Time Series Analysis - Introduction

ECON 722

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Introduction

- History popular in early 90 s, making comeback now.
- What is a Time Series? Y_t
 - The main difference between time series econometrics and cross-section is in dependence structure. Cross section econometrics mainly deals with i.i.d. observations, Y_i, while in time series each new arriving observation is stochastically depending on the previously observed, Y_t.
- The dependence is our best friend and a great enemy.
 - On one side, the dependence screw up your inferences (CLT you learned was for i.i.d. data). On the other side, the dependence allow us to do more by exploiting it. For example, we can make forecasts (which are almost non-sense in cross-section).

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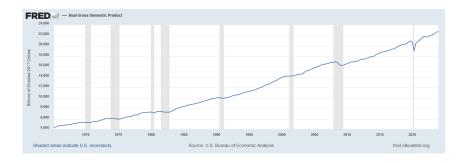
Introduction

- We are going to ask questions like:
 - What are the characteristics of time series data?
 - Which type of questions can we ask/answer?
 - Which type of questions we cannot answer? Or can we?
- We can roughly divide time series into macro and finance related stuff.
 - Macro Time series mostly focuses on means. Often limited by small number of observations available over long horizon (e.g. 20 years monthly is T=300).
 - Financial data usually high-frequency over short period of time. This
 allows us to model volatility and higher moments.

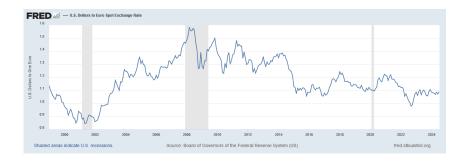
Introduction

- Which type of questions can we ask/answer?
 - In the first part of the class we will be using Regression Models for *Forecasting*.
 - Forecasting and estimation of causal effects are quite different objectives (more on this later).
 - In the second half of the class we will discuss if and when we can talk about "causality".
- Forecasting:
 - It is very important in macro and finance.
 - Omitted variable bias isn't a problem.
 - External validity is paramount: the model estimated using historical data must hold into the (near) future
 - Intepretation of the coefficients is not the goal.

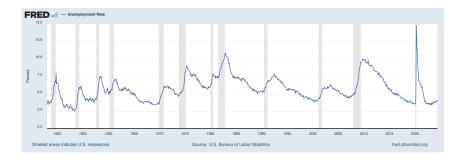
Real GDP, 1965:1-2024:1, Quarterly (GDP is only quarterly), Seasonally adjusted, billions of chained 2017 dollars. From FRED.



US Dollars to Euro Spot Exchange rate. Not Seasonally Adjusted. 1999:01:04-2024:07:12. From FRED.



Unemployment Rate (in %). Seasonally Adjusted. 1948:01-01 - 2024:06-01. From FRED.



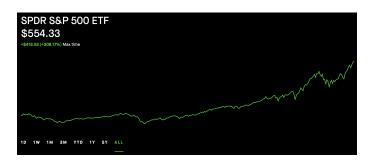
- U.S. Liquor Retail Sales: Beer, Wine, and Liquor Stores
- Millions of Dollars, Not Seasonally Adjusted
- 1992-01-01 to 2024-05-01, Monthly



- SP-500 returns
- Percentage Change, Not Seasonally Adjusted
- 2014-12-01 to 2024-07-23, Daily



SPDR S&P 500 ETF (Ticker: SPY). Not Seasonally Adjusted. 2000:01-30 - 2024:07-21. From Robinhood



Components of any Time Series

There are **three** main components in any Time Series:

Trend

Part of a series' movement that corresponds to long-term, slow evolution

Seasonality

Part of a series' movement that repeats each year

Cycles

A catch-all phrase for various forms of dynamic behavior that link the present to the past and the future to the present

Trends and Breaks

• Trend is a slow, long run evolution in the variable that we want to model

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- It is also called deterministic trend because evolves in predictable way. How do we model a deterministic trend?

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Stochastic trends are non predictable (e.g unit roots, random walk)

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• Linear Trend

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• Quadratic Trend

$$y_t = c + \beta_1 t + \beta_2 t^2$$

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Exponential Trend

$$y_t = ce^{\beta_1 t}$$
 or $\log(y_t) = \log(c) + \beta_1 t$

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Linear Trend

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• Quadratic Trend

$$y_t = c + \beta_1 t + \beta_2 t^2$$

• Exponential Trend

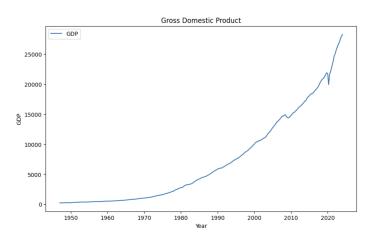
$$y_t = ce^{\beta_1 t}$$
 or $\log(y_t) = \log(c) + \beta_1 t$

There are other kinds of trends but those three approximate any trend fairly well!

Adding an explanatory variable that looks like a trend will explain the same property.

Looking Closer at GDP

Follow the Module 1 Python code file to graph this using the data 'GDP.csv'



The model we have in mind is

$$GDP_t = c + \beta t + u_t$$

We can run this using any statistical software using some OLS command and we get the following

			OLS Regr	ess:	ion Re	esults		
Dep. Variat			GE		D ca	uared:		0.865
Model:	ote.		01					0.865
						R-squared:		
Method:			t Square	-		stistic:		1973.
Date:		Tue, 23				(F-statistic):	1.03e-135
Time:			19:45:1	.0	Log-I	_ikelihood:		-2884.3
No. Observa	ations:		36	19	AIC:			5773.
Df Residua	ls:		36	7	BIC:			5780.
Df Model:				1				
Covariance	Type:		nonrobus	t				
	coe	f std	err		t	P> t	[0.025	0.975]
Constant	-4763.037	9 311	.942			0.000	-5376.853	-4149.223
Time	77.862	0 1	.753	44.	.422	0.000	74.413	81.311
Omnibus:			34.12	0	Durb:	in-Watson:		0.005
Prob(Omnibu	us):		0.00	00	Jarqu	ue-Bera (JB):		42.573
Skew:			0.96	12	Prob	(JB):		5.69e-10
Kurtosis:			3.23	86	Cond.	No.		355.

			0	LS F	Regress	ion Res	ults			
				====						
Dep. Varia	able:				GDP	R-squa	red:		0.86	
Model:					0LS	Adj. R	-squared:		0.86	
Method:			Least	Squ	ares	F-stat	istic:		1973	
Date:		Т	ue, 23	Jul	2024	Prob (F-statisti	c):	1.03e-135	
Time:				19:45:10			Log-Likelihood:			
No. Observ	ations:				309	AIC:			5773	
Df Residua	als:				307	BIC:			5780	
Df Model:					1					
Covariance	Type:		n	onro	bust					
				===:						
		coef	std	err		t	P> t	[0.025	0.975	
Constant	-4763.	0379	311.	942	-15	.269	0.000	-5376.853	-4149.22	
Time	77.	8620	1.	753	44	.422	0.000	74.413	81.31	
Omnibus:				34	1.120	Durbin	-Watson:		0.00	
Prob(Omnil	ous):			6	000	Jarque	-Bera (JB)	:	42.57	
Skew:				(9.902	Prob(J	B):		5.69e-1	
Kurtosis:				3	3.236	Cond.	No.		355	

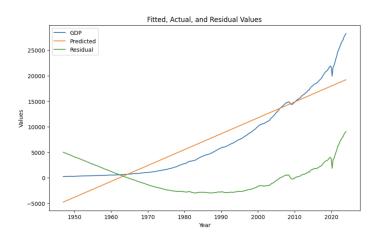
 \bullet Trend explains most of GDP variation. Typical of most time series'.

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Dep. Varia	able:				GDP	R-squa	red:		0.86	
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Date:		Т	ue, 23	Jul	2024	Prob (F-statisti	c):	1.03e-135	
Time:				19:45:10			Log-Likelihood:			
No. Observ	ations:				309	AIC:			5773	
Df Residua	als:				307	BIC:			5780	
Df Model:					1					
Covariance	Type:		n	onro	bust					
				===:						
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- Trend explains most of GDP variation. Typical of most time series'.
- R^2 is really high. Is this a good regression? Why or why not?

			0LS	Regre	ssion Re	sults		
Dep. Varia	able:			GDP	R-sau	ared:		0.865
Model:				OLS		R-squared:	:	0.865
Method:			Least S	quares		tistic:		1973.
Date:			Tue. 23 Ju	1 2024	Prob	(F-statis	tic):	1.03e-135
Time:			19	:45:10	Log-L	ikelihood	:	-2884.3
No. Observ	ations	:		309	AIC:			5773.
Df Residua	als:			307	BIC:			5780.
Df Model:				1				
Covariance	Type:		non	robust				
		coef	std er	r	t	P> t	[0.025	0.975]
Constant	-4763	.0379	311.94	2 -	15.269	0.000	-5376.853	-4149.223
Time	77	8620	1.75	3	44.422	0.000	74.413	81.311
Omnibus:				34,128	Durbi	n-Watson:		0.005
Prob(Omnil	ous):			0.000		e-Bera (J	3):	42,573
Skew:				0.902				5.69e-10
Kurtosis:				3.236				355.

- Trend explains most of GDP variation. Typical of most time series'.
- R^2 is really high. Is this a good regression? Why or why not?
- But how do we know if this is truly deterministic trend? We will spent A LOT of time talking about this and the difference between deterministic and stochastic trends

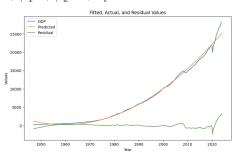


- Look at fitted values! (What do we learn?)
- Which pattern do we want the residuals to have?

Let us now try a quadratic deterministic trend model!

$$GDP_t = c + \beta_1 t + \beta_2 t^2 + u_t$$

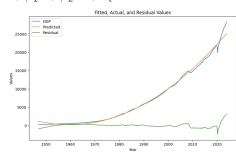
Dep. Varia	ble:			GDP	R-sq	ared:		0.99
Model:				0LS	Adj.	R-squared:		0.994
Method:		L	east Squ	ares	F-sta	tistic:		2.683e+84
Date:		Tue,	23 Jul	2024	Prob	(F-statistic):		0.00
Time:			20:1	9:05	Log-l	ikelihood:		-2395.6
No. Observ	ations:			309	AIC:			4796
Df Residua	ls:			306	BIC:			4807.
Df Model:				2				
Covariance	Type:		nonro	bust				
	coe	f :	std err		t	P> t	[0.025	0.975]
Constant	1174.292	8	95.795	12	258	0.000	985.793	1362.792
Time	-38.177	1	1.437	-26	570	0.000	-41.884	-35.356
Time_sq	0.376	8	0.005	83	425	0.000	0.368	0.386
Onnibus:			164	1.549	Durb	in-Watson:		0.096
Prob(Omnib	us):			.000	Jarq	ue-Bera (JB):		1611.955
Skew:			1	.969	Prob	(JB):		0.00
Kurtosis:			11	1.474	Cond	Mo		1.27e+85



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Dep. Varia	ble:		GDP	R-squ	ared:		0.994
Model:			OLS	Adj.	R-squared:		0.994
Method:		Least So	uares	F-sta	tistic:		2.683e+84
Date:		Tue, 23 Jul	2024	Prob	(F-statistic)	:	0.00
Time:		20:	19:05	Log-L	ikelihood:		-2395.0
No. Observ	ations:		309	AIC:			4796.
Df Residua	ls:		306	BIC:			4807.
Df Model:			2				
Covariance	Type:	non	robust				
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Skew:			1.969	Prob(JB):		0.00
Kurtosis:			13.474	Cond.	No.		1.27e+85

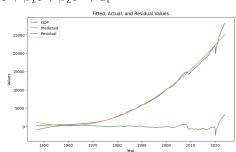


What can you say from the output? What about from the graph of the residuals?

Let us now try a quadratic deterministic trend model!

$$GDP_t = c + \beta_1 t + \beta_2 t^2 + u_t$$

Dep. Varia	ble:				GDP	R-squ	ared:		0.994
Model:					0LS	Adj.	R-squared:		0.994
Method:			Least	Squa	res	F-sta	tistic:		2.683e+84 8.88
Date:		T	ue, 23	Jul 2	2024	Prob	(F-statistic):		
Time:			20:19:05 309			Log-l	ikelihood:	-2395. 4796	-2395.6
No. Observ	ations:					AIC:			
Df Residua	ls:				306	BIC:			4807.
Df Model:					2				
Covariance	Type:		n	onrob	oust				
		coef	std	err		t	P> t	[0.025	0.975]
Constant	1174	2928	95.	795	12.	258	0.000	985.793	1362.792
Time	-38	1771	1.	437	-26	570	0.000	-41.884	-35.356
Time_sq	0.	3768	0.	805	83.	425	0.000	0.368	0.386
Onnibus:				164.	549	Durb:	n-Watson:		0.096
Prob(Omnib	us):			0.	808	Jarqu	e-Bera (JB):		1611.955
Skew:				1.	969	Prob	JB):		0.00
Kurtosis:				13.	474	Cond.	No.		1.27e+85



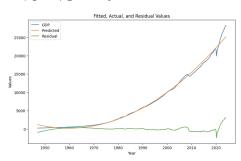
- What can you say from the output? What about from the graph of the residuals?
- In general we would like the residuals to be inside the interval in the graph.

Let us now try a quadratic deterministic trend model!

$$GDP_t = c + \beta_1 t + \beta_2 t^2 + u_t$$

		ULS P	egression			
Dep. Varia	ble:		GDP R-	squared:		0.994
Model:			OLS Ad	j. R-squared:		0.994
Method:		Least Squ	ares F-	statistic:		2.683e+84
Date:		Tue, 23 Jul	2024 Pr	ob (F-statist	ic):	0.00
Time:		20:1	9:05 Lo	g-Likelihood:		-2395.0
No. Observ	ations:		309 AI	C:		4796.
Df Residua	ls:		306 BI	C:		4807.
Df Model:			2			
Covariance	Type:	nonro	bust			
	coe	std err		t P> t	[0.025	0.975]
Constant	1174.292	95.795	12.25	8 0.000	985.793	1362.792
Time	-38.177	1.437	-26.57	0 0.000	-41.884	-35.350
Time_sq	0.376	0.005	83.42	5 0.000	0.368	0.386
Omnibus:		164	.549 Du	rbin-Watson:		0.090
Prob(Omnib	us):	9	.000 Ja	rque-Bera (JE	t):	1611.955
Skew:		1	.969 Pr	ob(JB):		0.00
Kurtosis:		13	.474 Co	nd. No.		1.27e+85

OLC Bearessian Beaults



Deterministic Trends do not have an economics interpretation, yet it is incredibly powerful for forecasting.

How do we select between different models?

• The model with the highest R^2 is not always the best model for out of sample forecast.

How do we select between different models?

- The model with the highest R^2 is not always the best model for out of sample forecast.
- Most ICs attempt to find the model with the smallest out-of-sample 1-step-ahead mean squared error.

$$MSE = \frac{\sum_{t=1}^{T} e_t^2}{T}$$
 where $e_t = y_t - \hat{y}_t = \hat{u}_t$

R² and Information Criteria

Note that

$$R^{2} = 1 - \frac{RSS}{TSS} = 1 - \frac{\sum e_{t}^{2}}{\sum (y_{t} - \bar{y})^{2}}$$

Where TTS depends only on the data.

Smallest MSE \Leftrightarrow highest R^2

We could look at the model with smallest MSE. We said that R^2 increases even if we include irrelevant variables so that we needed to correct for the degrees of freedom (d.f.). The same is true for MSE.

Different Information Criteria correct MSE for the degrees of freedom with different weights.

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• Akaike Information Criteria (AIC)

$$AIC = e^{\frac{2k}{T}}MSE$$

• Schwartz Information Criteria (SIC)

$$SIC = T^{\frac{k}{T}}MSE$$

Information Criteria

Different Information Criteria correct MSE for the degrees of freedom with different weights.

• Akaike Information Criteria (AIC)

$$AIC = e^{\frac{2k}{T}}MSE$$

• Schwartz Information Criteria (SIC)

$$SIC = T^{\frac{k}{T}}MSE$$

Between two models, pick the one with the smallest AIC and SIC.

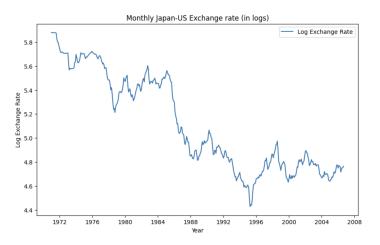
Breaks

Breaks

Often we may want to model a break in the mean or a break in trend.

- We can use dummy variables to model that.
- Break in the intercept or break in the slope of the trend? (think of an example)
- If we know the date of the break (new law, big event..) then we can impose that break. We should test for significance of that break.
- If we do not know the date of the break we can test for a break at unknown time.

Some time we may be interested in estimating a "broken trend". Often there are reason to expect that there is a structural break at a specific date.



As before, we can of course estimate a linear trend

$$yen_t = c + \beta t + u_t$$

		OLS B	earess	ion Re	sults		
		000 11					
Dep. Variable	:	Lo	g_ER	R-squ	ared:		0.83
Model:			0LS	Adj. I	R-squared:		0.83
Method:		Least Squ	ares	F-sta	tistic:		2096
Date:	1	ue, 23 Jul	2024	Prob	(F-statistic):		7.93e-167
Time:		20:5	4:44	Log-L	ikelihood:		165.73
No. Observati	ons:		429	AIC:			-327.5
Df Residuals:			427	BIC:			-319.3
Df Model:			1				
Covariance Ty	pe:	nonro	bust				
	coef	std err		t	P> t	[0.025	0.975
Constant	5.7447	0.016	361	.595	0.000	5.713	5.776
Time	-0.0029	6.43e-05	-45	.786	0.000	-0.003	-0.003
Omnibus:		27	.346	Durbi	n-Watson:		0.027
Prob(Omnibus)	:	0	.080	Jarque	e-Bera (JB):		17.74
Skew:		-0	.368	Prob(.	JB):		0.080140
Kurtosis:		2	. 328	Cond.	No.		494



It looks like there is a break in the trend around 1986. Before mid 1986 the trend is flatter and it is decreasing after that!

To model this we use Dummy variables!

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A **dummy variable** is a variables that only take the values of 0 and 1.

Define D_t as follows

$$D_t = \begin{cases} 0 \text{ if } t < \text{December 1985} \\ 1 \text{ if } t > \text{December 1985} \end{cases}$$

Let's see what model we are estimating if we include this variable in this regression.

$$yen_t = c + \beta_1 D_t + \beta_2 t + u_t$$

Clearly, before 1986, $D_t = 0$

$$yen_t = c + \beta_2 t + u_t$$

However, after 1986, $D_t = 1$

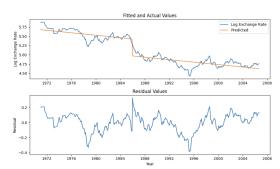
$$yen_t = (c + \beta_1) + \beta_2 t + u_t$$

Ergo, the coefficient on the dummy determines how the intercept has changed before and after 1986!

Kurtosis:

Dep. Variab	le:	Log	ER	R-squa	red:		8.917
Model:			OLS	Adj. F	-squared:		0.91
Method:		Least Squa	res	F-stat	istic:		2354.
Date:	T)	e, 23 Jul 2	824	Prob (F-statistic)		5.67e-23
Time:		20:59	:32	Log-Li	kelihood:		318.55
No. Observa	tions:		429	AIC:			-631.1
Df Residual	s:		426	BIC:			-618.9
Df Model:			2				
Covariance	Type:	nonrob	ust				
	coef	std err		t	P> t	[0.025	0.975
Constant	5.6766	0.012	489	.486	0.000	5.654	5.69
Post_85	-0.4583	0.022	-21	.038	0.000	-0.501	-8.41
Time	-0.0014	8.68e-05	-15	.912	0.000	-0.002	-8.80
Omnibus:		30.	978	Durbin	-Watson:		0.093
Prob(Omnibu	s):	0.	888	Jarque	-Bera (JB):		35.79
Skew:		_0	661	Prob()	D).		1.69e-86

3,507 Cond. No.



 How many data points are used to estimate the mean in each time period?

979.

 What does this tell us about how many data points we need in each period?

We can use the same idea to allow for the possibility of a change in the slope. Use an "interaction" term $(D_t \cdot t)$

$$yen_t = c + \beta_1 D_t + \beta_2 t + \beta_3 (D_t \cdot t) + u_t$$

Before 1986, $D_t = 0$

$$yen_t = c + \beta_2 t + u_t$$

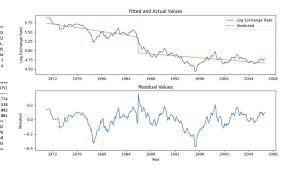
After 1986, $D_t = 1$

$$yen_t = (c + \beta_1) + (\beta_2 + \beta_3) t + u_t$$

The coefficient on the dummy determines how the intercept has changed before and after 1986, the coefficient on the interactive term determines how the slope has changed before and after 1986.

Dep. Variable:	Log_ER	R-squared:	0.922
Model:	0LS	Adj. R-squared:	0.922
Method:	Least Squares	F-statistic:	1679.
Date:	Tue, 23 Jul 2024	Prob (F-statistic):	3.64e-235
Time:	21:03:39	Log-Likelihood:	332.34
No. Observations:	429	AIC:	-656.7
Df Residuals:	425	BIC:	-640.4
Df Model:	3		
Covariance Type:	nonrobust		

	coef	std err	t	P> t	[0.025	0.97
Constant	5.7417	0.017	345.225	0.000	5.709	5.7
Post_85	-0.6069	0.035	-17.313	0.000	-0.676	-0.5
Time	-0.0021	0.000	-13.120	0.000	-0.002	-0.0
Time_Post_85	0.0010	0.000	5.312	0.000	0.001	0.0
Omnibus:		35.302	Durbin-V	latson:		8.892
Prob(Omnibus):		0.086	Jarque-E	Bera (JB):		43.676
Skew:		-0.671	Prob(JB)	:		3.28e-10
Kurtosis:		3.886	Cond. No			2.27e+03



Testing for Breaks

Why is important to detect breaks?

 If a break occurs in the population regression model in the sample, then the OLS will estimate a relationship that holds 'on average'.
 The average combines the two periods!

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Why is important to detect breaks?

- If a break occurs in the population regression model in the sample, then the OLS will estimate a relationship that holds 'on average'.
 The average combines the two periods!
 - Depending on the location and the size of the break the estimated 'average' relationship can be very different from the true population line
 - For this reason it is important that we test the presence of breaks in time series.

• Sometime we may suspect that there is a break at a certain date (e.g. Bretton Woods (1973)).

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We want to test the null hypothesis of no break. What is the null?
 Which test can we use? Answer: Let's try the Chow's Test!

Operationalizing the Chow's Test

Consider the model

$$yen_t = c + \beta_1 D_t + \beta_2 t + \beta_3 D_t \cdot t + u_t$$

We want to test the hypothesis

$$H_0: \beta_2 = \beta_3 = 0$$

 H_a : otherwise

The null hypothesis of joint insignificance of D can be run as an F-test with k and $N_1 + N_2 - 2k$ degrees of freedom where N_1 is $card(t_1, ..., t^*)$ and N_2 is $card(t^* + 1, ..., T)$ and k parameters.

$$F = \frac{(RSS_C - (RSS_1 + RSS_2))/k}{(RSS_1 + RSS_2)/(N_1 + N_2 - 2k)}$$

Note: The same result can be achieved with dummy variables!

Operationalizing the Chow's Test

```
# Run a Chow test to test for structural break
         # Define the two subsamples
         er_pre = er[er['DATE'] <= '1985-12-31']
         er post = er[er['DATE'] > '1985-12-31']
        # Run the regression for the two subsamples
         model_pre = sm.OLS(er_pre['Log_ER'], er_pre[['Constant', 'Time']])
         results pre = model pre.fit()
         model post = sm.OLS(er post['Log ER'], er post[['Constant', 'Time']])
         results_post = model_post.fit()
         # Compute the sum of squared residuals for the two subsamples
         SSR pre = np.sum(results pre.resid**2)
         SSR post = np.sum(results post.resid**2)
         # Compute the total sum of squared residuals
        SSR_total = np.sum((er['Log_ER'] - np.mean(er['Log_ER']))**2)
         # Compute the Chow test statistic
         Chow = ((SSR total - (SSR pre + SSR post)) / 2) / ((SSR pre + SSR post) / (len(er) - 4))
         Chow
[17] V 0.0s
                                                                                                                                                 Python
     2518 020081600110
```

Clearly, we find that we reject the null hypothesis! The same conclusion was reached through the use of dummy variables.

- More often we don't know when the break occurred but we suspect that it was sometime between date t₀ and t₁. This is a much more interesting question!
- The Chow test can be modified to to handle this by testing for breaks at all possible dates between t₀ and t₁ and then taking the largest of the resulting F-tests.
- This is called the Quandt test (or sup-Wald test)
- Because the Quandt test is the *largest* of an individual F-test, its distribution is not the same as the individual tests but it will have its own distribution.

Digression: Use of Dummy Variables

Dummy variables are very useful. Once you understand the role of the dummy variables and the interactive terms, we can apply the same principle to many questions:

- Test if the demand is on average different at different times of the years (i.e. seasonality, more examples later).
- Test if the elasticity of demand is different at different times of the year (car sales example etc).
- · Isolate specific times of the year

Remember the question before. How many data points do we use in each of the periods we isolate?

Seasonality

Seasonality

- Seasonality is a pattern that repeats every year.
- From a micro point of view, seasonality comes from links of technology, preferences and institutions to the calendar. (examples?).
- We will only look at deterministic seasonality = repetition is exact and predictable.
- Seasonality is a very typical component of Time Series.

How do we deal with Seasonality?

1. If we are interested in forecasting non seasonal fluctuations we may want to remove seasonality and work with seasonally adjusted series.

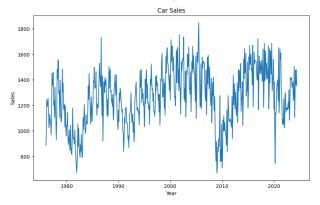
Do we really want to do this? In general we want to forecast all variations in the series.

2. We can take seasonality into account in our forecast and model seasonality.

Often data are already seasonally adjusted (SA), be careful when you download data which kind of data you really want. Data that is SA is passed through a complicated filter.

How to Model Seasonality

- To model seasonality, we use dummy variables!
- To motivate, let us use the cars.csv dataset which spans from 1967:1 to 2024:6
- It looks like the sales of cars are very different.



Reminder: Dummy Variable Trap

- If you run your regression with the constant, you need to define one less dummy than categories (in the example, 4 seasons => 3 dummy variables).
- Alternatively, you can define the same number of dummies as categories and then run the regression without the constant.

Intercept and Slope in a Dummy Model

We formulate a pure seasonal dummy model w/ one explanatory variable (for now)

$$cars_t = c + \beta_1 Q_1 + \beta_2 Q_2 + \beta_3 Q_3 + \beta_4 price_t + u_t$$

Clearly from here, we know that

$$\mathbb{E}(cars_t|winter) = c + \beta_1 + \beta_4 price$$

 $\mathbb{E}(cars_t|spring) = c + \beta_2 + \beta_4 price$
 $\mathbb{E}(cars_t|summer) = c + \beta_3 + \beta_4 price$
 $\mathbb{E}(cars_t|fall) = c + \beta_4 price$

The intercept is different for each season. What does this mean in economic terms?

Some Results

OLS Regression Results

				====			
Dep. Variab	ile:	Car Sa	les	R-sq	uared:		0.191
Model:			DLS	Adj.	R-squared:		0.185
Method:		Least Squa	res	F-st	atistic:		34.00
Date:		Tue, 20 Aug 20	324	Prob	(F-statistic)	:	1.69e-25
Time:		15:28	:13	Log-	ikelihood:		-3906.6
No. Observa	tions:		582	AIC:			7823.
Df Residual	s:		577	BIC:			7845.
Df Model:			4				
Covariance	Type:	nonrob	ıst				
	coef	std err		t	P> t	[0.025	0.975]
Constant	826.2518	46.051	17.	942	0.000	735.804	916.700
Q1	-11.4374	23.436	-0.	488	0.626	-57.467	34.592
Q2	140.0223	23.435	5.	975	0.000	93.995	186.056
Q3	69.5385	23.556	2.	952	0.003	23.272	115.805
Car Price	2.9879	0.330	9.	048	0.000	2.339	3.637
Omnibus:		2.0	571	Durb	in-Watson:		0.858
Prob(Omnibu	is):	0.3	263	Jarq	ue-Bera (JB):		2.658
Skew:		-0.	165	Prob	(JB):		0.265
Kurtosis:		2.9	973	Cond	No.		768.

Notes:

For a given car price, the average sales are higher in the second and third quarters than in the fourth quarter.

^[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

Some Extensions

- What is the economic intuition of our estimates?
- What other variable/s would you include in this model?
- Do you think that differences in the average demand for cars across seasons is the only kind of seasonality we need to take into account?

Interaction Terms

- It may also be that not only is the sales higher in Q2 for a given price but also that the demand for cars is more sensitive to the interest rate in the different seasons!
- We want to model possible shifts in the slope of the regression line for different seasons
- We can do this by looking at the interaction of the dummy variables with the price variable

Interactive Dummy Variables

Consider this simple model with intercept and slope dummy

$$\textit{cars}_t = c + \beta_1 Q 1 + \beta_2 Q_2 + \beta_3 Q_3 + \beta_4 p_t + \beta_5 (Q_1 \cdot p_t) + \beta_6 (Q_2 \cdot p_t) + \beta_7 (Q_3 \cdot p_t) + u_t$$

. Varia	le:		ar Sale	s R-squa	red:		0.195	
iel:			0L	S Adj. R	-squared:		0.185	
ate: Tue, 20		Least Squares Tue, 20 Aug 2024		s F-stat				
				4 Prob (F-statisti	:):	7.26e-24	
		16:13:2	3 Log-Li	kelihood:		-3905.1		
No. Observations:				2 AIC:			7826.	
Residua	s:		57	4 BIC:			7861.	
Model:				7				
variance	Type:		onrobus	t				
	coef	std	err	t	P> t	[0.025	0.975]	
stant	748.0431	. 89.	883	8.322	0.000	571.503	924.583	
	142.5440	123.	148	1.157	0.248	-99.332	384.420	
	279.7676	124.	202	2.253	0.025	35.822	523.713	
	72.6679	126	668	0.574	0.566	-176.122	321.457	
Price	3.5895	0.	679	5.283	0.000	2.255	4.924	
Price	-1.1878	0.	932	-1.274	0.203	-3.019	0.644	
Price	-1.0761	. 0.	939	-1.146	0.252	-2.921	0.769	
Price	-0.0206	0.	960	-0.021	0.983	-1.907	1.865	
ibus:			2.87	3 Durbin	-Watson:		0.850	
			Jarque-Bera (JB):					

What is the underlying model for each season?

Forecasting with Seasonality

- Given that we are only looking at deterministic seasonality, forecasting with seasonal dummy variables is very easy! (i.e. we know exactly how the dummy variable looks in the future)
- You can predict what the value for the dummy will be next quarter and you can use the estimated values for your parameters to forecast.
- We can re-estimate using only part of our sample and see how we are doing in forecasting.
- We don't have a good R² so most likely we will not do too well in forecasting.

Cycles

Cycles

- Cycles include any sort of dynamics that is not captured by trend and seasonality.
- This includes dynamic, persistence, and any way in which the present is linked to the past or the future.

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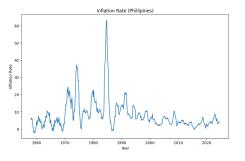
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- The economy is subject to shocks! These shocks move the economy up or down in the period that occurs, then persist for a period of time.
- The shocks are represented by the values of the error terms.
- Think of technology shocks slowly moving to different sectors.

Stationarity

- Before modeling cycles, we need to define when a time series is stationary since we can only deal (for the most part) with data that is stationary.
- Think about a time series. We observe a part of the path of the series. In theory, a time series begins in the infinite past and continues in the infinite future and we only observe a small part of it.
- Since we only observe a finite period, we would like the series to be stable over all periods.
- Stationarity is then very important!

Autocorrelation

When we think of cycles, we often think about *autocorrelation*, the correlation of a series with itself lagged.



	Inflation Rate	Change in the Inflation Rate
Lag 1	0.983451	0.443811
Lag 2	0.952210	0.311938
Lag 3	0.910634	0.233348
Lag 4	0.861315	0.186828

- The inflation rate is **highly serially correlated** since $\rho(1) = 0.98$
- Last month's inflation rate contains much information about this month's inflation. Moreover, the plot is dominated by multiyear swings but there are still surprise movements!

How do we model cycles?

- The best way to capture the fact that things move slowly is to include a lagged dependent variable as an explanatory variable (ala AR(1)). Why the name?
- I could also add two lags $\implies AR(2)$
- Or we can use *MA* or *ARMA* models! We will learn to see which model/s are best when
- You can really only talk about AR terms when you have stationary data. If you don't (unit roots) you have to transform the data so they are stationary (i.e. through differencing)

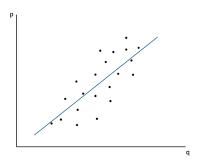
Cycles - in Practice

- Most time series will need at least an AR(1), often not more than an AR(2)
- Once you add lags and other variables, the coefficients may become insignificant. Intuition?
- R² will increase significantly!
- AR terms are your best friends when it comes to forecasting.
- They are at the core of time series analysis.
- Always look at the residuals or the ACF or PACF of the residuals to make sure there is no serial correlation in the errors. If there is, add one more lag.

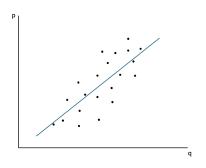
Other Topics

- What if there are breaks? → Testing
- How do we forecast and evaluate the forecasts?
- What if there is autocorrelation in the variance instead of the mean
 → ARCH and GARCH
- Extension to multivariate → VAR and SVAR
- \bullet What if the data is non stationary? \to Unit Root, Unit Root Testing, Cointegration, VECM
- Estimation in the Frequency Domain
- Kalman Filter
- ... We will see how much time we have!

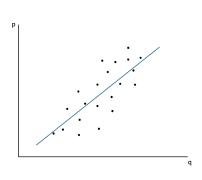
Structural vs Time Series Models



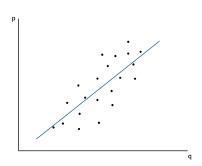
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- The estimator is not correct, it is biased! The slope is not estimating the elasticity. There is a simultaneity bias!
- The prices are a combined effect of supply and demand, because both are price contingent but also influence prices simultaneously!

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 - If you set prices based on the expectation that you will have an increase in demand...
 - You will not be estimating the elasticity by simply regressing demand on prices
- If we really want to estimate the slope of the demand function, we need to make sure that the RHS variable is
 - "exogenous/predetermined" (e.g. say weather)

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- In pure time series, we do not worry about giving econ interpretations to coefficients because we are only interested in forecasting and not the structural parameters
- What you do is a combination of these two which introduces difficulties in interpreting your coefficients.

The models you currently use somewhere in between: prices are in a way predetermined, gas cost is not. \rightarrow Your models are a mix of the two

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- We cannot always give a causal interpretation unless we are sure that p is exogenous
- This will not affect the quality of your forecasts!