

# Hypothesis Testing: Market Making on CSI 300 ETF via Signal-Gated DRL

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## 1 Hypothesis Testing Design

Following the criteria for a rigorous, systematic trading hypothesis, this section defines the statistical tests designed to validate the Signal-Gated DRL framework on the 510300 ETF. Given the high-frequency nature of the data, which typically exhibits fat tails and serial correlation, we move beyond simple *t*-tests to incorporate robust econometric evaluations.

### 1.1 H1: Predictive Efficacy and Forecast Robustness of SGUs

- **Subject:** The standalone Signal Generating Units (XGBoost for volatility range and LSTM for trend).
- **Independent Variables:** LOB-derived features including order flow imbalance, historical realized price ranges, and volume-weighted average price (VWAP).
- **Dependent Variables:** Actual modified realized price range ( $y_i^{RR}$ ) and financial pseudo-returns ( $y_i^{TR}$ ).
- **Anticipated Outcome:** SGUs will generate signals that are significantly more accurate than a naive drift forecast, even after adjusting for the high noise-to-signal ratio inherent in A-share ETF microstructure.
- **Validation Method:** To account for heteroscedasticity and autocorrelation (HAC) in high-frequency residuals, we employ **Newey-West adjusted standard errors**. The HAC covariance matrix estimator is defined as  $\hat{\Sigma}_{HAC} = (X'X)^{-1}\mathbf{S}(X'X)^{-1}$ , where  $\mathbf{S}$  is split to ensure proper alignment:

$$\begin{aligned} \mathbf{S} = & \sum_{t=1}^T \hat{\epsilon}_t^2 \mathbf{x}_t \mathbf{x}'_t + \\ & \sum_{j=1}^L w_j \sum_{t=j+1}^T \hat{\epsilon}_t \hat{\epsilon}_{t-j} (\mathbf{x}_t \mathbf{x}'_{t-j} + \mathbf{x}_{t-j} \mathbf{x}'_t) \end{aligned} \tag{1}$$

where  $w_j = 1 - \frac{j}{L+1}$  represents the Bartlett kernel weights for lag  $L$ . Additionally, we use the **Diebold-Mariano (DM) test**:

$$DM = \frac{\bar{d}}{\sqrt{\hat{V}(\bar{d})/T}} \sim N(0, 1) \quad (2)$$

where  $d_t = L(e_{t,\text{naive}}) - L(e_{t,\text{SGU}})$  is the loss differential between the naive baseline and the SGU prediction.

- **Significance:** We reject the null hypothesis of equal forecast accuracy if the DM statistic  $> 1.96$  ( $p < 0.05$ ).

## 1.2 H2: Nonlinear Execution Intensity and Microstructure Facts

- **Subject:** The Limit Order Book (LOB) execution dynamics near price limits.
- **Independent Variables:** Distance to the  $\pm 10\%$  price limit ( $D_{\text{limit}} = |M_t - P_{\text{limit}}|$ ).
- **Dependent Variables:** Realized fill rates  $\lambda(\delta)$ .
- **Anticipated Outcome:** The execution intensity  $\lambda(\delta)$  deviates from the standard exponential decay as  $D_{\text{limit}} \rightarrow 0$ , showing a cliff effect where liquidity vanishes.
- **Validation Method:** We use the **Durbin-Watson (DW) statistic** to detect first-order autocorrelation in the residuals of the fitted intensity model  $\lambda(\delta) = Ae^{-k\delta}$ :

$$d = \frac{\sum_{t=2}^T (e_t - e_{t-1})^2}{\sum_{t=1}^T e_t^2} \quad (3)$$

where  $e_t$  is the residual at time  $t$ .

- **Significance:** A value of  $d$  significantly lower than 2 indicates positive serial correlation, confirming that a static exponential model fails to capture the persistent microstructure dependencies near price limits.

## 1.3 H3: Comparative Superiority and Risk-Adjusted Returns

- **Subject:** DRL Agent performance vs. GLFT benchmark.
- **Independent Variables:** Strategy configuration.
- **Dependent Variables:** Terminal Wealth ( $W_T$ ) and Sharpe Ratio ( $SR$ ).
- **Anticipated Outcome:** The DRL agent achieves a significantly higher PnL-to-MAP ratio and Terminal Wealth than GLFT.

- **Validation Method:** Since HFT returns are non-Gaussian and exhibit volatility clustering, we employ **Moving Block Bootstrapping (MBB)** to construct robust confidence intervals. For a sequence of returns  $\{r_t\}$ , we compute the bootstrap distribution of the Sharpe Ratio difference  $\Delta SR^* = SR_{DRL}^* - SR_{GLFT}^*$ :

$$\text{Var}^*(\Delta SR) = \frac{1}{B-1} \sum_{b=1}^B (\Delta SR_b^* - \overline{\Delta SR}^*)^2 \quad (4)$$

We also report the asymptotic standard error for the estimated Sharpe Ratio:

$$SE(\widehat{SR}) = \sqrt{\frac{1 + \widehat{SR}^2 / 2}{T}} \quad (5)$$

- **Significance:** The DRL agent is deemed superior if the 95% MBB confidence interval for  $\Delta SR$  excludes zero.

#### 1.4 H4: Economic Rationality of Policy Skewing

- **Subject:** The learned reward-driven policy  $\pi_\theta$ .
- **Independent Variables:** Inventory level  $I_t$ .
- **Dependent Variables:** Quoting offsets  $A_t = [Q_{ask} - M, M - Q_{bid}]$ .
- **Anticipated Outcome:** The agent prioritizes spread capturing over trend chasing by enforcing a linear absolute inventory penalty:  $R_{t+1} = \text{PnL}_{t+1} - \lambda|I_{t+1}|$ .
- **Validation Method:** Analysis of Partial Dependence Plots (PDP) to estimate the marginal effect of  $I_t$  on  $A_t$ :

$$\hat{h}_S(x_S) = \frac{1}{n} \sum_{i=1}^n f(x_S, \mathbf{x}_{C,i}) \quad (6)$$

- **Significance:** The policy is verified if  $\frac{\partial \text{AskOffset}}{\partial I} < 0$  and  $\frac{\partial \text{BidOffset}}{\partial I} > 0$ , demonstrating that the agent has successfully learned to skew quotes to mitigate directional risk.

## 2 Statistical Robustness and HFT Considerations

To address the challenges of high-frequency data, we acknowledge that standard parametric tests may yield inflated  $t$ -statistics due to the high sample size ( $T$ ). Therefore, we prioritize the **Model Confidence Set (MCS)** to identify the superior strategy among a group of candidates. The MCS procedure sequentially eliminates inferior models based on a loss function  $L_{i,t}$ , until the remaining set

$\mathcal{M}^*$  contains only models with equal predictive power at a given confidence level:

$$t_{ij} = \frac{\bar{d}_{ij}}{\sqrt{\widehat{\text{Var}}(\bar{d}_{ij})}} \quad (7)$$

where  $d_{ij,t} = L_{i,t} - L_{j,t}$  represents the loss differential. This avoids the p-hacking pitfalls associated with massive datasets where even economically negligible differences appear statistically significant.

## 2.1 Microstructural Adaptation: BTC vs. CSI 300 ETF

While the core First-Passage Time (FPT) execution model is preserved, the transition from the original Bitcoin environment to the CSI 300 ETF requires specific adjustments to account for distinct microstructural regimes:

- **Finite Horizon and Session Boundary Effects:** Unlike the 24/7 continuous trading of BTC, the 510300 ETF operates within two discrete 2-hour sessions. We modify the reward logic to account for the heightened inventory risk as the market approaches the 15:00 CST close. This necessitates a time-dependent weighting of the absolute inventory penalty  $\lambda|I_{t+1}|$  to ensure the agent avoids a liquidity trap at the end of the trading day.
- **FPT Validity near Price Limits:** In the A-share market, the  $\pm 10\%$  price limit imposes a hard boundary on the FPT model. In the original BTC framework, execution intensity is assumed to decay with distance from the mid-price. For the ETF, we test the hypothesis that the FPT model remains valid but requires a volatility-gating mechanism; specifically, once the price is locked at a limit, the execution intensity  $\lambda$  drops to zero on the side of the limit, regardless of the quoting offset.
- **Tick Size and Spread Saturation:** The 510300 ETF often experiences spread saturation, where the bid-ask spread is locked at a single tick (0.001) for extended periods. This differs from the more volatile BTC spreads. Our testing framework evaluates whether the DRL agent’s continuous action space  $A_t$  allows it to exploit sub-tick order flow signals that are invisible to benchmarks utilizing crude discretization.

## 2.2 Robustness to Intraday Seasonality

To ensure that performance metrics are not skewed by high-volatility opening auctions or midday pauses, we apply **Newey-West adjustments** to all strategy regressions. This correction accounts for the significant intraday autocorrelation in returns:

$$\begin{aligned} \mathbf{S} = & \sum_{t=1}^T \hat{\epsilon}_t^2 \mathbf{x}_t \mathbf{x}'_t + \\ & \sum_{j=1}^L \left( 1 - \frac{j}{L+1} \right) \sum_{t=j+1}^T \hat{\epsilon}_t \hat{\epsilon}_{t-j} (\mathbf{x}_t \mathbf{x}'_{t-j} + \mathbf{x}_{t-j} \mathbf{x}'_t) \end{aligned} \quad (8)$$

This ensures that the reported  $p$ -values reflect the true significance of the DRL agent’s alpha beyond the expected intraday seasonality of the China A-share market.