



CLINICAL REVIEW

Cognitive behavioral therapy for insomnia in patients with chronic pain – A systematic review and meta-analysis of randomized controlled trials



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SUMMARY

Several randomized controlled trials have implemented cognitive behavioral therapy for insomnia (CBT-I) for patients with comorbid insomnia and chronic pain. This systematic review and meta-analysis investigated the effectiveness of CBT-I on patient-reported sleep, pain, and other health outcomes (depressive symptoms, anxiety symptoms, and fatigue) in patients with comorbid insomnia and chronic non-cancer pain. A systematic literature search was conducted using eight electronic databases. Upon duplicate removal, 6374 records were screened against the inclusion criteria. Fourteen randomized controlled trials were selected for the review, with twelve (N = 762 participants) included in the meta-analysis. At post-treatment, significant treatment effects were found on global measures of sleep (standardized mean difference = 0.89), pain (0.20), and depressive symptoms (0.44). At follow-up (up to 12 mo), CBT-I significantly improved sleep (0.56). Using global measures of sleep, we found a probability of 81% and 71% for having better sleep after CBT-I at post-treatment and final follow-up, respectively. The probability of having less pain after CBT-I at post-treatment and final follow-up was 58% and 57%, respectively. There were no statistically significant effects on anxiety symptoms and fatigue at either assessment point. Future trials with sufficient power, longer follow-up periods, and inclusion of CBT for pain components are warranted.

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Introduction

Chronic non-malignant pain is a long-term and debilitating condition that affects 50 million adults in the USA [1,2]. Types of chronic non-malignant pain include fibromyalgia, musculoskeletal pain, osteoarthritis, and complex regional pain syndrome. Insomnia, associated with disturbed sleep and problems initiating, maintaining, or returning to sleep, is frequently reported as a comorbidity in

the pain population who commonly suffer from sleep disturbance and problems initiating, maintaining, or returning to sleep [3,4]. Prior research suggests that clinical insomnia is prevalent in up to 53% of patients with chronic pain, which is higher than in the general population [4,5]. Further, up to 90% of patients with chronic pain report seeking medical attention for insomnia treatment [6,7].

Several studies report a reciprocal relationship between sleep and pain, such that sleep impairments exacerbate chronic pain, and patients with greater pain levels report more sleep complaints [8–11]. Interestingly, a growing body of research suggests that sleep is a stronger predictor of subsequent pain than vice versa [9,12]. Further, experimental studies have found that sleep disturbances contribute to the development of chronic pain, and

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Glossary of Terms

| | |
|----------|--|
| CBT-I | cognitive behavioral therapy for insomnia |
| CBT-P | cognitive behavioral therapy for pain |
| CINAHL | cumulative index of nursing and allied health literature |
| PRISMA | preferred reporting items for systematic reviews and meta-analyses |
| PROSPERO | international prospective register of systematic reviews |
| RCT | randomized controlled trial |
| SD | standard deviation |
| SE | sleep efficiency |
| SMD | standardized mean difference |
| SOL | sleep onset latency |
| TST | total sleep time |
| WASO | wake after sleep onset |

improving sleep reduces pain intensity [9,12,13]. These findings suggest that sleep could have a causal impact on chronic pain [9], and may be an important therapeutic target in treating insomnia in patients with chronic pain.

Cognitive behavioral therapy for insomnia (CBT-I) is a non-pharmacological treatment for insomnia comprising multiple components, including sleep hygiene education, stimulus control, sleep restriction, cognitive therapy, and relaxation training [14,15]. CBT-I is the recommended initial treatment for insomnia, works to strengthen the bed/bedroom environment as a cue for sleep, establish a regular sleep schedule, and minimize thoughts and behaviors that impair sleep [16]. An increasing number of studies have sought to use CBT-I to treat comorbid insomnia in patients with chronic pain, with significant improvements reported in sleep outcomes but mixed results on pain outcomes [15,17,18]. To address the inconsistency in pain outcomes, a small number of studies have developed hybrid interventions combining CBT-I and cognitive behavioral therapy for pain (CBT-P) to effectively address both pain and insomnia symptoms [19,20]. Given that insomnia is associated with an increased risk of developing other adverse outcomes including depressive and anxiety symptoms [21–23], CBT-I could potentially prevent exacerbation of these outcomes. However, studies addressing this question have reported inconsistent findings [24–26]. To date, it is uncertain whether better sleep with CBT-I also helps improve chronic pain and other health outcomes.

The role of CBT-I in patients with chronic pain has previously been examined in narrative reviews [15,17]. It was also evaluated in a meta-analysis within a group of broadly defining non-pharmacological interventions in patients with either chronic non-cancer or cancer pain [18]. To date, there is no meta-analysis conducted to specifically appraise CBT-I trials for individuals with chronic non-malignant pain. The objective of the current systematic review and meta-analysis is to assess the efficacy of CBT-I in improving patient-reported sleep and other health outcomes (e.g., pain, depressive symptoms, anxiety symptoms, and fatigue) among adults with chronic pain. While we hypothesized that CBT-I will improve sleep, our analyses were exploratory in nature with regards to the other health outcomes. We also conducted a meta-regression to determine the predictors of treatment effects.

Methods

The protocol of this systematic review was registered in the international prospective register of systematic reviews

(PROSPERO) (CRD42020201264). This study was performed in accordance with the preferred reporting items for systematic reviews and meta-analyses (PRISMA) guidelines [27].

Inclusion and exclusion criteria

The inclusion criteria for the study were: 1) randomized controlled trial (RCT), testing CBT-I as a stand-alone intervention or part of a hybrid intervention with CBT-P, in adults (>18 y), with chronic non-malignant pain (e.g., arthritis, fibromyalgia, musculoskeletal pain); 2) measures patient-reported sleep as a primary or secondary objective; 3) measures patient-reported pain as a primary or secondary objective, 4) the following components must be included to classify as CBT-I: at least one behavioral technique (e.g., stimulus control, sleep restriction) and cognitive therapy, and 5) English language. Studies exclusively containing patients with pain relating to rheumatoid arthritis, headache, cancer, palliative conditions, and traumatic brain injury were excluded to focus on the broader chronic pain population.

Literature search strategy

A systematic search was conducted by an information specialist (ME) of Medline, Medline In-Process/ePubs, Embase, Cochrane Controlled Register of Trials, Cochrane Database of Systematic Reviews, APA PsycINFO, CINAHL, and ClinicalTrials.Gov. The duration of the search was set from database inception to April 2020. The search terms included “chronic pain”, “insomnia”, and cognitive behavioral therapy”. All search terms are listed in Table S1.

Study selection

Fig. 1 illustrates the PRISMA flow diagram of the search and selection process. The literature search yielded 7772 records, and an additional record was identified from independently reviewing the citations of reviews and included articles. After removing 1399 duplicate records, 6374 records were screened. Records were

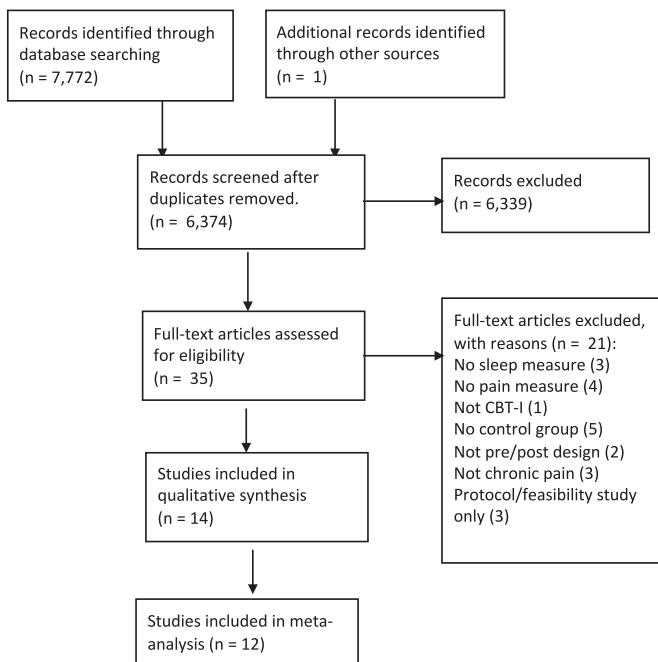


Fig. 1. PRISMA flowchart.

independently selected by two reviewers (JS, CP), who screened the titles and abstracts according to the inclusion and exclusion criteria. Thirty-five full text articles were retrieved for further examination. Disagreements between the two reviewers were resolved by consulting the senior author (FC). Upon full-text screening, 14 articles were included in the data extraction.

Data extraction and quality rating

A data extraction form was designed in Excel, and used to collect the following information: methodological (author, year of publication, country, participants, sample size, number of arms, outcome measures, quality ratings), and treatment details (components, duration and frequency, format). Data extraction was performed independently by two authors (JS, CP). Attempts were made to obtain any missing data from studies by contacting the corresponding author. Any differences in extracted data were reviewed by a third author (FC) and a consensus was reached via discussion.

For studies with multiple control arms, the least active arm was chosen. For studies with multiple active sleep intervention arms (e.g., CBT-I versus hybrid CBT-I/CBT-P), the arm comprising mostly of CBT-I components was chosen [19]. When multiple global measures (i.e., questionnaires) were used to evaluate the same outcome, data for the assessment used most commonly across the studies was extracted to reduce heterogeneity. The following outcomes were also extracted from studies using sleep diaries to measure sleep: sleep onset latency (SOL), wake after sleep onset (WASO), total sleep time (TST), and sleep efficiency (SE). Means and standard deviations (SD) of the selected outcome measures were obtained for the intervention and control groups at baseline, post-treatment, and follow-up (if applicable). For studies with multiple follow-ups, data from the final follow-up was extracted.

Quality assessments were conducted independently by two authors (JS, CP), and a third author (FC) was consulted to resolve disagreement. Additionally, both screeners (JS, CP) received training from the senior author (FC) on how to utilize these assessments. The Cochrane Risk of Bias tool [28] and Yates et al. [29] quality rating assessment were used to qualitatively and quantitatively determine study quality, respectively. For the Cochrane tool, studies received a "low", "high", or "unclear" rating for each risk category. The domain "blinding of participants and personnel" was removed from the quality assessment. Given that CBT-I and other nonpharmacological treatments require active involvement with the therapeutic materials, participants and therapists are never truly blinded to the group assignment [18]. The Cochrane tool has shown adequate inter-rater reliability for RCTs in patients with neck pain [30]. The Yates quality rating scale evaluates the treatment and methodological quality of RCTs for psychological treatments and has demonstrated adequate validity and inter-rater reliability [29]. The total score of the scale is 35 (nine for treatment + 26 for methodology), and two raters reported scores of 22.70, 18.71, and 12.10 as "excellent", "average", and "poor" psychological trials for chronic pain [29].

Data analysis

For trials with multiple publications, the article stating the main and most relevant analysis was included in the meta-analysis [26,31–33]. A random-effects model was used to calculate standardized mean differences (SMD) between scores at post-treatment (baseline - post-treatment) and final follow-up (baseline - follow-up) for each group. Then, these SMD values ($SD = 1$) were entered into a random-effects model to calculate SMD values between the CBT-I and control groups at post-treatment and follow-up. Confidence interval (CI) and prediction intervals (95%)

were determined for SMD. We used the number of studies, SMD, upper value of the 95% CI and Tau^2 was used to calculate the predictive interval. The probability of the benefit of the treatment was calculated using the Z table statistics.

Publication bias was assessed using the Egger's test and Begg's test. The I^2 statistic was used to examine heterogeneity for each outcome (excluding sleep diary variables). Sensitivity analysis or influential analysis was performed to explore the heterogeneity, where one study was excluded each time to recalculate I^2 and end estimate. Planned subgroup analyses were conducted to compare treatment effects between group versus individually delivered CBT-I for each outcome (excluding sleep diary variables) and assessment point. Meta-regression analyses were performed *a priori* to determine whether the following dichotomous factors predicted treatment effects on global measures of sleep: study quality by Cochrane risk assessment (some concerns vs. low concerns), Yates quality rating scale ≥ 25 (yes vs. no), loss of patients at post-treatment (yes vs. no), CBT-I format (group vs. individual), and sample size > 50 (yes vs. no). Meta-analyses were conducted using RevMan 5.4 and the meta-regression was performed with Comprehensive Meta-Analysis Software.

Results

Study characteristics

Table 1 presents the study characteristics of the 14 RCTs included in the systematic review. Studies were conducted in Canada ($n = 1$), Spain ($n = 3$), UK ($n = 1$), and the USA ($n = 9$). Twelve of these studies were included in the meta-analysis, comprising 762 participants (range of 20–367 participants, mean age: 45–73.1 y, female: 55%–100%) [19,20,24–26,33–39].

Four studies involved participants with different types of chronic non-malignant pain, with three utilizing a physician or physical examination to confirm chronic pain diagnosis [19,20,24,25]. Another five studies comprised of individuals with fibromyalgia who met the American College of Rheumatology criteria for fibromyalgia [26,32,33,36,37,40]. The remaining five studies were conducted in persons with osteoarthritis [31,33,35,38,39], and the diagnosis confirmed by radiographic or physician examination in three of these studies [35,38,39].

The diagnostic criteria for insomnia was based on the Diagnostic and Statistical Manual of Mental Disorders (DSM-III-R, DSM-IV and DSM-IV-TR) manual [41–43] in five studies [24,26,32,34,36], and the Research Diagnostic Criteria for Insomnia in four studies [31,33,35,38,44]. The remaining five RCTs utilized a combination of self-reported criteria that varied in duration, frequency, and included outcomes (e.g., sleep onset latency, wake after sleep onset, total sleep time) [19,20,25,37,39]. Two studies also included the Insomnia Severity Index in their insomnia diagnostic criteria (≥ 10 [35] and ≥ 15 [19]). Five studies utilized polysomnography to rule out for other sleep disorders (e.g., sleep apnea) [19,25,32,34,39].

All RCTs involved multimodal CBT-I, although our inclusion criteria only required at least one behavioral and one cognitive technique (**Table 2**). The common components include: sleep education, sleep hygiene, sleep restriction, stimulus control, cognitive therapy, and relaxation training. Four RCTs utilized a combined CBT-I and CBT-P (e.g., cognitive therapy for pain, pain education, reducing pain catastrophizing and safety seeking behavior, and relaxation) [20,31,33,36]. The frequency and duration of CBT-I ranged from 4 to 10 wk and 15–120 min per week, respectively. All CBT-I interventions were conducted in a face-to-face medium, with seven trials performed in a group (4–12 participants per group) [24,26,31–33,36,39] and seven trials used an individual format [19,20,25,34,35,37,38]. The control arms were comprised of

Table 1
Study characteristics.

| Author Year (Country) | Pain Condition and Diagnosis | Insomnia Diagnosis | N | Age, mean (years) | Female (%) | Assessment points | Global Measures | Yates Quality Rating |
|-------------------------------------|--|---|-----|-------------------|------------|--|--|---|
| Currie et al. (2000) [24] (Canada) | Chronic pain: back (72%), neck (20%), lower limbs (5%), pelvic (3%) Dx: specialist in physical medicine | DSM-IV | 60 | 45 | 55 | Baseline, post-Rx, 3-mo follow-up | Sleep: PSQI Pain: MPI-PS Depressive symptoms: BDI Anxiety symptoms: None Fatigue: None | Rx: 9 Methodology: 16.5 Overall: 25.5 |
| Edinger et al. (2005) [34] (USA) | FM (100%) Meet the American College of Rheumatology criteria for FM. Dx: rheumatologist | DSM-III-R; sleep diary; and PSG | 47 | 48.6 | 95.7 | Baseline, post-Rx, 6-mo follow-up | Sleep: ISQ Pain: MPQ Depressive symptoms: None Anxiety symptoms: None Fatigue: None | Rx: 8 Methodology: 17.5 Overall: 25.5 |
| Heffner et al. (2018) [35] (USA) | Knee OA (100%) Dx: radiographic evidence or physician confirmation | Research diagnostic criteria for insomnia; ISI | 30 | 61 | 60 | Baseline, post-Rx | Sleep: ISI Pain: WOMAC Index Depressive symptoms: None Anxiety symptoms: None Fatigue: None | Rx: 2.5 Methodology: 9.5 Overall: 12 |
| Jungquist et al. (2010) [25] (USA) | Chronic pain: lower back (64%), neck (32%), and thoracic spinal level (4%) Dx: physical examination, urinalysis, and bloodwork | SOL > 30 min and/or WASO > 30 min for > 3 d/wk for > 6 mo; PSG | 28 | 48.7 | 78 | Baseline, post-Rx | Sleep: ISI Pain: MPI-PS Depressive symptoms: BDI Anxiety symptoms: None Fatigue: None | Rx: 9 Methodology: 16.5 Overall: 25.5 |
| Lami et al. (2018) [36] (Spain) | FM (100%) Meet the American College of Rheumatology criteria for FM for > 6 mo | DSM-IV-TR | 113 | 50.2 | 100 | Baseline, post-Rx, 3-mo follow-up | Sleep: PSQI Pain: MPQ-SF-VAS Depressive symptoms: SCL-90-R-D Anxiety symptoms: SCL-90-R-A Fatigue: MFI | Rx: 7 Methodology: 19.5 Overall: 26.5 |
| Martinez et al. (2013) [26] (Spain) | FM (100%) Had to meet the American College of Rheumatology criteria for FM for > 6 mo | DSM-IV-TR | 59 | 47.6 | 100 | Baseline, post-Rx, 3- and 6-mo follow-up | Sleep: PSQI Pain: MPQ-SF-VAS Depressive symptoms: SCL-90-R-D Anxiety symptoms: SCL-90-R-A Fatigue: MFI | Rx: 7.5 Methodology: 19 Overall: 26.5 |
| McCrae et al. (2019) [37] (USA) | FM (100%) Diagnosis: tender point testing; Meet the American College of Rheumatology criteria for FM for > 6 mo | SOL > 30 min or WASO > 30 min for at least 3 d per wk for > 6 mo; Sleep diary | 113 | 53 | 97.3 | Baseline, post-Rx, 6-mo follow-up | Sleep: Not applicable Pain: MPQ Depressive symptoms: BDI Anxiety symptoms: STAI-Y1 Fatigue: None | Rx: 8.5 Methodology: 21.5 Overall: 30 |
| McCurry et al. (2014) [31] (USA)* | OA (100%) Medical care for OA in the prior 3 y; Significant arthritis pain as defined by Grade II, III, or IV pain on the Graded Chronic Pain Scale | Self-reported sleep difficulties for 3 or more nights per wk during the past month with at least one daytime sleep-related problem; Research diagnostic criteria for insomnia | 367 | 73.1 | 78.2 | Baseline, 18-mo follow-up | Sleep: ISI Pain: GCPS Depressive symptoms: GDS Anxiety symptoms: None Fatigue: None | Rx: 8 Methodology: 20 Overall: 28 |
| Miro et al. (2011) [32] (Spain)* | FM (100%) Meet the American College of Rheumatology criteria for FM; FM Dx by medical exam. | DSM-IV; Interview; Questionnaires; Neuropsychological test; PSG | 31 | 46.5 | 100 | Baseline, post-Rx | Sleep: PSQI Pain: MPQ-SF-VAS Depressive symptoms: HADS-D Anxiety symptoms: HADS-D Fatigue: None | Rx: 7 Methodology: 16 Overall: 23** |
| Pigeon et al. (2012) [19] (USA) | Chronic pain: spine, shoulders, hips, or limbs Pain for > 6 mo; patients had a physical examination with ECG, clinical chemistries and toxicology screens | SOL and/or WASO 30 min or longer for > 3 d per wk for 6 mo or longer; PSG | 21 | 50.7 | 66.7 | Baseline, post-Rx | Sleep: ISI Pain: MPI Depressive symptoms: CEDS-Revised Anxiety symptoms: None Fatigue: MFI | Rx: 4.5 Methodology: 11.5 Overall: 16** |
| Smith et al. (2015) [38] (USA) | Knee OA (100%) Meet the American College of Rheumatology criteria for knee OA; Dx: rheumatologist plus | Research diagnostic criteria for insomnia; Symptoms lasting >1 mo, > 2 nocturnal awakenings lasting | 100 | 59.4 | 79 | Baseline, post-Rx, 3- and 6-mo follow-up | Sleep: ISI Pain: WOMAC Index Depressive symptoms: None | Rx: 5.5 Methodology: 19.5 Overall: 25 |

| | | | | |
|-----------------------------------|--|---|---|--|
| | radiographic evidence of knee OA, knee pain ratings, and stable dose of NSAIDs. | >15 min, or WASO and/or SOL | Anxiety symptoms: None Fatigue: None | Rx: 6.5 Methodology: 17.5 Overall: 24** |
| Tang et al. (2012) [20] (UK) | Chronic pain: MSK pain (85%), OA (35%), RA (10%), complex regional pain syndrome (10%), headache (10%), and FM (5%) Pain for > 6 mo; moderate severity based on BPI | ISI; SOL or WASO 30 min or longer for 3 or more nights per wk for one mon or longer | Baseline, post-Rx, 1-mo, 6-mo follow-up | Sleep: ISI Pain: BPI-PI Depressive symptoms: HADS-D Anxiety symptoms: HADS-A Fatigue: MFI |
| Vitiello et al. (2009) [39] (USA) | OA (100%) Dx: radiograph or MRI; Receiving physician Rx; Moderate pain rating on SF-36 pain item, or pain item from the Arthritis Impact Measurement Scales-2 | At least 3 episodes of insomnia per wk for at least 6 mo (defined as SOL 30 min or longer, WASO 60 min or longer, total sleep time < 6.5 h per night); PSG | Baseline, post-Rx, 12-mo follow-up | Sleep: Not applicable Pain: MPQ-SF Depressive symptoms: GDS Anxiety symptoms: None Fatigue: None |
| Vitiello et al. (2013) [33] (USA) | OA (100%) Medical care for OA in the prior 3 y; Significant OA pain as defined by Grade II, III, or IV pain on the Graded Chronic Pain Scale | Self-reported sleep difficulties for 3 or more nights per wk during the past mon with at least one daytime sleep-related problem; research diagnostic criteria for insomnia | Baseline, post-Rx, 9-mo follow-up | Sleep: ISI Pain: GCPs Depressive symptoms: GDS Anxiety symptoms: None Fatigue: None |

Abbreviations: BDI, Beck Depression Inventory; BPI-PI, Brief Pain Inventory Pain Interference subscale; CEDS-Revised, Center for Epidemiologic Studies Depression Scale Revised; DBAS, Dysfunctional Beliefs and Attitudes about Sleep; DSM-III-R, American Psychiatric Association: Diagnostic and Statistical Manual of Mental Disorders Third Edition, revised; DSM-IV-TR, American Psychiatric Association: Diagnostic and Statistical Manual of Mental Disorder Fourth Edition, Text Revision; FM, fibromyalgia; GCPs, Graded Chronic Pain Scale; GDS, Geriatric Depression Scale; HADS-A, Hospital Anxiety and Depression Scale Anxiety subscale; HADS-D, Hospital Anxiety and Depression Scale Depression subscale; ISI, Insomnia Severity Index; MFI, Insomnia Symptom Questionnaire; MFI, Multidimensional Pain Inventory; MPI-PS, Multidimensional Pain Inventory; MPI-PSS, Multidimensional Pain Inventory Short Form; OA, osteoarthritis; PSC, polysomnography; PSQI; Pittsburgh Sleep Quality Index; SCL-90-R-A, Symptom Checklist-90-revised Anxiety subscale; SCL-90-R-D, Symptom Checklist-90-Revised Depression subscale; SE, sleep efficiency; SOL, sleep onset latency; STAI-Y1, State-Trait Anxiety Inventory-Form Y1; WASO, wake after sleep onset; VAS, visual analogue scale; WOMAC Index; Western Ontario and McMaster Universities Osteoarthritis Index. *Study was not included in the meta-analysis, results from main trial article was included in analyses. **Pilot trial did not calculate power and sample size.

passive (e.g., waitlist control, usual medical care, contact control) [19,24,33,35–37], and active interventions (e.g., sleep hygiene, education, contact control, behavioral desensitization, symptom monitoring, attention control) [20,25,26,31–33,38,39].

All but three studies reported sleep, pain, and at least one other health outcome at baseline and post-treatment [34,35,38]. The follow-up period ranged from 3 to 18 mo.

Quality assessment

The quantitative assessments using the Yates quality rating scale demonstrated a mean treatment score of 6.8 (SD: 1.83, range: 2.5 to 9), mean methodology score of 16.7 (SD: 3.23, range: 9.5 to 21.5), and mean overall score of 23.5 (SD: 4.77, range: 12 to 30) (Table 1).

Using the Cochrane tool, all studies had a low or unclear bias on all domains (Fig. S1). Random sequence generation was the most sufficiently addressed, with all but three studies reporting low bias in this domain [34,35,38]. Allocation concealment was the least sufficiently addressed, with only five studies reporting low bias in this domain [20,26,32,36,37].

Meta-analysis

Sleep

Ten studies contributed 644 participants to analyses on global measures of sleep [19,20,24–26,33–36,38]. The Insomnia Severity Index was the most commonly used measure of sleep [45]. Other global measures of sleep included the Pittsburgh Sleep Quality Index [46], and Insomnia Symptom Questionnaire [47]. Across all outcome measures of sleep, CBT-I caused significant large effects at post-treatment ($SMD = 0.89$, [95% Confidence Interval: 0.53, 1.25], $p < 0.00001$) (Fig. 2a). The 95% prediction interval was -0.24 to 2.02 . Both Egger's Test and Begg's test showed evidence of publication bias on post-treatment effects of sleep ($p < 0.005$), but no publication bias was apparent at follow-up. Given that significant heterogeneity was found between studies at post-treatment, a sensitivity analysis detected Jungquist et al. [25] and Vitiello et al. [33] as potential sources of heterogeneity. Excluding studies decreased the I^2 statistic from 73% to 37% and the overall treatment effect on sleep remained significant (0.86 [0.57, 1.15], $p < 0.00001$) (Table 3). Subgroup analyses demonstrated that the efficacy of CBT-I remained largely significantly for both group-based [24,26,33,36] and individual-based [19,20,25,34,35,38] interventions at post-treatment (Table 3).

Six studies contributed data from 567 participants on sleep at follow-up (range: three to nine months) [24,26,33,34,36,38]. There was a large significant treatment effect of CBT-I at final follow-up (0.56 [0.25, 0.87], $p = 0.0004$) (Fig. 2b). The 95% prediction interval was -0.38 to 1.50 . Significant heterogeneity was found with an I^2 statistic of 64%. Removing Currie et al. [24] decreased the heterogeneity to a nonsignificant I^2 statistic of 18%, and the overall treatment effect remained largely significant (0.40 [0.19, 0.61], $p = 0.0001$) (Table 3). Subgroup analyses showed that group-[24,26,33,36] and individual-based [34,38] CBT-I demonstrated significant effects at follow-up (Table 3). In comparison to controls, the probability of having better sleep after CBT-I was 81% and 71% at post-treatment and final follow-up, respectively.

Seven studies contributed to pooled analyses assessing the effects of CBT-I on SOL and WASO at post-treatment ($n = 358$) [20,24,25,34,37–39], and four studies involving 259 participants contributed to analyses at follow-up (range: three to six months) [24,34,37,38]. In comparison to controls, CBT-I elicited significant effects on SOL at post-treatment (1.06 [0.48, 1.64], $p = 0.0003$) and follow-up (1.08 [0.13, 2.04], $p = 0.03$) (Figs. S2a and b). The prediction intervals for SOL at post-treatment and follow-up

Table 2

CBT-I and control arm components.

| Author, Year | Number of Arms: Description of Arms | Frequency (Duration) | Medium | Format |
|------------------------------|---|--|--------------|--------------------------|
| Currie et al. (2000) [24] | Two arms: CBT-I vs. wait-list control | 7 weekly sessions (120 min) | Face-to-face | Group (5–7 individuals) |
| Edinger et al. (2005) [34] | Three arms: CBT-I vs. sleep hygiene vs. usual Care | 6 weekly sessions (1st lasted 45–50 min and subsequent ones 15–30 min) | Face-to-face | Individual |
| Heffner et al. (2018) [33] | Two arms: CBT-I vs. control | 6 weekly sessions (1st lasted 90 min and subsequent ones 60 min) | Face-to-face | Individual |
| Jungquist et al. (2010) [25] | Two arms: CBT-I vs. contact control | 8 weekly sessions (30–90 min) | Face-to-face | Individual |
| Lami et al. (2018) [36] | Three arms: CBT-P vs. CTP-IP vs. usual medical care | 9 weekly sessions (90 min) | Face-to-face | Group (5–7 individuals) |
| Martinez et al. (2013) [26] | Two arms: CBT-I vs. sleep hygiene | 6 weekly sessions (90 min) | Face-to-face | Group (5–6 individuals) |
| McCrae et al. (2019) [37] | Three arms: CBT-I vs. CTP-P vs. wait-list control | 8 weekly sessions (50 min) | Face-to-face | Individual |
| McCurry et al. (2014) [31] | Three arms: CBT-P vs. CTP-IP vs. education-only Control | 6 weekly sessions (90 min) | Face-to-face | Group (5–12 individuals) |
| Miro et al. (2011) [32] | Two arms: CBT-I vs. sleep hygiene | 6 weekly sessions (90 min) | Face-to-face | Group (5–6 individuals) |
| Pigeon et al. (2012) [19] | Four arms: CBT-P vs. CBT-I vs. CBT-IP vs. wait-list control | 10 weekly sessions (50–60 min) | Face-to-face | Individual |
| Smith et al. (2015) [38] | Two arms: CBT-I vs. behavioral desensitization | 8 weekly sessions (45 min) | Face-to-face | Individual |
| Tang et al. (2012) [20] | Two arms: CBT-IP vs. symptom monitoring | 4 weekly sessions (120 min) | Face-to-face | Individual |
| Vitiello et al. (2009) [39] | Two arms: CBT-I vs. attention control | 8 weekly sessions (120 min) | Face-to-face | Group (4–8 individuals) |
| Vitiello et al. (2013) [33] | Three arms: CBT-P vs. CTP-IP vs. education-only control | 6 weekly sessions (90 min) | Face-to-face | Group (5–12 individuals) |

Abbreviations: CBT, cognitive behavioral therapy; CBT-I, cognitive behavioral therapy for insomnia; CBT-IP, cognitive behavioral therapy for insomnia and pain; CBT-P, cognitive behavioral therapy for pain.

were –0.85 to 2.97 and –3.28 to 5.44 respectively. Additionally, significant treatment effects on WASO were observed at post-treatment (1.00 [0.53, 1.47], $p < 0.0001$) and follow-up (0.87 [0.13, 1.60], $p = 0.02$) (Figs. S2c and d). The prediction intervals for WASO at post-treatment and follow-up were –0.47 to 2.47 and –2.43 to 4.17 respectively. Eight studies contributed to pooled analyses assessing the effects of CBT-I on TST and SE at post-treatment ($n = 368$) [19,20,24,25,34,37–39], and four studies involving 259 participants contributed to analyses at follow-up (range: three to six months) [24,34,37,38]. No significant treatment effects were found on TST at post-treatment (–0.11 [–0.39, 0.16], $p = 0.41$) or follow-up (–0.43 [–0.88, 0.02], $p = 0.06$) (Figs. S2e and f). The prediction intervals for TST at post-treatment and follow-up were –0.75 to 0.53 and –0.43 to –2.26 respectively. Compared with controls, CBT-I resulted in significant improvements on SE at post-treatment (–1.25 [–1.76, –0.74], $p < 0.00001$) and follow-up (–1.10 [–1.90, –0.31], $p = 0.006$) (Figs. S2g and h). The prediction intervals for SE at post-treatment and follow-up were –2.85 to 0.35 and –4.65 to 2.45 respectively.

Pain

All studies contributed to pooled analyses assessing the effects of CBT-I on pain at post-treatment ($n = 783$). The McGill Pain Questionnaire (full and short form) was the most commonly used measure [48,49]. There were large significant treatment effects on pain at post-treatment (0.20 [0.06, 0.34], $p = 0.006$) (Fig. 2c). The 95% prediction interval was 0.04–0.35. There was no significant heterogeneity between studies. Subgroup analyses demonstrated a significant treatment effect for group-based [24,26,33,36,39] CBT-I ($p = 0.01$), but not individual-based [19,20,25,34,35,37,38] CBT-I (Table 3). Post-hoc analyses revealed significant post-treatment effects in studies assessing pain intensity/severity alone (0.19 [0.04, 0.34], $p = 0.01$) [24,25,33–39], and pain interference alone (0.75 [0.14, 0.37], $p = 0.02$) [20,25]. A post-hoc analysis comprising only of hybrid CBT-I/CBT-P studies also revealed a significant treatment effect at post-treatment (0.26 [0.05, 0.47], $p = 0.02$) [20,33,36].

Six RCTs contributed 583 participants to pooled analyses assessing the effectiveness of CBT-I on pain at follow-up (range:

three to nine months) (Fig. 2d) [24,33,34,36–38]. There was no significant overall treatment effect (0.17 [–0.11, 0.45], $p = 0.24$). The 95% prediction interval was –0.50 to 0.84. Significant heterogeneity was found between studies, and a sensitivity analysis identified Smith et al. [38] as a potential source of heterogeneity. Excluding this study reduced the I^2 statistic from 60% to 44%, and the overall treatment effect remained non-significant (0.26 [–0.01, 0.53], $p = 0.06$) (Table 3). Similar to post-treatment, subgroup analyses on follow-up revealed a significant improvement for group-based [24,33,36] CBT-I ($p = 0.005$), but not individual-based [34,37,38] CBT-I (Table 3). Further, a post-hoc analysis comprising only of hybrid CBT-I/CBT-P studies did not find a treatment effect at follow-up [33,36]. In comparison to controls, the probability of having better pain after CBT-I was 58% and 57% at post-treatment and final follow-up, respectively.

Depressive symptoms, anxiety symptoms, and fatigue

The Beck Depression Inventory and Geriatric Depression Scale were the most commonly used measures of depressive symptoms [50,51]. Eight RCTs contributed the data of 383 participants to examine the effects of CBT-I on depressive symptoms at post-treatment [19,20,24–26,36,37,39]. There was a significant improvement on depressive symptoms at post-treatment (0.44 [0.09, 0.79], $p = 0.01$) (Fig. S3a). Excluding Jungquist et al. [25] led to a nonsignificant I^2 statistic of 44%, and the overall treatment effect was borderline significant at post-treatment (0.30 [–0.00, 0.60], $p = 0.05$) (Table 3). Subgroup analyses indicated a significant improvement on depressive symptoms from individual-based [19,20,25,37] CBT-I ($p = 0.01$), but not group-based [24,26,36,39] CBT-I (Table 3). Five RCTs contributed 325 participants to the pooled analysis on the effects of CBT-I at follow-up [24,26,36,37,39]. There was no significant overall treatment effect on depressive symptoms (Fig. S3b). No significant heterogeneity was found between studies. A subgroup analysis on group-based CBT-I did not reveal a significant treatment effect (Table 3) [24,26,36,39].

Four RCTs assessed anxiety symptoms at post-treatment ($n = 234$) [20,26,36,37], three of which ($n = 214$) contributed to the pooled analysis at follow-up [26,36,37]. There was no

significant overall effect of CBT-I on anxiety symptoms at post-treatment or follow-up (Figs. S3c and d). Subgroup analyses did not reveal significant treatment effects of group-based [26,36] or individual-based [20,37] CBT-I (Table 3). A subgroup analysis did not reveal a significant treatment effect of group-based CBT-I (Table 3) [26,36]. No significant heterogeneity was found between studies at either assessment point. Four RCTs contributed 168 participants to the pooled analysis assessing fatigue at post-treatment [19,20,26,36]. There was no significant improvement on fatigue at post-treatment (Fig. S3e). After removing Martinez et al. [26] and Lami et al. [36], the I^2 statistic decreased from 88% to 0%, and the overall treatment effect became significant (1.49 [0.65, 2.33], $p = 0.0005$) (Table 3). Subgroup analyses indicated a significant improvement on fatigue from individual-based [19,20] CBT-I ($p = 0.0005$), but not group-based [26,36] CBT-I (Table 3). Two RCTs evaluated fatigue at follow-up and contributed 128 participants to the pooled analysis [26,36]. There was no significant overall effect of CBT-I on fatigue at follow-up (Fig. S3f). No significant heterogeneity was found between studies.

Meta-regression

Meta-regression analyses are presented in Table S2. The Yates quality rating scale ≥ 25 significantly predicted the effectiveness of CBT-I at post-treatment, such that higher scores were associated with greater post-treatment effects on global measures of sleep. Fewer lost to follow-up significantly predicted better treatment effects at post-treatment and follow-up. Individual CBT-I predicted greater CBT-I effects at post-treatment (probability: 89% vs. 75%) and at follow-up (probability: 75% vs. 70%). The Cochrane risk assessment and a sample size > 50 did not significantly predict treatment effects at either assessment points.

Discussion

Patient-reported health outcomes such as pain, depressive symptoms, anxiety symptoms, and fatigue may contribute to sleep impairment in patients with chronic pain [9,52–55]. The current study involves a focused meta-analysis on the effectiveness of CBT-I interventions on these multiple patient-reported health outcomes in adults with chronic non-malignant pain. Across 10 RCTs [19,20,24–26,33–36,38], we found that CBT-I largely improved global measures of sleep at post-treatment and follow-up, with a probability of 81% and 71% of having better sleep after CBT-I at post-treatment and final follow-up, respectively. Diary-derived SOL, WASO, and SE also improved at post-treatment and follow-up. We observed significant treatment effects on pain at post-treatment, with a probability of 58% and 57% of having less pain after CBT-I at post-treatment and final follow-up, respectively. While depressive symptoms significantly decreased at post-treatment, no significant effects were found on anxiety symptoms and fatigue.

In comparison to controls, CBT-I caused substantial improvements in global measures of sleep at post-treatment ($SMD = 0.89$) and follow-up ($SMD = 0.56$), which were maintained when studies were removed to eliminate heterogeneity [24,25,33]. This finding is consistent with Tang et al.'s [18] subgroup analysis demonstrating significant post-treatment and follow-up effects on sleep by non-pharmacological interventions (including behavioral therapy and psychoeducation) in patients with chronic non-malignant pain. Additionally, both group-based [24,26,33,36] and individual-based [19,20,25,34,35,38] CBT-I led to meaningful improvements ($p < 0.001$) in sleep at both assessment points.

Although it is well established that individual CBT-I is highly effective for improving sleep [56], our finding in chronic pain patients with comorbid insomnia is consistent with a prior meta-

analysis that demonstrated medium to large effects of group CBT-I on sleep outcomes in adults with primary or comorbid insomnia [57]. Interestingly, a prior meta-analysis comprising adults with comorbid insomnia found that individual CBT-I produced stronger treatment effects [58], and aligns with our meta-regression results that individual CBT-I significantly predicted better sleep at post-treatment and follow-up. Future trials should evaluate the comparative effectiveness of individual-versus group-based CBT-I in patients with chronic pain.

Given the significant prevalence of comorbid insomnia in the chronic non-cancer population, healthcare providers should assess for insomnia symptoms in pain clinics. Due to the high prevalence of comorbid insomnia in the chronic pain population and shortage of qualified CBT-I therapists, a stepped care approach could effectively distribute resources to individuals based on the severity and prognosis of their insomnia [56]. The effectiveness of CBT-I in self-help (e.g., digital, books/manuals, audiovisual) and remote (e.g., telephone, telemedicine) are emerging in the literature, and meta-analyses have reported their efficacy on several patient-reported sleep outcomes, including the Insomnia Severity Index and sleep diary variables (e.g., SOL, WASO, and SE) [59–62]. To our knowledge, no such interventions have been tested for efficacy in the chronic non-cancer pain population, although we are aware of the ongoing OsteoArthritis and Therapy for Sleep trial (Washington, USA) that will be examining the effectiveness of a telephone CBT-I intervention for co-morbid insomnia in patients with osteoarthritis [63].

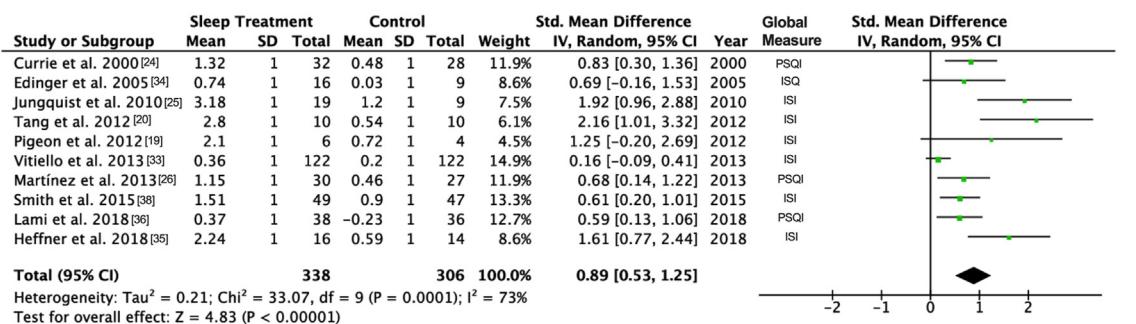
We also found a significant treatment effect on pain at post-treatment ($SMD = 0.20$), but not at follow-up. Individual studies reported mixed findings on changes in pain outcomes [19,20,24–26,31–39], although a recent meta-analysis found that nonpharmacological treatments caused near significant effects at post-treatment in adults with chronic non-malignant pain ($p = 0.07$) [18]. Our results are consistent with the speculation that sleep is a strong predictor of pain [9]. It is also well established that sleep deprivation can impact pain processing and experimentally inducing prolonged sleep restriction can affect pain habituation and sensitization, thereby increasing the susceptibility to chronic pain in otherwise healthy participants [8–13,64]. Further, sleep disturbances may worsen existing chronic pain conditions, while good sleep may predict pain resolution [9]. Given the mixed findings reported among individual trials, it is possible they were not adequately powered to detect a long-term effect on pain. Further, using a treatment protocol to address only insomnia and sleep habits may not be sufficient for targeting chronic pain [15].

It was suggested that the addition of CBT-P to a CBT-I protocol may effectively lead to improvement in pain [65]. As noted earlier, there is a reciprocal relationship between chronic pain and insomnia [8–11]. Longitudinal research has also shown that sleep disruption may be associated with development of a pain condition, lower physical functioning, and greater levels of inflammatory markers [66]. Many patients report pain as a primary reason for the onset of their insomnia [67–69]. Hence, hybrid treatments that combine CBT-I and CBT-P components may cause greater and more durable effects on both sleep and pain. Hybrid treatments may also allow clinicians with more flexibility and opportunities to personalize treatments to individual needs, and addressing multiple health concerns through one program can improve accessibility to services, target behavioral and cognitive processes that influence both insomnia and pain (e.g., catastrophizing or avoidant behaviors [70], and allows patients to manage comorbidities that contribute to their pain and disability [71].

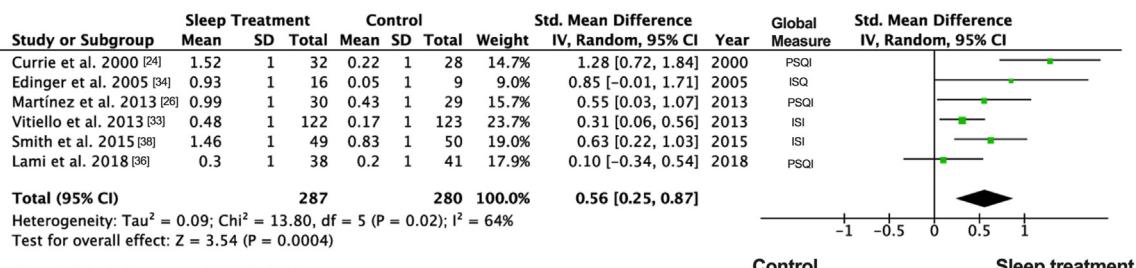
Interestingly, our post-hoc analysis comprising only of hybrid CBT-I and CBT-P trials showed similar effects on pain to traditional CBT-I at post-treatment only ($SMD = 0.26$) [20,33,36]. However, a secondary analysis of Vitiello et al.'s [33] trial found that

Sleep

a. Baseline vs. Post-treatment between sleep treatment and control group

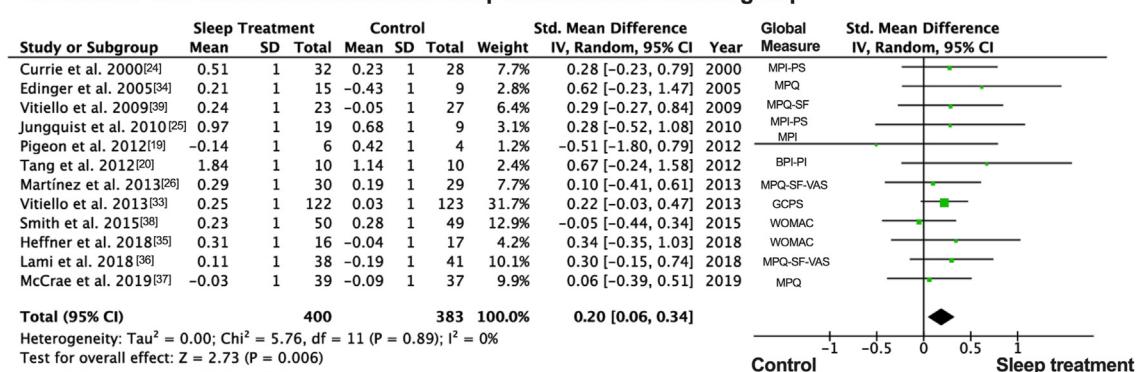


b. Baseline vs. Follow-up between sleep treatment and control group



Pain

c. Baseline vs. Post-treatment between sleep treatment and control group



d. Baseline vs. Follow-up between sleep treatment and control group

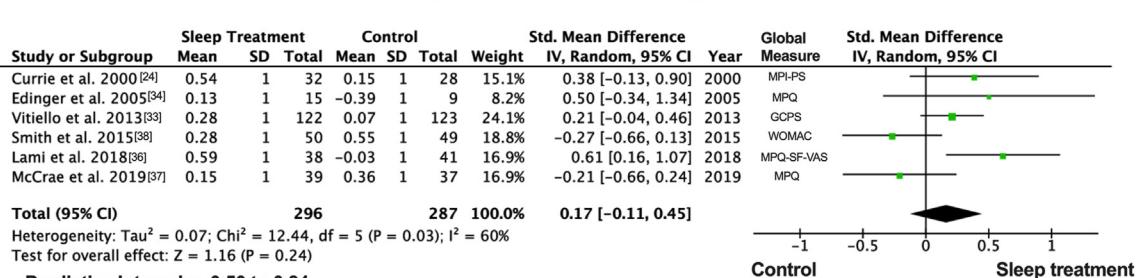


Fig. 2. Forest plots of treatment effects on global measures of a) sleep at post-treatment, b) sleep at follow-up, c) pain at post-treatment, and d) pain at follow-up. Abbreviations: BPI-PI, Brief Pain Inventory Pain Interference subscale; GCPS, Graded Chronic Pain Scale; ISI, Insomnia Severity Index; ISQ, Insomnia Symptom Questionnaire; MPI,

Table 3

Sensitivity and subgroup analyses.

| Outcome | | Sensitivity Analysis | | | Subgroup Analysis | |
|---------------------|-----------|--|---------------------------------|--------------------|--------------------------------|--------------------------------|
| | | Excluded Study | Modified Treatment Effect (SMD) | Modified I^2 (%) | Group-based CBT-I (SMD) | Individual-based CBT-I (SMD) |
| Sleep [§] | Post-Rx | Junquist et al. [25] & Vitiellio et al. [33] | 0.86 [0.57, 1.15]**** | 37% | 0.51 [0.17, 0.86]* | 1.28 [0.70, 1.86]*** |
| | Follow-Up | Currie et al. [24] | 0.40 [0.19, 0.61]** | 18% | 0.52 [0.09, 0.95] [†] | 0.67 [0.30, 1.03]** |
| Pain | Post-Rx | N/A | N/A | N/A | 0.23 [0.05, 0.41] [†] | 0.14 [-0.09, 0.38] |
| | Follow-Up | Smith et al. [38] | 0.26 [-0.01, 0.53] | 44% | 0.34 [0.10, 0.57]* | -0.13 [-0.47, 0.21] |
| Depressive Symptoms | Post-Rx | Jungquist et al. [25] | 0.30 [-0.00, 0.60] [‡] | 44% | 0.16 [-0.09, 0.41] | 1.09 [0.23, 1.94] [†] |
| | Follow-Up | N/A | N/A | N/A | 0.01 [-0.24, 0.26] | N/A |
| Anxiety Symptoms | Post-Rx | N/A | N/A | N/A | 0.04 [-0.29, 0.37] | 0.37 [-0.04, 0.77] |
| | Follow-Up | N/A | N/A | N/A | -0.22 [-0.56, 0.11] | N/A |
| Fatigue | Post-Rx | Martinez et al. [26] & Lami [36] | 1.49 [0.65, 2.33]** | 0% | -0.28 [-1.33, 0.76] | 1.49 [0.65, 2.33]** |
| | Follow-Up | N/A | N/A | N/A | N/A | N/A |

[§] Global measures of sleep included the Insomnia Severity Index, Pittsburgh Sleep Quality Index, and Insomnia Symptom Questionnaire.****p < 0.00001, ***p < 0.0001, **p < 0.001, *p < 0.01, [†]p < 0.05, [‡]p = 0.05.

participants in the hybrid treatment with more severe baseline insomnia and pain had significant improvements in pain when compared with CBT-P alone at 18 mo [31]. This finding suggests that the addition of insomnia treatment to CBT-P may have an added benefit on pain in patients with more severe symptoms. Given the few number of studies, larger RCTs incorporating long follow-up periods are required to determine whether utilizing a hybrid protocol could effectively target both insomnia and pain.

Few studies have investigated the use of other non-pharmacological treatments for comorbid insomnia with chronic non-cancer pain. These interventions, such as Tai Ji Quan (a martial art) [72], hydrotherapy [73], massage-myofacial therapy [74], and manual therapy [75], have shown to improve pain and sleep quality in patients with fibromyalgia. Additionally, the effects of pain medications on sleep have been explored. Specifically, administering pregabalin led to significant improvements in pain and sleep compared with placebo in patients with fibromyalgia [76–80]. By improving pain levels, opioid therapy may also improve sleep. However, a recent systematic review and meta-analysis reported inconsistent and small effects of opioid therapy on subjective sleep quality, with the presence of excessive daytime sleepiness in patients with chronic non-cancer pain [81]. Given the small number and methodological quality of trials, further research is needed to evaluate the effects of alternative nonpharmacological and pharmacological treatments on both objective and subjective measures of sleep and pain.

We found a significant therapeutic effect on depressive symptoms at post-treatment (SMD = 0.44) [19,20,24–26,33,37,39]. Contrary to our findings, a previous meta-analysis did not report significant treatment effects on depressive symptoms in patients with chronic non-malignant pain, possibly due to insufficient power with only five studies included in their analysis [18]. In spite of individual trials reporting significant effects on anxiety symptoms [26] and fatigue [19,20,26,36], we did not observe treatment effects on anxiety symptoms or fatigue, although the impact on fatigue became significant when potential sources of heterogeneity were removed [26,36] and individual CBT-I [19,20] was analyzed separately. Given the limited number of studies examining the effectiveness of CBT-I on anxiety symptoms and fatigue in patients with chronic pain, further research in this area is warranted.

The overall quality of trials generally ranged from moderate to high with the exception of one study [35], although several trials reported small sample sizes and unclear allocation concealment procedures. Having an overall score of ≥ 25 on the Yates quality rating scale was associated with greater treatment effects on sleep at post-treatment. Most studies provided a detailed description of the CBT-I protocol and used robust inclusion criteria for chronic pain and insomnia diagnoses. Given that none of the studies reported blinding of both participants and researchers (outcome assessors and therapists), performance bias and expectation bias could have impacted findings. Therefore, the positive effect of CBT-I on insomnia may partly be a result of expectations that the treatment will improve their sleep over time. Additionally, most trials (92%) included in the meta-analysis reported a final follow-up at less than one year after baseline. Hence, whether CBT-I can elicit durable effects on sleep and pain remains uncertain.

The current study has several limitations. Publication bias was detected for post-treatment results on sleep, indicating that published results may be different from unpublished studies. Our search was also restricted to English-language and published studies only. Of the 12 RCTs included in the analyses, only two studies used double-blinded procedures, whereby participants were not told which of the interventions were active and blinded to the nature of the control interventions, assessors were blinded to group assignment, and interventionists were blinded to group assignment and the purpose of the control interventions [33,38]. Potential sources of heterogeneity across studies could be differences in chronic pain types or treatment protocols (e.g., components, delivery, and format). To address this limitation, guidelines outlining standardized methodological and CBT-I treatment procedures should be created and followed by future RCTs. Although sensitivity analyses were conducted to reduce heterogeneity, moderate heterogeneity was found for pain at follow-up (44%) and depressive symptoms at post-treatment (44%). Additionally, post-treatment effects on fatigue became significant upon removal of these studies, suggesting a possible introduction of bias and underpowered results. Indeed, the current meta-analysis comprised 12 RCTs only, nine of which had less than 100 participants [19,20,24–26,34,35,38,39]. Some of our analyses had less than

Multidimensional Pain Inventory; MPI-PS, Multidimensional Pain Inventory Pain Severity subscale; MPQ, McGill Pain Questionnaire; MPQ-SF-VAS; McGill Pain Questionnaire Short Form Visual Analogue Scale; MPQ-SF; McGill Pain Questionnaire Short Form; PSQI; Pittsburgh Sleep Quality Index; WOMAC Index; Western Ontario and McMaster Universities Osteoarthritis Index.

three studies per cell, such as the subgroup analyses on group-versus individual-based CBT-I, our post-hoc analysis of hybrid trials at follow-up, and effects on fatigue at follow-up. Including higher powered studies in future meta-analyses could reduce the likelihood of introducing bias when conducting sensitivity and subgroup analyses. Due to concerns of major heterogeneity in the technology and methodological procedures used to collect objective data (e.g., actigraphy), the current meta-analysis included subjectively reported outcomes only. However, patient-reported assessments are subject to recall bias and may not accurately reflect the patient's health. Other non-sleep-related outcomes, such as quality of life and physical functioning, should also be included in future trials to determine whether CBT-I can affect overall health.

Conclusion

This systematic review and meta-analysis found that CBT-I is effective for eliciting short- and long-term improvements in sleep, along with short-term improvements in pain and depressive symptoms in patients with comorbid insomnia and chronic non-malignant pain. Across 10 RCTs [19,20,24–26,33–36,38], we found a probability of 81% and 71% for having better sleep after CBT-I at post-treatment and final follow-up, respectively. The probability of having less pain after CBT-I at post-treatment and final follow-up was 58% and 57%, respectively. No treatment effects were found on anxiety symptoms or fatigue. Since insomnia is often reported as a comorbidity in patients with chronic pain [12], CBT-I could be used as the first-line of treatment in this population with fewer side effects than pharmacological therapies [82,83]. Additionally, future trials should incorporate larger sample sizes, components from CBT-P to actively target pain, self-help formats, and longer follow-up periods to determine whether CBT-I can elicit sustainable effects in patients with chronic pain.

Practice points

1. There is a high prevalence of insomnia among the chronic non-cancer pain population, and sleep is a stronger predictor of subsequent pain than vice versa.
2. CBT-I elicits immediate and long-term improvements in patient-reported sleep.
3. Short-term treatment effects are evident in other health outcome measures, including pain and depressive symptoms.

Research agenda

1. Larger and adequately powered randomized controlled trials are needed to examine the effectiveness of CBT-I on comorbid insomnia and chronic pain.
2. Longer follow-up periods are recommended to further investigate the long-term effects of CBT-I.
3. Further research is needed to determine whether hybrid interventions combining CBT-I and CBT-P can better address both insomnia and chronic pain symptoms.
4. Future studies should include objective measures of sleep and pain, along with non-sleep-related outcomes such as quality of life and physical functioning.

Conflicts of interest

MN reports the Academic Medical Organization Southwestern Ontario (AMOSO) opportunity fund, Alternative Fund Plan (AFP) Innovation fund and Lawson Internal Research fund. CAE reports National Institute for Health Research: Oxford Biomedical Research Centre; shareholder in Big Health. CMM reported research grant from CIHR, Fonds de recherche du Québec, Idorsia and Canopy Health, Consultant/Advisory Board for Eisai, Merck, Pear Therapeutics, Sunovion, Weight Watchers. FC reports research support from the Ontario Ministry of Health and Long-Term Care, University Health Network Foundation, Up-to-date royalties, consultant to Takeda Pharma and Masimo, STOP-Bang proprietary to University Health Network. JS, CP, PWHP, and ME did not report conflicts of interest.

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Appendix A. Supplementary data

Supplementary data to this article can be found online at <https://doi.org/10.1016/j.smrv.2021.101460>.

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