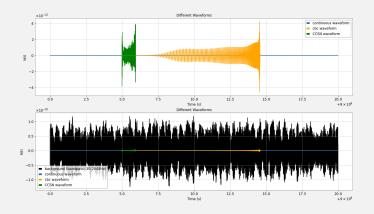
Parameter Estimation of Gravitational Waves

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The signal detected by a GW detector:

$$h(t)=D^{ij}h_{ij}(t),$$

where D^{ij} (detector tensor) depends on the detector geometry.

$$s(t)=h(t)+n(t),$$

with $|h(t)| \ll |n(t)|$ in reality.

If we have some knowledge of h(t), by multiplying s(t) and h(t):

$$\frac{1}{T}\int_0^T s(t)h(t)dt = \frac{1}{T}\int_0^T h^2(t)dt + \frac{1}{T}\int_0^T n(t)h(t)dt,$$

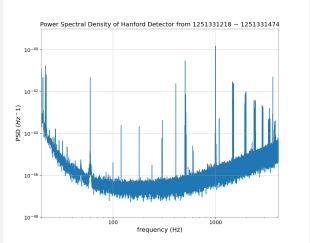
where

$$\frac{1}{T} = \int_0^T h^2(t) dt \sim h_0^2, \quad \frac{1}{T} \int_0^T n(t) h(t) dt \sim \left(\frac{\tau_0}{T}\right)^{1/2} n_0 h_0.$$

Optimal Signal-to-Noise Ratio:

$$\left(rac{S}{N}
ight)^2 = 4\int_0^\infty rac{| ilde{h}(f)|^2}{S_n(f)}df, \quad \langle | ilde{n}(f)|^2
angle = rac{1}{2}S_n(f)T.$$





Parameter Estimation of Gravitational Waves Signal

- There are models describing gravitational wave form the coalescence binary compact objects: IMRPhenom, SEOBNR, SpinTaylor, IMRSpinPrecEOB...
- From the strain data of an event, we estimate the most probable model and the parameters.
- Intrinsic Parameters: $m_1, m_2, a_1, a_2, \theta_1, \theta_2, \delta\phi, \phi_{jl}$.
- Extrinsic Parameters: $ra, dec, \theta_{jn}, \psi, d_L, \phi_c, t_c$.

Bayes' Theorem

Bayes' Theorem

$$p(\theta_i|d, H) = \frac{p(d|\theta_i, H)p(\theta_i|H)}{p(d|H)}$$

- H: model,
- θ_i : parameters for the model H,
- d: observed data,
- $p(\theta_i|H)$: prior,
- $p(d|\theta_i, H)$: likelihood,
- $p(\theta_i|d,H)$: posterior,
- p(d|H): evidence.



Bayes' Theorem

Marginalization

$$p(\theta_1|d,H) = \int_{\theta_2^{\min}}^{\theta_2^{\max}} \cdots \int_{\theta_N^{\min}}^{\theta_N^{\max}} p(\theta_1,\cdots,\theta_N|d,H) d\theta_2 \cdots d\theta_N.$$

Evidence

$$p(d|H) = \int_{\theta_1^{\min}}^{\theta_1^{\max}} \cdots \int_{\theta_N^{\min}}^{\theta_N^{\max}} p(\theta_1, \cdots, \theta_N | d, H) d\theta_1 \cdots d\theta_N.$$



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Evaluating Posterior

Bayes' Theorem

$$p(\theta_i|d,H) = \frac{p(d|\theta_i,H)p(\theta_i|H)}{p(d|H)} = \frac{L(\theta_i) \cdot \pi(\theta_i)}{Z}.$$

- We don't know the normalization constant Z.
- Use Markov Chain Monte Carlo algorithms to generate samples from $L(\theta_i) \cdot \pi(\theta_i)$ in the parameter space.

$$L(\theta_i) = \mathcal{N} \exp \left\{ -\frac{1}{2} \Big(s - h(\theta_i) | s - h(\theta_i) \Big) \right\},$$

where

$$(A|B) = 4\operatorname{Re}\int_0^\infty \tilde{A}^*(f)[S_n^{-1}(f)]\tilde{B}(f)df.$$



Markov Chain Monte Carlo Algorithm

Generate a Markov Chain:

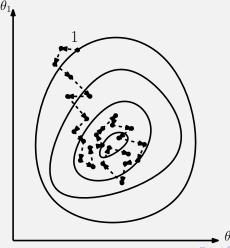
$$\left\{\theta_i^{(0)} \to \theta_i^{(1)} \to \theta_i^{(2)} \to \cdots \to \theta_i^{(N)}\right\},\,$$

The steps are stochastic and determined by the probabilities $T(\theta_i, \theta_i')$ associated with the trasition $\theta_i \to \theta_i'$.

- $T(\theta_i, \theta'_i) \geq 0$.
- The posterior is an invariant distribution of the chain: $p(\theta'_i|d,H) = \int p(\theta_i|d,H)T(\theta_i,\theta'_i)d\theta_i$.
- Detailed Balance: $p(\theta_i|d, H)T(\theta_i, \theta_i') = p(\theta_i'|d, H)T(\theta_i', \theta_i)$
- Ergodicity:
 - $\exists n$ such that $T^n(\theta_i, \theta'_i) > 0$ for all θ_i, θ'_i ,
 - $\exists \theta_i$ such that $T(\theta_i, \theta_i) > 0$.



Markov Chain Monte Carlo Algorithm



Metropolis-Hasting Sampling

- **①** Starting at a random $\theta_i^{(0)}$ in the parameter space.
- ② An update proposal θ_i' of $\theta_i^{(k)}$ is generated by sampling form a known proposal distribution $\tilde{p}(\theta_i')$ (e.g. Gaussian distribution centered at $\theta_i^{(k)}$).
- 3 Acceptance: $A(\theta_i^{(k)}, \theta_i') = \min\left(1, \frac{p(\theta_i')}{p(\theta_i^{(k)})} \frac{\tilde{p}(\theta_i^{(k)})}{\tilde{p}(\theta_i')}\right)$.
- If $A(\theta_i^{(k)}, \theta_i') \ge 1$ (accepted): record $\theta_i^{(k+1)} = \theta_i'$, else:
 - $\theta_i^{(k+1)} = \theta_i'$ with the probability: $\frac{p(\theta_i')}{p(\theta_i^{(k)})} \frac{\tilde{p}(\theta_i^{(k)})}{\tilde{p}(\theta_i')}$.
 - $\theta_i^{(k+1)} = \theta_i^{(k)}$ with the probability: $1 \frac{p(\theta_i')}{p(\theta_i^{(k)})} \frac{\tilde{p}(\theta_i^{(k)})}{\tilde{p}(\theta_i')}$.



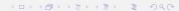
Metropolis-Hasting Sampling

- Now the Markov chain we get: $\left\{\theta_i^{(0)} \to \cdots \to \theta_i^{(N)}\right\}$ can be considered as a correction of the proposal distribution $\tilde{p}(\theta_i)$ to the posterior distribution $p(\theta_i|d,H)$.
- Now Drawing the histogram plot of $\left\{\theta_i^{(0)},\cdots,\theta_i^{(N)}\right\}$, we can see the distribution proportional to the posterior and find out the most probable parameters.

- MCMC methods: Generates samples proportional to the the posterior.
- Nested Sampling: Simultaneously estimates the evidence and the posterior.

Pros of Nested Sampling:

- well-defined stopping criteria for terminating sampling,
- generating a sequence of independent samples,
- flexibility to sample from complex, multi-modal distributions,
- the ability to derive how statistical and sampling uncertainties impact results from a single run,
- being trivially parallelizable.



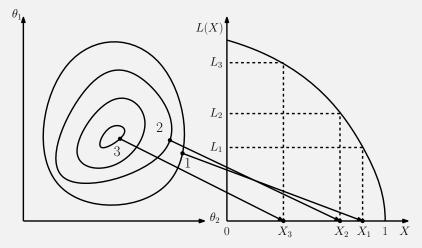
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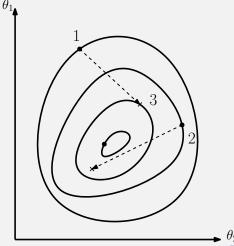
Prior Mass

$$X(\lambda) = \int_{\theta: L(\theta) > \lambda} \pi(\theta) d\theta.$$

Evidence

$$Z = \int_0^1 L(X)dX$$
, where $L(X(\lambda)) = \lambda$.





- **1** Sample N live points $\{\theta^{(1)} \cdots \theta^{(N)}\}$ from prior $\pi(\theta)$.
- While not termination condition:
 - record live point (i) with the lowest L_i as L_k ,
 - 2 assign $X_k = t_k X_{k-1}$ where t_k from $P(t_k) = Nt_k^{N-1}$,
 - **3** replace point (*i*) with sample from $\pi(\theta)$ subject to $L_i > L_k$.
- **3** Estimate evidence Z by integrating $\{L_k, X_k\}$.

Using Bilby for Parameter Estimation

- Get strain data,
- 2 Estimate Power Spectral Density,
- Operation of the state of th
- Set the prior distributions of the estimated parameters,
- Set likelihood,
- Set the sampler,
- Run the sampler!
- Opening Plot the results.





Thank you