HW6

2023-09-30

42*7.5. Refer to Patient satisfaction Problem 6.15.

7.5a Obtain the **analysis of variance table** that decomposes the regression sum of squares into **extra sums of squares** associated with **X2**; with **X1**, given X2; and with **X3**, given X2 and X1.

Note: For 7.5 (a) ANOVA table, follow the example Tables 7.3 & 7.4 in the textbook (pp.261-262), but pay attention that in Problem 7.5, X2 is entered first instead of X1.

	SS	MS	DF	
Regression (R)	9120.5	3040.167	3	
(X2)	4860.3	4860.3	1	n=46
(X1 X2)	3896	3896	1	p=4
(X3 X2,X1)	364.2	364.2	1	
Error	4248.8	101.1619	42	
Total	13369.3	297.0956	45	

```
144 - # -----
145 # 7.5 Fit the 2nd model x2, x1, x3.
146 m2 <- lm(Satisfaction ~ Illness + Age + Anxiety, data = Satisfaction)
147 summary(m2)
148 anova(m2)
149
150 n <- 46
151 p <- 4
152 SSR <- 4860.3 + 3896.0 + 364.2
153 SSE <- 4248.8
154 SSTO <- SSR + SSE
155 MSE <- 101.2
156 MSR <- SSR / (p-1)
157 MSTO <- SSTO / (n-1)
158 SSR;SSE;SSTO
159 MSR; MSE; MSTO
> anova(m2)
Analysis of Variance Table
Response: Satisfaction
          Df Sum Sq Mean Sq F value
                                     Pr(>F)
         1 4860.3 4860.3 48.0439 1.822e-08 ***
Illness
          1 3896.0 3896.0 38.5126 2.008e-07 ***
Age
Anxiety
         1 364.2
                    364.2 3.5997 0.06468 .
Residuals 42 4248.8
                    101.2
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
> SSR;SSE;SSTO
[1] 9120.5
[1] 4248.8
[1] 13369.3
> MSR; MSE; MSTO
[1] 3040.167
[1] 101.2
[1] 297.0956
```

7.5b. Test whether X3 can be dropped from the regression model given that X1, and X2 are retained.

Use the F* test statistic and level of significance 0.025. State the alternatives, decision rule, and conclusion. What is the P -yalue of the test?

Note: For 7.5 (b), an example is given in the textbook Section 7.2 (the first example), which shows you can use quantities available in the ANOVA table you created in part (a) to compute the test statistic. Alternatively, you can also use the "general linear test approach".

```
163 - ############ The F-test of extra sum of squares ###################
164 # b. whether X3 can be dropped from the reg. model given X1 & X2 being retained.
165 # HO: \beta3=0, X3 can be dropped.
166 # H1: \beta1≠0, X3 should be included in the model.
167 MSR.X3 <- 364.2
168 MSE <- 101.2
169 F.obs.X3 <- MSR.X3/MSE
170 F.obs.X3
171 1-pf(q=F.obs.x3, df1=1, df2=42)
172 pf(q = F.obs.x3, df1 = 1, df2 = 42, lower.tail=F)
173
     anova(m2)
# conclusion: With p=0.06471098 > \alpha, we failed to reject H0: β3=0.
175 # There is not enough evidence suggesting that X3 should be included.
176 # Thus, X3 can be dropped from the model.
> anova(m2)
                                                        > MSR.X3 <- 364.2
Analysis of Variance Table
                                                        > MSE <- 101.2
Response: Satisfaction
                                                       > F.obs.X3 <- MSR.X3/MSE
        Df Sum Sq Mean Sq F value
                                 Pr(>F)
                                                       > F.obs.X3
         1 4860.3 4860.3 48.0439 1.822e-08 ***
                                                       [1] 3.598814
Age 1 3896.0 3896.0 38.5126 2.008e-07 ***
Anxiety 1 364.2 364.2 3.5997 0.06468 .
Residuals 42 4248.8 101.2
                                                       > 1-pf(q=F.obs.X3, df1=1, df2=42)
                                                        [1] 0.06471098
179 # Effect size by X3, partial R-squared
180 # R2.Y3_12 <- SSR(X3|X1,2)/SSE(X1,2)
181 # R2.Y3_12 <- SSR3_12/SSE.m3
182 m3 <- lm(Satisfaction ~ Illness + Age, data = Satisfaction )
183 anova(m3)
184 SSE.m3 <- 4613.0
185 SSR3_12 <- 364.2
186 R2.Y3_12 <- SSR3_12/SSE.m3
187
188 # R2.Y3_12_1 <- (SSE.m3-SSE.m2)/SSE.m3
189 SSE.m2 <- 4248.8
190 R2.Y3_{12_1} \leftarrow (SSE.m3-SSE.m2)/SSE.m3
191 R2.Y3_12; R2.Y3_12_1
192
193 # Verify with what an add-on function gives:
194 install.packages("rsq")
195
     library(rsq)
196 rsq::rsq.partial(objF = m2, objR = m3)
197
198 # The effect size by addubg X3 is 7.89%.
199 # By addubg X3 to the model where X1 and X2 already exist, the variation in Y
200 # unexplained by the model reduces by R2y_X3|X1,2 is 7.89%.
> R2.Y3_12; R2.Y3_12_1
[1] 0.07895079
[1] 0.07895079
> rsq::rsq.partial(objF = m2, objR = m3)
$adjustment
[1] FALSE
$variables.full
[1] "Illness" "Age"
                      "Anxietv"
$variables.reduced
[1] "Illness" "Age"
$partial.rsq
[1] 0.07894201
```

```
203 # general linear test approach
204 m.F <- lm(Satisfaction ~ Illness + Age + Anxiety, data = Satisfaction)
205 m.R <- lm(Satisfaction ~ Illness + Age, data = Satisfaction )
# F-test comparing the two models
anova(m.R, m.F)
# Effect size: partial R-squared
rsq::rsq.partial(objF = m.F, objR = m.R)
# Impression: The anova comparison shows that the difference between two models
211 # are not significant with the p=0.0648. The effect size by the X3 is 7.89%.
> anova(m.R, m.F)
Analysis of Variance Table
Model 1: Satisfaction ~ Illness + Age
Model 2: Satisfaction ~ Illness + Age + Anxiety
Res.Df RSS Df Sum of Sq F Pr(>F)
     43 4613.0
      42 4248.8 1
                    364.16 3.5997 0.06468 .
2
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
> # Effect size: partial R-squared
> rsq::rsq.partial(objF = m.F, objR = m.R)
$adjustment
[1] FALSE
$variables.full
[1] "Illness" "Age"
                        "Anxiety"
$variables.reduced
[1] "Illness" "Age'
$partial.rsq
[1] 0.07894201
7.56. Test whether X's can be dropped prior X1, X2 are retained.
   Stop1. Assumption: € $\frac{\tau}{\times} N(0,\sigma^2)
            Diprostics are already completed
  Stop 2. Hypothesis
               Ho: B3=0, X3 can be dropped
               H: B2 $ 0, X3 Should be Included in the model.
   Step 3. Test Statistis.
                               Pralue = 0.0647 0.025-0
             Fobs = MSR (X3 | X1, X2) = 364.2 = 3.598814 = 3.599
   Stop 4. p-value.
             1-Pf( 2=3598814, df1=1, df2=46-4=42)=0.06471098
    Steps. With P=0.0647, greater than predetermined significance level (0.025), we fail to reject to.
             Thus, we conclude that X3(Austrely) can be dropped from the repressor model. The effect size by adding X3, where X12X2 are already exist, is 7.89%.
```

*7.6. Refer to Patient satisfaction Problem 6.15. Test whether both X2 and X3 can be dropped from the regression model given that X1 is retained. Use $\alpha = 0.025$. State the alternatives, decision rule, and conclusion. What is the P-value of the test?

Note: Solve 7.6 using the "general linear test approach".

For work done by applying the "general linear test approach", clearly state the full model and the reduced model in Step 2 (hypotheses). Then give H0 where you specify which parameters are constrained at what value, and then clearly state H1 as well.

```
215 # whether both X2 and X3 can be dropped from the regression model given that
216 # X1 is retained.
217 # HO: \beta2 = \beta3 = 0. Both X2 & X3 can be dropped. Reduced model is better.
218 # H1: \beta2\neq0 or \beta3\neq0. X2, X3 or both should be included. Full model is better.
219
220 m.F <- lm(Satisfaction ~ Age + Illness + Anxiety, data=Satisfaction)
221 m.R <- lm(Satisfaction ~ Age, data=Satisfaction)
222 anova(m.R, m.F)
223 F.obs <- 4.1768 # From anova result
224 SSR12_1 <- 845.07 # From anova result
225 SSE.m.R <- 5093.9 # From anova result
226 SSE.m.F <- 4248.8 # From anova result
227 MSE.m.F <- SSE.m.F /42 # From anova result
228 F_obs <- ((SSE.m.R-SSE.m.F)/2) / MSE.m.F # From anova result
229 F..obs <- ((SSR12_1)/2) / MSE.m.F # From anova result
230 F.obs; F_obs; F..obs
231 1-pf(q=F.obs, df1=2, df2=42) # From anova result
232 # Impression: With p=0.02216 < \alpha=0.025, we have enough evidence to reject H0.
233 # We conclude that the full model fits the data significantly better than
# the reduced model. At least one or more of X2 and X3 should be included in the model.
235
236 # Effect size
237 rsq::rsq.partial(objF = m.F, objR = m.R)
> anova(m.R, m.F)
 Analysis of Variance Table
 Model 1: Satisfaction ~ Age
 Model 2: Satisfaction ~ Age + Illness + Anxiety
              RSS Df Sum of Sq
   Res.Df
                                       F Pr(>F)
        44 5093.9
 1
 2
        42 4248.8 2
                          845.07 4.1768 0.02216 *
 signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' '1
 > F.obs <- 4.1768 # From anova result
 > SSR12_1 <- 845.07  # From anova result
> SSE.m.R <- 5093.9  # From anova result
> SSE.m.F <- 4248.8  # From anova result
 > MSE.m.F <- SSE.m.F /42 # From anova result
 > F_obs <- ((SSE.m.R-SSE.m.F)/2) / MSE.m.F # From anova result
> F..obs <- ((SSR12_1)) / MSE.m.F # From anova result
 > F.obs; F_obs; F..obs
 [1] 4.1768
 [1] 4.176968
 Γ11 4.176819
 > 1-pf(q=F.obs, df1=2, df2=42) # From anova result
 [1] 0.02216124
> rsq::rsq.partial(objF = m.F, objR = m.R)
$adjustment
[1] FALSE
$variables.full
[1] "Age"
              "Illness" "Anxiety"
$variables.reduced
[1] "Age"
$partial.rsq
[1] 0.1658989
```

```
Stepl. Assumption: E NO.072)
         1) Linear relationship between each X and Y.
         (2) No outher (stud resigned ~ V)
         (1) Not heteroskedostrc. (r.student~ predict)
         (a) Normal Error distribution N(0,52), by O.S. plot
         (5) Independent observation: randomly selected patracts.
 Stop2. Hypothesis
        Y= Bo+ B1 X, + B2 X2+ B2 X3
        Ho: \beta = \beta = om. R: Both X2, X3 can be dropped.
       H1: $2 to or $2 to > M.F : Albeat one of X20rX3
                                     Should be included to the model.
       Full model (M.F); Satisfaction= &+B, Age + bzillness + B. Annacty.
        Roduced model (m. R); Satisfaction: Bo + B, Ape
Stops. Test statistic
             SSR(F)-SSP(R)
                             SSE(P) - SSE(F)
              HelF)- dfp(R = dfe(R) - dfe(F)
                                  MSE.F
                MSE.F
               842.07
                               5093.9-4248.8
                                   44-42
                4248. 8
                                   4248.8
                                     48
          2 4. 1768
Step4. P-Value.
                                45.55e(F)
        1-pf ( 9=4.1768, Sf1= 44-42, Sf2=42
         = 0.02216124
Staps. Conclusion.
       With p=0.02216, less than 0x=0.025, we reject to.
       thus, we conclude that full model fits the data significantly better
       adactive at less one or more of Xz and X3 should be acluded.
```

*7.9. Refer to **Patient satisfaction Problem 6.15**. Test whether $\beta 1 = -1.0$ and, $\beta 2 = 0$; use $\alpha = 0.025$. State the alternatives, full and reduced models, decision rule, and conclusion.

```
Assumption: E E NOOD
      (1) tomas relationship
      (2) Outtree (no)
      (3) Not heteroskedastru
      (4) Normal sonor distribution (0.0 pts)
      (5) Independent observation (random sapty)
Styp2. Hypothesis
        Y= Bo + Bo X, +Bo X2 + Bo x3.
Test Bo = -1. Bo =0
       Full model M.F.; Y=B.+(B,X,+B2X2+B3X3
       Reduced model M.R; Y+ X,=
        Ho: BI=-1 and Bz=0
        H1: $1+1 or $2+0, The full model fits data better than the restrotes model. (01=0.025)
Stops. Test Statistics.
                          P 20.420B
                     Fobs 20.884
                    SSE(F)-SSE(F)
                                            4427.7-4248.8
                     Dfsse(R)-Dfsse(F)
                                               44-42
                                                             =0.8842
                    SSR(F)-SSR(R) = ASSR
                                                178-81
                     Of spa(F) - Of spa(R) = 2
                                                         - = 0.8837813
                        MSE.F
               ≈ 0.884
 Stop 1. P- value
           1- pf ( g=0.884, df=2, df=42)=0.4207535
Steps. Conclusion.
           With P=0.4208 > x=0.025, we fail to reject Ho.
           There is Not enough evidence supporting the restricted model is supported and defended to the full model.
          The construction model with \beta_1 = 1 and \beta_2 = 0 fets the dots
             about as well as the fall mall
           thus, we may choose the smaller model (restricted model).
```

```
243 # Test whether \beta 1 = -1.0 and, \beta 2 = 0; use \alpha = 0.025.
244 # Y = \beta 0 -1*x1 + 0*x2 + \beta 3*x3
245 # Y + X1 = \beta0 + \beta3X3
246 # Full model
247 data.full <- Satisfaction
248 data.full$Y <- data.full$Satisfaction
249 m.F <- lm(Y ~ Age + Illness + Anxiety, data=data.full)
250 # Reduced model
251 data.reduced <- Satisfaction
252 data.reduced$Y <- with(Satisfaction, Satisfaction + Age) # Y + X1
253 m.R <- lm(Y \sim Anxiety, data=data.reduced) # \beta0 + \beta3x3 254 # F-test comparing the two models
255 anova(m.R, m.F)
256 F_obs <- 0.8838
257 # F = {SSE(R-F)/DFe(R-F)}/ (SSE.F/DFe.F)
258 F.obs <- ((4427.7 - 4248.8)/(44-42))/(4248.8/42)
259 # F = {SSR.delta/Df}/MSE
260 SSR.delta <- 178.81
261 F..obs <- (SSR.delta/2)/(4248.8/42)
262 F_obs; F.obs; F..obs
263
264 1-pf(q=0.8838, df1=44-42, df2=42)
265 # with p=0.4207535 > \alpha= 0.025, we fail to reject H0.
266 # There isn't enough evidence suggesting that the constraints \beta 1=-1 and \beta 2=0 267 # are significantly different to the full model.
268 # The constrained model fits the data about as well as the full model.
> anova(m.R, m.F)
Analysis of Variance Table
Model 1: Y ~ Anxiety
Model 2: Y ~ Age + Illness + Anxiety
Res.Df RSS Df Sum of Sq F F
                                         F Pr(>F)
        44 4427.7
 1
        42 4248.8 2
                            178.81 0.8838 0.4208
> F_obs <- 0.8838
> # F = {SSE(R-F)/DFe(R-F)}/ (SSE.F/DFe.F)
 > F.obs <- ((4427.7 - 4248.8)/(44-42))/(4248.8/42)
 > # F = {SSR.delta/Df}/ MSE
 > SSR.delta <- 178.81
 > F..obs <- (SSR.delta/2)/(4248.8/42)
 > F_obs; F.obs; F..obs
 [1] 0.8838
 [1] 0.8842261
[1] 0.8837813
> 1-pf(q=0.8838, df1=44-42, df2=42)
[1] 0.4207535
```

R code and the results

Import data

```
rm(list=ls()) # Clean up the workspace for the new analysis
# Load packages: Every time opining a new R session, need to Load packages.
library(Hmisc) # describe functions of Hmisc and Psych are same, thus rename
it.

Hmisc.describe <- describe
library(psych) # describe functions of Hmisc and Psych are same, thus rename
it.

Psych.describe <- describe
library(rgl)</pre>
```

```
library(rsq)
# Set the following to your own folder
setwd("C:/Users/jyang/OneDrive - Arizona State University/10
Classes OneDrive/2023_STP530_Regression/HW6")
############ Read data, check data #############
Satisfaction <- read.table("CH06PR15.txt")</pre>
head(Satisfaction)
##
    V1 V2 V3 V4
## 1 48 50 51 2.3
## 2 57 36 46 2.3
## 3 66 40 48 2.2
## 4 70 41 44 1.8
## 5 89 28 43 1.8
## 6 36 49 54 2.9
colnames(Satisfaction) <- c("Satisfaction", "Age", "Illness", "Anxiety")</pre>
str(Satisfaction) #Displays a concise structure of an R object
                   46 obs. of 4 variables:
## 'data.frame':
## $ Satisfaction: int 48 57 66 70 89 36 46 54 26 77 ...
            : int 50 36 40 41 28 49 42 45 52 29 ...
## $ Age
                : int 51 46 48 44 43 54 50 48 62 50 ...
## $ Illness
## $ Anxiety : num 2.3 2.3 2.2 1.8 1.8 2.9 2.2 2.4 2.9 2.1 ...
attach(Satisfaction) #Adds a database to R's search path,
#allowing direct variable referencing. Use with caution.
```

7.5a.Obtain the analysis of variance table that decomposes the regression sum of squares into extra sums of squares associated with X2; with X1, given X2; and with X3, given X2 and X1.

```
# Residual diagnostics are done in the previous HW5
# 7.5 Fit the 2nd model X2, X1, X3.
m1 <- lm(Satisfaction ~ Age + Illness + Anxiety, data = Satisfaction)
m2 <- lm(Satisfaction ~ Illness + Age + Anxiety, data = Satisfaction)
summary(m2)
##
## lm(formula = Satisfaction ~ Illness + Age + Anxiety, data = Satisfaction)
## Residuals:
      Min
##
               1Q
                    Median
                               3Q
                                      Max
## -18.3524 -6.4230
                    0.5196 8.3715 17.1601
##
```

```
## Coefficients:
##
             Estimate Std. Error t value Pr(>|t|)
-0.4420 0.4920 -0.898 0.3741
## Illness
## Age -1.1416 0.2148 -5.315 3.81e-06 ***
## Anxiety -13.4702 7.0997 -1.897 0.0647 .
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 10.06 on 42 degrees of freedom
## Multiple R-squared: 0.6822, Adjusted R-squared: 0.6595
## F-statistic: 30.05 on 3 and 42 DF, p-value: 1.542e-10
anova(m2)
## Analysis of Variance Table
##
## Response: Satisfaction
          Df Sum Sq Mean Sq F value Pr(>F)
1 3896.0 3896.0 38.5126 2.008e-07 ***
## Age
## Anxiety 1 364.2 364.2 3.5997
                                    0.06468 .
## Residuals 42 4248.8 101.2
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
n <- 46
p <- 4
SSR <- 4860.3 + 3896.0 + 364.2
SSE <- 4248.8
SSTO <- SSR + SSE
MSE <- 101.2
MSR \leftarrow SSR / (p-1)
MSTO <- SSTO / (n-1)
SSR;SSE;SSTO
## [1] 9120.5
## [1] 4248.8
## [1] 13369.3
MSR; MSE; MSTO
## [1] 3040.167
## [1] 101.2
## [1] 297.0956
```

7.5b. whether X3 can be dropped from the reg. model given X1 & X2 being retained.

```
# H0: 63=0, X3 can be dropped.
# H1: 61≠0, X3 should be included in the model.
MSR.X3 <- 364.2
MSE <- 101.2
F.obs.X3 <- MSR.X3/MSE
F.obs.X3
## [1] 3.598814
1-pf(q=F.obs.X3, df1=1, df2=42)
## [1] 0.06471098
pf(q = F.obs.X3, df1 = 1, df2 = 42, lower.tail=F)
## [1] 0.06471098
anova(m2)
## Analysis of Variance Table
##
## Response: Satisfaction
         Df Sum Sq Mean Sq F value
                                   Pr(>F)
## Age 1 3896.0 3896.0 38.5126 2.008e-07 ***
## Anxiety 1 364.2 3.5997
                                  0.06468 .
## Residuals 42 4248.8 101.2
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
# Conclusion: With p=0.06471098 > \alpha, we failed to reject H0: 63=0.
# There is not enough evidence suggesting that X3 should be included.
# Thus, X3 can be dropped from the model.
# Effect size by X3, partial R-squared
\# R2.Y3 12 \leftarrow SSR(X3|X1,2)/SSE(X1,2)
# R2.Y3 12 <- SSR3 12/SSE.m3
m3 <- lm(Satisfaction ~ Illness + Age, data = Satisfaction )
anova(m3)
## Analysis of Variance Table
## Response: Satisfaction
     Df Sum Sq Mean Sq F value
## Age
           1 3896.0 3896.0 36.317 3.348e-07 ***
## Residuals 43 4613.0 107.3
```

```
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
SSE.m3 <- 4613.0
SSR3_12 <- 364.2
R2.Y3 12 <- SSR3 12/SSE.m3
# R2.Y3 12 1 <- (SSE.m3-SSE.m2)/SSE.m3
SSE.m2 <- 4248.8
R2.Y3_12_1 <- (SSE.m3-SSE.m2)/SSE.m3
R2.Y3_12; R2.Y3_12_1
## [1] 0.07895079
## [1] 0.07895079
# Verify with what an add-on function gives:
rsq::rsq.partial(objF = m2, objR = m3)
## $adjustment
## [1] FALSE
##
## $variables.full
## [1] "Illness" "Age" "Anxiety"
## $variables.reduced
## [1] "Illness" "Age"
##
## $partial.rsq
## [1] 0.07894201
# The effect size by addubg X3 is 7.89%.
# By addubg X3 to the model where X1 and X2 already exist, the variation in Y
# unexplained by the model reduces by R2y X3/X1,2 is 7.89%.
# general linear test approach
m.F <- lm(Satisfaction ~ Illness + Age + Anxiety, data = Satisfaction)
m.R <- lm(Satisfaction ~ Illness + Age, data = Satisfaction )
# F-test comparing the two models
anova(m.R, m.F)
## Analysis of Variance Table
##
## Model 1: Satisfaction ~ Illness + Age
## Model 2: Satisfaction ~ Illness + Age + Anxiety
## Res.Df
              RSS Df Sum of Sq F Pr(>F)
## 1
        43 4613.0
        42 4248.8 1 364.16 3.5997 0.06468 .
## 2
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

```
# Effect size: partial R-squared
rsq::rsq.partial(objF = m.F, objR = m.R)
## $adjustment
## [1] FALSE
##
## $variables.full
## [1] "Illness" "Age"
                          "Anxiety"
##
## $variables.reduced
## [1] "Illness" "Age"
##
## $partial.rsq
## [1] 0.07894201
# Impression: The anova comparison shows that the difference between two
models
# are not significant with the p=0.0648. The effect size by the X3 is 7.89%.
```

7.6. whether both X2 and X3 can be dropped from the regression model given that X1 is retained.

```
# Use "general linear test approach"
# H0: 62 = 63 = 0. Both X2 \& X3 can be dropped. Reduced model is better.
# H1: 62≠0 or 63≠0. X2, X3 or both should be included. Full model is better.
m.F <- lm(Satisfaction ~ Age + Illness + Anxiety, data=Satisfaction)</pre>
m.R <- lm(Satisfaction ~ Age, data=Satisfaction)
anova(m.R, m.F)
## Analysis of Variance Table
## Model 1: Satisfaction ~ Age
## Model 2: Satisfaction ~ Age + Illness + Anxiety
## Res.Df
              RSS Df Sum of Sq F Pr(>F)
## 1
        44 5093.9
## 2
        42 4248.8 2
                      845.07 4.1768 0.02216 *
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
F.obs <- 4.1768 # From anova result
SSR12_1 <- 845.07 # From anova result
SSE.m.R <- 5093.9 # From anova result
SSE.m.F <- 4248.8 # From anova result
MSE.m.F <- SSE.m.F /42 # From anova result
F obs <- ((SSE.m.R-SSE.m.F)/2) / MSE.m.F # From anova result
F..obs <- ((SSR12_1)/2) / MSE.m.F # From anova result
F.obs; F_obs; F..obs
## [1] 4.1768
## [1] 4.176968
```

```
## [1] 4.176819
1-pf(q=F.obs, df1=2, df2=42) # From anova result
## [1] 0.02216124
# Impression: With p=0.02216 < \alpha=0.025, we have enough evidence to reject H0.
# We conclude that the full model fits the data significantly better than
# the reduced model. At least one or more of X2 and X3 should be included in
the model.
# Effect size
rsq::rsq.partial(objF = m.F, objR = m.R)
## $adjustment
## [1] FALSE
## $variables.full
## [1] "Age" "Illness" "Anxiety"
##
## $variables.reduced
## [1] "Age"
##
## $partial.rsq
## [1] 0.1658989
# The effect size by both X2 & X3 is 16.59%. By adding both X2 and X3 to the
# model given X1, the Y unexplained reduces by 16.59%.
7.9. Test whether \beta 1 = -1.0 and, \beta 2 = 0; use \alpha = 0.025.
# Test whether \theta1 = -1.0 and, \theta2 = 0; use \alpha = 0.025.
# Y = 60 -1*X1 + 0*X2 + 63*X3
# Y + X1 = 60 + 63X3
# Full model
data.full <- Satisfaction
data.full$Y <- data.full$Satisfaction</pre>
m.F <- lm(Y ~ Age + Illness + Anxiety, data=data.full)
# Reduced model
data.reduced <- Satisfaction</pre>
data.reduced$Y <- with(Satisfaction, Satisfaction + Age) # Y + X1</pre>
m.R <- lm(Y ~ Anxiety, data=data.reduced) # 60 + 63X3
# F-test comparing the two models
anova(m.R, m.F)
## Analysis of Variance Table
##
## Model 1: Y ~ Anxiety
## Model 2: Y ~ Age + Illness + Anxiety
```

Res.Df RSS Df Sum of Sq F Pr(>F)

```
## 1 44 4427.7
## 2
         42 4248.8 2 178.81 0.8838 0.4208
F_obs <- 0.8838
\# F = \{SSE(R-F)/DFe(R-F)\}/(SSE.F/DFe.F)
F.obs <- ((4427.7 - 4248.8)/(44-42))/(4248.8/42)
# F = {SSR.delta/Df}/ MSE
SSR.delta <- 178.81
F..obs <- (SSR.delta/2)/(4248.8/42)
F_obs; F.obs; F..obs
## [1] 0.8838
## [1] 0.8842261
## [1] 0.8837813
1-pf(q=0.8838, df1=44-42, df2=42)
## [1] 0.4207535
# With p=0.4207535 > \alpha= 0.025, we fail to reject H0.
# There isn't enough evidence suggesting that the constraints 61=-1 and 62=0
# are significantly different to the full model.
# The constrained model fits the data about as well as the full model.
# Effect size: partial R-squared
rsq::rsq.partial(objF = m.F, objR = m.R)
## $adjustment
## [1] FALSE
##
## $variables.full
## [1] "Age" "Illness" "Anxiety"
##
## $variables.reduced
## [1] "Anxiety"
##
## $partial.rsq
## [1] 0.5647479
# The effect size is 56.47%
```