 2.3: Posterior distributions of different models for A , with added anomalies, and calculated with independent Bayes model

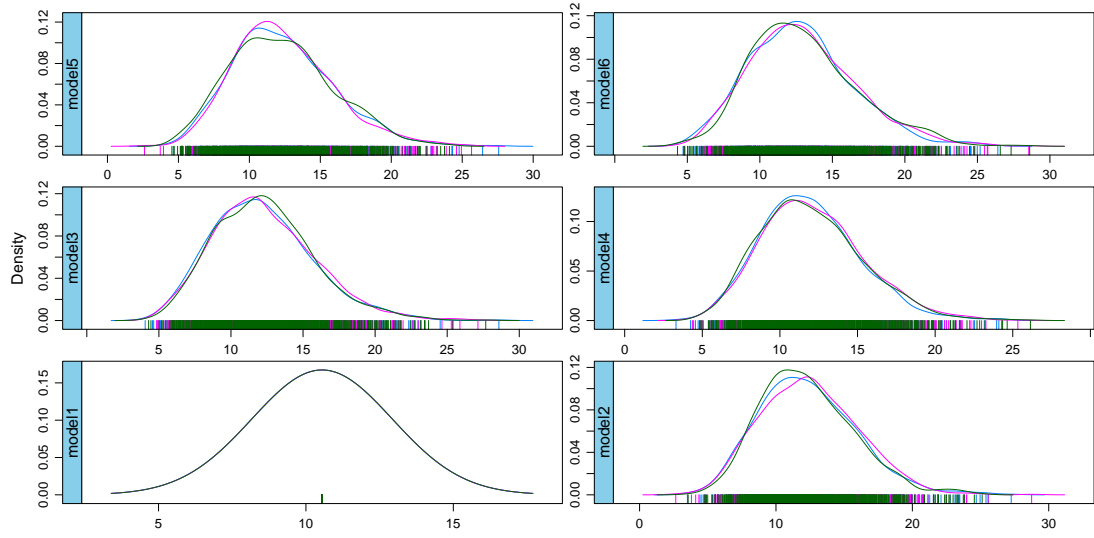


Figure 2.7

	mean	sd	2.5%	50%	97.5%	Rhat	n.eff
model1	5.2477	0.0000	5.2477	5.2477	5.2477		0.0000
model2	7.4081	2.7444	2.8634	7.1136	13.3165	1.0000	2273.0000
model3	7.3234	2.7423	2.8910	7.0234	13.3408	1.0000	2372.0000
model4	6.9192	2.4111	3.0404	6.6744	12.3585	1.0000	3000.0000
model5	7.6640	2.8689	3.0550	7.2466	14.0954	1.0000	1886.0000
model6	7.8524	2.7397	3.4434	7.5279	14.0425	1.0000	2856.0000

Table 2.4: Posterior distributions of different models for AA, with added anomalies, and calculated with independent Bayes model

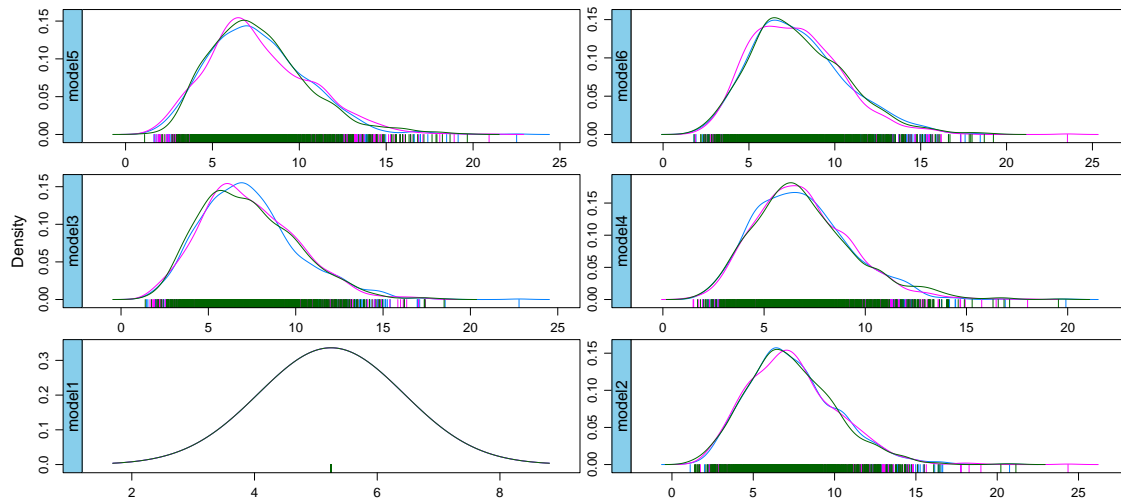


Figure 2.8

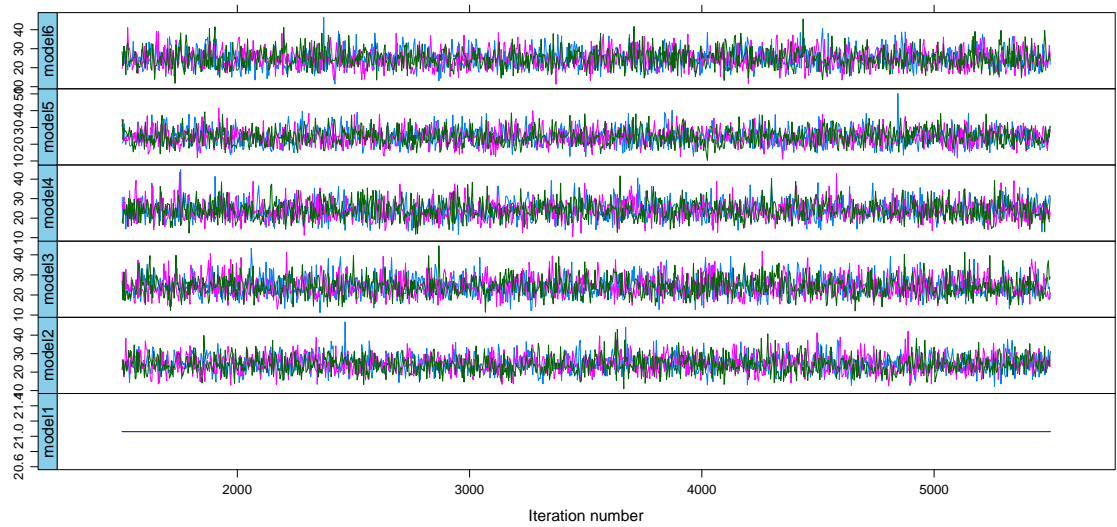


Figure 2.9: Trace plot for category total of model 1 to model 6

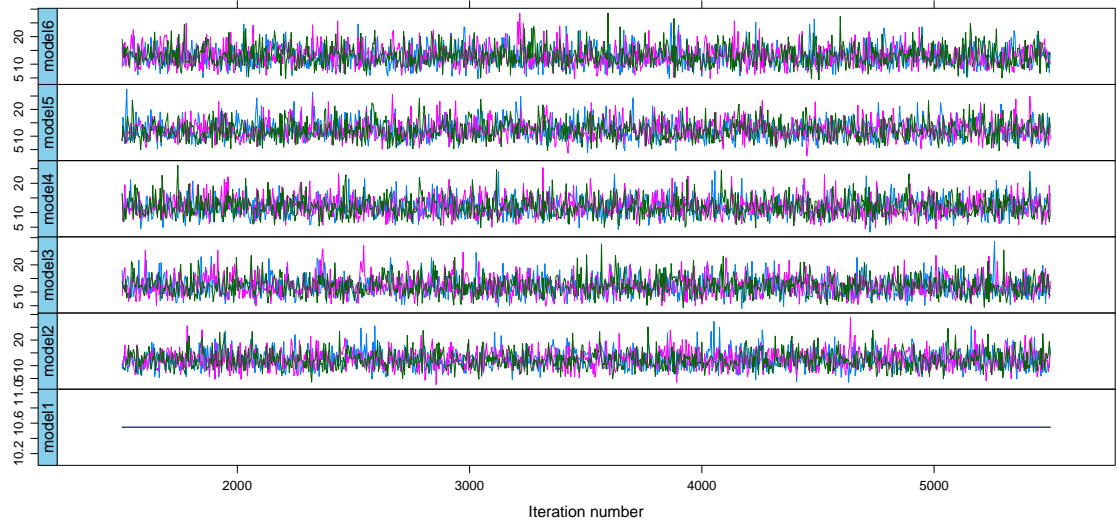


Figure 2.10: Trace plot for category A of model 1 to model 6

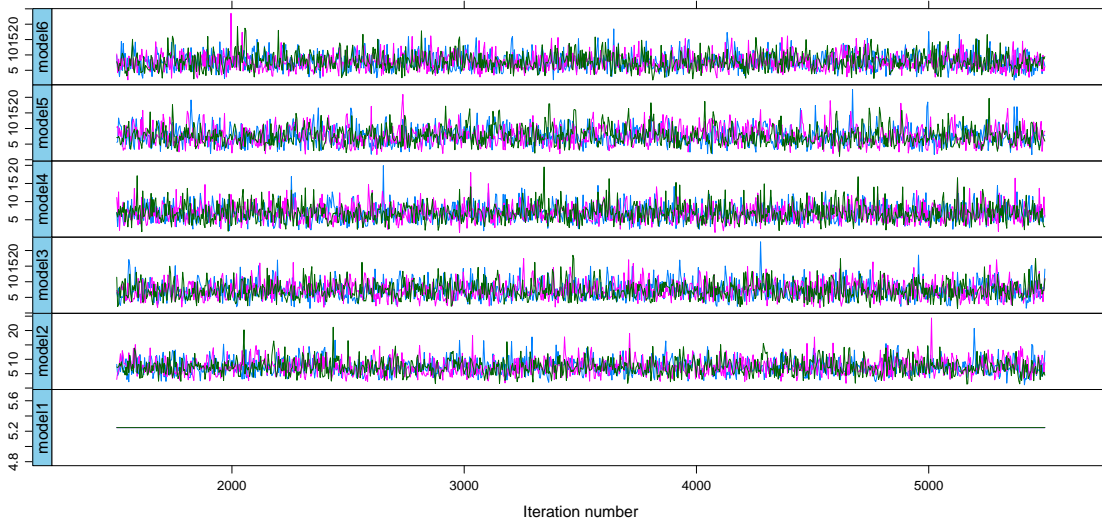


Figure 2.11: Trace plot for category AA of model 1 to model 6

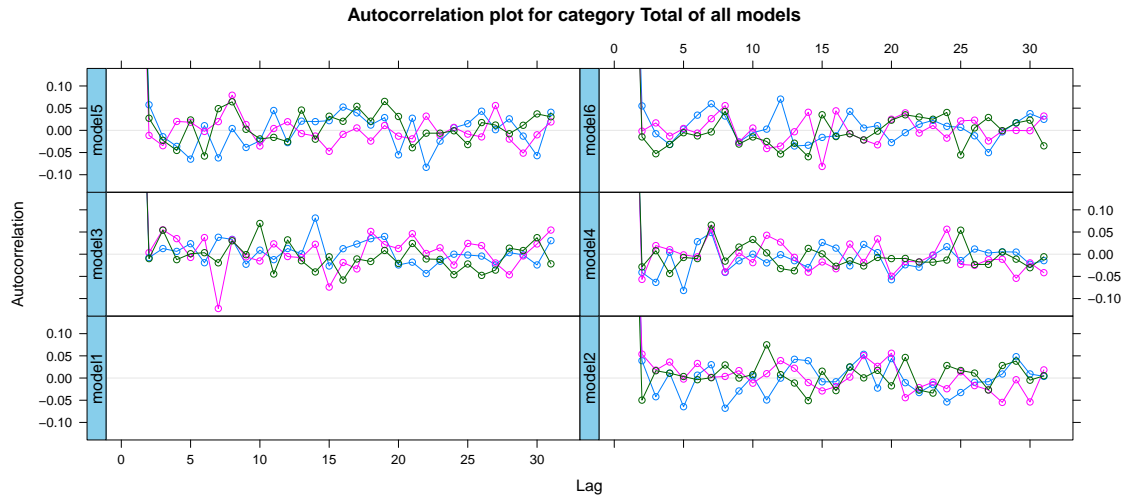


Figure 2.12: Trace plot for category AA of model 1 to model 6

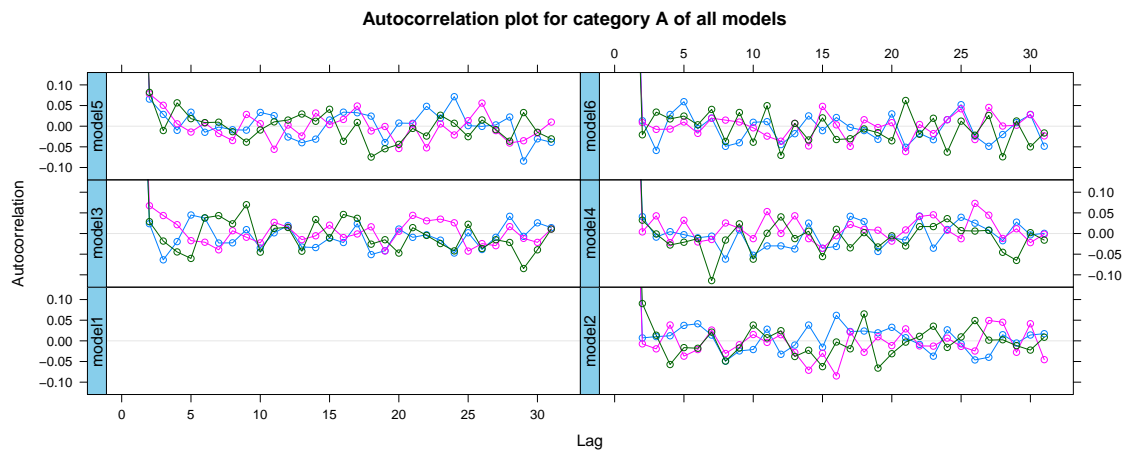


Figure 2.13: Trace plot for category AA of model 1 to model 6

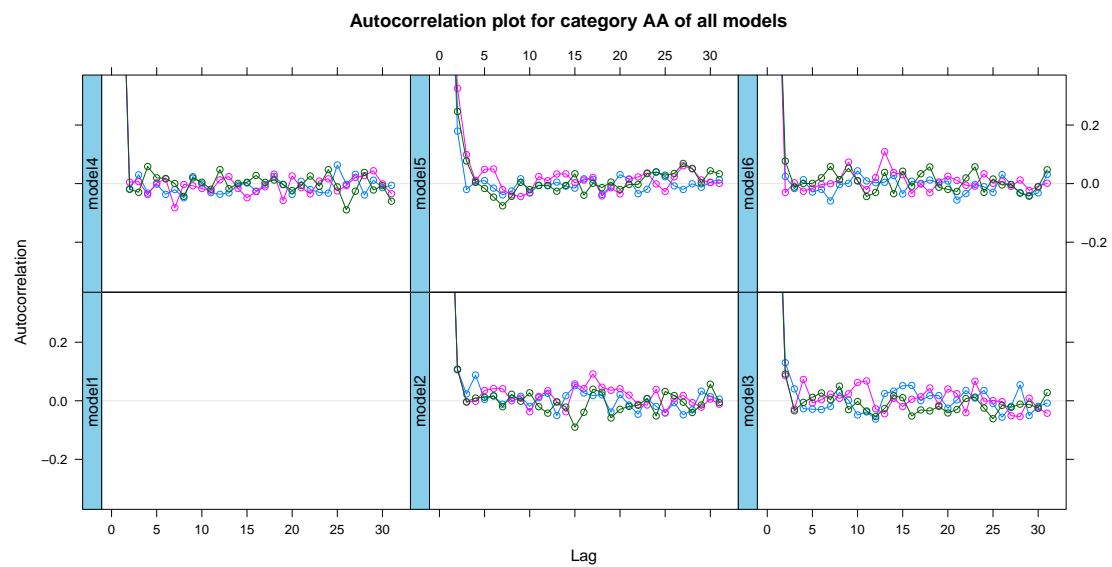


Figure 2.14: Trace plot for category AA of model 1 to model 6

2.3.4 Discussion

...suggest non-informative prior is

And because all of other models tried are not backed by purely theoretical there are no reason to believe one model is better than another.

However I knew that the random number generation used during synthesizing process generated normal distribution for all categories. Therefore using a normal prior seem to be the simplest option. And it is decided that model 2 is our best prior to use for all other simulation testings.

chandola2009anomaly

There are some differences in Anomaly detection techniques, such as supervision and parametrisation.

(add these term? important concept but might take too long)

Therefore, the goal of supervised learning is to learn a function that, given a sample of data and desired outputs, best approximates the relationship between input and output observable in the data. Unsupervised learning, on the other hand, does not have labeled outputs, so its goal is to infer the natural structure present within a set of data points.

<https://towardsdatascience.com/supervised-vs-unsupervised-learning-14f68e32ea8d>

<http://home.cc.umanitoba.ca/~gersteia/MBIO7040/06-introStats.html>

Supervised vs unsupervised

later on we will go into the concept of bayesian, however here what we need to know is that bayesian is a form of supervised/unsupervised.....

Parametric vs non-parametric

In the literal meaning of the terms, a parametric statistical test is one that makes assumptions about the parameters (defining properties) of the population distribution(s) from which one's data are drawn, while a non-parametric test is one that makes no such assumptions. In this strict sense, "non-parametric" is essentially a null category, since virtually all statistical tests assume one thing or another about the properties of the source population(s).

Strength and weakness

no acf, true for simulation as each day is generated randomly and assume complete independence

2.4 Simulation 2: Size of anomaly

2.4.1 Simulation setups

Examining the Counts on the day that we add anomaly using R codes, we can see that anomaly Different anomalies is added to category AA of the synthetic data....etc etc...

```
> daily1.df[anomalydateloc,]
day total A  B AA AB BA BB
3606 2015-11-15      19 7 12  2  5  5  7
> daily1.S10.df[anomalydateloc,]
day total A  B AA AB BA BB
3606 2015-11-15      21 9 12  4  5  5  7
> daily1.S25.df[anomalydateloc,]
day total A  B AA AB BA BB
3606 2015-11-15      24 12 12  7  5  5  7
> daily1.S50.df[anomalydateloc,]
day total A  B AA AB BA BB
3606 2015-11-15      30 18 12 13  5  5  7
> daily1.S100.df[anomalydateloc,]
day total A  B AA AB BA BB
3606 2015-11-15      40 28 12 23  5  5  7
> daily1.S250.df[anomalydateloc,]
day total A  B AA AB BA BB
3606 2015-11-15      72 60 12 55  5  5  7
> daily1.S500.df[anomalydateloc,]
day total A  B AA AB BA BB
3606 2015-11-15     124 112 12 107  5  5  7
```

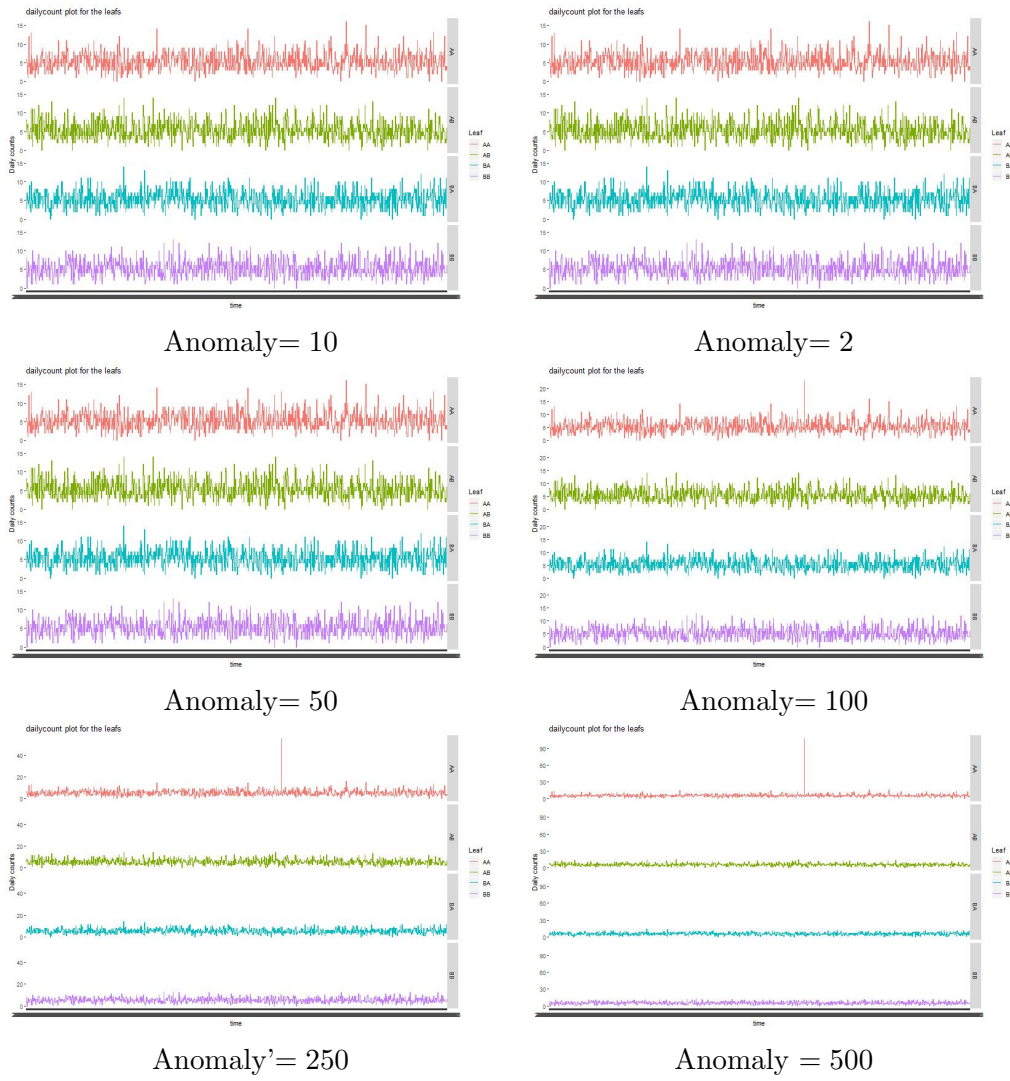


Figure 2.15: Visual inspection of different increment of anomalies at level 2 of hierarchy

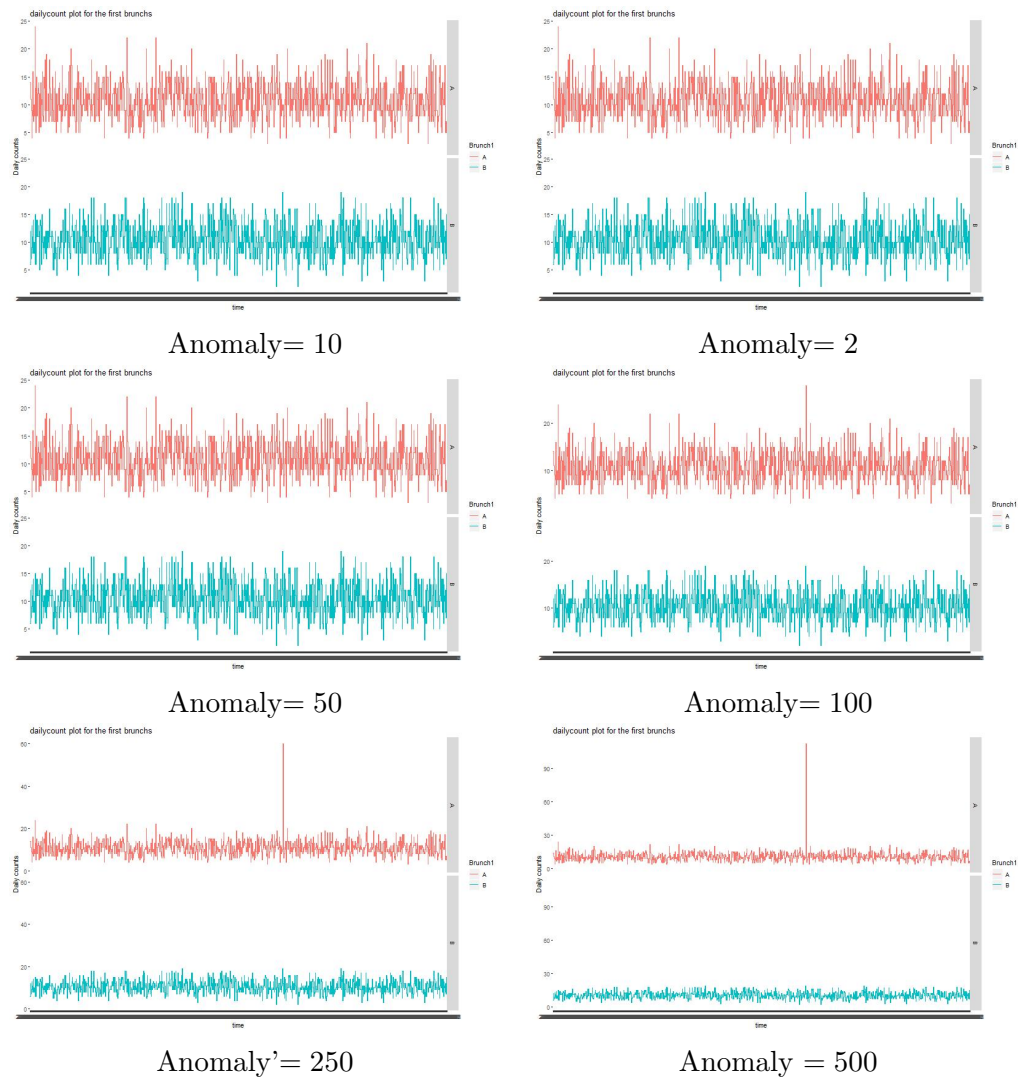


Figure 2.16: Visual inspection of different increment of anomalies at level 1 of hierarchy

2.4.2 Results

Anomaly (% of total)	Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
0	-0.466	0.188	0.437	0.545	0.955	1.556
10	-1.910	-0.693	0.876	1.144	1.894	6.640
25	-6.274	-3.079	-1.966	-1.984	-0.658	1.825
50	-4.352	-1.185	-0.038	-0.385	0.823	2.420
100	-5.074	-1.096	0.323	-0.383	1.194	1.873
250	-1.813	-1.657	-0.917	-0.181	0.546	3.681
500	-4.416	-1.678	-0.613	-0.875	0.430	1.401

Table 2.5: DIC comparisons between Independent Bayes model and Hierarchical Bayes model (HBM -IBM) for different anomalies sizes

see appendix

	mean	sd	2.5%	50%	97.5%	Rhat	n.eff
Anomaly0	19.1481	4.2840	11.7152	18.8136	28.3845	1.0000	3000.0000
Anomaly10	21.4084	4.7272	13.1528	21.0041	31.6607	1.0000	3000.0000
Anomaly25	24.4356	4.9374	16.0687	24.0399	34.7248	1.0000	2975.0000
Anomaly50	30.2932	5.3613	20.8844	29.8973	41.9325	1.0000	2912.0000
Anomaly100	40.4338	6.1531	29.3330	40.0987	53.8314	1.0000	2889.0000
Anomaly250	71.9683	8.4372	56.5531	71.6763	89.4346	1.0000	3000.0000
Anomaly500	122.4365	10.7971	101.9952	122.2718	144.5457	1.0000	3132.0000

Table 2.6: Posterior distributions of different models for Total, with different increments of added anomalies, and calculated with independent Bayes model

	mean	sd	2.5%	50%	97.5%	Rhat	n.eff
Anomaly0	19.3114	4.4361	11.5960	18.9555	28.9686	1.0000	2934.0000
Anomaly10	21.3883	4.5971	13.3033	21.1152	31.2900	1.0000	2799.0000
Anomaly25	24.3800	4.9083	15.8281	24.1057	34.9855	1.0000	3000.0000
Anomaly50	30.5664	5.4857	20.7071	30.2800	42.3242	1.0000	2713.0000
Anomaly100	40.5794	6.2480	29.1623	40.2459	53.0399	1.0000	2704.0000
Anomaly250	72.0177	8.2580	57.1126	71.8129	88.7105	1.0000	2723.0000
Anomaly500	121.9641	10.3188	102.7605	121.6646	142.4118	1.0000	2173.0000

Table 2.7: Posterior distributions of different models for Total, with different increments of added anomalies, and calculated with Hierarchical Bayes model

	mean	sd	2.5%	50%	97.5%	Rhat	n.eff
Anomaly0	7.1758	2.7278	3.0586	6.7391	13.6749	1.0000	2844.0000
Anomaly10	9.3215	3.1367	4.1286	8.9552	16.3998	1.0000	2630.0000
Anomaly25	12.3410	3.5655	6.4488	11.9857	20.3604	1.0000	3038.0000
Anomaly50	18.4421	4.3162	10.8431	18.1170	27.7122	1.0000	2839.0000
Anomaly100	28.3032	5.3432	19.1588	27.8652	40.0187	1.0000	3008.0000
Anomaly250	58.6906	7.3628	45.1079	58.6335	73.7809	1.0000	2118.0000
Anomaly500	104.7585	9.5672	87.1657	104.5580	123.4846	1.0000	2235.0000

Table 2.8: Posterior distributions of different models for A , with different increments of added anomalies, and calculated with Independent Bayes model

	mean	sd	2.5%	50%	97.5%	Rhat	n.eff
Anomaly0	7.1740	2.7263	2.8494	6.8467	13.2835	1.0000	2559.0000
Anomaly10	9.1731	3.0292	4.1612	8.8307	15.8837	1.0000	2961.0000
Anomaly25	12.4274	3.6049	6.3125	12.0656	20.2620	1.0000	2703.0000
Anomaly50	18.5048	4.3162	10.9271	18.2437	28.0090	1.0000	2580.0000
Anomaly100	28.6123	5.2341	19.2090	28.1992	39.9182	1.0000	2370.0000
Anomaly250	59.6236	7.2375	46.4370	59.4535	74.8256	1.0000	1940.0000
Anomaly500	107.9043	9.7689	89.7498	107.4554	129.0404	1.0000	1338.0000

Table 2.9: Posterior distributions of different models for A , with different increments of added anomalies, and calculated with Hierarchical Bayes model

	mean	sd	2.5%	50%	97.5%	Rhat	n.eff
Anomaly0	2.3062	1.4859	0.4009	2.0034	6.0143	1.0000	2401.0000
Anomaly10	4.2976	2.1384	1.2137	3.8992	9.5420	1.0000	2461.0000
Anomaly25	7.4243	2.8041	2.8709	7.0744	13.8003	1.0000	2214.0000
Anomaly50	13.3455	3.5920	7.3160	13.0175	21.0238	1.0000	2428.0000
Anomaly100	22.4374	4.5194	14.6914	22.0898	31.9623	1.0000	2029.0000
Anomaly250	48.8502	6.1544	37.2238	48.5964	61.2842	1.0000	1568.0000
Anomaly500	100.5322	11.0472	79.0267	100.3259	122.1376	1.0200	1115.0000

Table 2.10: Posterior distributions of different models for AA, with different increments of added anomalies, and calculated with Independent Bayes model

	mean	sd	2.5%	50%	97.5%	Rhat	n.eff
Anomaly0	2.1538	1.4960	0.3505	1.7838	6.1010	1.0000	2073.0000
Anomaly10	4.3301	2.1448	1.1920	3.9651	9.5508	1.0000	2046.0000
Anomaly25	7.6847	2.8131	3.0945	7.3379	13.8861	1.0000	2149.0000
Anomaly50	13.6061	3.6369	7.2589	13.3337	21.4732	1.0000	2054.0000
Anomaly100	23.0253	4.5067	14.9764	22.7288	32.6668	1.0000	1679.0000
Anomaly250	51.9695	6.3876	40.2516	51.8076	65.5484	1.0100	951.0000
Anomaly500	96.0807	9.1704	79.5825	95.8225	115.7111	1.0000	951.0000

Table 2.11: Posterior distributions of different models for AA, with different increments of added anomalies, and calculated with Hierarchical Bayes model

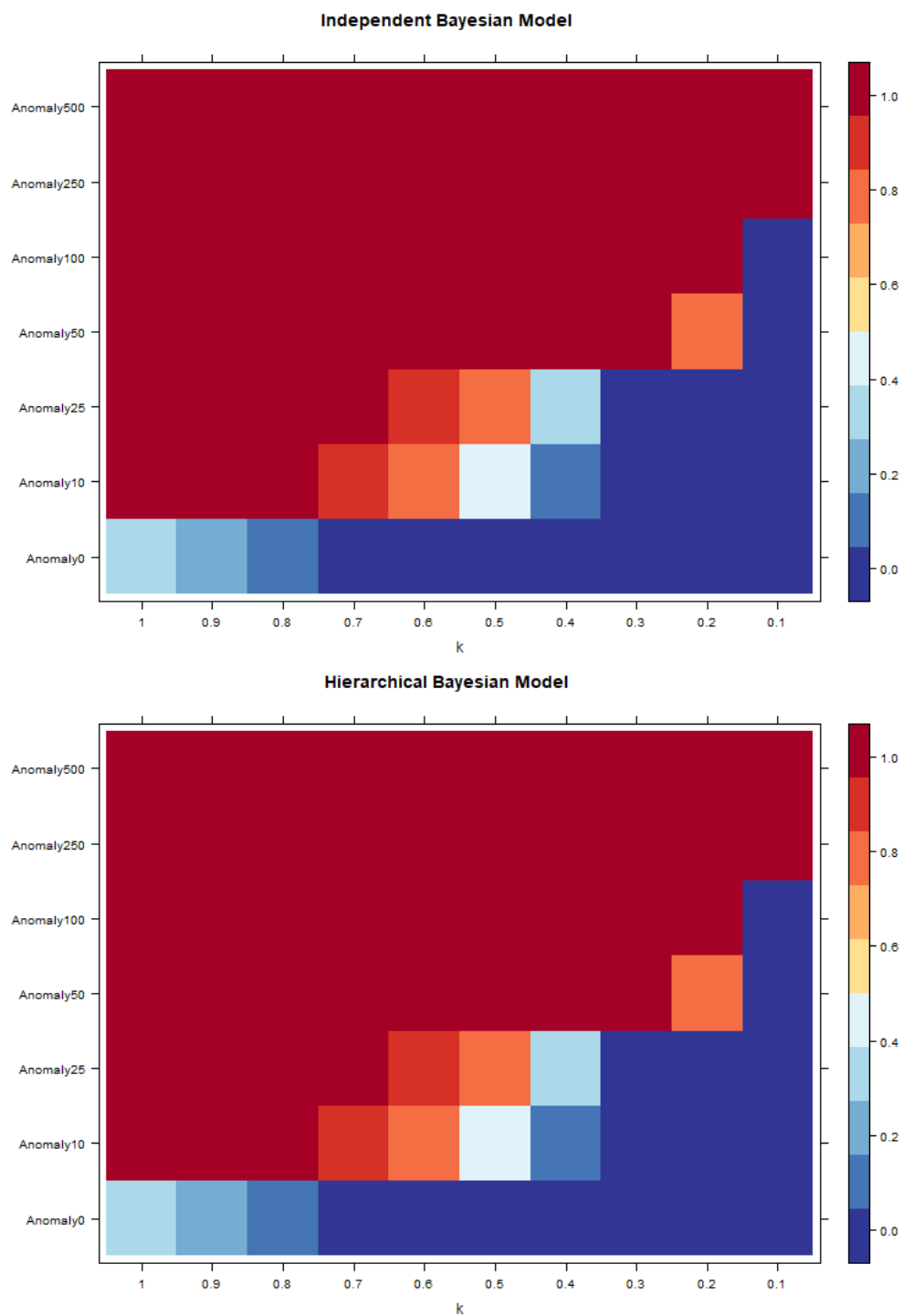


Figure 2.17: total

2.4.3 Discussion

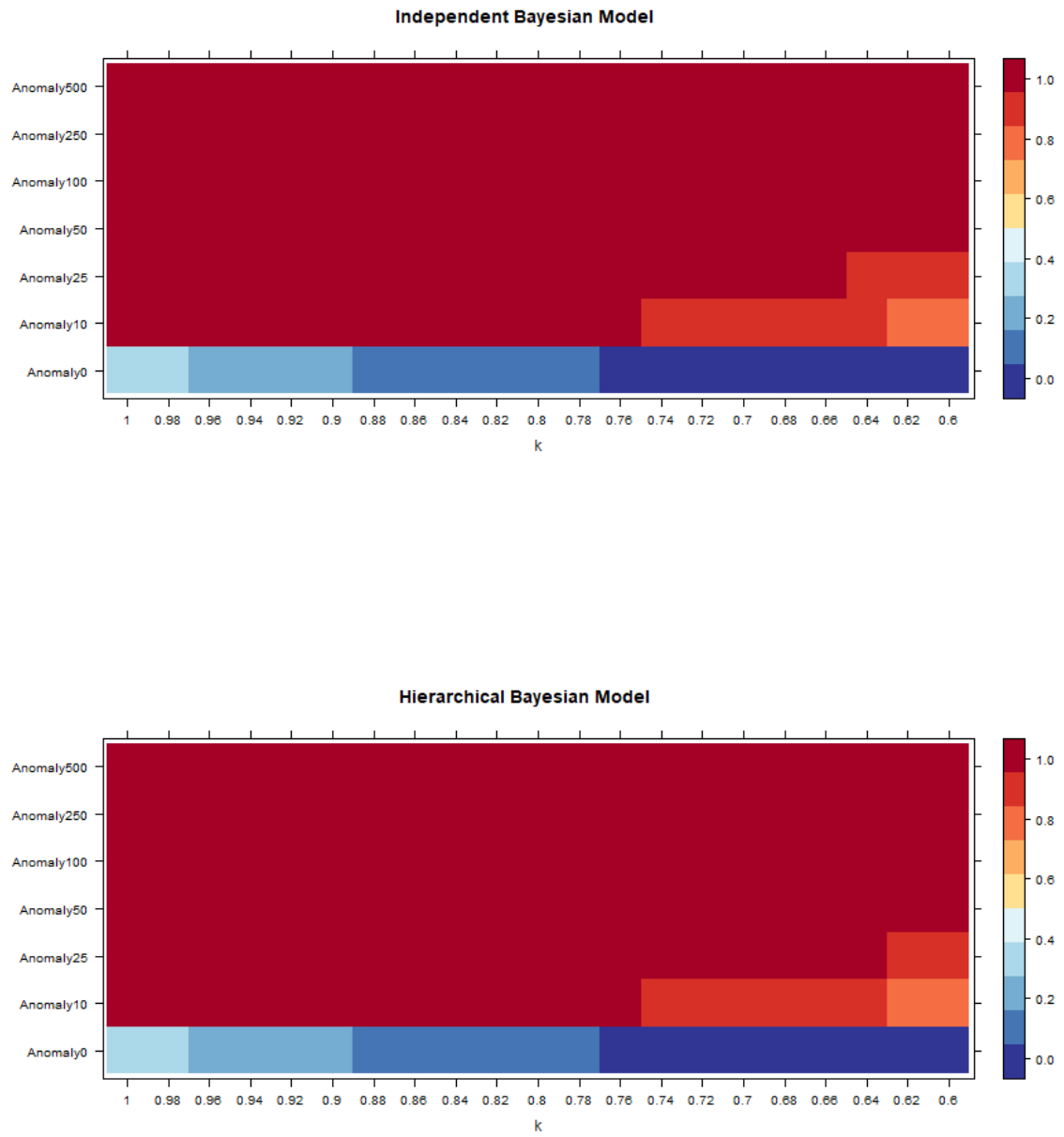


Figure 2.18: total

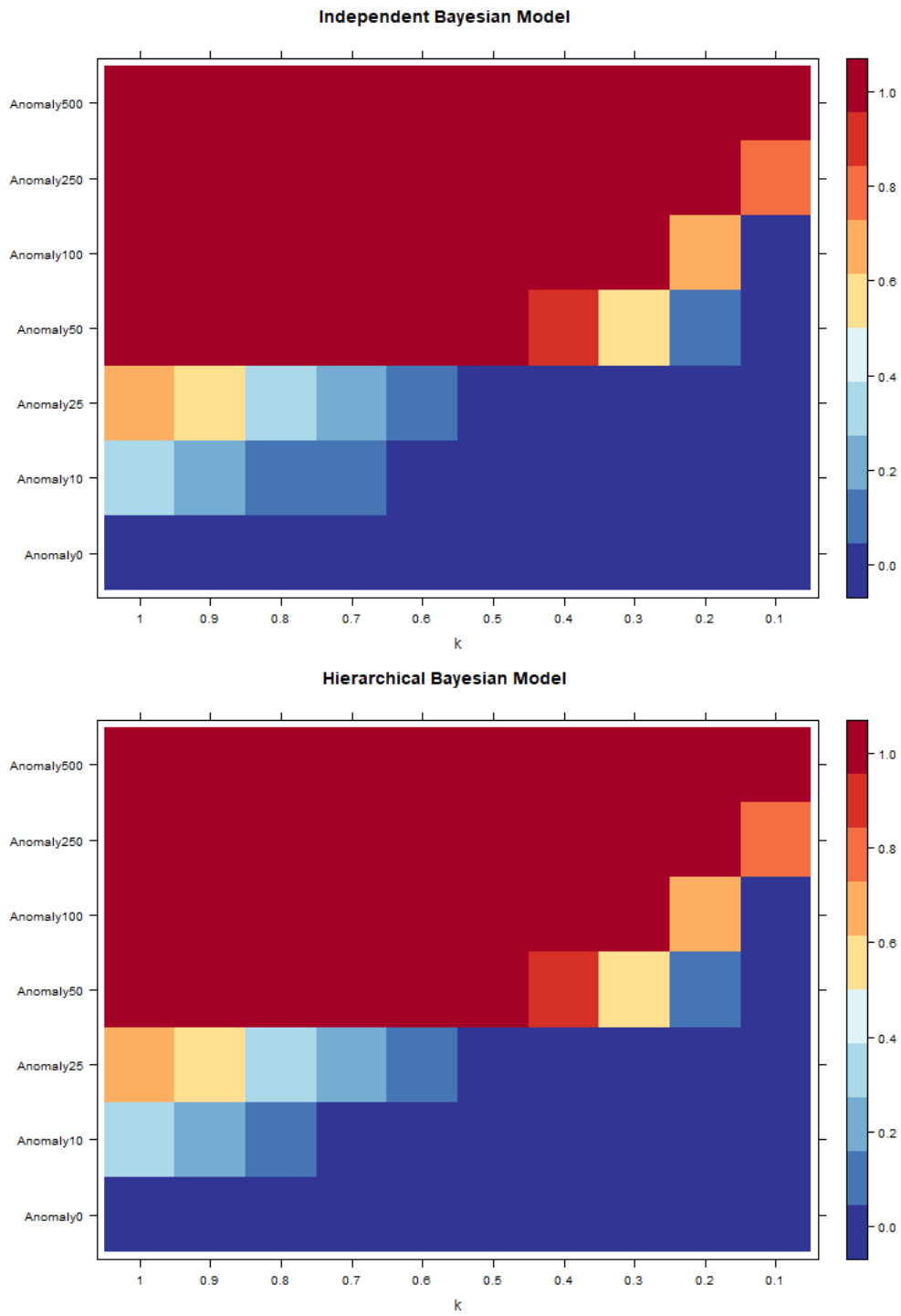


Figure 2.19: A

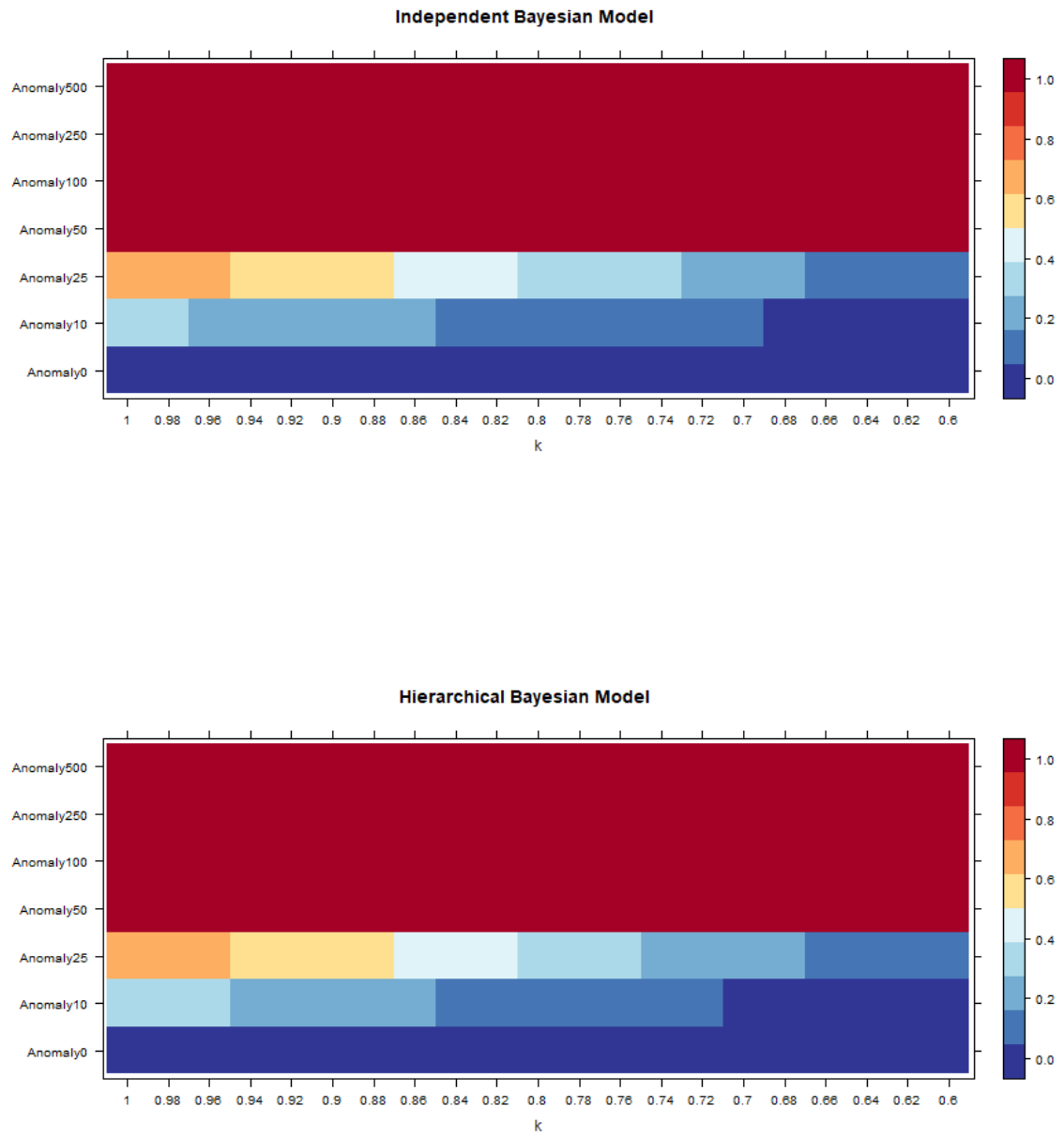


Figure 2.20: A

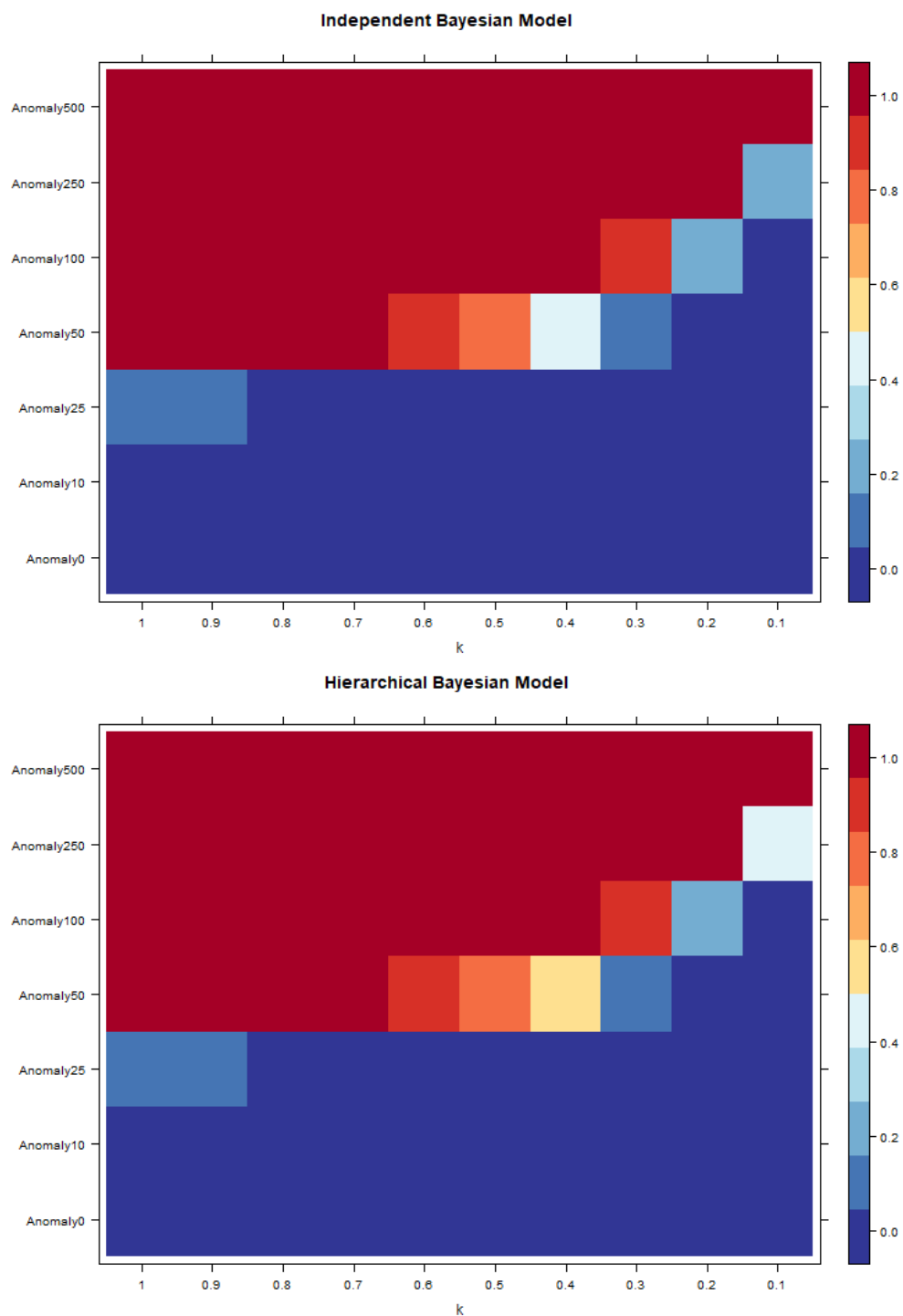


Figure 2.21: AA

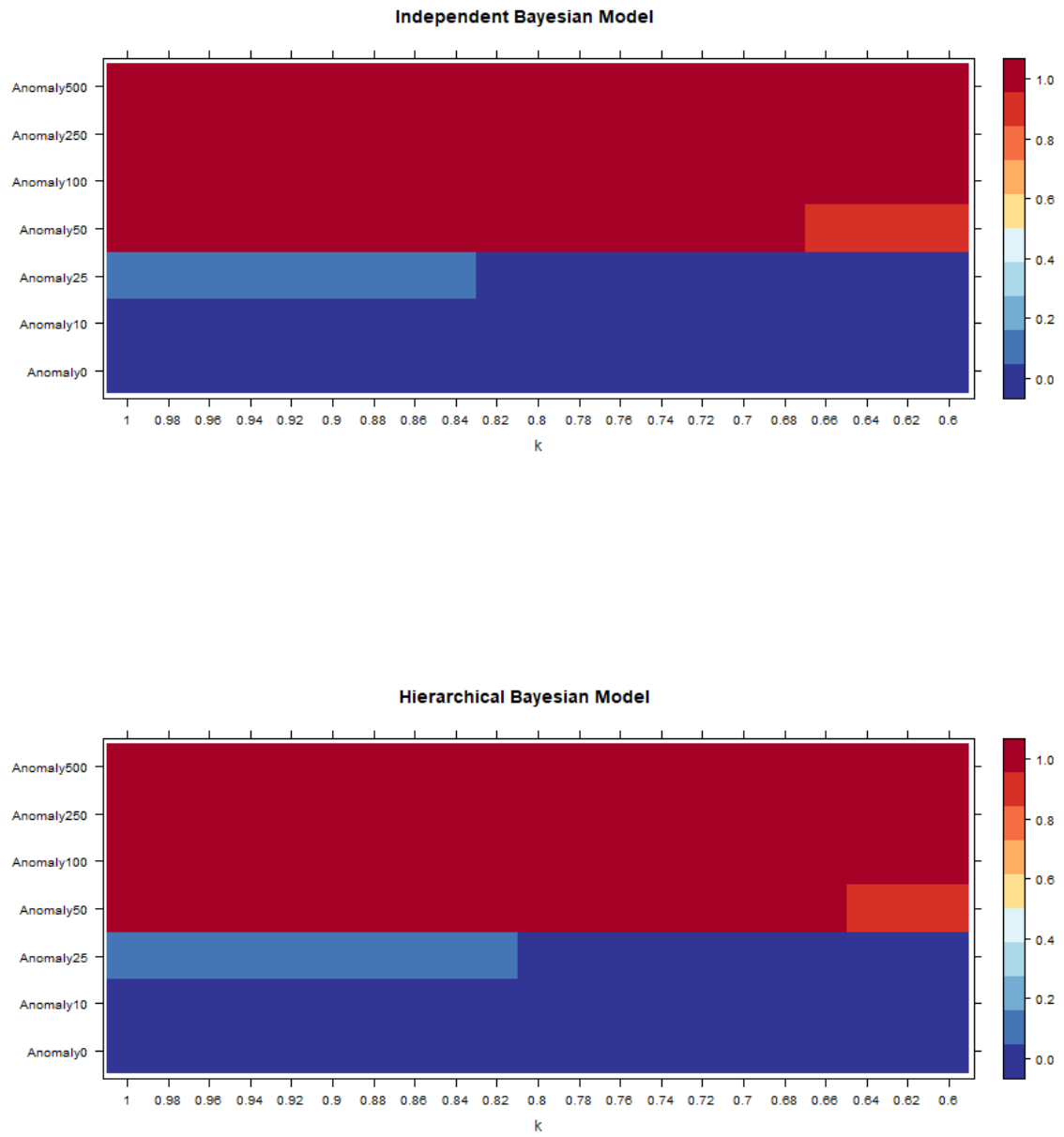


Figure 2.22: AA

2.5 Simulation 3:

2.5.1 Simulation setups

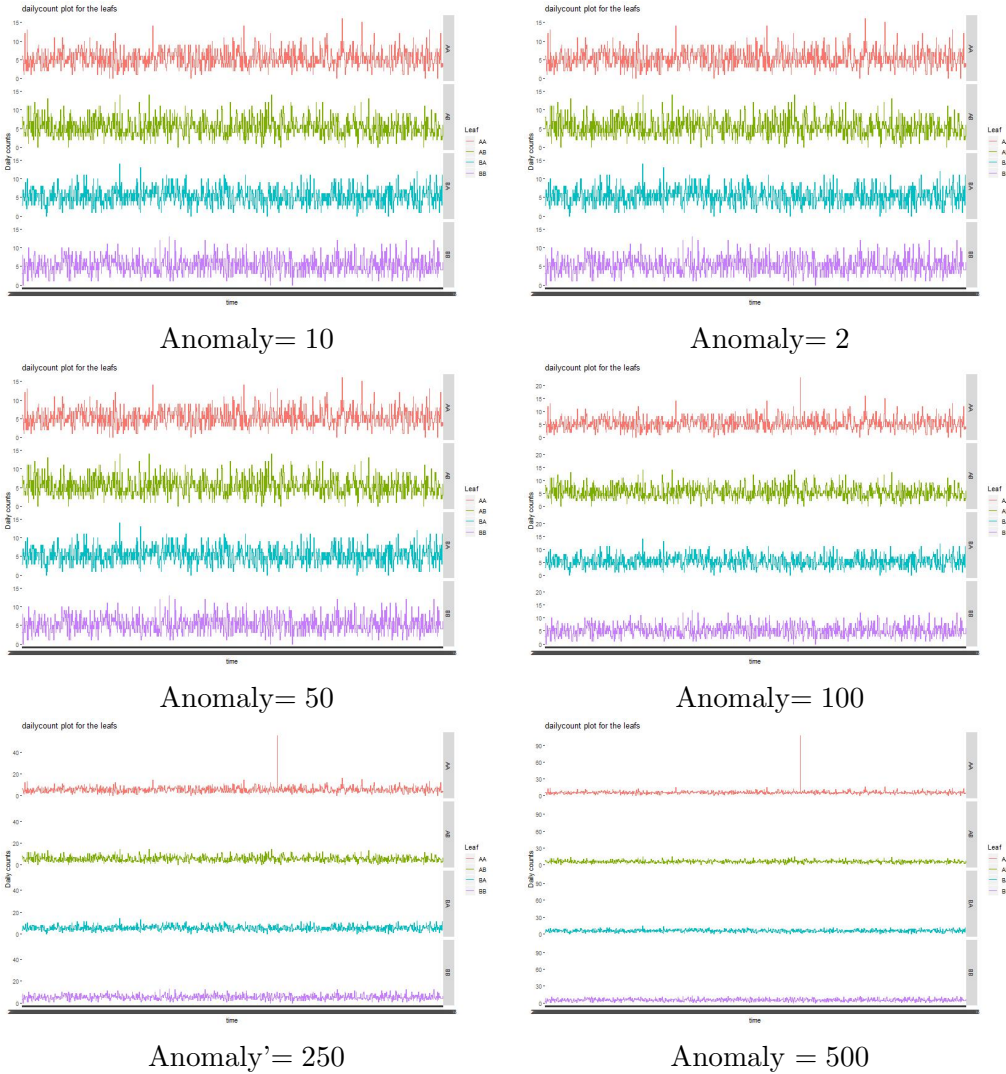


Figure 2.23: Visual inspection of different increment of anomalies at level 2 of hierarchy

2.5.2 Results

still unstable, but multiple testing shows that mean and median mostly negative numbers, which meant hierarchical is better

```

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\begin{table}
\caption{postsumA ,digits = 4,type = "latex",

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	Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
1	-7.033	-1.960	-0.809	-1.527	0.294	1.172
2	-4.435	-2.722	-0.994	-0.591	0.214	6.308
3	-2.710	-1.013	0.554	0.521	1.440	4.575
4	-3.418	-0.451	0.370	0.305	1.422	3.356
5	-2.715	-1.447	-1.065	-0.318	1.022	2.791
6	-3.082	-0.733	-0.129	0.121	1.042	4.216
7	-4.280	-2.843	-1.286	-1.581	-0.213	0.790
8	-4.349	-2.215	-0.967	-0.422	1.065	7.971
9	-4.931	-1.649	-0.335	-0.303	1.014	4.038
10	-5.247	-1.578	-0.234	-0.668	0.618	3.247
11	-2.597	-1.126	1.002	0.466	1.655	3.751
12	-3.340	-1.786	-0.102	0.223	1.838	4.623

Table 2.12: DIC comapreisons for brunching number at A

```

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= cap5,label = "tab:pstprototal") ; xtable(postsumAA,digits = 4,type = "latex",
file = "plots/sim2/MCMCsumcAA.tex",caption = cap3,label = "tab:pstprototal") ;
xtable(postsumAAh,digits = 4,type = "latex", file = "plots/sim2/MCMCsumcAAh.tex",caption
= cap6,label = "tab:pstprototal")

```

	mean	sd	2.5%	50%	97.5%	Rhat	n.eff
brunch1	21.2367	4.7008	12.9239	20.9013	31.1209	1.0000	2787.0000
brunch2	21.4198	4.6439	13.5516	20.9515	31.6895	1.0000	2904.0000
brunch3	30.4154	5.6673	20.1950	30.2022	42.1275	1.0000	2877.0000
brunch4	13.1540	3.5731	6.9948	12.7660	21.1359	1.0000	3000.0000
brunch5	26.4296	5.0584	17.2806	26.1034	37.3137	1.0000	2815.0000
brunch6	22.3157	4.7457	14.0816	22.0346	32.1219	1.0000	2881.0000
brunch7	17.2269	4.1264	10.2189	16.7961	26.4545	1.0000	2587.0000
brunch8	19.1497	4.3697	11.4783	18.8391	28.4958	1.0000	3000.0000
brunch9	27.4028	5.2818	18.0539	26.9926	38.6413	1.0000	2854.0000
brunch10	18.2805	4.2982	10.6334	18.0157	27.3239	1.0000	3028.0000
brunch11	22.2082	4.7100	13.9406	21.8632	32.7639	1.0000	3000.0000
brunch12	23.2141	4.8762	14.6652	22.8895	33.4077	1.0000	3090.0000

Table 2.13: Posterior distributions of different models for Total, with different brunch number at A, and calculated with independent Bayes model

2.5.3 Discussion

	mean	sd	2.5%	50%	97.5%	Rhat	n.eff
mu[1]	21.1342	4.6356	13.0372	20.8723	31.0156	1.0000	2899.0000
mu[2]	21.4399	4.6658	13.3765	21.0743	31.3564	1.0000	2888.0000
mu[3]	30.4796	5.5918	20.4537	30.1087	42.6573	1.0000	3069.0000
mu[4]	13.0428	3.6218	6.9370	12.7050	21.1728	1.0000	2798.0000
mu[5]	26.4792	5.0783	17.5450	26.1769	37.5123	1.0000	2903.0000
mu[6]	22.5513	4.7242	14.3816	22.2968	32.5958	1.0000	3000.0000
mu[7]	17.2175	4.1172	9.9473	16.9176	25.9870	1.0000	3000.0000
mu[8]	19.2860	4.3666	11.7869	19.0001	28.9699	1.0000	2712.0000
mu[9]	27.4363	5.2930	18.1801	27.1123	38.8600	1.0000	2719.0000
mu[10]	18.3153	4.2436	10.9404	18.0915	27.1782	1.0000	2961.0000
mu[11]	22.2699	4.7638	13.7242	21.8257	32.6838	1.0000	3000.0000
mu[12]	23.3112	4.9101	14.5832	22.9914	33.6660	1.0000	3006.0000

Table 2.14: Posterior distributions of different models for Total, with different brunch number at A, and calculated with Hierarchical Bayes model

	mean	sd	2.5%	50%	97.5%	Rhat	n.eff
brunch1	10.2222	3.2257	4.9491	9.9313	17.3630	1.0000	3000.0000
brunch2	7.2101	2.5820	3.0653	6.9089	13.0457	1.0000	3135.0000
brunch3	13.3965	3.5903	7.1940	13.0422	21.4547	1.0000	2843.0000
brunch4	5.2344	2.2966	1.7544	4.9212	10.8095	1.0100	2659.0000
brunch5	12.2653	3.5442	6.3801	11.9468	20.0809	1.0000	2831.0000
brunch6	12.4534	3.5617	6.4566	12.0511	20.3078	1.0000	2826.0000
brunch7	11.3345	3.3140	5.6904	11.0780	18.6164	1.0000	2763.0000
brunch8	11.3301	3.4455	5.6405	10.9504	18.8542	1.0000	3000.0000
brunch9	15.2998	3.9278	8.5925	15.0144	24.1089	1.0000	3000.0000
brunch10	10.2282	3.2404	4.9712	9.8590	17.4283	1.0000	2719.0000
brunch11	12.2384	3.5230	6.2765	11.9505	19.9642	1.0000	3000.0000
brunch12	16.2605	4.1472	9.1019	15.9211	25.5411	1.0000	2614.0000

Table 2.15: Posterior distributions of different models for A , with different brunch number at A, and calculated with independent Bayes model

	mean	sd	2.5%	50%	97.5%	Rhat	n.eff
mu[1]	10.3658	3.2745	4.9563	9.9946	17.5591	1.0000	2480.0000
mu[2]	7.1699	2.7092	2.8318	6.7969	13.3853	1.0000	2763.0000
mu[3]	13.5106	3.6263	7.2349	13.2756	21.5433	1.0000	2721.0000
mu[4]	4.9821	2.1601	1.7617	4.6571	9.9815	1.0000	2848.0000
mu[5]	12.2966	3.5260	6.2470	11.9827	20.1420	1.0000	2691.0000
mu[6]	11.9993	3.3669	6.1917	11.7362	19.3448	1.0000	2515.0000
mu[7]	11.1468	3.3194	5.5362	10.8140	18.2800	1.0000	2791.0000
mu[8]	10.9205	3.2059	5.5478	10.6286	17.9246	1.0000	2596.0000
mu[9]	15.1743	3.7615	8.6417	14.8581	23.3316	1.0000	2371.0000
mu[10]	9.4961	2.9241	4.6206	9.1667	16.0461	1.0000	2730.0000
mu[11]	11.7279	3.3009	5.9251	11.4201	18.8472	1.0000	2389.0000
mu[12]	15.4642	3.7186	9.0081	15.1212	23.5432	1.0000	2361.0000

Table 2.16: Posterior distributions of different models for A , with different brunch number at A, and calculated with Hierarchical Bayes model

	mean	sd	2.5%	50%	97.5%	Rhat	n.eff
brunch1	10.3390	3.3313	4.9841	9.9711	18.0423	1.0000	3015.0000
brunch2	2.2864	1.4796	0.4436	1.9694	5.9125	1.0000	2383.0000
brunch3	4.3819	2.1286	1.2499	4.0705	9.5828	1.0000	2021.0000
brunch4	2.2971	1.5500	0.3511	1.9390	6.0469	1.0000	1776.0000
brunch5	4.3729	2.0611	1.2799	4.0873	9.4255	1.0000	1568.0000
brunch6	5.1595	2.0347	1.8054	4.9425	9.5605	1.0000	1448.0000
brunch7	3.2759	1.6852	0.7437	3.0322	7.3844	1.0000	1259.0000
brunch8	3.2901	1.6666	0.7674	3.0726	7.0195	1.0000	1255.0000
brunch9	3.1983	1.5386	0.8286	3.0263	6.7680	1.0000	1424.0000
brunch10	4.3662	1.5571	1.6687	4.1853	7.6161	1.0100	1110.0000
brunch11	2.1852	1.3714	0.2855	1.9361	5.4025	1.0000	1077.0000
brunch12	2.9119	1.4034	0.7174	2.6996	6.1571	1.0000	947.0000

Table 2.17: Posterior distributions of different models for AA, with different brunch number at A, and calculated with independent Bayes model

	mean	sd	2.5%	50%	97.5%	Rhat	n.eff
mu[1]	10.3563	3.2788	4.9509	10.0156	17.7970	1.0000	2540.0000
mu[2]	2.1951	1.5255	0.3203	1.8362	5.9132	1.0000	1978.0000
mu[3]	4.5592	2.2061	1.2605	4.2379	9.5741	1.0000	1910.0000
mu[4]	2.2663	1.5396	0.3203	1.9350	5.9029	1.0000	1540.0000
mu[5]	4.4613	1.9943	1.3097	4.2350	8.8343	1.0000	1354.0000
mu[6]	5.2129	2.0225	1.8571	5.0210	9.6588	1.0000	1023.0000
mu[7]	3.3193	1.6606	0.7695	3.0899	7.1315	1.0000	1053.0000
mu[8]	3.3147	1.6493	0.7960	3.0817	7.1798	1.0200	1023.0000
mu[9]	3.2724	1.6032	0.7633	3.0402	6.8838	1.0000	711.0000
mu[10]	4.4763	1.6416	1.7625	4.3287	8.0546	1.0100	737.0000
mu[11]	2.2327	1.2427	0.3545	2.0587	4.9807	1.0000	952.0000
mu[12]	3.0880	1.4240	0.8039	2.9188	6.2747	1.0000	816.0000

Table 2.18: Posterior distributions of different models for AA, with different brunch number at A, and calculated with Hierarchical Bayes model

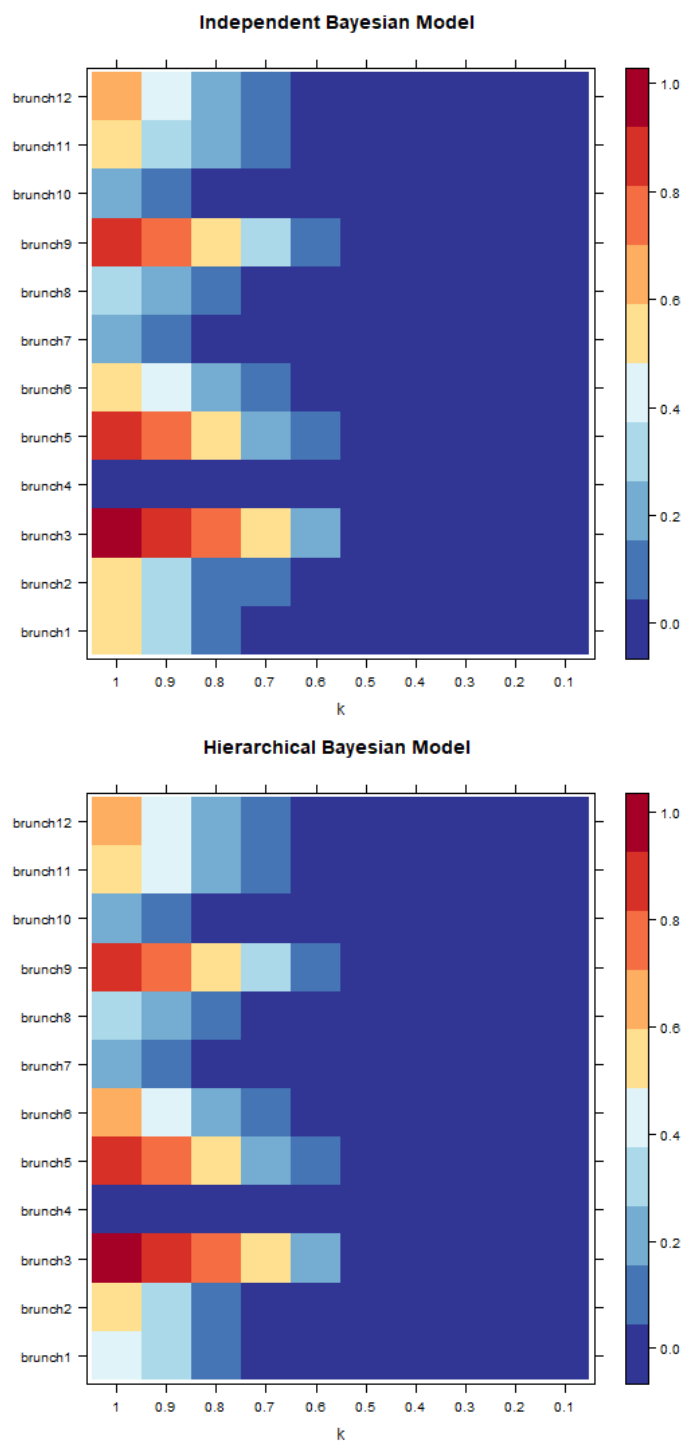


Figure 2.24

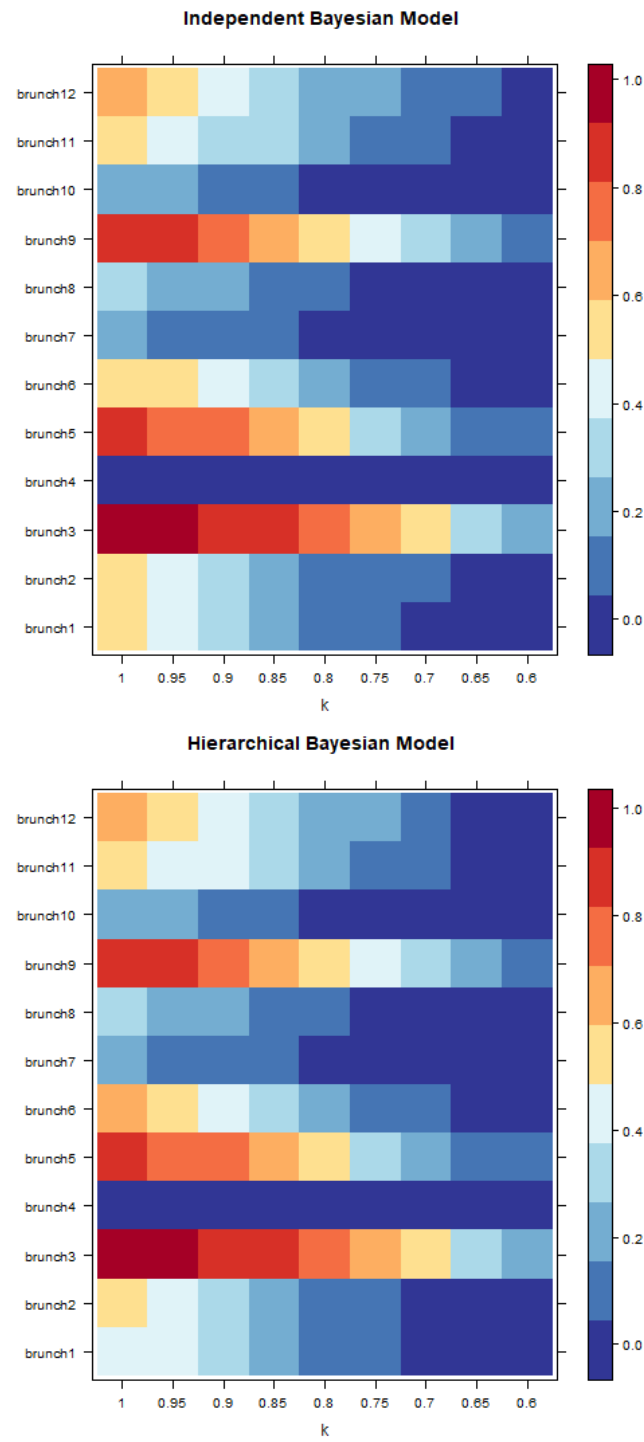


Figure 2.25

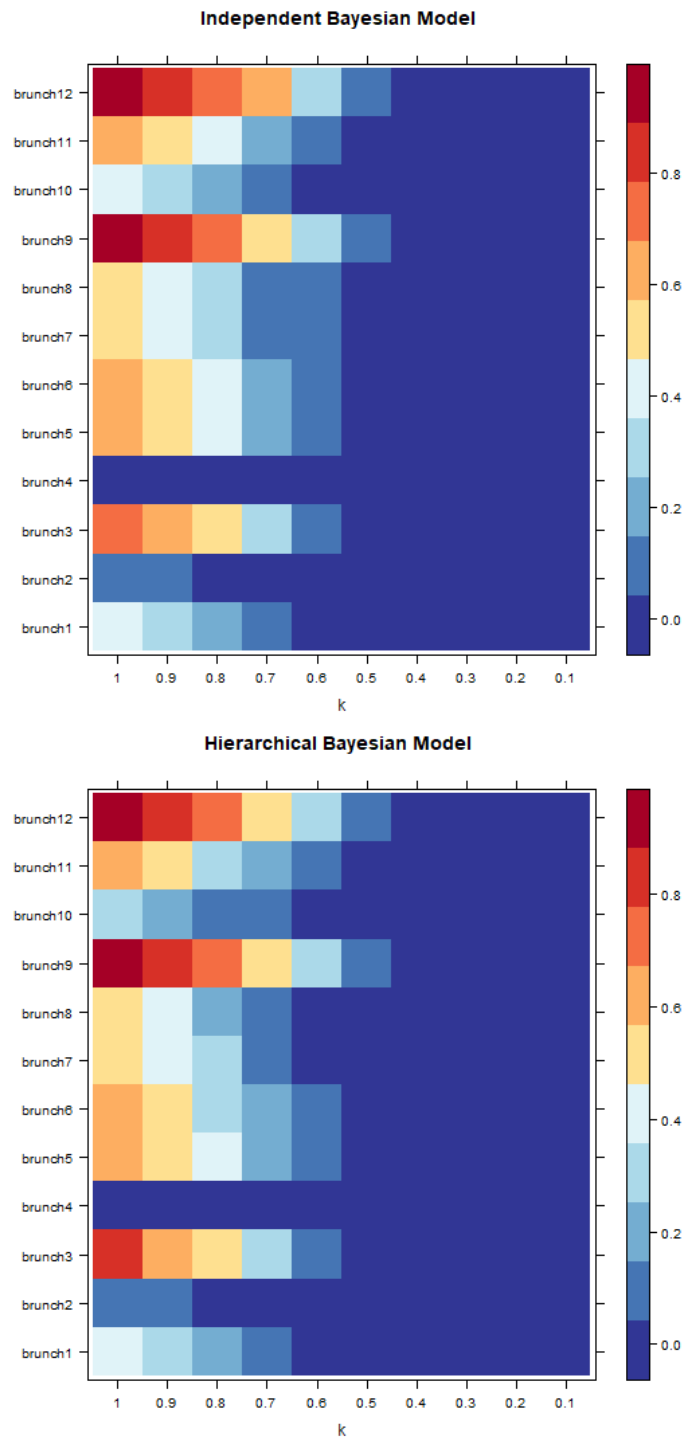


Figure 2.26

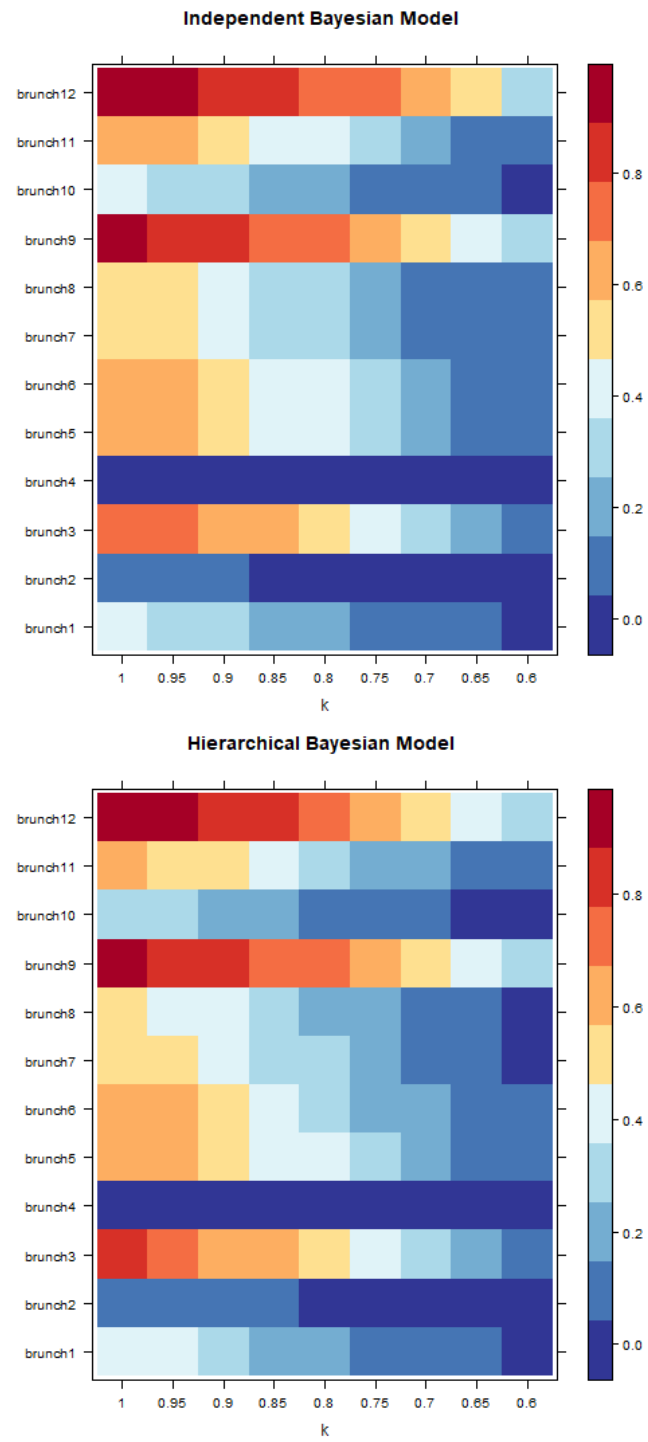


Figure 2.27

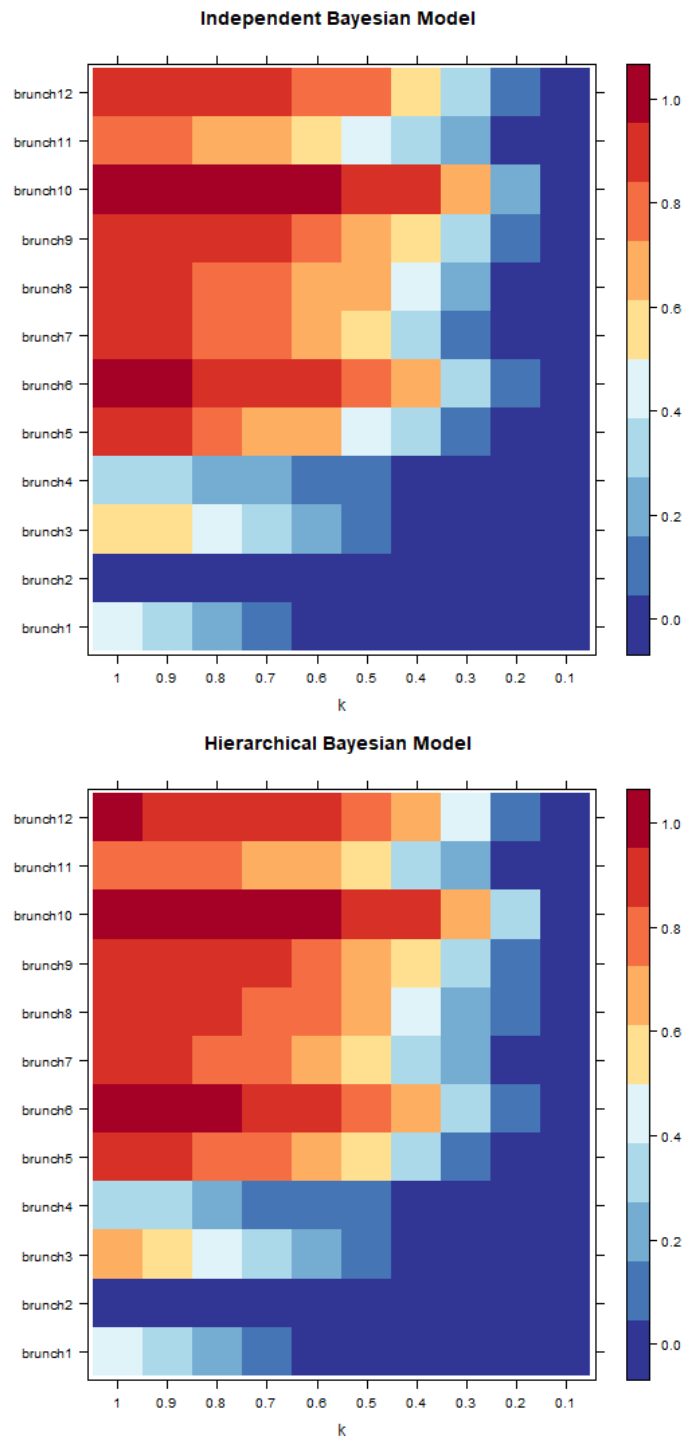


Figure 2.28

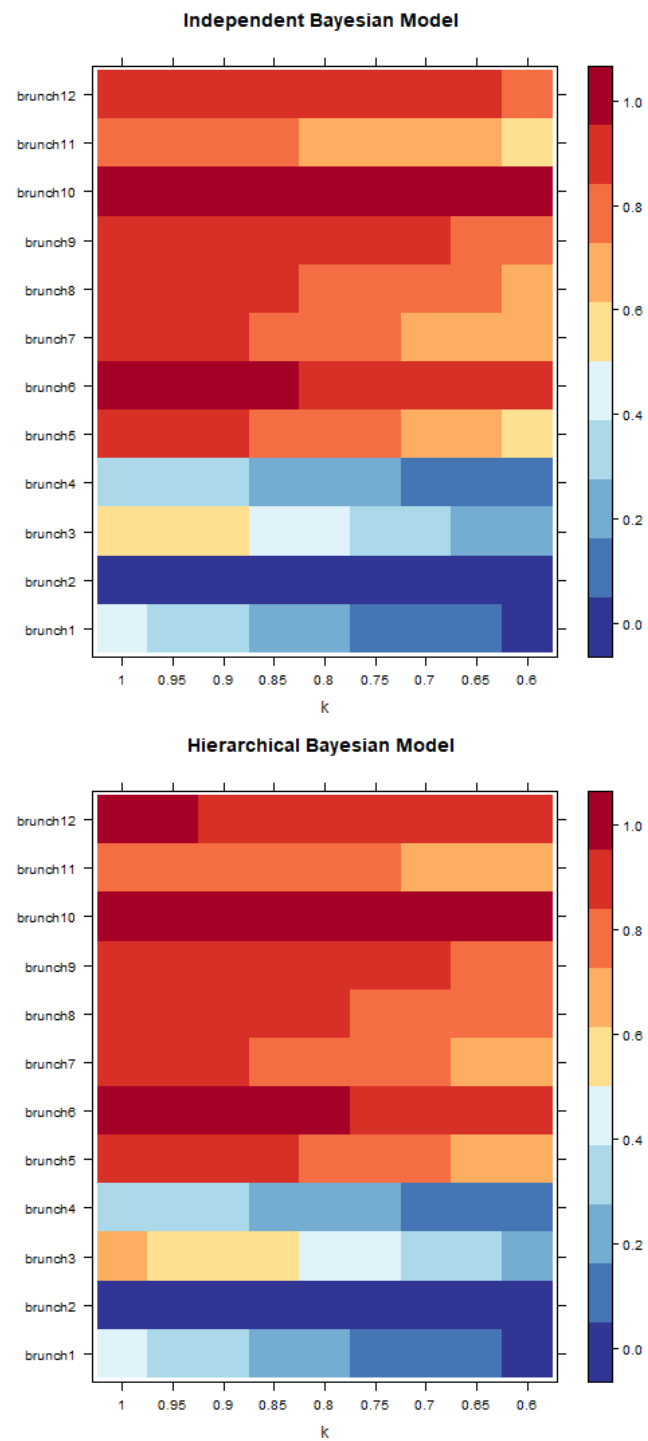


Figure 2.29