A Derivation of Non-central Chi-square Density

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Lemma 1. Suppose $X \sim N(\mu, 1)$. Then the characteristic function of X^2 is

$$h(t;\mu) = \frac{\exp\left(\frac{i\mu^2 t}{1-2it}\right)}{(1-2it)^{1/2}}, \ t \in \mathbb{R}.$$
 (1)

Proof. By definition, the characteristic function of X^2 is

$$h(t) = \mathbb{E}\left[e^{itX^{2}}\right] = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} \exp\left\{itx^{2} - \frac{(x-\mu)^{2}}{2}\right\} dx$$

$$= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} \exp\left\{-\left[(1-2it)^{1/2}x - \frac{\mu}{(1-2it)^{1/2}}\right]^{2} + \frac{i\mu^{2}t}{1-2it}\right\} dx$$

$$= \exp\left(\frac{i\mu^{2}t}{1-2it}\right) \underbrace{\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} \exp\left\{-\left[(1-2it)^{1/2}x - \frac{\mu}{(1-2it)^{1/2}}\right]^{2}\right\} dx}_{(a)},$$
(2)

where the term (a) is $(1-2it)^{-1/2}$ as the integral of a shifted Gaussian density.

Lemma 2. Suppose V is a non-central chi-square variable with degree of freedom p > 0 and non-centrality parameter $\lambda > 0$. Then V can be represented as

$$V = Z_1^2 + Z_2^2 + \dots + Z_p^2, \ Z_1 \sim N(\sqrt{\lambda}, 1), \ Z_2, \dots, Z_p \sim N(0, 1),$$
(3)

where Z_1, \dots, Z_p are independent.

Proof. Let $X_i \sim N(\mu_i, 1), i = 1, \dots, p$, with $\lambda = \mu_1^2 + \dots + \mu_p^2 > 0$. Denote by **X** the random vector composed of X_1, \dots, X_p . By definition,

$$V = X_1^2 + X_2^2 + \dots + X_p^2 = ||\mathbf{X}||_2^2, \ \mathbf{X} \sim N(\boldsymbol{\mu}, \mathbf{I}_p), \tag{4}$$

where $\boldsymbol{\mu} = (\mu_1, \cdots, \mu_p)^{\top} \in \mathbb{R}^p$. Then we can expand $\lambda^{-1/2}\boldsymbol{\mu}$ to an orthogonal matrix $\mathbf{Q} = \begin{bmatrix} \lambda^{-1/2}\boldsymbol{\mu}^{\top} \\ * \end{bmatrix}$ of which the rows form an orthonormal basis on \mathbb{R}^p . Let $\mathbf{Z} = \mathbf{Q}\mathbf{X}$. Then

$$Z_1^2 + Z_2^2 + \dots + Z_p^2 = \|\mathbf{Z}\|_2^2 = \|\mathbf{Q}\mathbf{X}\|_2^2 = \|\mathbf{X}\|_2^2 = V.$$
 (5)

Moreover, Z_1, Z_2, \dots, Z_p are independent Gaussian variables characterized by (4).

Lemma 3 (Convolution theorem). Let X, Y be independent random variables. Then the characteristic function of X + Y is the pointwise product of the characteristic functions of X and Y.

Theorem 4 (Density of non-central chi-square distributions). Suppose V is a non-central chi-square variable with degree of freedom p > 0 and non-centrality parameter $\lambda > 0$. Then the probability density function of V is

$$f_{NC}(v; p, \lambda) = \sum_{k=0}^{\infty} \frac{e^{-\lambda/2} (\lambda/2)^k}{k!} f_{\chi^2}(v; p+2k),$$
 (6)

where $f_{\chi^2}(\cdot; p+2k)$ is the probability density function of $\chi^2(p+2k)$:

$$f_{\chi^2}(v; p+2k) = \frac{v^{p/2+k-1}}{2^{p/2+k}\Gamma(p/2+k)} e^{-v/2}.$$
 (7)

Proof. We use the representation of V given by Lemma 2. Then applying Lemmas 1 and 3 yields the characteristic function of V:

$$\varphi_V(t) = h(t; \sqrt{\lambda}) \prod_{i=2}^p h(t; 0) = \frac{\exp\left(\frac{i\lambda t}{1 - 2it}\right)}{(1 - 2it)^{p/2}}, \ t \in \mathbb{R}.$$
(8)

We can expand the numerator as follows:

$$\exp\left(\frac{\mathrm{i}\lambda t}{1-2\mathrm{i}t}\right) = \mathrm{e}^{-\lambda/2}\exp\left(\frac{\lambda/2}{1-2\mathrm{i}t}\right) = \sum_{k=0}^{\infty} \frac{\mathrm{e}^{-\lambda/2}}{k!} \left(\frac{\lambda/2}{1-2\mathrm{i}t}\right)^k. \tag{9}$$

Then

$$\varphi_V(t) = \sum_{k=1}^{\infty} \frac{e^{-\lambda/2} (\lambda/2)^k}{k!} \underbrace{\frac{1}{(1-2it)^{p/2+k}}}_{CF \text{ of } \chi^2(p+2k)}.$$
(10)

Applying Fourier transform on both sides of (10) yields the result of (6).

Remark. This theorem proposes another method of generating non-central chi-square variables. Fix $p, \lambda > 0$.

- Generate $k \sim \text{Poisson}(\lambda/2)$.
- Generate i.i.d. $X_1, \dots, X_{p+2k} \sim N(0,1)$, and set $V = X_1^2 + \dots + X_{p+2k}^2$.

Then V is a non-central chi-square variable with degree of freedom p and non-centrality parameter λ .