Spatiotemporal Modeling for HRF

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Bias-Variance Trade-off

Case I: i.i.d errors

Let's assume Canonical + Derivative (DD) model for HRF estimation and assume that error terms follow iid distribution with unknown variance of σ^2 .

For independent (Non-spatial Model), all timecourse for each voxel(coordinates) are fitted independently.

$$Y(x_1, x_2) = X\beta(x_1, x_2) + \epsilon, \quad \epsilon \sim N(0, \sigma^2 I)$$

Then,

$$\hat{\sigma}^{2} = \frac{1}{T - p} \hat{\epsilon}'(x_{1}, x_{2}) \hat{\epsilon}(x_{1}, x_{2}) \text{ where } \hat{\epsilon}(x_{1}, x_{2}) = Y(x_{1}, x_{2}) - X \hat{\beta}(x_{1}, x_{2})$$
$$\hat{\beta}(x_{1}, x_{2}) = (X'(x_{1}, x_{2})X(x_{1}, x_{2}))^{-1} X'(x_{1}, x_{2})Y(x_{1}, x_{2})$$

For indepdent model, we know fitted β is unbiased and sample variance for β is estimated using sample errors.

$$\hat{Var}(\hat{\beta}) = (X'X)^{-1}\hat{\sigma}^2 = (X'X)^{-1}\frac{1}{T-n}\hat{\epsilon}'\hat{\epsilon}$$

For Gaussian Kernel smoothing, we have biased but smaller variance of estimated β .

Gaussian kernel smoothing is used to capture spatial dependencies across neighboring voxels. The estimated smoothed time course data at time t at coordinates (x_1, x_2) , denoted as $\tilde{Y}(t, x_1, x_2)$, is given by:

$$\tilde{Y}(t, x_1, x_2) = \frac{\sum_i K(x_{1i}, x_{2i} | x_1, x_2, \sigma) Y_i(t, x_{1i}, x_{2i})}{\sum_i K(x_{1i}, x_{2i} | x_1, x_2, \sigma)},$$

where the Gaussian kernel $K(x_1, x_2 | \mu_1, \mu_2, \sigma)$ is defined as:

$$K(x_1, x_2 | \mu_1, \mu_2, \sigma) = \frac{1}{2\pi\sigma^2} \exp\left(-\frac{(x_1 - \mu_1)^2 + (x_2 - \mu_2)^2}{2\sigma^2}\right).$$

Researchers typically use a fixed standard deviation for the Gaussian kernel such that the Full Width at Half Maximum (FWHM) of the kernel is between 4 mm and 8 mm—commonly set to 6 mm. Assuming the coordinates are in millimeters, we calculate:

$$FWHM = 2\sqrt{2 \ln 2} \ \sigma = 6 \implies \sigma = \frac{3}{\sqrt{2 \ln 2}} \approx 2.55.$$

Let $\tilde{Y}(x_1, x_2)$ represent the time course at voxel (x_1, x_2) :

$$\tilde{Y}(x_1, x_2) = \begin{pmatrix} \tilde{Y}(1, x_1, x_2) \\ \vdots \\ \tilde{Y}(T, x_1, x_2) \end{pmatrix}.$$

For each voxel, the fitted β using Gaussian kernel smoothed time course data is expressed as:

$$\hat{\tilde{\beta}}(x_1, x_2) = (X' \Sigma^{-1} X)^{-1} X' \Sigma^{-1} \tilde{Y}(x_1, x_2).$$

The expected value of the fitted β is:

$$\mathbb{E}(\hat{\tilde{\beta}}(x_1, x_2)) = (X'X)^{-1}X'\mathbb{E}(\tilde{Y}(x_1, x_2)).$$

We know that $\mathbb{E}(Y(x_1, x_2)) = X\beta(x_1, x_2)$. Thus:

$$\mathbb{E}(\tilde{Y}(t, x_1, x_2)) = \mathbb{E}\left(\frac{\sum_{i} K(x_{1i}, x_{2i} | x_1, x_2, \sigma) Y(t, x_{1i}, x_{2i})}{\sum_{i} K(x_{1i}, x_{2i} | x_1, x_2, \sigma)}\right)$$

$$= \frac{\sum_{i} K(x_{1i}, x_{2i} | x_1, x_2, \sigma) \mathbb{E}(Y(t, x_{1i}, x_{2i}))}{\sum_{i} K(x_{1i}, x_{2i} | x_1, x_2, \sigma)}$$

$$\mathbb{E}(\tilde{Y}(x_1, x_2)) = \frac{\sum_{i} K(x_{1i}, x_{2i} | x_1, x_2, \sigma) X \beta(x_{1i}, x_{2i})}{\sum_{i} K(x_{1i}, x_{2i} | x_1, x_2, \sigma)}$$

$$\therefore \mathbb{E}(\tilde{\beta}(x_1, x_2)) = \frac{\sum_{i} K(x_{1i}, x_{2i} | x_1, x_2, \sigma) \beta(x_{1i}, x_{2i})}{\sum_{i} K(x_{1i}, x_{2i} | x_1, x_2, \sigma)}.$$

Thus, the fitted β using the Gaussian kernel smoothed time course data is biased, and its expectation is the Gaussian kernel smoothed true β .

For variance, we also using kernel smoothed residuals to estimate variance.

$$\begin{split} \hat{Var}(\hat{\tilde{\beta}}) &= (\boldsymbol{X}'\boldsymbol{X})^{-1}\hat{\tilde{\sigma}}^2 \\ \text{where } \hat{\tilde{\sigma}}^2 &= \frac{1}{T-p}(\tilde{\boldsymbol{\epsilon}})'(\tilde{\boldsymbol{\epsilon}}) \end{split}$$

Also, we know that residuals from gaussian kernel smoothed timecourse data are gaussian kernel smoothed residuals from independent model. i.e,

$$\begin{split} \tilde{\epsilon}(x_1, x_2) &= \tilde{Y}(x_1, x_2) - X \tilde{\beta}(x_1, x_2) \\ &= \frac{\sum_i K(x_{1i}, x_{2i} | x_1, x_2, \sigma) Y(x_{1i}, x_{2i})}{\sum_i K(x_{1i}, x_{2i} | x_1, x_2, \sigma)} \\ &- X(x^{'}X)^{-1} X^{'} \frac{\sum_i K(x_{1i}, x_{2i} | x_1, x_2, \sigma) Y(x_{1i}, x_{2i})}{\sum_i K(x_{1i}, x_{2i} | x_1, x_2, \sigma)} \\ &= \frac{\sum_i K(x_{1i}, x_{2i} | x_1, x_2, \sigma) (Y(x_{1i}, x_{2i}) - X \hat{\beta}(x_{1i}, x_{2i}))}{\sum_i K(x_{1i}, x_{2i} | x_1, x_2, \sigma)} \\ &= \frac{\sum_i K(x_{1i}, x_{2i} | x_1, x_2, \sigma) \hat{\epsilon}(x_{1i}, x_{2i})}{\sum_i K(x_{1i}, x_{2i} | x_1, x_2, \sigma)} \end{split}$$

Then, we can conclude that expectation of sample variance for gaussian kernel smoothed beta is smaller than for independent model which shows bias-variance trade off.

$$\begin{split} \mathbb{E}[\hat{Var}(\hat{\beta}) - \hat{Var}(\hat{\hat{\beta}})] &= \mathbb{E}[(X^{'}X)^{-1}(\hat{\sigma}^{2} - \hat{\tilde{\sigma}^{2}})] \\ &= (X^{'}X)^{-1}\frac{1}{T-p}\mathbb{E}[\sum_{t=1}^{T}[(\hat{\epsilon}(t))^{2} - (\tilde{\epsilon}(t))^{2}]] \\ &= (X^{'}X)^{-1}\frac{1}{T-p}\sum_{t=1}^{T}[(1 - \frac{\sum_{i}K_{i}^{2}}{(\sum_{i}K_{i})^{2}})\sigma^{2}(X(X^{'}X)^{-1}X^{'})_{tt}] \\ &> 0 \end{split}$$

Case II: AR(p) errors

Let's assume error terms are following AR(1) model for independent (non-spatial) model. We will estimate ρ by iterative Yule-Walker methods so we assume we have estimator for ρ which is consistent and asymptotically efficient.

Then, set

$$Y^* = DY, \ X^* = DX$$
 where $D = \begin{pmatrix} 1 & 0 & 0 & \cdots & 0 \\ -\rho & 1 & 0 & \cdots & 0 \\ 0 & -\rho & 1 & \cdots & 0 \\ \vdots & & & \ddots & \vdots \\ 0 & 0 & \cdots & -\rho & 1 \end{pmatrix}$
$$DY = DX\beta + D\epsilon$$

$$Y^* = X^*\beta + u$$
 where $u \sim N(0, \sigma^2)$

For independently fitted beta, with σ unknown, it's also unbiased.

$$\mathbb{E}[\hat{\beta}] = \mathbb{E}[(X^{*'}X^{*})^{-1}X^{*'}Y^{*}]$$

= β

$$\begin{split} \hat{Var}[\hat{\beta}] &= (X^{*'}X^{*})^{-1}\hat{\sigma}^2 \\ &= (X^{'}D^{'}DX)^{-1}\hat{\sigma}^2 \\ \text{where } \hat{\sigma}^2 &= \frac{1}{T-p}\hat{u}^{'}\hat{u} \\ \text{where } \hat{u} &= DY - DX\hat{\beta} \end{split}$$

For gaussian kernel smoothing beta we have biased beta but reduced variance of estimated sample variance.

$$\mathbb{E}[\hat{Var}(\hat{\beta}) - \hat{Var}(\hat{\hat{\beta}})] = (X^{*'}X^{*})^{-1} [\mathbb{E}(\hat{\sigma}^{2} - \frac{\sum_{i}[K_{i}^{2}]}{[\sum_{i}K_{i}]^{2}}\tilde{\sigma}^{2})]$$

$$\geq 0$$