INTRODUCTORY APPLIED MACHINE LEARNING

Yan-Fu Kuo

Dept. of Bio-industrial Mechatronics Engineering National Taiwan University

Today:

Linear regression

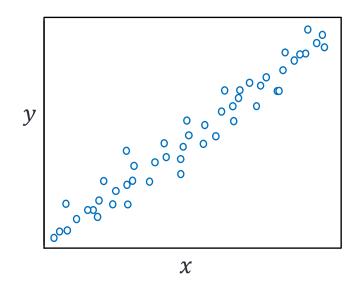
Outline

- Goal of the lecture
- Data dependency
- Simple linear regression
- Least squares
- Coefficient of determination
- Residual analysis
- Multiple regression

Goals

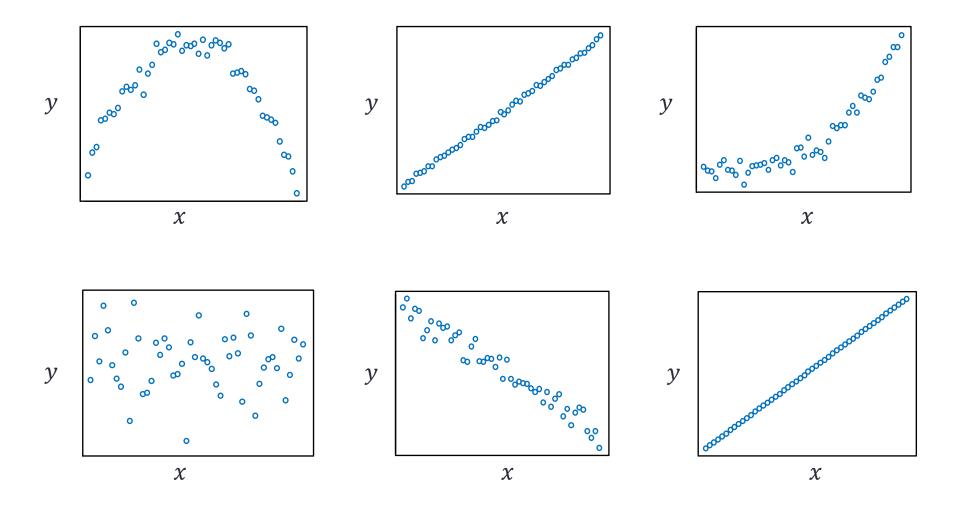
- After this, you should be able to:
 - Calculate and interpret the simple correlation between two variables
 - Calculate and interpret the simple linear regression equation for a set of data
 - Understand the assumptions behind regression analysis
 - Calculate the confidence interval for regression slope
 - Recognize some potential problems if regression analysis is used incorrectly

Scatter Plot



- The best way to view the relationship between two variables
- In some situations, we want to measure the dependency of one variable against another
- In other situations, we want to assess how the observed property matches the predicted property
- In all cases we will measure multiple samples or work with a population of subjects

Example Scatter Plots



Correlation Analysis

- Linear relationship between two variables x and y
- Correlation coefficient:

$$-1 \le r = \frac{\sum_{i=1}^{N} (x_i - \bar{x})(y_i - \bar{y})}{\sqrt{\sum_{i=1}^{N} (x_i - \bar{x})^2 \sum_{i=1}^{N} (y_i - \bar{y})^2}} \le 1$$

Low Correlation Medium Correlation High Correlation

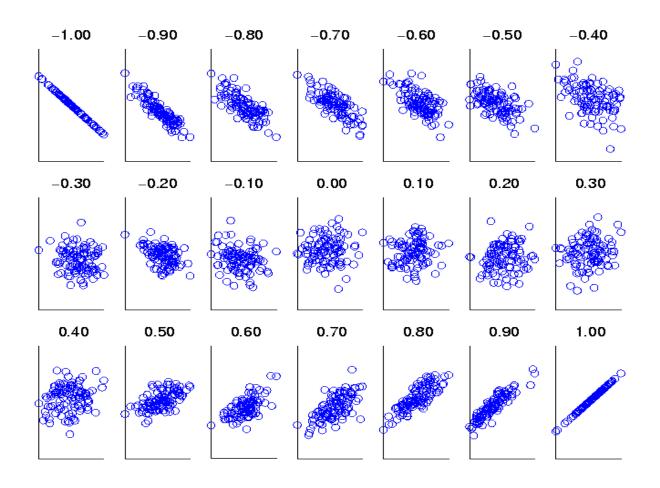
y

y

x

High Correlation

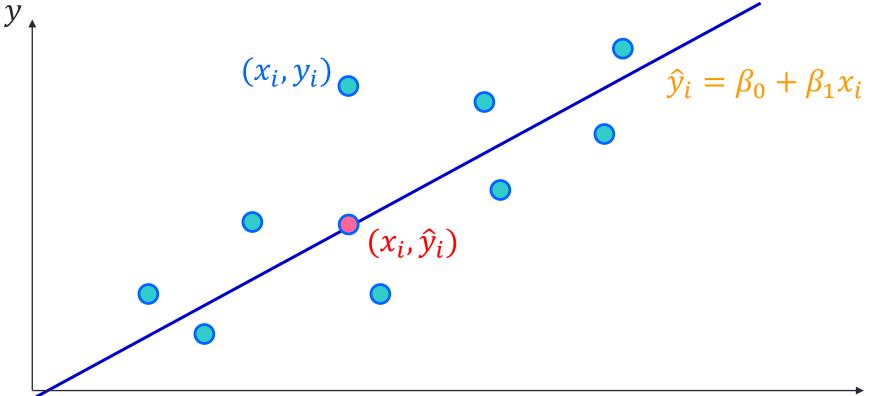
Illustration of Correlation Coefficient



Scatter plots showing the correlation coefficients ranged from –1 to 1

Simple Linear Regression

- A bunch of data points (x_i, y_i) are collected
- Assume x and y are linearly correlated



Simple Linear Regression Analysis

- Regression analysis is used to:
 - Predict the value of a response variable y based on the value of at least one explanatory variable x
 - Explain the impact of changes in an explanatory variable *x* on the response variable *y*

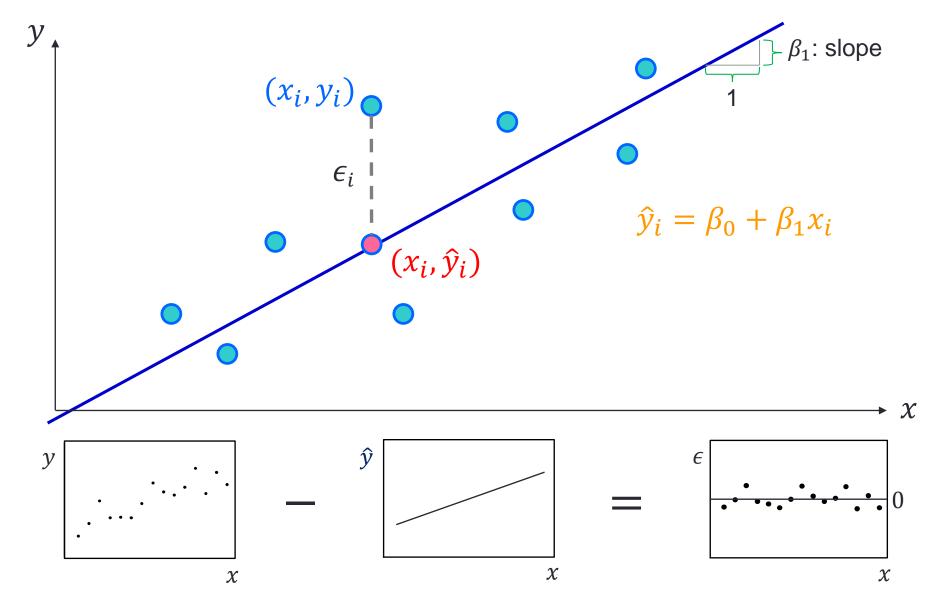
Nominal relationship between x and y:

$$\hat{y} = \beta_0 + \beta_1 x$$

Actual relationship between x and y:

$$y = \beta_0 + \beta_1 x + \epsilon$$

The Error Term ϵ



Assumption – i.i.d.

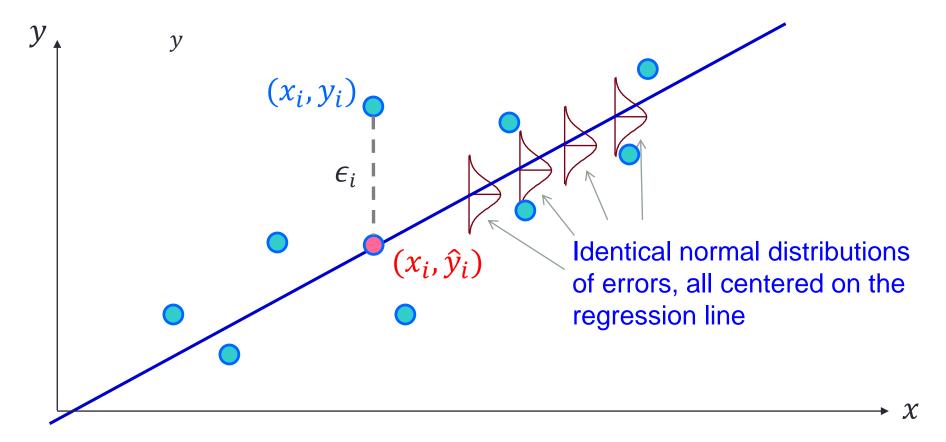
Relationship between x and y:

$$y = \beta_0 + \beta_1 x + \epsilon$$

where ϵ is assumed to be independently and identically distributed (i.i.d.)

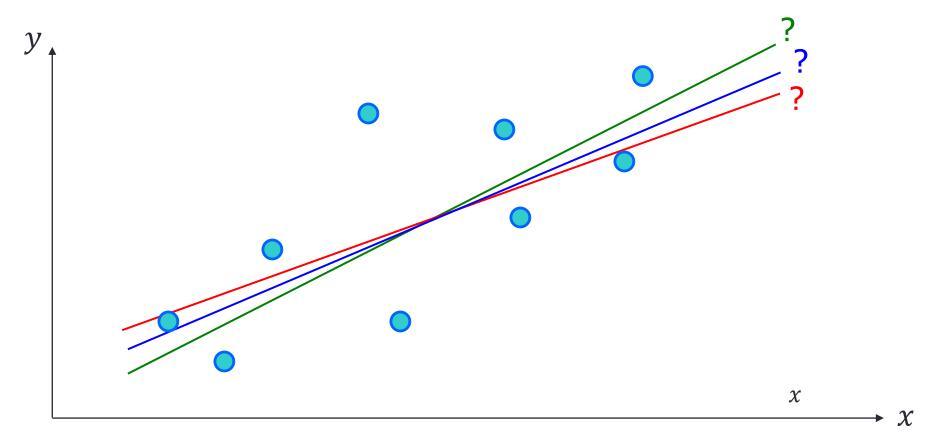
- More specifically,
 - 1. ϵ and x are independent, i.e., $P_{x,\epsilon}(x,\epsilon) = P_x(x)P_{\epsilon}(\epsilon)$
 - 2. The probability distribution of the errors ϵ is normal

Observations



- The errors ϵ are uncorrelated in successive observations
- The errors ϵ are normally distributed, i.e., $\epsilon \sim N(0, \sigma^2)$

Which Line Best Fit to the Data?



Through optimization!

Loss Function

- How does one mathematically define "best"?
- One has to define the "error" (or "loss") first
- The loss may be, for example, the squared loss

$$loss(y, \hat{y}) = \sum_{i} (y_i - \hat{y}_i)^2 = \sum_{i} \epsilon_i^2$$

The goal is to minimize the error/loss on data points

Least-squares Method

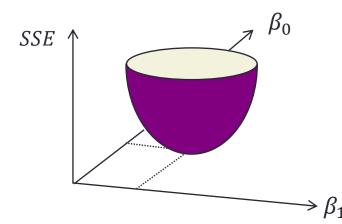
• Estimation of a simple linear regression relationship involves finding estimated values of the intercept β_0 and slope β_1 of the model

$$y = \beta_0 + \beta_1 x + \epsilon \leftrightarrow \hat{y} = \beta_0 + \beta_1 x$$

• Define sum of squared errors (SSE):

$$SSE = \sum_{i} \epsilon_{i}^{2} = \sum_{i} (y_{i} - \hat{y}_{i})^{2} = \sum_{i} (y_{i} - \beta_{0} - \beta_{1}x_{i})^{2}$$

• Identify the β_0 and β_1 that minimize the SSE



Solving Least-squares Regression

 SSE is minimized when its gradient with respect to each parameter is equal to zero:

$$SSE = \sum_{i} \epsilon_{i}^{2} = \sum_{i} (y_{i} - \hat{y}_{i})^{2} = \sum_{i} (y_{i} - \beta_{0} - \beta_{1}x_{i})^{2}$$

$$\frac{\partial SSE}{\partial \beta_0} = -2\sum_{i} (y_i - \beta_0 - \beta_1 x_i) = 0$$

$$\frac{\partial SSE}{\partial \beta_1} = -2\sum_{i} (y_i - \beta_0 - \beta_1 x_i) x_i = 0$$

Solving Least-squares Regression

Suppose there exists N data points:

$$\sum_{i=1}^{N} y_i = \beta_0 \cdot N + \beta_1 \sum_{i=1}^{N} x_i$$

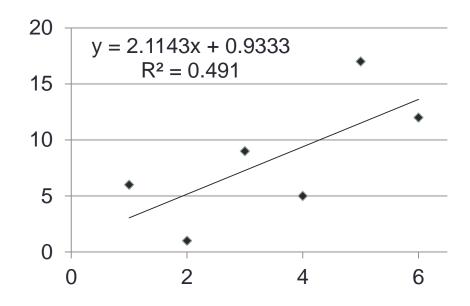
$$\sum_{i=1}^{N} y_i x_i = \beta_0 \sum_{i=1}^{N} x_i + \beta_1 \sum_{i=1}^{N} x_i^2$$

$$\Rightarrow \begin{bmatrix} \sum y_i \\ \sum y_i x_i \end{bmatrix} = \begin{bmatrix} N & \sum x_i \\ \sum x_i & \sum x_i^2 \end{bmatrix} \begin{bmatrix} \beta_0 \\ \beta_1 \end{bmatrix}$$

Simple Regression Example

Suppose we have:

χ	у
1	6
2	1
3	9
4	5
5	17
6	12



Least Squares Regression Properties

- The sum of the squared residuals is a minimum, i.e. minimal $\sum (y_i \hat{y}_i)^2$
- The sum of the residuals from the least squares regression line is zero, i.e. $\sum (y_i \hat{y}_i) = \sum \epsilon_i = 0$
- The simple regression line always passes through the mean of the response variable \bar{y} and the mean of the explanatory variable \bar{x}
- The least squares coefficients are <u>unbiased</u> estimates of β_0 and β_1

Assessing the Model

- The least squares method will always produce a straight line
- Determining regression model coefficients is easy, but...
- 1. How does one access the model and know how well it fits the data?
 - One common method check how much <u>variation</u> is explained by the model
- 2. Particularly, what if the relationship between the variables is NOT linear?

Explained and Unexplained Variation

Total variation is made up of two parts:

$$SST = SSE + SSR$$

Sum of squares total

Sum of squares error

Sum of squares regression

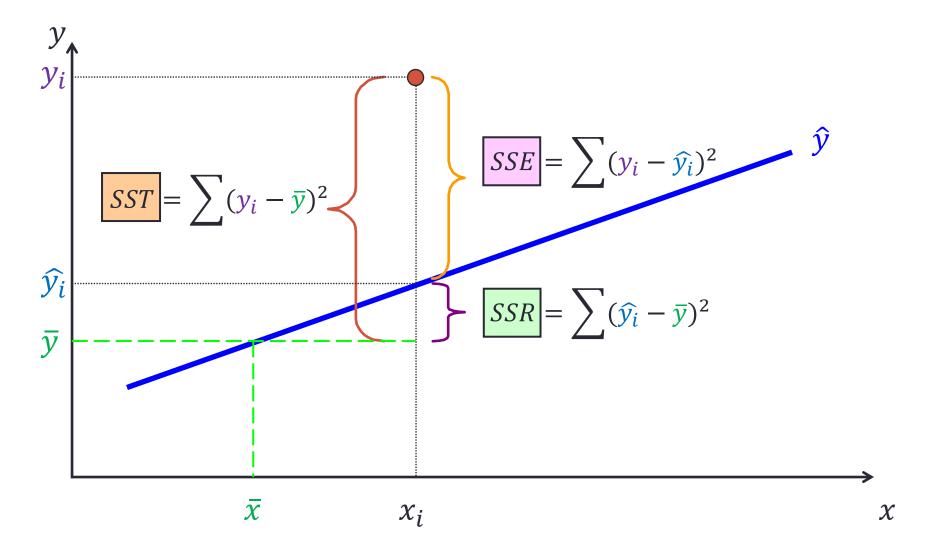
$$SST = \sum (y_i - \bar{y})^2$$

$$SSE = \sum (y_i - \widehat{y}_i)^2$$

$$SSR = \sum (\widehat{y}_i - \overline{y})^2$$

- SST measures the variation of the y_i values around their mean \bar{y}
- SSE represents the variation attributable to factors other than the relationship between x and y
- SSR explains variation attributable to the relationship between x and y

Explained and Unexplained Variation (Cont'd)



Proof of SST = SSE + SSR

• Starting from $(y_i - \overline{y}) = [(y_i - \hat{y}_i) + (\hat{y}_i - \overline{y})]$:

$$\sum_{i} (y_{i} - \bar{y})^{2} = \sum_{i} [(y_{i} - \hat{y}_{i}) + (\hat{y}_{i} - \bar{y})]^{2}$$

$$(y_{i} - \hat{y}_{i})^{2} + 2\sum_{i} (y_{i} - \hat{y}_{i})(\hat{y}_{i} - \bar{y}) + \sum_{i} (\hat{y}_{i} - \bar{y}_{i})(\hat{y}_{i} - \bar{y}_{i}) + \sum_{i} (\hat{y}_{i} - \bar{y}_{i})(\hat{y}_{i} - \bar{y}_{i})(\hat{y}_{i} - \bar{y}_{i}) + \sum_{i} (\hat{y}_{i} - \bar{y}_{i})(\hat{y}_{i} - \bar{y}_{i})($$

$$= \sum_{i} (y_i - \hat{y}_i)^2 + 2 \sum_{i} (y_i - \hat{y}_i)(\hat{y}_i - \bar{y}) + \sum_{i} (\hat{y}_i - \bar{y})^2$$

• The middle expression:

$$2\sum_{i} (y_{i} - \hat{y}_{i})(\hat{y}_{i} - \bar{y}) = 2\sum_{i} [\hat{y}_{i}(y_{i} - \hat{y}_{i}) - \bar{y}(y_{i} - \hat{y}_{i})]$$

$$= 2\sum_{i} \hat{y}_{i}\epsilon_{i} - 2\sum_{i} \bar{y}\epsilon_{i} = 2\sum_{i} (\beta_{0} + \beta_{1}x_{i})\epsilon_{i} - 2\bar{y}\sum_{i} \epsilon_{i}$$

$$= 2\beta_{0}\sum_{i} \epsilon_{i} + 2\beta_{1}\sum_{i} x_{i}\epsilon_{i} - 2\bar{y}\sum_{i} \epsilon_{i} = 0$$

Coefficient of Determination

- The coefficient of determination R^2 is a measure of how well the regression line fits the data
- The coefficient of determination is the portion of the total variation in the response variable that is explained by variation in the explanatory variable:

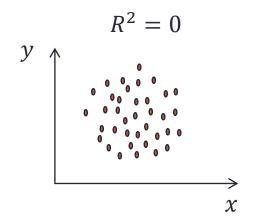
$$R^2 = \frac{SSR}{SST} = 1 - \frac{SSE}{SST} = \frac{SS \text{ explained by regression}}{\text{total SS}}$$

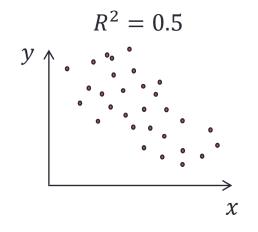
where $0 \le R^2 \le 1$

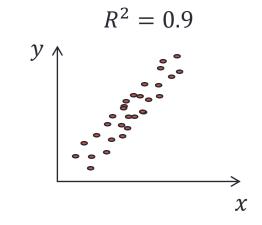
• In the simple regression, the coefficient of determination is equal to the square of correlation coefficients, i.e., $R^2=r^2$

Example Coefficient of Determination

Illustration of R²

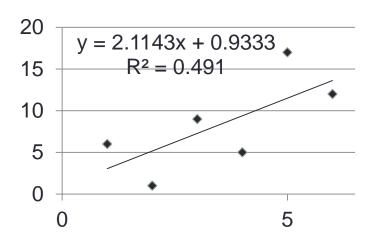




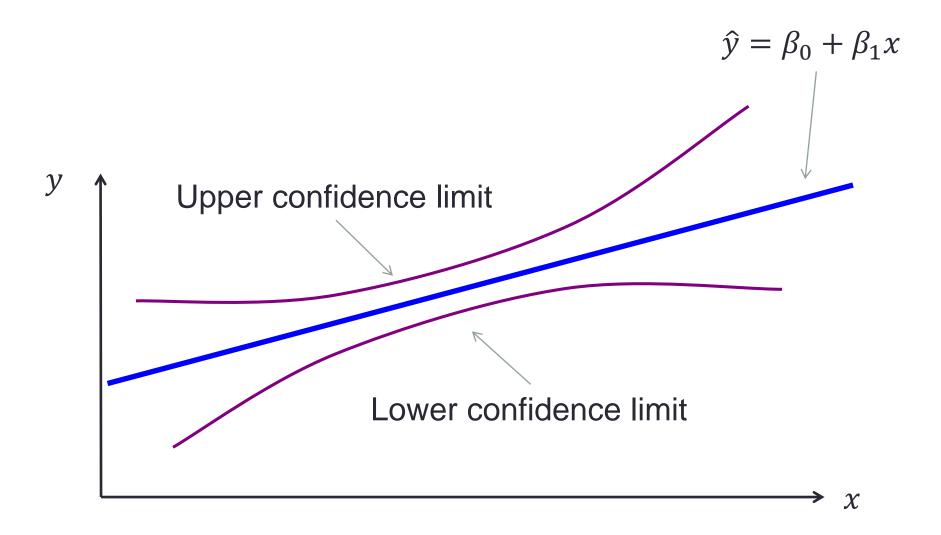


Examples of coefficient of determination:

$\boldsymbol{\chi}$	y
1	6
2	1
3	9
4	5
5	17
6	12



Interval Estimates



Standard Error and Confidence Interval of β_1

Standard error of the regression line slope is defined as:

$$s_{\beta_1} = \frac{s_e}{\sqrt{\sum x_i^2 - \frac{(\sum x_i)^2}{N}}}$$

• The $(1 - \alpha)$ % confidence interval of the regression line slope is defined as:

$$s_{\beta_1} \cdot t_{\underline{\alpha}, N-2} \le \beta_1 \le s_{\beta_1} \cdot t_{\underline{\alpha}, N-2}$$

where $t_{\frac{\alpha}{2},N-2}$ is the $\frac{\alpha}{2}$ th percentile t-distribution for N-2 degrees of freedom

Standard Deviation of the Residuals

- Sample statistics are point estimates for the population parameters, which is unknown
- The standard deviation of the residuals s_e , for all points in the population, is estimated by the standard deviation of the residuals:

$$s_e = \sqrt{\frac{\sum residual^2}{n-2}} = \sqrt{\frac{\sum y_i^2 - \beta_0 \sum y_i - \beta_1 \sum x_i y_i}{n-2}},$$

where n-2 is the degree of freedom

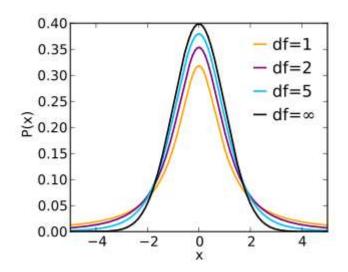
Student's t-distribution

- A continuous probability distribution for estimating the mean of a normally distributed population in situations where the <u>sample size is small</u> and <u>population standard</u> <u>deviation</u> is unknown
- Published in 1908 by William Sealy Gosset using the pseudonym "student"

 The shape of the t-distribution is similar to that of the normal distribution

Student's t-distribution (Cont'd)

- There are many different t-distributions, one for each degree of freedom
- For small degrees of freedom, the t-distribution is very dispersed
- The limiting distribution for the t distribution is the normal distribution



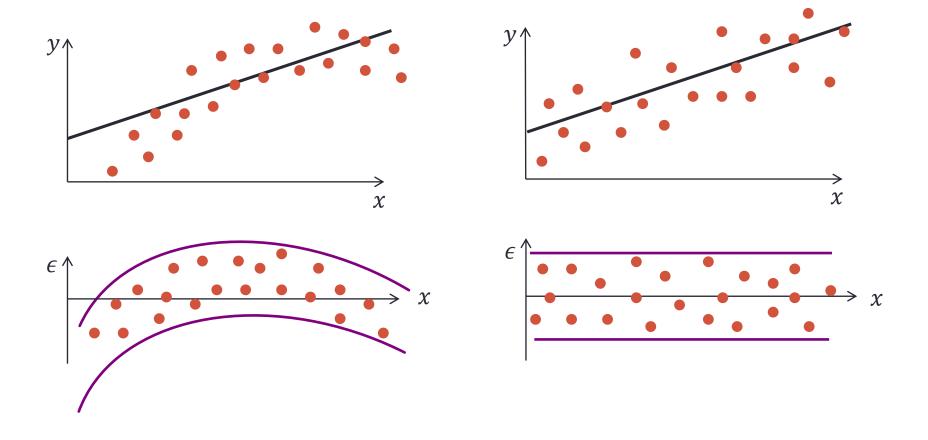
Linear Regression Assumptions

- 1. Error values ϵ are independent to the explanatory variable x
- 2. The probability distribution of the errors ϵ is normal
- 3. The probability distribution of the errors ϵ has constant variance
- 4. The underlying relationship between the explanatory variable x and the response variable y is linear

Residual Analysis

- Perform <u>residual analysis</u> to check any violation of the assumption
- The residual is the difference between its observed and predicted value, i.e. $\epsilon_i = y_i \hat{y}_i$
- Check the following assumptions:
 - Linearity
 - 2. Homoscedasticity (constant variance)
 - Normal distribution
 - 4. Independence

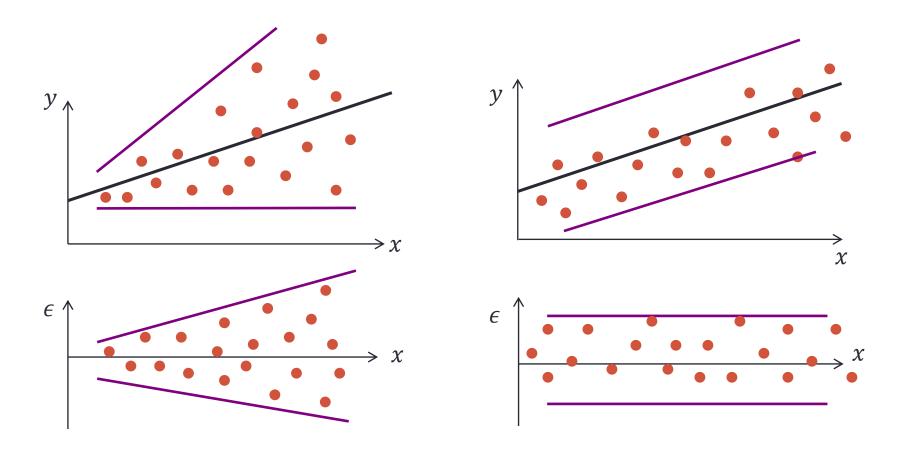
Residual Analysis for Linearity



Remedy for Violating Linearity

- 1. Piecewise linear regression:
 - Break down (x_i, y_i) into j sets, i.e., $(x_i^{(j)}, y_i^{(j)})$ where $\{x_i\} = \sum_j \{x_i^{(j)}\}$ and $\{y_i\} = \sum_j \{y_i^{(j)}\}$
 - Perform regression for each set $y_i^{(j)} = \beta_0 + \beta_1 x_i^{(j)}$
- 2. Variable transformation
 - Polynomial: $y = \beta_0 + \beta_1 x + \beta_2 x^2 + \cdots$
 - Logarithm: $y = \beta_0 + \beta_1 \log(x)$ or $\log(y) = \beta_0 + \beta_1 x$
 - Exponential: $y = \beta_0 + \beta_1 e^x$
 - Inverse: plus $y = \beta_0 + \beta_1 \frac{1}{x}$

Residual Analysis for Homoscedasticity

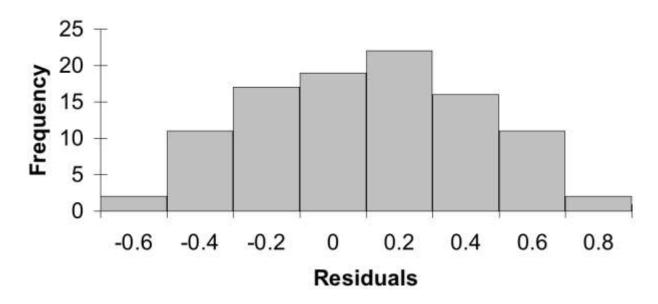


Remedy for Violating Homoscedasticity

- Divide the entire equation by x, e.g., $y = \beta_0 + \beta_1 x$ will become $\frac{y}{x} = \frac{\beta_0}{x} + \beta_1$
- Notice that for large values of x the new error (ϵ/x) will be smaller

Normality

Residual histogram

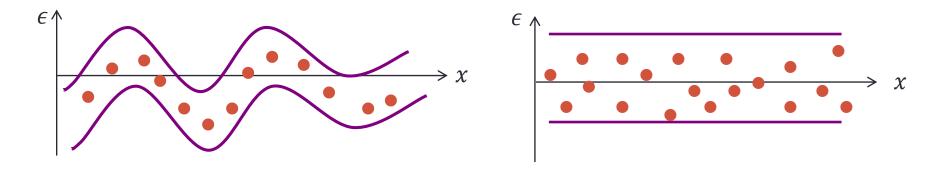


 Traditional way to test for normality – looking for a bell shaped histogram with the mean close to zero

Remedy for Violating Normality

- Transform the response variable to make the distribution of the random errors approximately normal
- Three transformations that are often effective for making the distribution of the random errors approximately normal:
 - \sqrt{y}
 - ln(y)
 - $\frac{1}{y}$

Residual Analysis for Independence



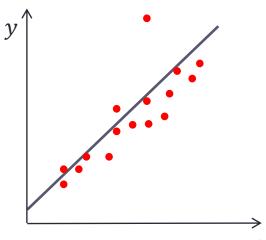
- A mathematical representation of the degree of periodical similarity between variable x over successive intervals
- This is called autocorrelation
- If a pattern emerges, it is likely that the independence requirement is violated

Remedy for Violating Independence

- For serial (temporal) correlation, include new variables in the equation, e.g., the value of y at moment t-1 as an independent variable
- For spatial correlation, model the relationships by introducing an weighting matrix

Outlier

- An outlier is a data point that is unusually small or large
- Possible reasons for the existence of outliers include:
 - There was an error in recording the value
 - The point should not have been included in the sample
- Outliers can be easily identified from a scatter plot
- Outliers need to be dealt with since they can easily influence the regression model



Procedure for Regression Analysis

- Gather data for the two variables in the model
- 2. Draw the scatter diagram to determine whether a linear model appears to be appropriate
- 3. Determine the regression equation
- Assess the model's fit
- Calculate the residuals and check the required conditions

Multiple Regression

- What if we have multiple explanatory variables $x = [x_1, x_2, ..., x_M]$ and one response variable y?
- The multiple regression model:

$$y = f(x) = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_M x_M$$

- For example: polynomial model
- Denote the i^{th} observation as $\mathbf{x}_i = [x_{i1}, ..., x_{iM}]^T$

Parameter

• Observation
$$\begin{bmatrix} 1 & x_{11} & \cdots & x_{1M} \\ \vdots & \vdots & \ddots & \vdots \\ 1 & x_{N1} & \cdots & x_{NM} \end{bmatrix} = X \implies \begin{bmatrix} y_1 \\ \vdots \\ y_N \end{bmatrix} = y$$

Formulating Multiple Regression

- Let $\boldsymbol{\beta} = [\beta_0 \quad \cdots \quad \beta_M]^{\mathrm{T}}$
- Data: $y_i = \begin{bmatrix} 1 & x_i \end{bmatrix}^T \boldsymbol{\beta} + \epsilon, \ \epsilon \sim N(0, \sigma^2)$
- Matrix form: $y = X\beta + \epsilon$

$$\begin{bmatrix} y_1 \\ \vdots \\ y_N \end{bmatrix} = \begin{bmatrix} 1 & x_{11} & \cdots & x_{1M} \\ \vdots & \vdots & \ddots & \vdots \\ 1 & x_{N1} & \cdots & x_{NM} \end{bmatrix} \begin{bmatrix} \beta_0 \\ \vdots \\ \beta_M \end{bmatrix}$$

Problem statement:

Given the vector \mathbf{y} and matrix \mathbf{X} above, find the coefficient $\boldsymbol{\beta}$ of the regression model $\hat{y} = \begin{bmatrix} 1 & \mathbf{x} \end{bmatrix}^T \boldsymbol{\beta}$ that most accurately predicts \mathbf{y}

Solving Multiple Regression

Sum of squares error:

$$SSE = \sum_{i=1}^{N} \epsilon_i^2 = \epsilon^T \epsilon = (y - X\beta)^T (y - X\beta)$$

$$= (y^T - \beta^T X^T) (y - X\beta)$$

$$= y^T y - y^T X\beta - \beta^T X^T y + \beta^T X^T X\beta$$

$$= y^T y - 2\beta^T X^T y + \beta^T X^T X\beta$$

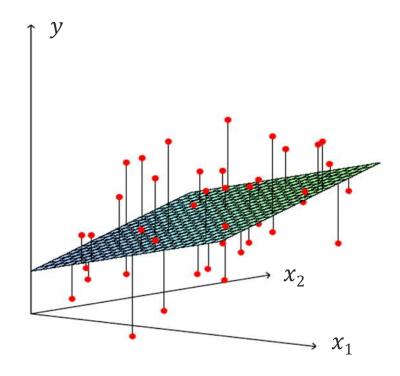
$$\frac{\partial SSE}{\partial \beta} = -2X^T y + 2X^T X\beta = 0$$

Multiple Regression

- Model coefficient $\beta = (X^TX)^{-1}X^Ty$
- Recall regression model $\hat{y} = \begin{bmatrix} 1 & x \end{bmatrix}^T \beta$
- Geometric explanation
- What about multivariate regression?

$$Y = XB$$

where **B** is the coefficient matrix



Summary

- Scatter plot is a useful diagnostic tool for determining association between variables
- Coefficient of determination provides a measure of how well future outcomes are likely to be predicted by the model
- There is an confidence interval for every statistic estimation
- The less the data samples, the smaller the degree of freedom, and the larger the confidence interval
- Always check the residual after performing regression analysis

References

- T. Hastie, R. Tibshirani, J. Friedman, The Elements of Statistical Learning: Data Mining, Inference, and Prediction
- M. Dodge and C. Stinson, Microsoft® Office Excel® 2007 Inside Out